

XSMLE - A Command to Estimate Spatial Panel Models in Stata

Federico Belotti^{*},
Gordon Hughes^o, Andrea Piano Mortari^{*}

^{*} CEIS, University of Rome Tor Vergata
^o School of Economics, University of Edinburg.

2013 German Stata Users Group Meeting
Potsdam, June 7 2013

- 1 Motivation
- 2 The `xsmle` command
 - Options common to all spatial models
 - The weighting matrix
 - Models' specific options
 - Postestimation command after `xsmle`
- 3 Examples
 - No missing data
 - Postestimation
 - Testing
 - Missing data
 - Testing with missing data

A general specification for Spatial Panel models:

$$y_{it} = \alpha + \tau y_{it-1} + \rho \sum_{j=1}^n w_{ij} y_{jt} + \sum_{k=1}^K x_{itk} \beta_k + \sum_{k=1}^K \sum_{j=1}^n w_{ij} x_{jtk} \theta_k + \mu_i + \gamma_t + \nu_{it} \quad (1)$$

$$\nu_{it} = \lambda \sum_{j=1}^n m_{ij} \nu_{jt} + \epsilon_{it} \quad i = 1, \dots, n \quad t = 1, \dots, T \quad (2)$$

Static Models ($\tau = 0$) and Dynamic Models ($\tau \neq 0$, Yu et al. (2008))

- if $\theta = 0$ and $\tau \neq 0$ → Spatial Autoregressive Model with Auto Regressive disturbances (SAC)
- if $\lambda = 0$ → Spatial Durbin Model (SDM)
- if $\lambda = 0$ and $\theta = 0$ → Spatial Autoregressive Model (SAR)
- if $\rho = 0$ and $\theta = 0$ → Spatial Error Model (SEM)
- if $\rho = 0$, $\theta = 0$, and $\mu_i = \phi \sum_{j=1}^n w_{ij} \mu_j + \eta_i$ → Generalised Spatial Panel Random Effects model (GSPRE)

A number of spatial-related routines have been written by users and available through SSC. A non-comprehensive list includes:

Data management and visualization

- `shp2dta` by K. Crow
- `spmat` by D.M. Drukker *et al*
- `spwmatrix` by P.W. Jeanty
- `spmap` by M. Pisati
- `geocode3` by S. Bernhard

Cross sectional data

- `spreg`: SAR, SEM, SAC *via* ML or GS2SLS by D. M. Drukker *et al*
- `spivreg`: SAC *via* GS2SLS by D. M. Drukker *et al*
- `spmlreg`: SAR, SEM, SDM, SAC *via* ML by P.W. Jeanty
- `spatreg`: SAR, SEM *via* ML by M. Pisati
- `spautoreg`: SAR, SEM, SDM, SAC *via* ML or GS2SLS by E.A. Shehata

Panel data

- `spreg*xt` suite SAR, SEM, SDM, SAC *via* LS, GLS, GMM or GS2SLS by E.A. Shehata (Lee, 2002)

DGP - 250 replications

$$y_{it} = \rho \sum_{j=1}^n w_{ij} y_{jt} + 0.3x_{1it} + 0.7x_{2it} + \mu_i + \gamma_t + \epsilon_{it} \quad n = 1, \dots, 188 \quad t = 1, \dots, 5$$

where the nuisance parameters μ_i ($i = 1, \dots, n$) are drawn from an iid standard Gaussian random variable. To allow for dependence between the unit-specific effects and the regressors, we generate the latter as follows

$$x_{kit} = 0.4\mu_i + (1 - 0.4^2)^{1/2} z_{kit},$$

where $k = 1, 2$ and the z_{kit} is an iid standard Gaussian random variable.

