

# **Handling interactions in Stata, especially with continuous predictors**

**Patrick Royston & Willi Sauerbrei**

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# Interactions – general concepts

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- General idea of a (two-way) interaction in multiple regression is **effect modification**:
  - $\eta(x_1, x_2) = f_1(x_1) + f_2(x_2) + f_3(x_1, x_2)$
- Often,  $\eta(x_1, x_2) = E(Y | x_1, x_2)$ , with obvious extension to GLM, Cox regression, etc.
- Simplest case:  $\eta(x_1, x_2)$  is **linear** in the  $x$ 's and  $f_3(x_1, x_2)$  is the **product** of the  $x$ 's:
  - $\eta(x_1, x_2) = \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2$
- Can extend to more general, **non-linear** functions

# The simplest type of interaction

- Binary x binary
- E.g. in the MRC RE01 trial in kidney cancer
- 12 month % survival since randomisation
- Substantial treatment effect in patients with low WCC
- Little or no treatment effect in those with high WCC
- But really, WCC is a continuous variable ...

Treatment group	White cell count low (<=10)	White cell count high (>10)
MPA	34% (se 4)	24% (se 4)
Interferon	49% (se 4)	21% (se 7)

# Overview

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- Interactions and factor variables (Stata 11/12)
  - Note: I am not an expert on factor variables! I sometimes use them.
- General interactions between continuous covariates in observational studies
  - Focus on **continuous** covariates ...
  - ... because people don't appear to know how to handle them!
- Special case: interactions between treatment and continuous covariates in randomized controlled trials

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# Interactions and factor variables

# Scope

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- We introduce the topic with a brief introduction to factor variables
- In this part, we consider only **linear** interactions:
  - Binary x binary (2 x 2 table)
  - Binary x continuous
  - Continuous x continuous

# Factor variables: brief notes

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- Implemented via prefixes (unary operators) and binary interaction operators
  - see `help fvvarlist`
- There are four factor-variable operators:

Operator	Description
i.	unary operator to specify indicators (dummies)
c.	unary operator to treat as continuous
#	binary operator to specify interactions
##	binary operator to specify factorial interactions
- Dummy variables are ‘virtual’ – not created *per se*
- Names of regression parameters easily found by inspecting the post-estimation result matrix `e(b)`

## Factor variables: i. prefix

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- Example from Stata manual [U]11.4.3:

```
. list group i.group in 1/5  
group 1b.group 2.group 3.group  
1.    1      0      0      0  
2.    1      0      0      0  
3.    2      0      1      0  
4.    2      0      1      0  
5.    3      0      0      1
```

## Example dataset

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- MRC RE01 trial in advanced kidney cancer
- Of 347 patients, only 7 censored, the rest died
- For simplicity, as a continuous response variable,  $Y$ , we use months to death,  $_t$ 
  - Ignore the small amount of censoring
- There are several **prognostic factors** that may influence time to death
- Some are binary, some categorical, some continuous

# Example: factor variable parameters

---

```
. regress _t i.who
```

Source	SS	df	MS	Number of obs	=	347
Model	7780.81413	2	3890.40707	F( 2, 344)	=	15.72
Residual	85126.5686	344	247.460955	Prob > F	=	0.0000
Total	92907.3828	346	268.518447	R-squared	=	0.0837
				Adj R-squared	=	0.0784
				Root MSE	=	15.731

_t	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
who					
1	-4.2783	2.026215	-2.11	0.035	-8.263629 -.2929707
2	-12.94365	2.354542	-5.50	0.000	-17.57476 -8.312534
_cons	19.17358	1.622518	11.82	0.000	15.98227 22.36488

```
. matrix list e(b)
```

```
e(b) [1,4]
```

	0b.	1.	2.	
	who	who	who	_cons
y1	0	-4.2782996	-12.943646	19.173577

## Basic analysis to understand binary x binary: the 2 x 2 table of means

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- Example:

```
. table rem sex, contents(mean _t) format(%6.2f)
```

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		x2: sex	
		male	female
x6: kidney removed (Y/N)	0	13.50	9.34
	1	13.77	18.19

---

# Fitting an interaction model

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- Consider 3 methods:
  - Method 1: binary operator to specify interactions
    - `regress _t rem#sex`
  - Method 2: binary operator to specify factorial interactions
    - `regress _t rem##sex`
  - Method 3: create multiplicative term(s) yourself
    - `gen byte remsex = rem * sex`
    - `regress _t rem sex remsex`
- Models are identical - all give the same fitted values
- Parameterisation of method 1 is different

# Method 1: Binary operator #

---

```
. regress _t rem#sex
```

Source	SS	df	MS	Number of obs	=	347
Model	2168.10384	3	722.701279	F( 3, 343)	=	2.73
Residual	90724.0188	343	264.501513	Prob > F	=	0.0437
Total	92892.1227	346	268.474343	R-squared	=	0.0233
				Adj R-squared	=	0.0148
				Root MSE	=	16.264

_t	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
rem#sex					
0 1	-4.160995	2.888457	-1.44	0.151	-9.842314 1.520323
1 0	.2724983	2.137064	0.13	0.899	-3.930903 4.475899
1 1	4.68417	2.581164	1.81	0.070	-.3927323 9.761072
_cons	13.49939	1.610327	8.38	0.000	10.33204 16.66675

## Notes on parameters:

rem#sex 0 1 (-4.16) is (sex=1) - (sex=0) at rem=0

rem#sex 1 0 (+0.27) is (rem=1) - (rem=0) at sex=0

rem#sex 1 1 (+4.68) is [(rem=1) | sex=1] - [(rem=0) | sex=0]

\_cons (13.50) is intercept (mean of Y at rem=0 & sex=0)

I don't recommend this parameterisation!

