

PROBLEM SET 4: SOLUTIONS

Point Distribution:  
 1), 2), 3) : 10 points each

1)

```
. capt program drop ps41

. program ps41, rclass
  1. drop _all
  2. quietly set obs 500
  3. generate x = rt($df)
  4. summarize x, detail
  5. return scalar mu = r(mean)
  6. return scalar med = r(p50)
  7. return scalar df = $df
  8. end

.

. global nsim 1000

. global df 100

. scalar theor = c(pi)/2

. ps41
```

x				
-----				
	Percentiles	Smallest		
1%	-2.178496	-3.30755		
5%	-1.660841	-2.662812		
10%	-1.322605	-2.465379	Obs	500
25%	-.7442768	-2.28337	Sum of Wgt.	500
50%	.0044037		Mean	-.037566
		Largest	Std. Dev.	.9865858
75%	.6558092	2.229074		

90%	1.236853	2.24305	Variance	.9733515
95%	1.585368	2.838025	Skewness	-.0096165
99%	2.126817	3.050234	Kurtosis	2.79699

```
. set seed 20110406

. foreach df in 3 6 10 100 {
2.     global df `df'
3.     qui simulate mu = r(mu) med = r(med) df = r(df), reps($nsim) nodots ///
>     saving(ps41_`df', replace): ps41
4.
. // Calculate mean squares of median and mean, given both are unbiased estimators.
.     g double med2 = med^2
5.     qui summ med2
6.     scalar numer = r(mean)
7.     g double mu2 = mu^2
8.     qui summ mu2
9.     scalar denom = r(mean)
10.    scalar ratio = numer / denom
11.    di _n "For `df' d.f., variance ratio = " ratio " vs. theoretical = " theor
12. }
```

For 3 d.f., variance ratio = .63639509 vs. theoretical = 1.5707963

For 6 d.f., variance ratio = 1.1881244 vs. theoretical = 1.5707963

For 10 d.f., variance ratio = 1.3032442 vs. theoretical = 1.5707963

For 100 d.f., variance ratio = 1.6338293 vs. theoretical = 1.5707963

Note as the degrees of freedom increase, the variance ratio approaches the theoretical ratio of  $\frac{\pi}{2} \approx 1.57$ . This makes sense since the t-distribution approximates the normal distribution as the degrees of freedom increase.

2)

```
. prog ps42, rclass
1. drop _all
2. qui set obs 50
3. g x = rnormal()
4. g z = rnormal()
5. if ($case == 1) {
6.     g eps = rnormal()
7. }
8. else if ($case == 2) {
```

```

9.          g eps = rt(5)
10. }
11. else if ($case == 3) {
12.          g eps = exp(0.2 * x) * rnormal()
13. }
14. g y = 1 + x + $gamma * z + eps
15. reg y x z
16. scalar tgamma = _b[z] / _se[z]
17. return scalar rej = tgamma^2 > 3.84
18. end

```

```
. global nsim 1000
```

```
. global gamma 0.9
```

```
. global case 1
```

```
. ps42
```

Source	SS	df	MS	Number of obs =	50
Model	74.3673906	2	37.1836953	F( 2, 47) =	38.60
Residual	45.2785697	47	.963373823	Prob > F =	0.0000
				R-squared =	0.6216
				Adj R-squared =	0.6055
Total	119.64596	49	2.44175429	Root MSE =	.98152

	y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
	x	1.09329	.1450482	7.54	0.000	.8014903	1.385089
	z	.65102	.154897	4.20	0.000	.3394075	.9626325
	_cons	1.033639	.1426087	7.25	0.000	.7467476	1.320531

```
. ret li
```

```
scalars:
```

```
    r(rej) = 1
```

```
. mat res = J(12, 3, .)
```

```
. local gamma -0.5 0.2 0.6 0.9
```

```
. local i 0
```

```

. foreach g of local gamma {
2.     forv c = 1/3 {
3.         global case 'c'
4.         global gamma 'g'
5.         local ++i
6.         mat res['i',1] = 'g'
7.         mat res['i',2] = 'c'
8.         qui simulate rej = r(rej), reps($nsim) nodots: ps42
9.         su rej, mean
10.        mat res['i',3] = r(mean)
11.    }
12. }