	$\rho = 0.3$		$\rho = 0.5$		$\rho = 0.7$	
	bias	MSE	bias	MSE	bias	MSE
xsmle	-0.0013	0.0020	-0.0016	0.0014	-0.0016	0.0007
spregfext	0.1473	0.0255	0.1972	0.0408	0.1859	0.0352
xtivreg2	0.0174	0.0091	0.0153	0.0063	0.0112	0.0033

`xsmle` fits (balanced) Spatial Panel data models *via* maximum likelihood (ML)

Requirements:

- (At least) Stata Version 10
- The $n \times n$ matrix of spatial weights. `xsmle` will deal with the longitudinal dimension automatically
- Data must be `tsset` or `xtset`

The basic `xsmle` syntax is the following

```
xsmle depvar [indepvars] [if] [in] [weight] [, options]
```

- The default model is the random-effects SAR model
- Only `aweight` are allowed but the declared `weight` variable must be constant within each panel unit
- The `mi` prefix is allowed
- Factor variables are allowed

Options common to all spatial models

- `model(name)` specifies the spatial model to be estimated. May be *sar* for the Spatial-AutoRegressive model, *sdm* for the Spatial Durbin Model, *sem* for the Spatial-Error Model, *sac* for the Spatial-Autoregressive with Spatially Autocorrelated Errors Model, *gspre* for the Generalised Spatial Random Effects Model.
- `re` use the random effects estimator; the default. This option cannot be specified when `model(sac)`.
- `fe` use the fixed effects estimator. This option cannot be specified when `model(gspre)`.
- `type(type_options [, leeyu])` specifies fixed-effects type; only for `fe` estimators. May be `ind` for spatial fixed effects, `time` for time fixed effects or `both` for both spatial and time fixed effects. Suboption `leeyu` allows to transform the data according to Lee and Yu (2010) approach and can be used only when `type(ind)`.

- `noconstant` suppresses the constant term in the model. Only for re-estimators.
- `noeffects` suppresses the computation of direct, indirect and total effects.
- `nsim(#)` sets the number of simulations for the LeSage and Pace (2009) procedure to compute the standard errors of the direct, indirect and total effects.
- `constraints(constraints_list)` applies specified linear constraints.
- `from(init_specs)` specifies initial values for the coefficients.
- `level(#)` sets confidence level for confidence intervals; default is `level(95)`.
- `postscore` save observation-by-observation scores in the estimation results list.
- `posthessian` save the Hessian corresponding to the full set of coefficients in the estimation results list.
- `hausman` performs the Hausman test.

Variance estimation

This section describes the arguments of the `vce(vcetype)` option.

- *oim* observed information matrix.
- *opg* outer product of the gradient vectors.
- *robust* clustered sandwich estimator where *clustvar* is the *panelvar*.
- *cluster clustvar* clustered sandwich estimator.
- *dkraay(#)* Driscoll-Kraay robust estimator. Where *#* is the maximum lag used in the calculation.

In `xsmle` the spatial weighting matrix can be

- a Stata matrix
- a `spmat` object

In both cases the matrix can be standardized or not.

e.g.

- a Stata matrix can be created using `matrix define`, imported from Mata using `st_matrix("string scalar name", real matrix)` or imported from GIS softwares like GeoDa using

```
spwmatrix gal using path_to_gal_file, wname(name_of_the_matrix)
```

- `spmat` objects are created by `spmat`

```
spmat import name_of_the_object using path_to_file
```

SAR model

- `wmatrix(name)` specifies the weight matrix for the spatial-autoregressive term.
- `dlag` includes (time) lagged dependent variable in the model.

SDM model

- `wmatrix(name)` specifies the weight matrix for the spatial-autoregressive term.
- `dmatrix(name)` specifies the weight matrix for the spatially lagged regressors; default is to use the matrix specified in `wmat(name)`.
- `durbin(dvarlist)` specifies the regressors that have to be spatially lagged; default is to lag all independent variables specified in `varlist`.
- `dlag` includes (time) lagged dependent variable in the model.

SEM model

- `ematrix(name)` specifies the weight matrix for the spatial-autocorrelated error term.

SAC model

- `wmatrix(name)` specifies the weight matrix for the spatial-autoregressive term.
- `ematrix(name)` specifies the weight matrix for the spatial-autocorrelated error term.