## Method 2: Binary operator ##

---

```
. regress _t rem##sex
```

Source	SS	df	MS	Number of obs = 347			
Model	2168.10384	3	722.701279	F( 3, 343) = 2.73			
Residual	90724.0188	343	264.501513	Prob > F = 0.0437			
Total	92892.1227	346	268.474343	R-squared = 0.0233			
				Adj R-squared = 0.0148			
				Root MSE = 16.264			
_t	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]		
1.rem	.2724983	2.137064	0.13	0.899	-3.930903	4.475899	
1.sex	-4.160995	2.888457	-1.44	0.151	-9.842314	1.520323	
rem#sex							
1 1	8.572667	3.792932	2.26	0.024	1.112333	16.033	
_cons	13.49939	1.610327	8.38	0.000	10.33204	16.66675	

### Notes on parameters:

1.rem (0.27) is (rem=1) - (rem=0) at sex=0

1.sex (-4.16) is (sex=1) - (sex=0) at rem=0

rem#sex (+8.57) is [(rem=1) - (rem=0) at sex=1] - [(rem=1) - (rem=0) at sex=0]

\_cons (13.50) is intercept (mean of Y at rem=0 & sex=0)

This is a 'standard' parameterisation with P-value as given above

# Method 3: DIY multiplicative term

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- `generate byte remsex = rem * sex`
- `regress _t rem sex remsex`

Source	SS	df	MS	Number of obs	=	347
Model	2168.10384	3	722.701279	F( 3, 343)	=	2.73
Residual	90724.0188	343	264.501513	Prob > F	=	0.0437
Total	92892.1227	346	268.474343	R-squared	=	0.0233
				Adj R-squared	=	0.0148
				Root MSE	=	16.264

_t	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
rem	.2724983	2.137064	0.13	0.899	-3.930903 4.475899
sex	-4.160995	2.888457	-1.44	0.151	-9.842314 1.520323
remsex	8.572667	3.792932	2.26	0.024	1.112333 16.033
_cons	13.49939	1.610327	8.38	0.000	10.33204 16.66675

## Notes on parameters:

rem is 1.rem in Method 2

sex is 1.sex in Method 2

remsex is rem#sex in Method 2

This is the same parameterisation as Method 2

# Interactions in non-normal errors models

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Key points:

1. In a  $2 \times 2$  table, an interaction is a 'difference of differences'
2. Tabulate the  $2 \times 2$  table of mean values of the linear predictor
3. May back-transform values via the inverse link function
  - e.g. exponentiation in hazards models

# Binary x continuous interactions

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- Use c. prefix to indicate continuous variable
  - Use the ## operator
- . regress \_t trt##c.wcc

Source	SS	df	MS	Number of obs	=	347
Model	5678.62935	3	1892.87645	F( 3, 343)	=	7.44
Residual	87228.7534	343	254.311234	Prob > F	=	0.0001
Total	92907.3828	346	268.518447	R-squared	=	0.0611
				Adj R-squared	=	0.0529
				Root MSE	=	15.947
_t	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
1.trt	12.81405	4.124167	3.11	0.002	4.702208	20.92589
wcc	-.2867831	.2741174	-1.05	0.296	-.8259457	.2523796
trt#c.wcc						
1	-1.034239	.4327233	-2.39	0.017	-1.885365	-.1831142
_cons	14.45292	2.712383	5.33	0.000	9.117919	19.787 <sup>16</sup> <sub>91</sub>

## Binary x continuous interactions (cont.)

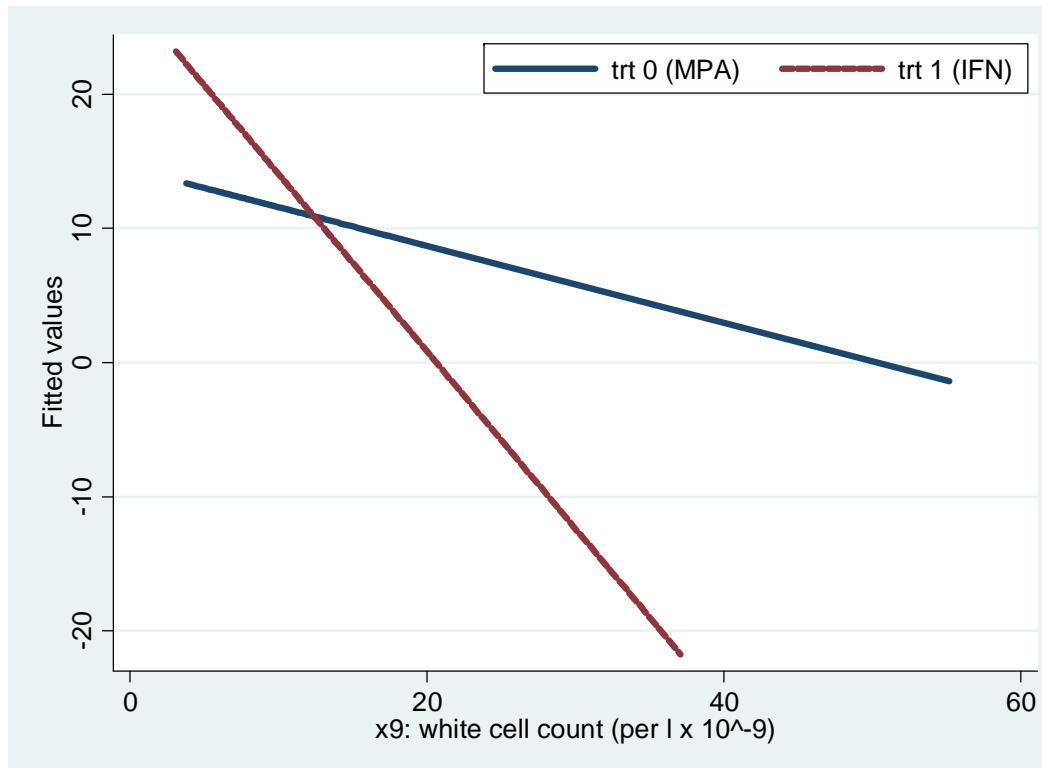
---

- The main effect of `wcc` is the slope in group 0
- The interaction parameter is the difference between the slopes in groups 1 & 0
- Test of `trt#c.wcc` provides the interaction parameter and test
- Results are nicely presented graphically
  - Predict linear predictor `xb`
  - Plot `xb` by levels of the factor variable
  - Also, ‘treatment effect plot’ (*coming later*)

