

Double-Hurdle Models with Dependent Errors and Heteroscedasticity

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Essen, April 2nd 2007

Tobit model

Model

Censoring of the dependent variable is traditionally dealt with using the Tobit model

$$y_i = y_i^* = \mathbf{x}_i\beta + \varepsilon_1 \quad \text{if} \quad y_i^* > 0 \quad \text{otherwise} \quad y_i = 0$$

Likelihood function

$$L = \prod_0 \left(1 - \Phi \left(\frac{\mathbf{x}_i\beta}{\sigma_{\varepsilon_1}} \right) \right) + \prod_+ \left(\frac{1}{\sigma_{\varepsilon_1}} \phi \left(\frac{y_i - \mathbf{x}_i\beta}{\sigma_{\varepsilon_1}} \right) \right)$$

Cragg's formulation

Model

Cragg (1971) proposed the extension that the probability of a zero realisation, $1 - \Phi(\cdot)$, is not directly to the density for a continuous realisation, $\phi(\cdot)$, but instead governed by some other process

$$\begin{aligned} y_i &= y_i^* = x_i\beta + \varepsilon_{1i} & \text{if } & x_i\beta + \varepsilon_{1i} > 0 \text{ and } z_i\alpha + \varepsilon_{2i} > 0 \\ &= 0 & \text{if } & x_i\beta + \varepsilon_{1i} \leq 0 \text{ and } z_i\alpha + \varepsilon_{2i} > 0 \\ & & \text{or } & x_i\beta + \varepsilon_{1i} > 0 \text{ and } z_i\alpha + \varepsilon_{2i} \leq 0 \\ & & \text{or } & x_i\beta + \varepsilon_{1i} \leq 0 \text{ and } z_i\alpha + \varepsilon_{2i} \leq 0 \end{aligned}$$

Cragg's formulation - errors

Independent

The original model made the assumption that the two error terms were jointly normal,

$$\begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \end{pmatrix} \sim N(0, \Sigma), \text{ and uncorrelated, } \Sigma = \begin{pmatrix} \sigma_{\varepsilon_1}^2 & 0 \\ 0 & 1 \end{pmatrix}$$

Likelihood function

$$L = \prod_0 \left(1 - \Phi \left(\frac{x_j \beta}{\sigma_{\varepsilon_1}} \right) \cdot 1 - \Phi(z_j \alpha) \right) \prod_+ \left(\Phi(z_j \alpha) \frac{1}{\sigma_{\varepsilon_1}} \phi \left(\frac{y_j - x_j \beta}{\sigma_{\varepsilon_1}} \right) \right)$$

Estimation

Separability

Similar to that demonstrated by McDowell (2003) for count models in Stata, the separability of the likelihood function permits the use of a combination of Stata command to estimate this model: `truncreg` and `probit`

Jones' extension

Correlated errors

This assumption has been relaxed in later work, e.g. Jones (1992),

$$\text{where } \Sigma = \begin{pmatrix} \sigma_{\varepsilon_1}^2 & \sigma_{\varepsilon_1} \rho \\ \sigma_{\varepsilon_1} \rho & 1 \end{pmatrix}$$

Likelihood function

$$L = \prod_0 [1 - F_2(\mathbf{z}_i \alpha, \mathbf{x}_i \beta / \sigma, \rho)] \prod_+ \Phi \left(\frac{\mathbf{z}_i \alpha + \frac{\rho}{\sigma} (\mathbf{y} - \mathbf{x}_i \beta)}{\sqrt{1 - \rho^2}} \right) \frac{1}{\sigma} \phi \left(\frac{(\mathbf{y} - \mathbf{x}_i \beta)}{\sigma} \right)$$

Estimation

Non-separable

Both parts of this likelihood function must, however, be maximised simultaneously; there is no two-step equivalent. This has been available in Stata, on an *ad hoc* basis since 2004 using the `dhurdle` command written for Stata 7.

Syntax

```
dhurdle y x1 x2, sel(d x1 t1)
```

Comparison to Flood and Gråsjö

| | True value | Stata | | Gauss | |
|------------|------------|---------|------|---------|-------|
| | | Bias(%) | RMSE | Bias(%) | RMSE |
| β_0 | -0.2 | -23.3 | .665 | -18.0 | .836 |
| β_1 | 0.2 | -17.6 | .152 | 3.7 | .085 |
| β_2 | 1 | .34 | .081 | 2.7 | .182 |
| α_0 | 0.7 | 49.0 | 1.25 | 175 | 8.065 |
| α_1 | 0.2 | 24.6 | .881 | -44.7 | 1.325 |
| α_2 | -0.2 | 27.3 | .392 | -80.7 | .570 |
| σ | 2 | 1.38 | .256 | -.2 | .227 |
| ρ | -0.5 | -26.3 | .261 | 31.2 | .447 |

Assumptions on the error terms

There is, by now, a wide variety of literature demonstrating that if the assumption of homoscedastic, normally-distributed, errors is violated then ML parameter estimates are inconsistent.

Solutions

Two extensions

1. Heteroscedastic errors
2. Non-normal errors

`dhurdle` now can incorporate variance dependent on a set of independent variables.

Syntax

```
dhurdle y x1 x2, sel(d x1 t1) het(.)
```

Non-normality

Non-normality

Robinson (1982) showed that ML estimation of LDV models leads to inconsistent parameter estimates if the assumption of normally distributed errors does not hold.

Syntax

```
dhurdle y x1 x2, sel(d x1 t1) ihs
```

Pipeline

Work in progress

The final steps to complete the estimation package that are currently underway are

1. Finalise `predict` options for the double hurdle.
2. Code a series of LR tests to test model specification post-estimation.

Testing 1, 2, 3

Using the IHS as the general form, the imposition of the following restrictions is feasible

1. If $\gamma = 0$ then conventional formulation without transformation.
2. If σ is constant then homoscedastic errors.
3. If $\rho = 0$ then independent double hurdle.
4. If $\prod_{+} \Phi(z_i \alpha)$ then no censoring present and model simplifies to a Heckman.
5. If $\Phi(z_i \alpha)$ and $\rho = 0$ then no censoring or selection present and model simplifies to a Tobit.