

nwxtregress: Network regressions in Stata

Swiss Stata Conference 2022

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November 18, 2022

Many applications of interactions are best represented using networks I

- Empirical analysis in social sciences (nearly) invariably relies on the assumption of cross-sectional independence.
- Most real-world applications involve interactions between units of observation.
 - ▶ E.g., companies buy and sell from one another, individuals share information with family and friends, etc.
- A key question remains: how do we analyze outcomes in a regression framework in the context of networks?
 - ▶ cross-sectional independence cannot be assumed!

Many applications of interactions are best represented using networks II

- Spatial econometrics provide an answer:
 - ▶ Models dependence across cross-sectional units
 - ▶ Initially used in regional science to model neighbouring regions
 - ▶ Empirical models and estimation techniques with a priori knowledge of relationship between units (LeSage and Pace, 2009; Kelejian and Piras, 2017)

Interactions pose identification challenges

- Consider a traditional panel model with 2 units:

$$y_{1t} = X_{1t}\beta + \epsilon_{1t}$$

$$y_{2t} = X_{2t}\beta + \epsilon_{2t}$$

- The independence assumption implies: $E[\epsilon_1\epsilon_2] = E[\epsilon_1]E[\epsilon_2]$
- This rules out the possibility that units 1 and 2 interact.
- Thus, for many applications, a more appropriate model is:

$$y_{1t} = \rho y_{2t} + X_{1t}\beta + \epsilon_{1t}$$

$$y_{2t} = \rho y_{1t} + X_{2t}\beta + \epsilon_{2t}$$

- This clearly violates independence (endogenous outcome y on RHS)
- Simultaneity invalidates inferences based on direct estimation

A parsimonious model of interactions

- Generalizing the panel model to N units gives:

$$y_{it} = \sum_{j \neq i} \rho_{ij} y_{jt} + X_{it} \beta + \epsilon_{it}$$

- Considering all interactions ($\approx N^2$) is impractical
- Ord (1975) proposed the parsimonious parameterization:

$$y_{it} = \rho \sum_{j \neq i} w_{ij,t} y_{jt} + X_{it} \beta + \epsilon_{it}$$

- $w_{ij,t}$ represents a priori link between i and j

We must invert the model to solve it

- It is more convenient to use matrix notation
- If we stack all elements in conforming vectors/matrices:

$$y = \rho W y + X\beta + \epsilon$$

- This is known as the Spatial Autoregressive (SAR) model
- Estimating the model “as is” poses various challenges (Manski, 1993; Angrist, 2014)
- Solving for a reduced-form data generating process is more useful:

$$y = (I - \rho W)^{-1}(X\beta + \epsilon)$$

- Note y s only appear on LHS, but model is nonlinear in parameters

The Model implies geometrically-decaying propagation

- $(I - \rho W)^{-1}$ can be “difficult” to calculate.
- Given mathematical restrictions on ρ and W :

$$(I - \rho W)^{-1} = I + \rho W + \rho^2 W^2 + \dots \quad (1)$$

$$\Rightarrow y = (I + \rho W + \rho^2 W^2 + \dots) (X\beta + \epsilon) \quad (2)$$

- Interpret outcome as geometric sum of:
 - ▶ Own effect (I term)
 - ▶ Immediate peers' effect (W term)
 - ▶ Peers of peers effect (W^2 term)
 - ▶ etc

Partial derivatives are no longer β s

- In traditional model:

$$\frac{\partial y_i}{\partial x_i} = \beta, \text{ and } \frac{\partial y_i}{\partial x_j} = 0, i \neq j$$

- In the model with interactions:

$$\frac{\partial y_i}{\partial x_j} = (I - \rho W)_{ij}^{-1} \beta, \forall i, j$$

- Listing all partial derivatives is impractical.
- LeSage and Pace (2009) propose summarizing partial derivative estimates into direct and indirect effect averages:
 - ▶ Direct: $\frac{1}{N} \sum_i \frac{\partial y_i}{\partial x_i}$
 - ▶ Indirect: $\frac{1}{N} \sum_i \sum_{j \neq i} \frac{\partial y_i}{\partial x_j}$

The SDM adds contextual effects to SAR

- The Spatial Durbin Model (SDM) is given by:

$$y = \rho Wy + X\beta + WX\theta + \epsilon$$

- The values in WX represent the covariates of peers
- The effect of these covariates is often referred to a contextual effect
- These values are assumed exogenous and do not materially change the estimation

