
How to assess the fit of multilevel logit models with Stata?

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“Models should not be true but it is important that they are applicable.”

John W. Tukey

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1. What is the problem ?

Current situation in applied research:

- An increasing number of people uses multilevel logistic models for qualitative dependent variables with binary and ordinal outcome
- But users often complain that there are no fit measures for these models
- Neither Stata 14 / 15 nor SPSS 24 offer any fit measure for these models
- Let me demonstrate how to generalize the Pseudo R^2 s for binary and ordinal logit model for the multilevel analysis

Which solutions does Stata provide?

- Indeed Stata estimates multilevel logit models for binary, ordinal and multinomial outcomes (melogit, meologit, gllamm) but it does not calculate any Pseudo R^2 . It provides only the information criteria AIC and BIC (estat ic)
- Stata provides a Wald-test for the fixed-effects and a Likelihood-Ratio- χ^2 test for the random effects of the exogenous variables
- Even special purpose programs like HLM, MlwiN, MPLUS or SuperMix do not calculate any Pseudo R^2

What can we learn from multilevel literature?

- Raudenbush & Bryk (2002), Heck & Thomas (2009) and Rabe-Hesketh & Skrondal (2013) do not mention Pseudo R^2 s
- Snijder & Bosker(2012) propose a variation of McKelvey & Zavoina Pseudo R^2 for random-intercept- and intercept-as-outcome logit models. It is not implemented in any program
- Hox (2010) discusses the McFadden, Cox & Snell, Nagelkerke and McKelvey & Zavoina Pseudo R^2 . He recommends the last one to assess the model fit

2. Summary of the econometric Monte-Carlo studies for testing Pseudo R^2 s

- Econometricians made a lot of Monte-Carlo studies in the early 90s:
 - ▶ Hagle & Mitchell 1992
 - ▶ Veall & Zimmermann 1992, 1993, 1994
 - ▶ Windmeijer 1995
 - ▶ DeMaris 2002
- They tested systematically the most common Pseudo- R^2 s for binary and ordinal probit / logit models

Which Pseudo R²s were tested in these studies?

- Likelihood-based measures:
 - ▶ Maddala / Cox & Snell Pseudo R² (1983 / 1989)
 - ▶ Cragg & Uhler / Nagelkerke Pseudo R² (1970 / 1992)
- Log-Likelihood-based measures:
 - ▶ McFadden Pseudo-R² (1974)
 - ▶ Aldrich & Nelson Pseudo R² (1984)
 - ▶ Aldrich & Nelson Pseudo R² with the Veall & Zimmermann correction (1992)
- Basing on the estimated probabilities:
 - ▶ Efron / Lave Pseudo R² (1970 / 1978)
- Basing on the variance decomposition of the estimated Probits / Logits:
 - ▶ McKelvey & Zavoina Pseudo R² (1975)

Results of the Monte-Carlo-Studies for binary and ordinal logits or probits

- The McKelvey & Zavoina Pseudo R^2 is the best estimator for the “true R^2 ” of the OLS regression
- The Aldrich & Nelson Pseudo R^2 with the Veall & Zimmermann correction is the best approximation of the McKelvey & Zavoina Pseudo R^2
- Lave / Efron, Aldrich & Nelson, McFadden and Cragg & Uhler Pseudo R^2 severely underestimate the “true R^2 ” of the OLS regression
- My personal advice:
 - ▶ Use the McKelvey & Zavoina Pseudo R^2 to assess the fit of binary and ordinal logit models

3. The generalization of the McKelvey & Zavoina Pseudo R^2 for the binary and ordinal multilevel logit model

- The multilevel logit model is a systematic extension of the classical binary and ordinal logit model for clustered subsamples (contextual units j)
 - ▶ The variance of the estimated logits is decomposed into ▶ Fixed effects, ▶ Random effects and ▶ Level-1 Error variance $\sigma^2(r_{ij})$
 - ▶ Because of its own heteroscedasticity the variance of level 1 residua $\sigma^2(r_{ij})$ can not be estimated. It is replaced by the variance of the logistic density function $(\pi^2 / 3)$

Let's have a short look at the lucky winner

● McKelvey & Zavoina Pseudo R^2 (M & Z Pseudo R^2)

$$M \ \& \ Z \ Pseudo \ R^2 = \frac{Var(\hat{y}^*)}{Var(\hat{y}^*) + Var(\varepsilon)} = \frac{\frac{\sum_{i=1}^n (\hat{y}_i^* - \overline{\hat{y}^*})^2}{n}}{\frac{\sum_{i=1}^n (\hat{y}_i^* - \overline{\hat{y}^*})^2}{n} + \frac{\pi^2}{3}}$$

Range: $0 \leq M \ \& \ Z\text{-Pseudo } R^2 \leq 1$

Legend:

$Var(\hat{y}^*)$: Variance of the estimated logits (latent variable Y^*)