GSPRE model

- `wmatrix(name)` specifies the weight matrix for the spatial-autocorrelated random-effects.
- `ematrix(name)` specifies the weight matrix for the spatial-autocorrelated error term.
- `error(#)` defines the structure of the model. $\#$ is equal to 1 when $\lambda \neq \phi \neq 0$, $\#$ is equal to 2 when $\lambda = 0$, $\#$ is equal to 3 when $\phi = 0$, $\#$ is equal to 4 when $\lambda = \phi$.

Postestimation command allows to post-estimate spatial fixed or random effects. The methods implemented in this command are the panel data extensions of those available in Drukker, Prucha, and Raciborski (2011)

```
predict [type] newvar [if] [in] [, statistic]
```

where `statistic` includes:

- `rform` the default, calculates predicted values from the reduced-form equation: $y_{it} = (I_n - \rho W)^{-1}(x_{it}\beta + \alpha_i)$
- `limited` predicted values based on the limited information set. This option is available only when `model(sac)`.
- `naive` predicted values based on the observed values of $y_{it} = \rho W y_{it} + x_{it}\beta + \alpha_i$
- `xb` calculates the linear prediction including the fixed or random effect $x_{it}\beta + \alpha_i$.
- `a` estimates α_i , the fixed or random-effect. In the case of fixed-effects models, this statistic is allowed only when `type(ind)`

DGP - Fixed effects SDM

$$\begin{aligned}
 y_{it} &= 0.3 \sum_{j=1}^n w_{ij} y_{jt} + 0.5x_{1it} - 0.3x_{2it} - 0.2x_{3it} + 0.3 \sum_{j=1}^n w_{ij} x_{1it} + \\
 &+ 0.6 \sum_{j=1}^n w_{ij} x_{2it} + 0.9 \sum_{j=1}^n w_{ij} x_{3it} + \mu_i + \gamma_t + \epsilon_{it} \quad n = 1, \dots, 188 \quad t = 1, \dots, 5
 \end{aligned}$$

where the nuisance parameters μ_i ($i = 1, \dots, n$) are drawn from an iid standard Gaussian random variable. To allow for dependence between the unit-specific effects and the regressors, we generate the latter as follows

$$x_{kit} = 0.4\mu_i + (1 - 0.4^2)^{1/2} z_{kit},$$

where $k = 1, 2, 3$, z_{1it} is standard Gaussian, z_{2it} is $N(0, 1.5^2)$ and z_{3it} is $N(0, 2^2)$.

```
.. *** load a dta dataset containing the spatial contiguity matrix
. use ASL_contiguity_mat_ns.dta, clear

. *** get an spmat objects from dta
. spmat dta W W*, replace

. *** Summarize the spmat obj
. spmat summarize W, links
```

Summary of spatial-weighting object W

Matrix	Description
Dimensions	188 x 188
Stored as	188 x 188
Links	
total	906
min	1
mean	4.819149
max	13


```
. ** Fixed-effects Durbin model (correctly specified, row normalized W)
. xsmle y x1 x2 x3, wmat(W) model(sdm) fe type(ind) nsim(500) nolog
Warning: All regressors will be spatially lagged
```

SDM with spatial fixed-effects Number of obs = 940

Group variable: id Number of groups = 188

Time variable: t Panel length = 5

R-sq: within = 0.5727
 between = 0.3663
 overall = 0.4554

Mean of fixed-effects = -0.0137

Log-likelihood = -1230.7734

	y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----							
Main							
	x1	.5186041	.0364303	14.24	0.000	.4472019	.5900062
	x2	-.2946314	.0236541	-12.46	0.000	-.3409925	-.2482702
	x3	-.1923373	.0192912	-9.97	0.000	-.2301474	-.1545272
-----+-----							
Wx							
	x1	.3772047	.075502	5.00	0.000	.2292235	.5251859
	x2	.5765484	.0449332	12.83	0.000	.4884809	.6646159
	x3	.8692021	.0372769	23.32	0.000	.7961408	.9422634
-----+-----							
Spatial							
	rho	.2519025	.0374278	6.73	0.000	.1785454	.3252596
-----+-----							
Variance							
	sigma2_e	.7915998	.0366863	21.58	0.000	.7196959	.8635037
-----+-----							

[CONTINUES]