A short primer on estimation

- Focusing on one cross-section (for notational convenience), the likelihood function of the model is:

$$f(Y, X; \rho, \beta, \sigma^2) = |I_N - \rho W| (2\pi\sigma^2)^{-N/2} \exp\left(-\frac{e'e}{2\sigma^2}\right)$$
$$e = (I - \rho W)Y - X\beta$$

- If ρ is known (say ρ_0), then β (and σ^2) can be integrated out in a maximum likelihood estimation (MLE).
- The problem becomes an optimization w.r.t. ρ only.
- The estimation proceeds with an MCMC sampler using the above likelihood over a grid of different values for ρ .
- $|I - \rho W|$ is the determinant, usually calculate via LU decomposition and challenging to calculate for large matrices.

How to estimate the model then?

nwxtregress

- estimates SAR and SDM models with a mix of a MLE and MCMC sampling (LeSage and Pace, 2009)
- allows the estimation of spatial/network models with
 - ▶ unbalanced datasets
 - ▶ time varying spatial weights/network dependencies
 - ▶ several formats to define the spatial weights/network dependencies
- calculates direct, indirect and total effects.

nwxtregress¹

Syntax

Spatial Autocorrelation Model (SAR)

```
nwxtregress depvar indepvars [ if ] , dvarlag(W1[,options1] )  
[ mcmc_options options2 ]
```

Spatial Durbin Model (SDM)

```
nwxtregress depvar indepvars [ if ] , dvarlag(W1[,options1] )  
ivarlag(W2[,options1] [ mcmc_options options2 ]
```

- W1 and W2 define spatial weight matrices, default is Sp object.

¹This command is work in progress. Options, functions and results might change.

nwxtregress

Spatial Weight Options

```
nwxtregress depvar indepvars [if] , dvarlag(W1[,options1])  
[ ivarlag(W2[,options1]) mcmc_options options2 ]
```

- options1 controls the spatial weight matrices:
 - ▶ mata declares weight matrix is mata matrix. [Details](#)
 - ▶ sparse if weight matrix is sparse. [Details](#)
 - ▶ timesparse weight matrix is sparse and varying over time. [Details](#)
 - ▶ frame(name) use spatial weight from frame.
 - ▶ id(string) vector of IDs if W is a non sparse mata matrix.
 - ▶ zero(real) how to treat zeros in spatial weight matrix when using (time) sparse matrices.

nwxtregress

Further Options

```
nwxtregress depvar indepvars [if] , dvarlag(W1[options1]) ]  
[ ivarlag(W2[options1]) mcmc_options options2 ]
```

- *options2* are:

- ▶ *nospars* do not convert weight matrix internally to a sparse matrix.
- ▶ *noconstant* suppress constant.
- ▶ *fe* add fixed effects.

- *mcmc_options* control the Markov Chain Monte Carlo [Details](#)

- ▶ *python* use Python to calculate $|I - \rho W|$
- ▶ *usebp* use BarryPace trick instead of LUD for $|I - \rho W|$.

Example: BEA I/O Tabela I

Data

- We collect USE/MAKE table data from the BEA's website
- These data represent the goods that were used (USE) and made (MAKE) by each industry in the US
- To construct links between industries, we convert into flows between industries
- Loaded data as S_p matrix using `spmatrix fromdata W = sam*` , `replace`, but only for year 1998.
- We also collect key variables about each industry: capital consumption, compensation, and net surplus.

Example: BEA I/O Tabels II

Data

- We are estimating:
 - ▶ SAR:

$$\begin{aligned} cap_cons &= \beta_0 + \rho W_1 cap_cons \\ &\quad + \beta_1 compensation + \beta_3 net_surplus + \epsilon \end{aligned}$$

- ▶ SDM:

$$\begin{aligned} cap_cons &= \beta_0 + \rho W_1 cap_cons + \gamma_1 W_2 compensation \\ &\quad + \beta_1 compensation + \beta_3 net_surplus + \epsilon \end{aligned}$$