\hat{y}_i^* : Estimated logit of case i

$\overline{\hat{y}^*}$: Expected value of the estimated logits

$\frac{\pi^2}{3}$: Variance of the logistic density function

Generalization to the 2-level logit model 2

- Prediction of the latent variable Y^* (estimated binary or cumulative logit) in two ways
 - ▶ Population-Average Prediction with the fixed effects of the exogenous variables (all random effects hold at zero)
 - Stata-command: `predict newvar1 if e(sample), xb`
 - ▶ Unit-Specific Prediction of the fixed and random effects of the exogenous variable
 - Stata-command: `predict newvar2 if e(sample), eta`
- Therefore, the variance of the estimated logits (Y^*) can be calculated in two different ways
 - ▶ Only for the fixedeffects of the exogenous variables
 - ▶ For the fixed and random effects of the exogenous variables

Generalization to the 2-level logit model 3

- Therefore we get two different McKelvey & Zavoina Pseudo R^2 s
 - ▶ “Population-Average” M & Z Pseudo R^2 (fixed effects)
 - ▶ “Unit-Specific” M & Z Pseudo R^2 (fixed- & random effects)
- For the “Unit-Specific” M & Z Pseudo R^2 uses estimated fixed and random effects for prediction, it assesses the fit more realistically as its “Population-Average” counterpart

Let's have a short look at the lucky loser

- McFadden-Pseudo R^2 (1974)

$$McFadden\ Pseudo\ R^2\ (\rho^2) = 1 - \left[\frac{\log L_A}{\log L_0} \right]$$

Range: $0 \leq McFadden\ Pseudo\ R^2 < 1$

but ρ^2 does not reach the maximum of 1.0

Rule of thumb: $0.20 \leq McFadden\ Pseudo\ R^2 \leq 0.40$ marks an excellent fit (McFadden 1979: 307)

Legend: $\log L_A$: Log-Likelihood of the actual model
 $\log L_0$: Log-Likelihood of the zero model

Generalization to the 2-level logit model 4

- Conditions of application
 - ▶ Maximum-Likelihood estimation of the fixed and random effects of the exogenous variables
 - ▶ Actual and zero model has to use the same sample
 - ▶ Choice of the “appropriate zero model” (M_0) depends on our knowledge to which context the respondent belongs
 - **Membership known:** Random-Intercept-Only Logit model estimates the proportion of Y^* which can be maximally explained by the context (= ANOVA model)
 - **Membership unknown:** Fixed-Intercept-Only Logit model estimates only the marginal distribution of Y^* (= true zero model)

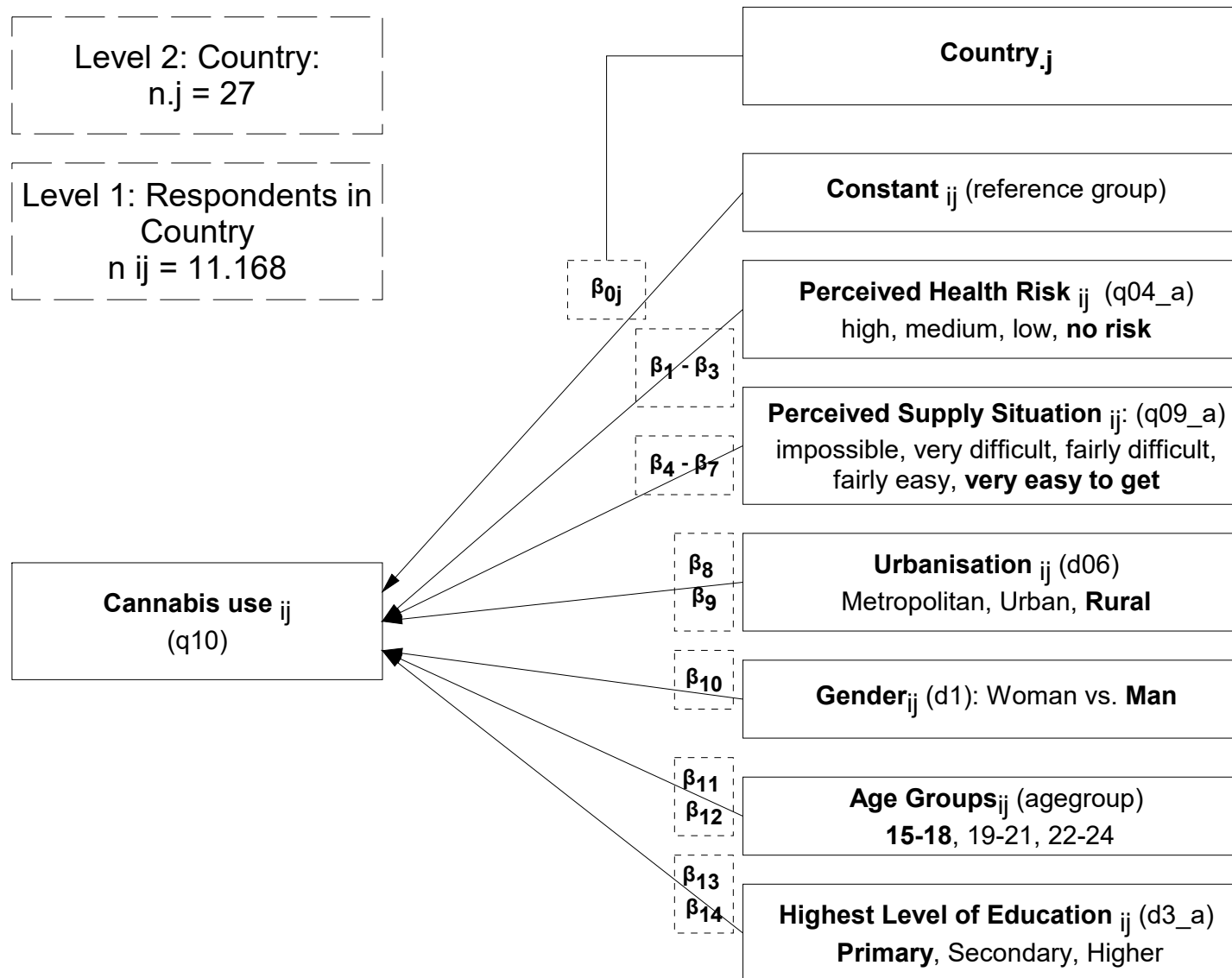
Generalization to the 2-level logit model 5

