

BOSTON COLLEGE  
 Department of Economics  
 EC 228 Econometrics, Prof. Baum, Ms. Yu, Fall 2003

**Problem Set 2 Solutions**

Problem sets should be your own work. You may work together with classmates, but if you're not figuring this out on your own, you will eventually regret it.

1.

```
. use http://fmwww.bc.edu/ec-p/data/wooldridge/ceosal2,clear

. type http://fmwww.bc.edu/ec-p/data/wooldridge/ceosal2.des
CEOSAL2.DES

salary   age      college  grad      comten   ceoten   sales    profits
mktval   lsalary  lsales  lmktval  comtensq ceotensq profmarg

Obs:    177

1. salary           1990 compensation, $1000s
2. age              in years
3. college          =1 if attended college
4. grad             =1 if attended graduate school
5. comten           years with company
6. ceoten           years as ceo with company
7. sales            1990 firm sales, millions
8. profits          1990 profits, millions
9. mktval           market value, end 1990, mills.
10. lsalary         log(salary)
11. lsales          log(sales)
12. lmktval         log(mktval)
13. comtensq        comten^2
14. ceotensq        ceoten^2
15. profmarg        profits as % of sales
```

a.

```
. summarize salary ceoten
```

Variable	Obs	Mean	Std. Dev.	Min	Max
salary	177	865.8644	587.5893	100	5299
ceoten	177	7.954802	7.150826	0	37

The average salary of CEOs' is 865.8644. The average tenure of them is 7.954802.

b.

```
. ttest salary=1000
```

One-sample t test

Variable	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
salary	177	865.8644	44.16591	587.5893	778.7015	953.0274

Degrees of freedom: 176

Ho: mean(salary) = 1000

Ha: mean < 1000	Ha: mean ~ = 1000	Ha: mean > 1000
t = -3.0371	t = -3.0371	t = -3.0371
P < t = 0.0014	P >  t  = 0.0028	P > t = 0.9986

We reject the null hypothesis that the average CEO salary is a million dollars, at a 95% confidence level.

c.

```
. ttest salary, by( college)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	5	1096.2	283.2906	633.4569	309.6593	1882.741
1	172	859.1686	44.74565	586.8337	770.8437	947.4936
combined	177	865.8644	44.16591	587.5893	778.7015	953.0274
diff		237.0314	266.7294		-289.389	763.4518

Degrees of freedom: 175

Ho: mean(0) - mean(1) = diff = 0

Ha: diff < 0	Ha: diff ~ = 0	Ha: diff > 0
t = 0.8887	t = 0.8887	t = 0.8887
P < t = 0.8123	P >  t  = 0.3754	P > t = 0.1877

We cannot reject the null hypothesis that CEOs that went to college make as much money as those who did not, at a 95% confidence level.

d.