# Plotting a binary x continuous interaction

---

```
. regress _t trt##c.wcc  
. predict fit  
. twoway (line fit wcc if trt==0, sort) (line fit wcc  
if trt==1, sort lp(-)), legend(lab(1 "trt 0 (MPA)")  
lab(2 "trt 1 (IFN)") ring(0) pos(1))
```



# Continuous x continuous interaction

---

- Just use c. prefix on each variable

```
. regress _t c.age##c.t_mt
```

Source	SS	df	MS	Number of obs	=	347
Model	7714.26052	3	2571.42017	F( 3, 343)	=	10.35
Residual	85193.1223	343	248.37645	Prob > F	=	0.0000
Total	92907.3828	346	268.518447	R-squared	=	0.0830
				Adj R-squared	=	0.0750
				Root MSE	=	15.76

_t	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0719063	.0876542	0.82	0.413	-.1005011 .2443137
t_mt	.0659781	.0128802	5.12	0.000	.040644 .0913122
c.age#c.t_mt	-.0008783	.0001861	-4.72	0.000	-.0012443 -.0005124
_cons	8.055213	5.256114	1.53	0.126	-2.28306 18.39349

## Continuous x continuous interaction

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- Results are best explored graphically
- Consider in more detail next

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# Continuous x continuous interactions

## Motivation: continuous x continuous intn.

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- Many people only consider linear by linear interactions
- Not sensible if main effect of either variable is **non-linear**
- Mismodelling the main effect may introduce spurious interactions
  - E.g. false assumption of linearity can create a spurious linear x linear interaction
- Or they categorise the continuous variables
  - Many problems, including loss of power

# The MFPIgen approach (1)

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- MFP = multivariable fractional polynomials
- I = interaction
- gen = general
- Fractional polynomials (FPs) can be used to model relationships that may be non-linear
- In Stata, FPs are implemented through the standard `fracpoly` and `mfp` commands
- MFPIgen is implemented through a user-written command, `mfpigen`

## The MFPIgen approach (2)

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- MFPIgen aims to identify non-linear main effects and their two-way interactions
- Assume  $x_1, x_2$  continuous and  $z$  confounders
- Apply MFP to  $x_1$  and  $x_2$  and  $z$ 
  - Force  $x_1$  and  $x_2$  into the model
  - FP functions  $FP_1(x_1)$  and  $FP_2(x_2)$  are selected for  $x_1$  and  $x_2$
  - Linear functions could be selected
- Add term  $FP_1(x_1) \times FP_2(x_2)$  to the model chosen
- Apply likelihood ratio test of interaction

# The MFPIgen approach in practice

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- Start with a list of covariates
- Check all pairs of variables for an interaction
- Simultaneously, apply MFP to adjust for confounders
- Use a low significance level to detect interactions, e.g. 1%
- Present interactions graphically
- Check interactions for artefacts graphically
- Use forward stepwise if more than one interaction remains

## Example: Whitehall 1

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- Prospective cohort study of 17,260 Civil Servants in London
- Studied various standard risk factors for common causes of death
- Also studied social factors, particularly job grade
- We consider 10-year all-cause mortality as the outcome
- Logistic regression analysis

## Example: Whitehall 1 (2)

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- Consider weight and age
- **mfpigen: logit all10 age wt**