In a spatial setting, the effect of an explanatory variable change in a particular unit affects not only that unit but also its neighbors (LeSage and Pace, 2009).

$$\left[\frac{\partial Y}{\partial x_{nk}} \right] = (I - \rho W)^{-1} \begin{bmatrix} \beta_k & w_{12}\theta_k & \cdot & w_{1n}\theta_k \\ w_{21}\theta_k & \beta_k & \cdot & w_{2n}\theta_k \\ \cdot & \cdot & \cdot & \cdot \\ w_{n1}\theta_k & w_{n2}\theta_k & \cdot & \beta_k \end{bmatrix}$$

If we have only 2 units and 1 regressor:

- SAR and SAC $\rightarrow (I - \rho W)^{-1} \begin{bmatrix} \beta_1 & 0 \\ 0 & \beta_1 \end{bmatrix}$
- SEM $\rightarrow \begin{bmatrix} \beta_1 & 0 \\ 0 & \beta_1 \end{bmatrix}$
- SDM $\rightarrow (I - \rho W)^{-1} \begin{bmatrix} \beta_1 & w_{12}\theta_1 \\ w_{21}\theta_1 & \beta_1 \end{bmatrix}$

[CONTINUES]

Direct								
	x1		.5481382	.0362326	15.13	0.000	.4771237	.6191527
	x2		-.2642811	.0231199	-11.43	0.000	-.3095953	-.2189669
	x3		-.1422518	.0176968	-8.04	0.000	-.1769369	-.1075668

Indirect								
	x1		.6480929	.090572	7.16	0.000	.470575	.8256108
	x2		.6450951	.0599307	10.76	0.000	.5276331	.7625571
	x3		1.050599	.058257	18.03	0.000	.9364176	1.164781

Total								
	x1		1.196231	.1038425	11.52	0.000	.9927034	1.399759
	x2		.380814	.0677252	5.62	0.000	.2480751	.513553
	x3		.9083474	.0660288	13.76	0.000	.7789334	1.037761

. estimates store sdm_fe

```
** Fixed-effects Durbin model (correctly specified, row normalized W)
. xsmle y x1 x2 x3, wmat(W) model(sdm) re type(ind) nsim(500) nolog noeff
Warning: Option type(ind) will be ignored
Warning: All regressors will be spatially lagged
```

```
SDM with random-effects                Number of obs =    940

Group variable: id                      Number of groups =   188
Time variable: t                        Panel length =     5

R-sq:  within = 0.5666
       between = 0.4543
       overall = 0.4936
```

Log-likelihood = -1513.7006

	y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----+-----						
Main						
	x1	.6230976	.0408605	15.25	0.000	.5430126 .7031826
	x2	-.2439834	.0264129	-9.24	0.000	-.2957518 -.192215
	x3	-.1688081	.0211584	-7.98	0.000	-.2102778 -.1273385
	_cons	-.0169191	.0811545	-0.21	0.835	-.1759791 .1421409
-----+-----						
Wx						
	x1	.3706183	.0824133	4.50	0.000	.2090911 .5321454
	x2	.557779	.0493092	11.31	0.000	.4611347 .6544234
	x3	.8845199	.0411496	21.50	0.000	.8038681 .9651717
-----+-----						
Spatial						
	rho	.2472432	.0376366	6.57	0.000	.1734769 .3210096
-----+-----						
Variance						
	lgt_theta	-.3920581	.1040247	-3.77	0.000	-.5959428 -.1881735
	sigma_e	1.005536	.0528831	19.01	0.000	.9018867 1.109185
-----+-----						

```
hausman sdm_fe sdm_re, eq(1:1 2:2 3:3)
```

		---- Coefficients ----			
		(b)	(B)	(b-B)	sqrt(diag(V_b-V_B))
		sdm_fe	sdm_re	Difference	S.E.
-----+-----					
comp1					
	x1	.5186041	.6230976	-.1044935	.
	x2	-.2946314	-.2439834	-.050648	.
	x3	-.1923373	-.1688081	-.0235292	.
-----+-----					
comp2					
	x1	.3772047	.3706183	.0065864	.
	x2	.5765484	.557779	.0187694	.
	x3	.8692021	.8845199	-.0153178	.
-----+-----					
comp3					
	rho	.2519025	.2472432	.0046593	.
-----+-----					

b = consistent under Ho and Ha; obtained from xsmle
 B = inconsistent under Ha, efficient under Ho; obtained from xsmle