SAR

Time constant spatial weights

```
. nwxregress cap_cons compensation net_surplus , ///
> dvarlag(W) seed(1234)
```

```
Order Spatial Weights (22)
```

```
—|— 10 —|— 20 —|— 30 —|— 40 —|— 50 %
..... 50
..... 100
```

```
Initialise Grid (100)
```

```
—|— 10 —|— 20 —|— 30 —|— 40 —|— 50 %
..... 50
..... 100
```

```
Griddy Gibbs (2000)
```

```
—|— 10 —|— 20 —|— 30 —|— 40 —|— 50 %
..... 50
..... 100
```

```
Panel Variable (i): ID
Time Variable (t): Year
```

```
Number of obs   =   1358
Number of groups =    62
Obs. of group:
    min =    19
    avg =    22
    max =    22
R-squared       =    0.77
Adj. R-squared  =    0.77
```

	cap_cons	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
	compensation	-.7195389	.0123628	-58.20	0.000	-.767359 - .6791014
	net_surplus	-.7595604	.014257	-53.28	0.000	-.8046761 - .708425
	_cons	.7214415	.0117927	61.18	0.000	.6828125 .7663866
W	cap_cons	.120138	.0247156	4.86	0.000	.038 .2
	/sigma_u	.0029666	.0001126			.0026178 .0033713

SAR

Direct Indirect Effects

```
. estat impact
```

Average Impacts Number of obs = 1358

cap_cons	dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]	
direct						
compensation	-1.89144	.5196662	-3.64	0.000	-4.290424	-.7961761
net_surplus	-1.994237	.5481843	-3.64	0.000	-4.478126	-.8297429
indirect						
compensation	1.080422	.4975491	2.17	0.030	.0388582	3.365773
net_surplus	1.139151	.5246808	2.17	0.030	.0404964	3.513023
total						
compensation	-.8110181	.0253925	-31.94	0.000	-.9246504	-.7340081
net_surplus	-.8550866	.0276329	-30.94	0.000	-.9651031	-.7663825

Example

Time varying spatial weight

- Network data in *timesparse* format as mata matrix W .²
- Especially for large datasets gains in speed and memory are considerable. Example
- The first column identifies the year, second and third the IDs and the last one the value of the weight.
- Non standardized timesparse W :

```
. mata W[1..4,.]
```

	1	2	3	4
1	1997	1	1	120.445105
2	1997	1	2	2646.806067
3	1997	1	3	0
4	1997	1	4	1594.653373

²Often called Coordinate (COO) list format.

SAR

```
. nwxtregress cap_cons compensation net_surplus , ///
> dvarlag(W,mata timesparse) seed(1234)
```

Order Spatial Weights (22)

```
—|— 10 —|— 20 —|— 30 —|— 40 —|— 50 %
..... 50
..... 100
```

Initialise Grid (100)

```
—|— 10 —|— 20 —|— 30 —|— 40 —|— 50 %
..... 50
..... 100
```

Griddy Gibbs (2000)

```
—|— 10 —|— 20 —|— 30 —|— 40 —|— 50 %
..... 50
..... 100
```

```

Panel Variable (i): ID          Number of obs   =      1358
Time Variable (t): Year        Number of groups =       62
                                Obs. of group:   =       22
                                min =                  19
                                avg =                   22
                                max =                   22
                                R-squared                =       0.77
                                Adj. R-squared          =       0.77
```

	cap_cons	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	compensation	-.7271403	.011947	-60.86	0.000	-.7724468	-.688848
	net_surplus	-.7666499	.0139404	-54.99	0.000	-.809483	-.7163076
	_cons	.7299168	.0111695	65.35	0.000	.6930421	.7725614
W	cap_cons	.1027875	.0239002	4.30	0.000	.023	.18
	/sigma_u	.002978	.000113			.0026278	.0033843

SDM

```
. nwxregress cap_cons compensation net_surplus , python ///
> dvarlag(W,mata timesparse) ///
> ivarlag(W: compensation,mata timesparse ) seed(1234)
```

Order Spatial Weights (22)

```
—|— 10 —|— 20 —|— 30 —|— 40 —|— 50 %
..... 50
..... 100
```

Initialise Grid using Python (22)

```
—|— 10 —|— 20 —|— 30 —|— 40 —|— 50 %
..... 50
..... 100
```

Griddy Gibbs (2000)

```
—|— 10 —|— 20 —|— 30 —|— 40 —|— 50 %
..... 50
..... 100
```

```

Number of obs      =      1358
Panel Variable (i): ID      Number of groups =      62
Time Variable (t): Year    Obs. of group:
                               min =      19
                               avg  =      22
                               max  =      22
R-squared          =      0.77
Adj. R-squared     =      0.77
```

	cap_cons	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	compensation	-.7015208	.0123799	-56.67	0.000	-.7482582	-.6556848
	net_surplus	-.7472091	.0141113	-52.95	0.000	-.7920688	-.6965197
	_cons	.7730435	.0143202	53.98	0.000	.7288873	.8187064
W	cap_cons	.031389	.0283132	1.11	0.268	-.054	.119
	compensation	-.0895239	.0181129	-4.94	0.000	-.1565037	-.0227881
	/sigma_u	.0029277	.0001141			.0025583	.0033638