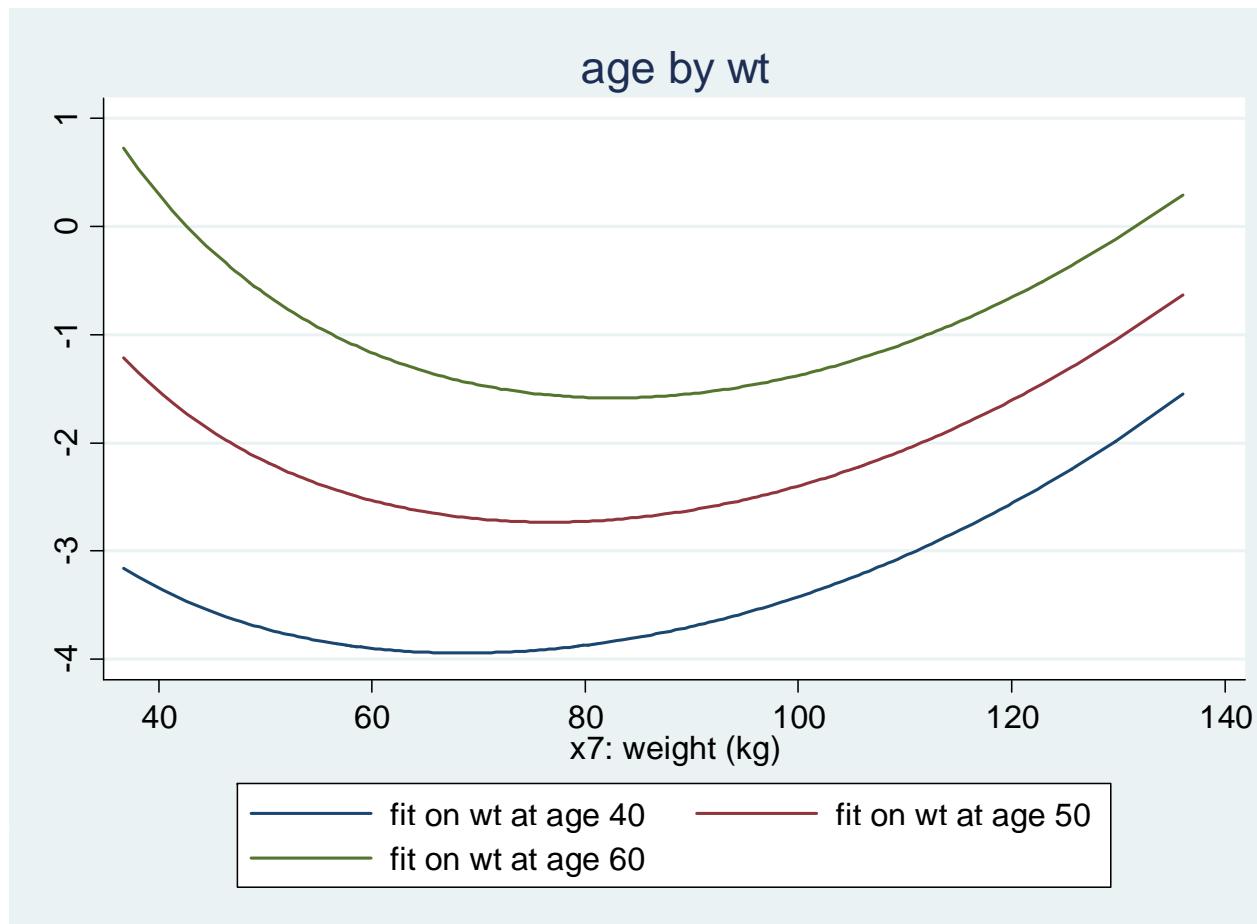
```
MFPGEN - interaction analysis for dependent variable all10
-----
variable 1    function 1    variable 2    function 2    dev. diff.   d.f.      P      Sel
-----
age          Linear        wt           FP2(-1 3)     5.2686       2    0.0718   0
-----
Sel = number of variables selected in MFP adjustment model
```

- Age function is linear, weight is FP2(-1, 3)
- No strong interaction ( $P = 0.07$ )

# Plotting the interaction model

---

```
. mfpigen, fplot(40 50 60): logit all10 age wt
```



# Mis-specifying the main effects function(s)

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- Assume age and weight are linear
  - The `dfdefault(1)` option imposes linearity
- . `mfpigen, dfdefault(1): logit all10 age wt`

```
MFPGEN - interaction analysis for dependent variable all10
-----
variable 1   function 1   variable 2   function 2   dev. diff.   d.f.   P   Sel
-----
age          Linear       wt           Linear        8.7375      1   0.0031   0
-----
Sel = number of variables selected in MFP adjustment model
```

- There appears to be a highly significant interaction ( $P = 0.003$ )