Test: Ho: difference in coefficients not systematic
 $\chi^2(7) = (b-B)'[(V_b-V_B)^{-1}](b-B)$
 = -47.08 $\chi^2 < 0 \implies$ model fitted on these
 data fails to meet the asymptotic
 assumptions of the Hausman test;
 see suest for a generalized test

```
. ** Fixed-effects Durbin model (correctly specified, row normalized W)
. xsmle y x1 x2 x3, wmat(W) model(sdm) fe type(ind) hausman noeff nolog
Warning: All regressors will be spatially lagged
... estimating random-effects model to perform Hausman test
SDM with spatial fixed-effects          Number of obs =      940

Group variable: id                      Number of groups =     188
Time variable: t                        Panel length =        5

R-sq:  within = 0.5727
       between = 0.3663
       overall = 0.4554

Mean of fixed-effects = -0.0137

Log-likelihood = -1230.7734
-----+-----
      y |      Coef.   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
Main   |
      x1 |   .5186041   .0364303    14.24  0.000   .4472019   .5900062
      x2 |  -.2946314   .0236541   -12.46  0.000  -.3409925  -.2482702
      x3 |  -.1923373   .0192912    -9.97  0.000  -.2301474  -.1545272
-----+-----
Wx     |
      x1 |   .3772047   .075502    5.00  0.000   .2292235   .5251859
      x2 |   .5765484   .0449332   12.83  0.000   .4884809   .6646159
      x3 |   .8692021   .0372769   23.32  0.000   .7961408   .9422634
-----+-----
Spatial |
      rho |   .2519025   .0374278    6.73  0.000   .1785454   .3252596
-----+-----
Variance |
      sigma2_e |   .7915998   .0366863   21.58  0.000   .7196959   .8635037
-----+-----
Ho: difference in coeffs not systematic chi2(7) = 89.58   Prob>=chi2 = 0.0000
-----+-----
```

```
. estat ic
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	940	.	-1230.773	8	2477.547	2516.314

Note: N=Obs used in calculating BIC; see [R] BIC note