```

```
. mat colnames res = Gamma EpsModel PrRej
```

```
. mat list res
```

```
res[12,3]
```

	Gamma	EpsModel	PrRej
r1	-.5	1	.92
r2	-.5	2	.749
r3	-.5	3	.88
r4	.2	1	.28
r5	.2	2	.205
r6	.2	3	.303
r7	.6	1	.974
r8	.6	2	.897
r9	.6	3	.975
r10	.9	1	1
r11	.9	2	.987
r12	.9	3	1

Note that the probability of rejecting the null of  $\gamma = 0$  using the Wald statistic increases as the distance of the true  $\gamma$  from zero increases. Also notice that the probability of rejection is lower when the error process is t-distributed. The distribution in this case has fatter tails and leads to less precise tests.

3) Note: You did not need to use Mata to do this problem.

```
. capt prog drop ps43
```

```
. mata: mata clear
```

```
. mata:
```

```
----- mata (type end to exit) -----
: void function lm(string scalar vname, string scalar vname2)
```

```

> {
>   vars = tokens(vname)
>   v = vars[[1,.]]
>   st_view(X,.,v)
>   vars = tokens(vname2)
>   v = vars[[1,.]]
>   st_view(e,.,v)
>   LM = (e' * X * invsym(X' * X) * X' * e) / ((e' * e) ./ rows(e))
>   st_numscalar("lm", LM)
> }

```

note: variable X may be used before set

note: variable e may be used before set

: end

```

. prog ps43, rclass
1. drop _all
2. qui set obs 50
3. g x = rnormal()
4. g z = rnormal()
5. g iota = 1
6. if ($case == 1) {
7.     g eps = rnormal()
8. }
9. else if ($case == 2) {
10.    g eps = rt(5)
11. }
12. else if ($case == 3) {
13.    g eps = exp(0.2 * x) * rnormal()
14. }
15. g y = 1 + x + $gamma * z + eps
16. // Run the wrong regression and get the residuals.
. reg y x
17. predict double e, resid
18. // Calculate the LM statistic for this regression.
. mata: lm("x z iota","e")
19. ret scalar lm = lm
20. return scalar rej = lm > 3.84
21. end

. global nsim 1000

. global gamma 0.9

. global case 1

```

```
. ps43
```

Source	SS	df	MS	Number of obs =	50
Model	38.9568037	1	38.9568037	F( 1, 48) =	23.06
Residual	81.0804436	48	1.68917591	Prob > F =	0.0000
				R-squared =	0.3245
				Adj R-squared =	0.3105
Total	120.037247	49	2.44973974	Root MSE =	1.2997

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
x	.8544895	.1779314	4.80	0.000	.4967345 1.212244
_cons	1.12147	.1844813	6.08	0.000	.7505456 1.492395

```
. ret li
```

```
scalars:
```

```
    r(rej) = 1  
    r(lm) = 15.15231059228623
```

```
.
```

```
. mat res = J(12, 3, .)
```

```
. local gamma -0.5 0.2 0.6 0.9
```

```
. local i 0
```

```
. foreach g of local gamma {  
2.     forv c = 1/3 {  
3.         global case 'c'  
4.         global gamma 'g'  
5.         local ++i  
6.         mat res['i',1] = 'g'  
7.         mat res['i',2] = 'c'  
8.         qui simulate rej = r(rej), reps($nsim) nodots: ps43  
9.         su rej, mean  
10.        mat res['i',3] = r(mean)  
11.     }  
12. }
```

```
. mat colnames res = Gamma EpsModel PrRej
```

```
. mat list res
```

res[12,3]			
	Gamma	EpsModel	PrRej
r1	-.5	1	.918
r2	-.5	2	.767
r3	-.5	3	.886
r4	.2	1	.296
r5	.2	2	.208
r6	.2	3	.286
r7	.6	1	.976
r8	.6	2	.872
r9	.6	3	.975
r10	.9	1	.998
r11	.9	2	.985
r12	.9	3	.999

Similar to Question 2, the probability of rejecting the null of  $\gamma = 0$  using the LM statistic increases as the distance of the true  $\gamma$  from zero increases. Note that we should not be very surprised that the results in Question 2 and Question 3 are similar as the two tests are equivalent asymptotically.