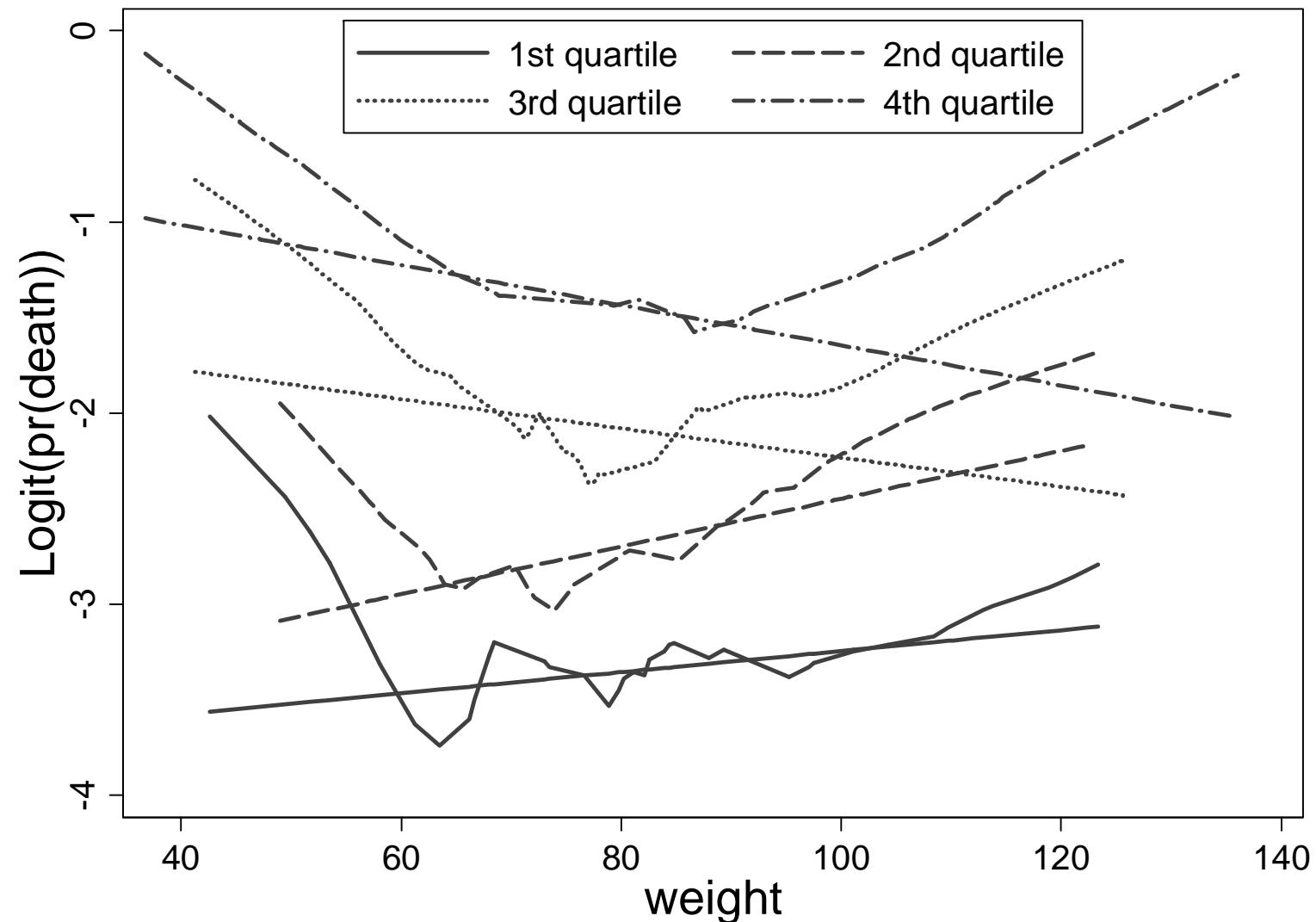
# Checking the interaction model

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- Linear age  $\times$  weight interaction seems important
- Check if it's real, or the result of mismodelling
- Categorize age into (equal sized) groups
  - for example, 4 groups
- Compute running line smooth of the binary outcome on weight in each age group, transform to logits
- Plot results for each group
- Compare with the functions predicted by the interaction model

# Whitehall 1: Check of age x weight linear interaction

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## Interpreting the plot

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- Running line smooths are roughly parallel across age groups  $\Rightarrow$  no (strong) interactions
- Erroneously assuming that the effect of weight is linear  $\Rightarrow$  estimated slopes of weight in age-groups indicate strong interaction between age and weight
- We should have been more careful when modelling the main effect of weight

# Whitehall 1: 7 variables, any interactions?

---

```
. mfpigen, select(0.05): logit all10 cigs  
sysbp age ht wt chol i.jobgrade
```

MFPIGEN - interaction analysis for dependent variable all10

variable 1	function 1	variable 2	function 2	dev. diff.	d.f.	P	Sel
cigs	FP1(.5)	sysbp	FP2(-2 -2)	0.7961	2	0.6716	5
	FP1(.5)	age	Linear	0.0028	1	0.9576	5
	FP1(.5)	ht	Linear	2.1029	1	0.1470	5
	FP1(.5)	wt	FP2(-2 3)	0.1560	2	0.9249	5
	FP1(.5)	chol	Linear	1.7712	1	0.1832	5
	FP1(.5)	i.jobgrade	Factor	4.3061	3	0.2303	5
sysbp	FP2(-2 -2)	age	Linear	3.1169	2	0.2105	5

*(remaining output omitted)*

# What `mfpigen` is doing

---

- FP functions for each pair of continuous variables are selected
  - Functions are simplified if possible
  - Closed test procedure in `mfp`
  - Controlled by the `alpha()` option
- The `select(0.05)` option tests confounders for inclusion in each interaction model at the 5% significance level
- The `Sel` column in the output shows how many variables are actually included in each confounder model

# Results: P-values for interactions

---

Variable	cigs*	sysbp*	age	height	weight*	chol
cigs*	—					
sysbp*	0.7	—				
age	0.9	0.2	—			
height	0.1	0.5	1.0	—		
weight*	0.9	0.5	0.1	0.4	—	
chol	0.2	0.07	0.001	0.8	0.2	—
grade	0.2	0.2	0.2	0.2	0.04	0.4

\*FP transformations were selected; otherwise, linear

## Graphical presentation of age x chol interaction

---

```
. fracgen cigs .5, center(mean)
. fracgen sysbp -2 -2, center(mean)
. fracgen wt -2 3, center(mean)

. mfpigen, linadj(cigs_1 sysbp_1 sysbp_2
> wt_1 wt_2 ht i.jobgrade) df(1)
> fplot(%10 35 65 90): logit all10 age chol
```