```
. estimates store sdm
```

```
. ***** Postestimation
. predict yhat, rform

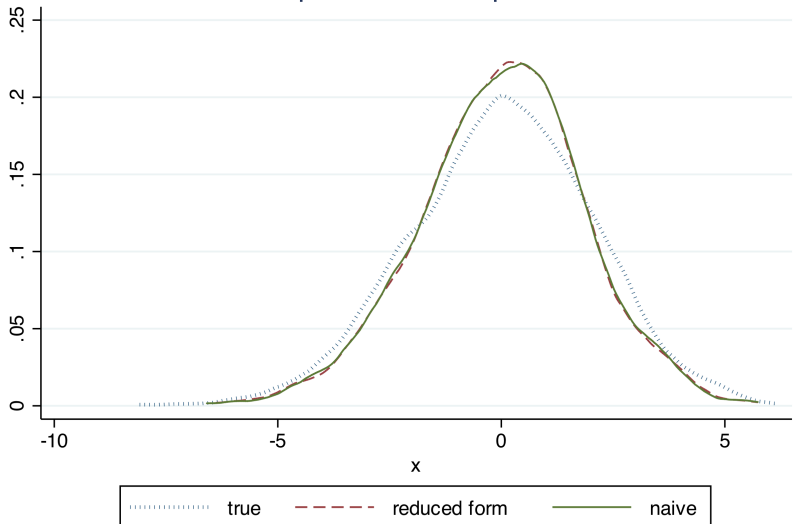
. predict yhat1, naive

. predict alphahat, a

. sum alpha alphahat
```

Variable	Obs	Mean	Std. Dev.	Min	Max
alpha	940	.037577	1.058726	-2.261747	3.343453
alphahat	940	-.013692	1.182919	-2.688471	4.028156

xsmle postestimation prediction



Using matrix notation the SDM ($\lambda = 0$) may be derived from a SEM model

$$\begin{cases} \mathbf{y} &= \mathbf{X}\boldsymbol{\beta} + \mathbf{u} \\ \mathbf{u} &= \lambda \mathbf{W}\mathbf{u} + \boldsymbol{\epsilon} \end{cases}$$

hence

$$\begin{aligned} \mathbf{u}(1 - \lambda \mathbf{W}) &= \boldsymbol{\epsilon} \\ \mathbf{y}(1 - \lambda \mathbf{W}) &= \mathbf{X}\boldsymbol{\beta}(1 - \lambda \mathbf{W}) + \boldsymbol{\epsilon} \\ \mathbf{y} &= \lambda \mathbf{W}\mathbf{y} + \mathbf{X}\boldsymbol{\beta} - \lambda \mathbf{W}\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon} \\ \mathbf{y} &= \lambda \mathbf{W}\mathbf{y} + \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\theta}\mathbf{W}\mathbf{X} + \boldsymbol{\epsilon} \end{aligned}$$

and test the following constraints

- ① $\boldsymbol{\theta} = 0$ and $\lambda \neq 0 \Rightarrow$ the model is a SAR
- ② $\boldsymbol{\theta} = -\boldsymbol{\beta}\lambda \Rightarrow$ the model is a SDM.

```

** Test for SAR
. test [Wx]x1 = [Wx]x2 = [Wx]x3 = 0

( 1) [Wx]x1 - [Wx]x2 = 0
( 2) [Wx]x1 - [Wx]x3 = 0
( 3) [Wx]x1 = 0

      chi2( 3) = 740.80
    Prob > chi2 = 0.0000

.
. ** Test for SEM
. testnl ([Wx]x1 = -[Spatial]rho*[Main]x1) ([Wx]x2 = -[Spatial]rho*[Main]x2) ([
> [Wx]x3 = -[Spatial]rho*[Main]x3)

(1) [Wx]x1 = -[Spatial]rho*[Main]x1
(2) [Wx]x2 = -[Spatial]rho*[Main]x2
(3) [Wx]x3 = -[Spatial]rho*[Main]x3

      chi2(3) = 545.31
    Prob > chi2 = 0.0000
  
```

```

** Test for SAC
. xsmle y x1 x2 x3, wmat(W) emat(W) model(sac) fe type(ind) noeff nolog

SAC with spatial fixed-effects                Number of obs =      940

Group variable: id                            Number of groups =    188
Time variable: t                              Panel length =       5

R-sq:    within = 0.2652
         between = 0.0011
         overall = 0.0912

Mean of fixed-effects = -0.0117

Log-likelihood = -1386.0860
-----+-----
          y |      Coef.   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
Main      |
   x1 |   .3212791   .0341734     9.40   0.000   .2543005   .3882577
   x2 |  -.3135993   .0232111    -13.51  0.000  -.3590923  -.2681064
   x3 |  -.2997975   .0178884    -16.76  0.000  -.334858  -.2647369
-----+-----
Spatial   |
   rho |  -.6676721   .0542468    -12.31  0.000  -.7739939  -.5613504
   lambda | .8426981   .0209346     40.25  0.000   .801667   .8837293
-----+-----
Variance  |
   sigma2_e |  1.001782   .0440957    22.72  0.000   .9153562   1.088208
-----+-----

```

```
. estat ic
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	940	.	-1386.086	6	2784.172	2813.247

Note: N=Obs used in calculating BIC; see [R] BIC note

```
.
```

DGP - Fixed effects SDM with missing values

$$\begin{aligned}
 y_{it} = & 0.5 \sum_{j=1}^n w_{ij} y_{jt} - 1.5x_{1it} - 0.7x_{2it} - 0.3x_{3it} - 0.9x_{4it} + \\
 & + 0.75 \sum_{j=1}^n w_{ij} x_{1it} + 0.35 \sum_{j=1}^n w_{ij} x_{2it} + .15 \sum_{j=1}^n w_{ij} x_{3it} + \\
 & + .45 \sum_{j=1}^n w_{ij} x_{4it} + \mu_i + \gamma_t + \epsilon_{it} \quad n = 1, \dots, 188 \quad t = 1, \dots, 5
 \end{aligned}$$