SDM

Direct Indirect Effects

```
. estat impact
```

Average Impacts Number of obs = 1358

cap_cons	dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]	
direct						
compensation	-.7002594	.0124331	-56.32	0.000	-.7490398	-.6526457
net_surplus	-.7472824	.0141128	-52.95	0.000	-.7922827	-.6965584
indirect						
compensation	.7741688	.014668	52.78	0.000	.7288423	.8191533
net_surplus	-.0248066	.0225161	-1.10	0.271	-.1018671	.0382383
total						
compensation	.0739094	.0151583	4.88	0.000	.0152978	.1373628
net_surplus	-.772089	.027054	-28.54	0.000	-.8611195	-.6896129

Conclusion

- `nwxtregress` extends `spxtregress`:
 - ▶ Allows for unbalanced datasets and time varying spatial weight matrices
 - ▶ Spatial weights can be directly loaded from datasets, frames, mata matrices or `spmatrix` objects.
- Available on GitHub (<https://janditzen.github.io/nwxtregress/>) or directly in Stata:

```
net install nwxtregress ,  
from(https://janditzen.github.io/nwxtregress/)
```

- Please, help us by providing feedback

References I

- Kelejian, H., and G. Piras. 2017. Spatial Econometrics. Academic Press.
- LeSage, J. P., and R. K. Pace. 2009. Introduction to Spatial Econometrics. Florida CRC Press.

Weight Matrices back

Square

Square matrix format

- The spatial weights are a matrix with dimension $N_g \times N_g$. It is time constant. An Example for a 5×5 matrix is:

	1	2	3	4		
	+-----+					
1		0	.1	.2	0	
2		0	0	.1	.2	
3		.3	.1	0	0	
4		.2	0	.2	0	
	+-----+					

Weight Matrices [back](#)

Sparse format

- The sparse matrix format is a $v \times 3$ matrix, where v is the number of non-zero elements in the spatial weight matrix.
- The weight matrix is time constant. The first column indicates the destination, the second the origin of the flow. A sparse matrix of the matrix from above is:

Destination	Origin	Flow
1	2	0.1
1	3	0.2
2	3	0.1
2	4	0.2
3	1	0.3
3	2	0.1
4	1	0.2
4	3	0.2

Weight Matrices back

Time-Sparse format

- The time sparse format can handle time varying spatial weights.
- The first column indicates the time period, the remaining are the same as for the sparse matrix. For example, if there are two time periods and we have the matrix from above for the first and the square for the second period:

Time	Destination	Origin	Flow
1	1	2	0.1
1	1	3	0.2
1	2	3	0.1
1	2	4	0.2
1	3	1	0.3
1	3	2	0.1
1	4	1	0.2
1	4	3	0.2
	<i>(next time period)</i>		
2	1	2	0.1
2	1	3	0.4
2	2	3	0.1
2	2	4	0.4
2	3	1	0.9
2	3	2	0.1
2	4	1	0.4
2	4	3	0.4

Example Sparse Matrix [back](#)

- Contiguity matrix for 100 units.
- Total number of elements in W : 10,000 (100×100).
- Non-zero elements: 200, 9800 elements are zero.
- Square matrix will consume 80,000bytes (in mata).
- COO matrix with 3×200 elements will consume 2,400bytes.
- Calculation ρW on square matrix implies 10,000 mathematical operations, but only 200 lead to a non-zero result!
- Using COO matrix will limit the mathematical operations to 200!

mcmc_options

[back](#)

Control the Markov Chain Monte Carlo:

- `python` use Python to calculate $|I - \rho W|$
- `draws(integer 2000)` number of griddy gibbs draws.
- `gridlength(integer 1000)` grid length
- `nomit(integer 500)` number of omitted draws
- `barrypace(numlist)` settings for BarryPace Trick, iterations, maxorder default: 50 100
- `usebp` use BarryPace trick instead of LUD for $|I - \rho W|$.
- `seed(#)` sets the seed.