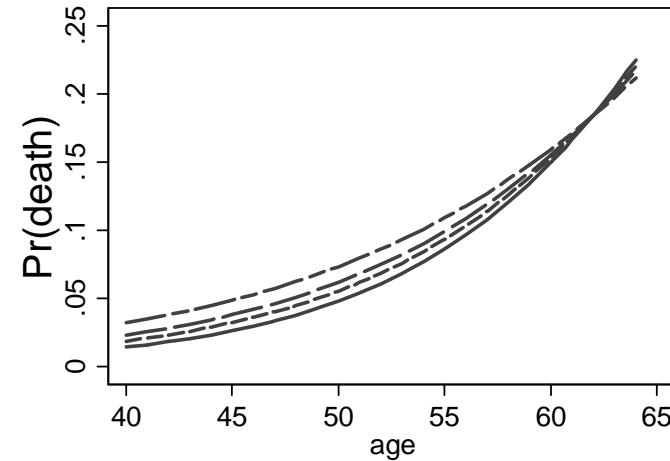
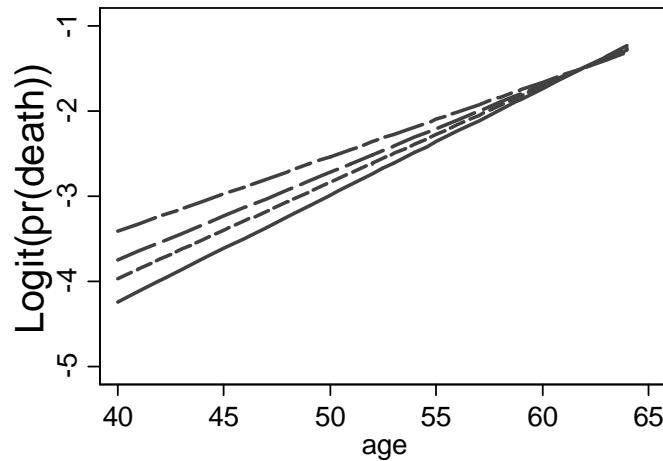
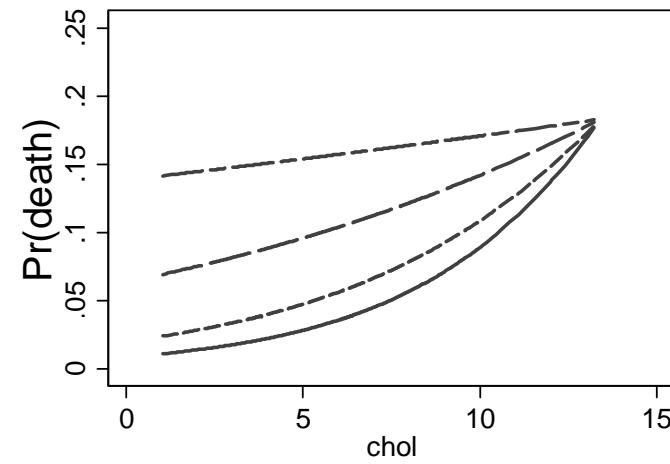
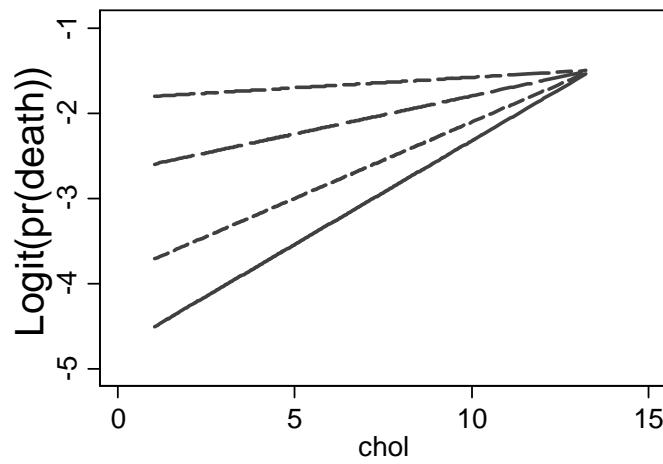
Alternatively:

```
. logit all10 c.age##c.chol cigs_1 sysbp_1 sysbp_2
> wt_1 wt_2 ht i.jobgrade
. sliceplot age chol, sliceat(10 35 65 90) percent
```

- **sliceplot** is a new user-written command

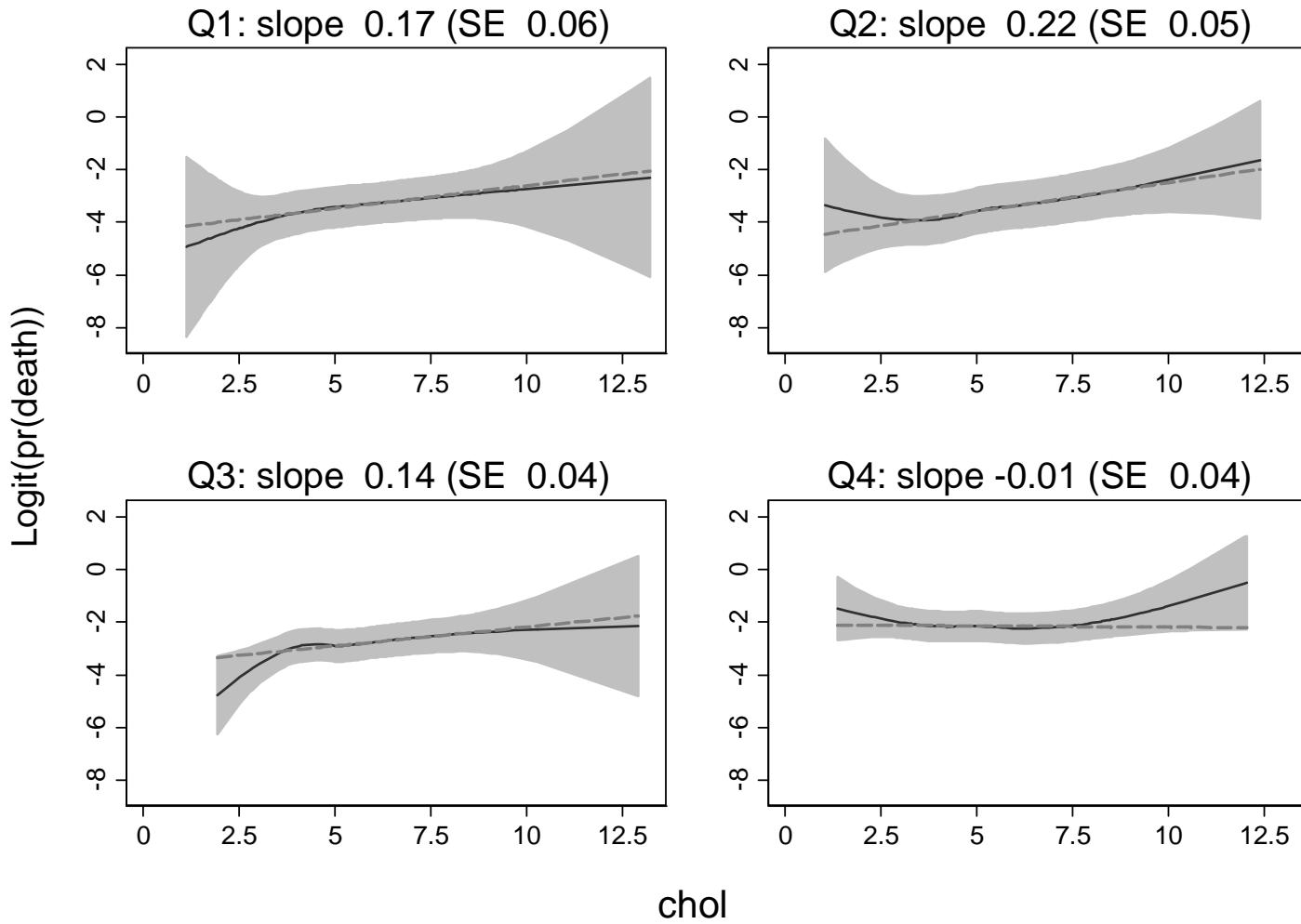
# Graphical presentation of age x chol intn.