where the nuisance parameters μ_i ($i = 1, \dots, n$) are drawn from an iid standard Gaussian random variable. To allow for dependence between the unit-specific effects and the regressors, we generate the latter as follows

$$x_{kit} = 0.4\mu_i + (1 - 0.4^2)^{1/2} z_{kit},$$

where $k = 1, 2, 3, 4$, z_{1it} and z_{3it} are $N(0, 1.5^2)$ and z_{2it} and z_{4it} are standard Gaussian. 5% missing values are randomly assigned to x_{1it} and x_{3it} to generate xm_{1it} and xm_{3it} .

```
. sum y x1 x1m x2 x3 x3m x4
```

Variable	Obs	Mean	Std. Dev.	Min	Max
y	940	.0392334	2.649594	-7.947968	7.694182
x1	940	.0447793	1.456754	-5.412737	4.76092
x1m	901	.039974	1.452938	-5.412737	4.76092
x2	940	.0209358	1.03554	-3.141913	3.0158
x3	940	-.0236917	1.426842	-4.23553	4.351037
x3m	884	-.0292849	1.433476	-4.23553	4.351037
x4	940	-.0297022	1.020829	-3.721222	3.210468

```
. ** Fixed-effects SDM model (correctly specified)
. xsmle y x1 x2 x3 x4, wmat(W) model(sdm) fe type(ind) nolog noeff
Warning: All regressors will be spatially lagged
```

[OUTPUT OMITTED]

```
. estimates store sdm_nomissing
```

```
mi set wide
. mi register imputed x1m x3m
. mi impute mvn x1m x3m = x2 x4, add(50) rseed(12345)
```

Performing EM optimization:

```
note: 7 observations omitted from EM estimation because of all imputation
      variables missing
      observed log likelihood = -1518.1423 at iteration 5
```

Performing MCMC data augmentation ...

```
Multivariate imputation           Imputations =      50
Multivariate normal regression      added =      50
Imputed: m=1 through m=50          updated =       0
```

```
Prior: uniform                     Iterations =     5000
                                   burn-in =      100
                                   between =      100
```

Variable	Observations per m			total
	complete	incomplete	imputed	
x1m	901	39	39	940
x3m	884	56	56	940

(complete + incomplete = total; imputed is the minimum across m of the number of filled in observations.)


```
. ***** SDM estimates using multiple imputed data
. mi estimate (coeff1: [Wx]x1m + [Spatial]rho*[Main]x1m) ///
> (coeff2: [Wx]x2 + [Spatial]rho*[Main]x2) ///
> (coeff3: [Wx]x3m + [Spatial]rho*[Main]x3m) ///
> (coeff4: [Wx]x4 + [Spatial]rho*[Main]x4), ///
> dots post saving(sdm_imputed, replace): ///
> xsmle y x1m x2 x3m x4, wmat(w) model(sdm) ///
> fe type(ind) nolog noeff
```

Imputations (50):

.....10.....20.....30.....40.....50 done

```
Multiple-imputation estimates      Imputations      =      50
SDM with spatial fixed-effects     Number of obs    =     940
                                   Average RVI       =     0.3120
DF adjustment:  Large sample       DF:  min         =     168.38
                                   avg         =    1679.48
                                   max         =    4655.10
Model F test:      Equal FMI       F( 10, 8545.3) =     332.30
Within VCE type:   OIM             Prob > F         =     0.0000
```