```
. tabulate grad
```

grad	Freq.	Percent	Cum.
0	83	46.89	46.89

1	94	53.11	100.00
Total	177	100.00	

. ttest salary, by( grad)

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
0	83	867.7349	74.11516	675.2212	720.2963	1015.174
1	94	864.2128	51.71468	501.3924	761.5177	966.9079
combined	177	865.8644	44.16591	587.5893	778.7015	953.0274
diff		3.522174	88.75501		-171.6458	178.6902

Degrees of freedom: 175

Ho: mean(0) - mean(1) = diff = 0

Ha: diff < 0	Ha: diff ~ = 0	Ha: diff > 0
t = 0.0397	t = 0.0397	t = 0.0397
P > t = 0.5158	P >  t  = 0.9684	P > t = 0.4842

We cannot reject the null hypothesis that CEOs attended grad school make as much money as those who did not, at a 95% confidence level.

e.

. correlate profmarg salary  
(obs=177)

	profmarg	salary
profmarg	1.0000	
salary	-0.0289	1.0000

## 2. (2.1)

- (i) Income, age, and family background (such as number of siblings) are just a few possibilities. It seems that each of these could be correlated with years of education. (Income and education are probably positively correlated; age and education may be negatively correlated because women in more recent cohorts have, on average, more education; and number of siblings and education are probably negatively correlated.)
- (ii) Not if the factors we listed in part (i) are correlated with *educ*. Because we would like to hold these factors fixed, they are part of the error term. But if *u* is correlated with *educ* then  $E(u|educ) \neq 0$ , and so SLR.3 fails.

**3. (2.2)** In the equation  $y = \beta_0 + \beta_1 x + u$ , add and subtract  $\alpha_0$  from the right hand side to get  $y = (\alpha_0 + \beta_0) + \beta_1 x + (u - \alpha_0)$ . Call the new error  $e = u - \alpha_0$ , so that  $E(e) = 0$ . The new intercept is  $\alpha_0 + \beta_0$ , but the slope is still  $\beta_1$ .

**4. (2.3)**

Note that it would be easiest to do this problem in Stata, typing the data into the Data Editor, and using the `regress` and `predict` commands to generate the desired solutions.

- (i) Let  $y_i = GPA_i$ ,  $x_i = ACT_i$ , and  $n = 8$ . Then  $\bar{x} = 25.875$ ,  $\bar{y} = 3.2125$ ,  $\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}) = 5.8125$ , and  $\sum_{i=1}^n (x_i - \bar{x})^2 = 56.875$ . From equation (2.9), we obtain the slope as  $\hat{\beta}_1 = 5.8125/56.875 \approx .1022$ , rounded to four places after the decimal. From (2.17),  $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x} \approx 3.2125 - (.1022)25.875 \approx .5681$ . So we can write

$$\begin{aligned} \widehat{GPA} &= .5681 + .1022ACT \\ n &= 8. \end{aligned}$$

The intercept does not have a useful interpretation because  $ACT$  is not close to zero for the population of interest. If  $ACT$  is 5 points higher,  $\widehat{GPA}$  increases by  $.1022(5) = .511$ . The effect of five units' increase of the  $x$  variable can be calculated after the regression as `display 5*.b[ACT]`.

- (ii) The fitted values and residuals — rounded to four decimal places — are given along with the observation number  $i$  and  $GPA$  in the following table:

$i$	$GPA$	$\widehat{GPA}$	$\hat{u}$
1	2.8	2.7143	.0857
2	3.4	3.0209	.3791
3	3.0	3.2253	-.2253
4	3.5	3.3275	.1725
5	3.6	3.5319	.0681
6	3.0	3.1231	-.1231
7	2.7	3.1231	-.4231
8	3.7	3.6341	.0659

You can verify that the residuals, as reported in the table, sum to  $-.0002$ , which is pretty close to zero given the inherent rounding error. These could be calculated with the two commands `predict GPAhat` and `predict GPares, resid`.

- (iii) When  $ACT = 20$ ,  $\widehat{GPA} = .5681 + .1022(20) \approx 2.61$ . This can be calculated by adding a 9th observation on  $ACT$  in the Data Editor and then doing `predict GPA20 in 9/9`.
- (iv) The sum of squared residuals,  $\sum_{i=1}^n \hat{u}_i^2$  is about  $.4347$  (rounded to four decimal places), and is given in the ANOVA table regression output as

the Residual SS. The total sum of squares,  $\sum_{i=1}^n (y_i - \bar{y})^2$ , is about 1.0288, and is given as the Total SS. So the  $R$ -squared from the regression is

$$R^2 = 1 - SSR/SST \approx 1 - (.4347/1.0288) \approx .577.$$

Therefore, about 57.7% of the variation in  $GPA$  is explained by  $ACT$  in this small sample of students.

### 5. (2.5)

- (i) The intercept implies that when  $inc = 0$ ,  $cons$  is predicted to be negative \$124.84. This, of course, cannot be true, and reflects that fact that this consumption function might be a poor predictor of consumption at very low-income levels. On the other hand, on an annual basis, \$124.84 is not so far from zero.
- (ii) Just plug 30,000 into the equation:  $\widehat{cons} = -124.84 + .853(30,000) = 25,465.16$  dollars.
- (iii) The MPC and the APC are shown in the following graph. Even though the intercept is negative, the smallest APC in the sample is positive. The graph starts at an annual income level of \$1,000 (in 1970 dollars).

### 6. (2.8)

- (i) From equation (2.66),

$$\tilde{\beta}_1 = \left( \sum_{i=1}^n x_i y_i \right) / \left( \sum_{i=1}^n x_i^2 \right).$$

Plugging in  $y_i = \beta_0 + \beta_1 x_i + u_i$  gives

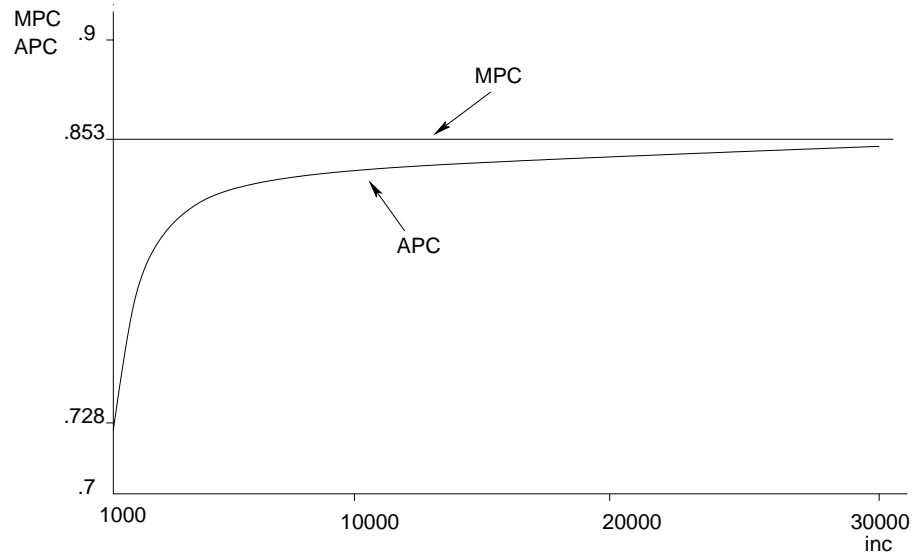
$$\tilde{\beta}_1 = \left( \sum_{i=1}^n x_i (\beta_0 + \beta_1 x_i + u_i) \right) / \left( \sum_{i=1}^n x_i^2 \right).$$

After standard algebra, the numerator can be written as

$$\beta_0 \sum_{i=1}^n x_i + \beta_1 \sum_{i=1}^n x_i^2 + \sum_{i=1}^n x_i u_i.$$

Putting this over the denominator shows we can write  $\tilde{\beta}_1$  as

$$\tilde{\beta}_1 = \beta_0 \left( \sum_{i=1}^n x_i \right) / \left( \sum_{i=1}^n x_i^2 \right) + \beta_1 + \left( \sum_{i=1}^n x_i u_i \right) / \left( \sum_{i=1}^n x_i^2 \right).$$



Conditional on the  $x_i$ , we have

$$E(\tilde{\beta}_1) = \beta_0 \left( \frac{\sum_{i=1}^n x_i}{\sum_{i=1}^n x_i^2} \right) + \beta_1$$

Because  $E(u_i) = 0$  for all  $i$ . Therefore, the bias in  $\tilde{\beta}_1$  is given by the first term in this equation. The bias is obviously zero when  $\beta_0 = 0$ . It is also zero when  $\sum_{i=1}^n x_i = 0$ , which is the same as  $\bar{x} = 0$ . In the later case, regression through the origin is identical to regression with an intercept.

**7. (2.11)**

(i)