---



# Check of chol x age interaction

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# Interactions with continuous covariates in randomized trials

## MFPI method (Royston & Sauerbrei 2004)

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- Continuous covariate  $x$  of interest, binary treatment variable  $t$  and other covariates  $z$
- Independent of  $x$  and  $t$ , use MFP to select an 'adjustment' (confounder) model  $z^*$  from  $z$
- Find best FP2 function of  $x$  (in all patients) adjusting for  $z^*$  and  $t$
- Test  $\text{FP2}(x) \times t$  interaction (2 d.f.)
  - Estimate  $\beta$ 's in each treatment group
  - Standard test for equality of  $\beta$ 's
- May also consider simpler FP1 or linear functions – choose e.g. by min AIC

# MFPI in Stata

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- MFPI is implemented as a user command,  
`mfpi`
- `mfpi` is available on SSC
- Details are given by Royston & Sauerbrei,  
*Stata Journal* 9(2): 230-251 (2009)
- Program was updated in 2012 to support  
factor variables

# Treatment effect function

---

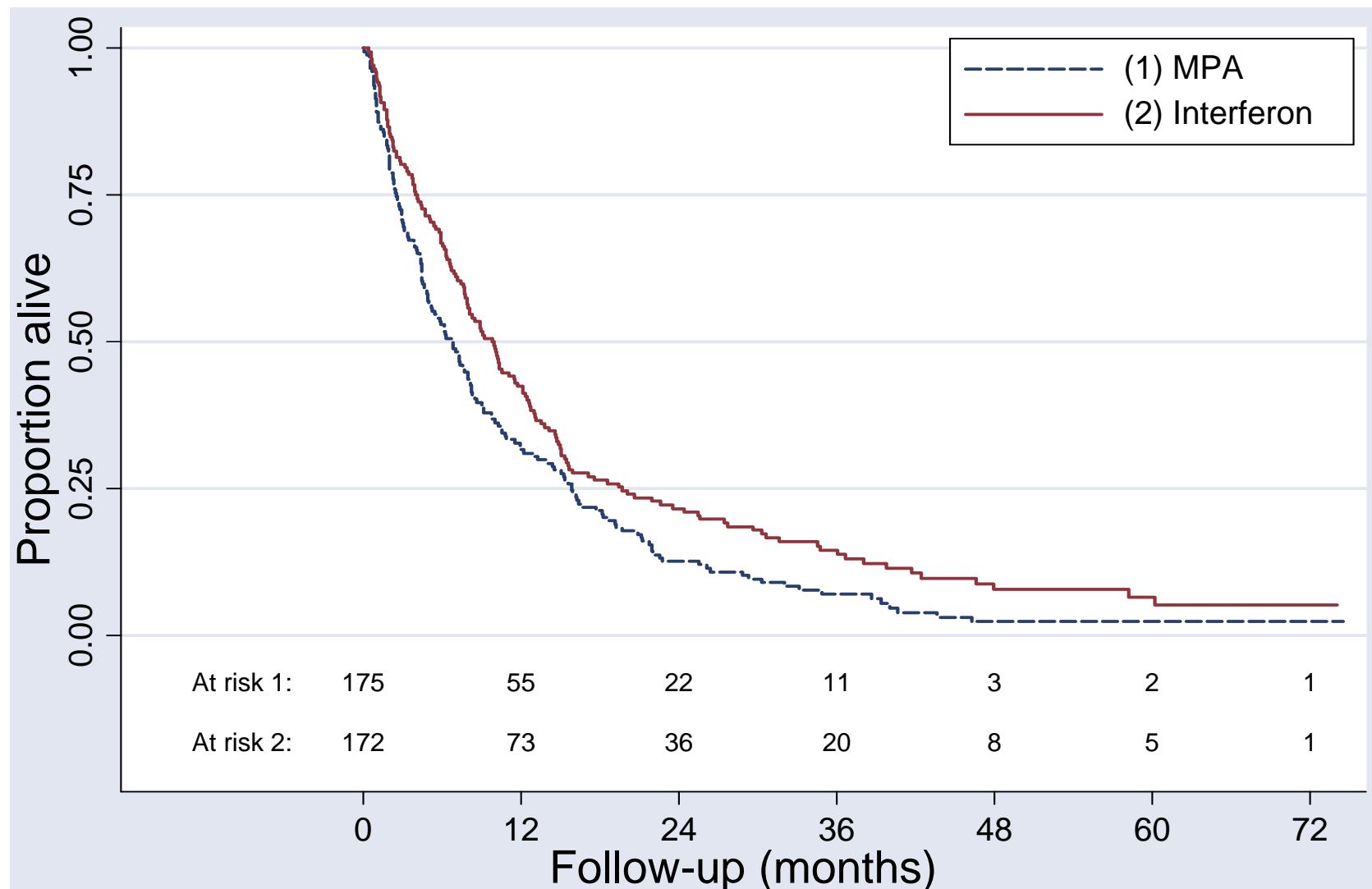
- Have estimated two FP2 functions – one per treatment group
- Plot the difference between functions against  $x$  to show the interaction
  - i.e. the treatment effect at different  $x$
- Pointwise 95% CI shows how strongly the interaction is supported at different values of  $x$ 
  - i.e. variation in the treatment effect with  $x$