	y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----							
Main							
	x1m	-1.430412	.0337574	-42.37	0.000	-1.496734	-1.36409
	x2	-.7335296	.0466659	-15.72	0.000	-.825132	-.6419272
	x3m	-.274662	.0313065	-8.77	0.000	-.3361047	-.2132194
	x4	-.9359276	.044795	-20.89	0.000	-1.023766	-.8480894
-----+-----							
Wx							
	x1m	.5188195	.0864699	6.00	0.000	.3491775	.6884615
	x2	.1321088	.0904967	1.46	0.144	-.0453216	.3095391
	x3m	.0546988	.0643793	0.85	0.396	-.0716161	.1810137
	x4	.3360411	.1018539	3.30	0.001	.1363592	.535723
-----+-----							

```
-----+-----
Spatial rho | .3869484 .0429191 9.02 0.000 .3027466 .4711502
-----+-----
Variance sigma2_e | 1.068042 .0737025 14.49 0.000 .9225424 1.213542
-----+-----
```

```
Transformations Average RVI = 0.1161
DF adjustment: Large sample DF: min = 1606.62
              avg = 7184.40
Within VCE type: OIM max = 12514.96
```

```
coeff1: [Wx]x1m + [Spatial]rho*[Main]x1m
coeff2: [Wx]x2 + [Spatial]rho*[Main]x2
coeff3: [Wx]x3m + [Spatial]rho*[Main]x3m
coeff4: [Wx]x4 + [Spatial]rho*[Main]x4
```

```
-----+-----
      y |      Coef.   Std. Err.      t    P>|t|     [95% Conf. Interval]
-----+-----
coeff1 |  -.0347924   .0575259    -0.60   0.545    - .147552   .0779673
coeff2 |  -.1516852   .0829808    -1.83   0.068    - .314351   .0109806
coeff3 |  -.0516019   .0636324    -0.81   0.418    - .1764131   .0732093
coeff4 |  -.0261852   .0942367    -0.28   0.781    - .210917   .1585465
-----+-----
```

```
. estimates store sdm_imputed
```

```
. ** Test for SAR
. mi test [Wx]x1m [Wx]x2 [Wx]x3m [Wx]x4
note: assuming equal fractions of missing information
```

- (1) [Wx]x1m = 0
- (2) [Wx]x2 = 0
- (3) [Wx]x3m = 0
- (4) [Wx]x4 = 0

```
F( 4,8383.2) = 10.09
Prob > F = 0.0000
```

```
. ** Test for SEM
. mi testtr coeff1 coeff2 coeff3 coeff4
note: assuming equal fractions of missing information
```

```
coeff1: [Wx]x1m + [Spatial]rho*[Main]x1m
coeff2: [Wx]x2 + [Spatial]rho*[Main]x2
coeff3: [Wx]x3m + [Spatial]rho*[Main]x3m
coeff4: [Wx]x4 + [Spatial]rho*[Main]x4
```

- (1) coeff1 = 0
- (2) coeff2 = 0
- (3) coeff3 = 0
- (4) coeff4 = 0

```
F( 4,17424.6) = 1.13
Prob > F = 0.3395
```

```
. estout sdm_nomissing sdm_imputed, c(b se) ren(x1m x1 x3m x3)
```

```
-----
```

	sdm_nomiss ^g b/se	sdm_imputed b/se

Main		
x1	-1.49692 .023471	-1.430412 .0337574
x2	-.7440764 .0340518	-.7335296 .0466659
x3	-.297554 .0233121	-.274662 .0313065
x4	-.9451319 .0349799	-.9359276 .044795

Wx		
x1	.6416344 .070266	.5188195 .0864699
x2	.1982546 .0729333	.1321088 .0904967
x3	.057889 .0494267	.0546988 .0643793
x4	.4091564 .0829535	.3360411 .1018539

Spatial		
rho	.4530939 .0356911	.3869484 .0429191

Variance		
sigma2_e	.7564478 .0355401	1.068042 .0737025

- Lee, L. F. (2002). Consistency and efficiency of least squares estimation for mixed regressive, spatial autoregressive models. *Econometric Theory*, 18(02):252–277.
- Lee, L. F. and Yu, J. (2010). Estimation of spatial autoregressive panel data models with fixed effects. *Journal of Econometrics*, 154(2):165–185.
- LeSage, J. P. and Pace, R. K. (2009). *Introduction to Spatial Econometrics*. Taylor & Francis.
- Yu, J., de Jong, R., and Lee, L.-F. (2008). Quasi-maximum likelihood estimators for spatial dynamic panel data with fixed effects when both n and t are large. *Journal of Econometrics*, 146:118–134.