```
. use http://fmwww.bc.edu/ec-p/data/wooldridge/ceosal2,clear
```

```
. summarize salary ceoten
```

Variable	Obs	Mean	Std. Dev.	Min	Max
salary	177	865.8644	587.5893	100	5299
ceoten	177	7.954802	7.150826	0	37

Average salary is about 865.864, which means \$865,864 because *salary* is in thousands of dollars. Average *ceoten* is about 7.95.

(ii)

```
. tabulate ceoten
```

ceoten	Freq.	Percent	Cum.
0	5	2.82	2.82
1	19	10.73	13.56
2	10	5.65	19.21
3	21	11.86	31.07
4	21	11.86	42.94
5	10	5.65	48.59
6	11	6.21	54.80
7	6	3.39	58.19
8	11	6.21	64.41
9	8	4.52	68.93
10	8	4.52	73.45
11	4	2.26	75.71
12	7	3.95	79.66
13	7	3.95	83.62
14	5	2.82	86.44
15	2	1.13	87.57
16	2	1.13	88.70
17	2	1.13	89.83
18	1	0.56	90.40
19	2	1.13	91.53
20	4	2.26	93.79
21	1	0.56	94.35
22	1	0.56	94.92
24	3	1.69	96.61
26	2	1.13	97.74
28	1	0.56	98.31
34	1	0.56	98.87
37	2	1.13	100.00
Total	177	100.00	

There are five CEOs with *ceoten* = 0. The longest tenure is 37 years.

(iii)