## Example: MRC RE01 trial in kidney cancer

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- Main analysis: Interferon improves survival
- HR: 0.76 (0.62 - 0.95),  $P = 0.015$
- Is the treatment effect similar in all patients?
- Nine possible covariates available for the investigation of treatment-covariate interactions – only one is significant (WCC)

# Kaplan-Meier showing treatment effect



# The mfpi command

---

```
. mfpi, select(0.05) fp2(wcc) with(trt) gendiff(d): stcox  
(whod1 whod2) t_dt t_mt rem mets haem
```

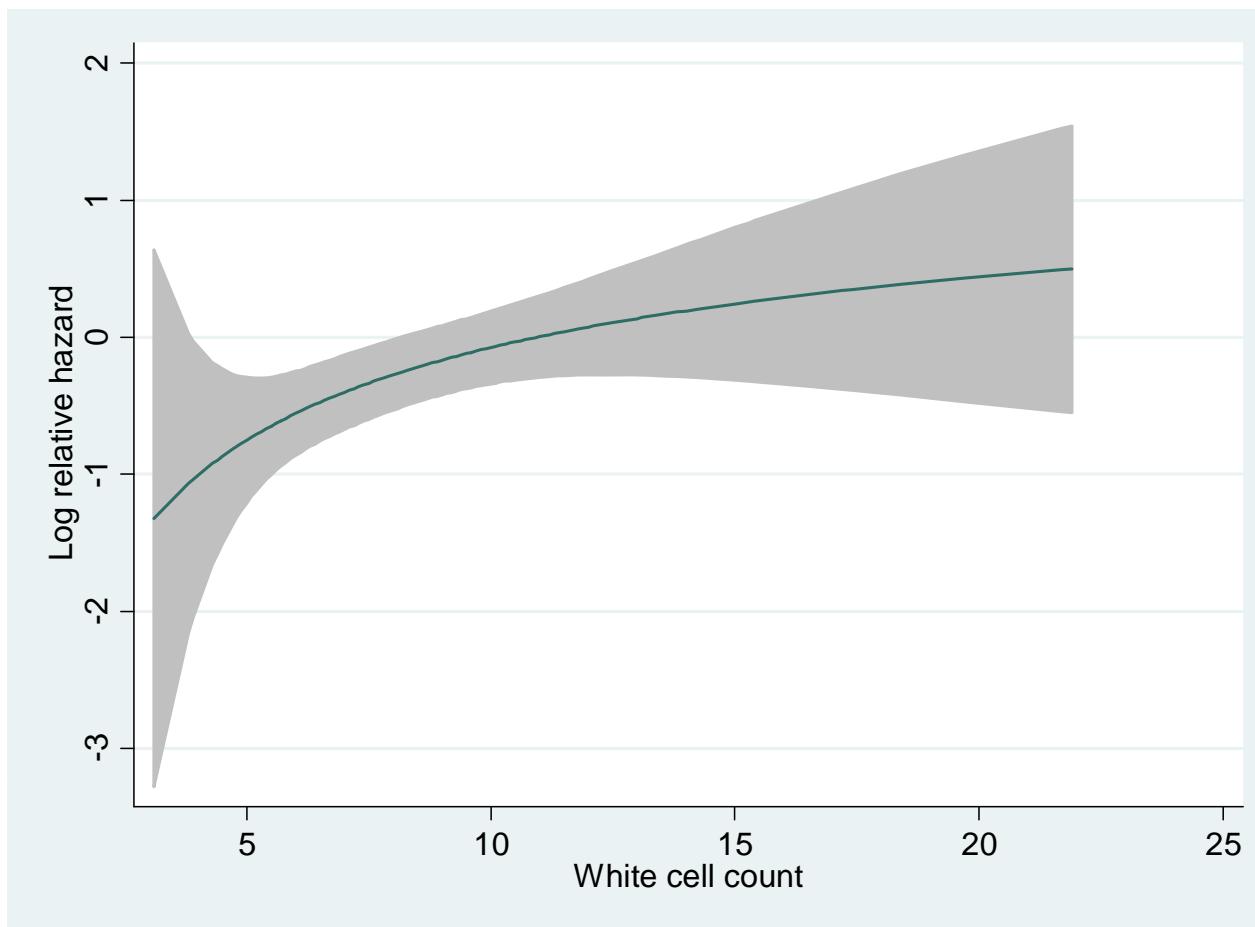
```
Interactions with trt (347 observations). Flex-1 model (least flexible)
```

Var	Main	Interact	idf	Chi2	P	Deviance	tdf	AIC
wcc	FP2(-1 -.5)	FP2(-1 -.5)	2	6.91	0.0316	3180.194	7	3194.194

```
-----  
idf = interaction degrees of freedom; tdf = total model degrees of freedom
```

```
. mfpi_plot wcc  
[using variables created by gendiff(d)]
```

# Treatment effect plot for wcc



About 25% of patients, those with WCC > 10 seem not to benefit from interferon

# Concluding remarks

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- MFPIgen and MFPI should help researchers detect, model and visualize interactions with continuous covariates
- Usually, we are **searching** for interactions, so small P-values are required
- Other methods not considered
  - STEPP – mainly graphical
  - ...

Thank you.

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## Cox (1984) paper: *Interaction*

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- Cox identifies 3 types of variable that might appear in interactions:
- Treatment variables
  - Can be modified or imposed
  - Treatments, e.g. chemotherapy, surgery
  - Behaviours, e.g. smoking, drinking
- Intrinsic variables
  - Cannot be modified
  - Often demographic, e.g. sex, age
- Unspecific variables
  - e.g. structural blocks, 'random' factors