```
. regress lsalary ceoten
```

Source	SS	df	MS	Number of obs =	177
Model	.850907685	1	.850907685	F( 1, 175) =	2.33
Residual	63.7953139	175	.364544651	Prob > F	= 0.1284
Total	64.6462215	176	.367308077	R-squared	= 0.0132
				Adj R-squared	= 0.0075
				Root MSE	= .60378

lsalary	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
ceoten	.0097236	.0063645	1.53	0.128	-.0028374 .0222847
_cons	6.505498	.0679911	95.68	0.000	6.37131 6.639686

The estimated equation is

$$\log(\widehat{salary}) = 6.51 + 0.097ceoten$$

$$n = 177, R^2 = .013$$

We obtain the approximate percentage change in *salary* given  $\Delta ceoten = 1$  by multiplying the coefficient on *ceoten* by 100,  $100(.0097) = .97\%$ . Therefore one more year as CEO is predicted to increase salary by almost 1%.

## 8. (2.13)

(i)

```
. use http://fmwww.bc.edu/ec-p/data/wooldridge/wage2,clear
. summarize wage IQ
```

Variable	Obs	Mean	Std. Dev.	Min	Max
wage	935	957.9455	404.3608	115	3078
IQ	935	101.2824	15.05264	50	145

Average salary is about \$957.95 and average IQ is about 101.28. The sample standard deviation of IQ is about 15.05, which is pretty close to the population value of 15.

(ii)

```
. regress wage IQ
```

Source	SS	df	MS	Number of obs =	935
--------	----	----	----	-----------------	-----



Model		14589782.6	1	14589782.6	F( 1, 933) =	98.55
Residual		138126386	933	148045.429	Prob > F =	0.0000
Total		152716168	934	163507.675	R-squared =	0.0955
					Adj R-squared =	0.0946
					Root MSE =	384.77

wage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
IQ		8.303064	.8363951	9.93	0.000	6.661631 9.944498
_cons		116.9916	85.64153	1.37	0.172	-51.08078 285.0639

This calls for a level-level model:

$$\widehat{wage} = 116.99 + 8.30IQ$$

$$n = 935, R^2 = .096.$$

An increase in  $IQ$  of 15 increases predicted monthly salary by  $8.30(15) = \$124.50$  (in 1980 dollars).  $IQ$  score does not even explain 10% of the variation in  $wage$ .

(iii)

. regress lwage IQ

Source		SS	df	MS	Number of obs =	935
Model		16.4150981	1	16.4150981	F( 1, 933) =	102.62
Residual		149.241196	933	.15995841	Prob > F =	0.0000
Total		165.656294	934	.177362199	R-squared =	0.0991
					Adj R-squared =	0.0981
					Root MSE =	.39995

lwage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
IQ		.0088072	.0008694	10.13	0.000	.007101 .0105134
_cons		5.886994	.0890206	66.13	0.000	5.71229 6.061698

This calls for a log-level (single-log) model:

$$\log(\widehat{wage}) = 5.89 + .0088IQ$$

$$n = 935, R^2 = .099.$$

If  $\Delta IQ = 15$  then  $\Delta \log(\widehat{wage}) = .0088(15) = .132$ , which is the (approximate) proportionate change in predicted wage. The percentage increase is therefore approximately 13.2.

**9. (3.1)**

- (i) *hsperc* is defined so that the smaller it is, the higher the student's standing in high school. Everything else equal, the worse the student's standing in high school, the lower is his/her expected college GPA.
- (ii) Just plug these value into the equation:

$$\widehat{colgpa} = 1.392 = .0135(20) + .00148(1050) = 2.676.$$

- (iii) The difference between A and B is simply 140 times the coefficient on *sat*, because *hsperc* is the same for both students. So A is predicted to have a score  $.00148(140) \approx .207$  higher.
- (iv) With *hsperc* fixed,  $\Delta\widehat{colgpa} = .00148\Delta sat$ . Now we want to find  $\Delta sat$  such that  $\Delta\widehat{colgpa} = .5$ , so  $.5 = .00148(\Delta sat)$  or  $\Delta sat = .5/ (.00148) \approx 338$ . Perhaps not surprisingly, a large ceteris paribus difference in SAT score – almost two and one-half standard deviations – is needed to obtain a predicted difference in college GPA or a half a point.