- Calculation of McFadden Pseudo R^2 is possible in two different ways using the following as a zero model
 - ▶ Random-Intercept-Only Logit-Model
 - It measures the proportional reduction of the log likelihood of the actual model caused by the fixed effects of the exogenous variables in comparison to the RIOM
 - Its Likelihood-Ratio χ^2 test refers to all fixed effects of the exogenous level 1 and level 2 variables
 - ▶ Fixed-Intercept-Only Logit-Model
 - It measures the proportional reduction of the log likelihood of the actual model caused by fixed and random effects of all exogenous variables in comparison to the FIOM
 - Its Likelihood-Ratio χ^2 test refers to all fixed and random effects of the exogenous level 1 and level 2 variables

4. Example of application

- Flash Eurobarometer No 330 about youth attitudes on drugs (2011)
 - ▶ WebCATI-Survey of $n_{ij} = 12.313$ respondents (aged 15 - 24) in $n_j = 27$ EU member states (contextual units j)
 - ▶ My focus:
 - prevalence of cannabis use by juveniles and young adults (q10): Have you used cannabis by yourself?
 - 1) never
 - 2) more than 12 months ago
 - 3) less than 12 months ago
 - 4) in the last 30 days
 - ▶ Let us have a look at the exogenous variables in the following diagram

Theoretical 2-level-model: RIM



Stata-Output Version 14

Mixed-effects ologit regression
Group variable: country

Number of obs = 11,168
Number of groups = 27

Obs per group:
min = 211
avg = 413.6
max = 490

Integration method: mvaghermite

Integration pts. = 7

Log likelihood = -7410.7117

Wald chi2(14) = 2142.78
Prob > chi2 = 0.0000

q10ord	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	

q4_a						
high risk	-2.670499	.1092326	-24.45	0.000	-2.884591	-2.456407
medium risk	-1.696693	.0730464	-23.23	0.000	-1.839861	-1.553525
low risk	-.7425748	.0611709	-12.14	0.000	-.8624676	-.622682
q9_a						
impossible	-3.006983	.1899514	-15.83	0.000	-3.379281	-2.634685
very difficult	-2.191986	.1207629	-18.15	0.000	-2.428677	-1.955295
fairly difficult	-1.555672	.0870857	-17.86	0.000	-1.726357	-1.384987
fairly easy	-.6291072	.0553719	-11.36	0.000	-.7376341	-.5205803
d6						
metropolitan zone	.3536598	.0713306	4.96	0.000	.2138545	.4934652
other town/urban centre	.196061	.0606935	3.23	0.001	.0771039	.315018
d1						
female	-.4654088	.0504709	-9.22	0.000	-.5643298	-.3664877
agegroup						
19 - 21	.4924681	.073827	6.67	0.000	.3477699	.6371663
22 - 24	.6847313	.0797637	8.58	0.000	.5283974	.8410652
d3_a						
secondary education	-.0345302	.0753855	-0.46	0.647	-.1822832	.1132227
higher education	-.0415283	.099673	-0.42	0.677	-.2368837	.1538271

/cut1	-.4269461	.1329725	-3.21	0.001	-.6875674	-.1663248
/cut2	.6715688	.133064	5.05	0.000	.4107681	.9323695
/cut3	1.857033	.1357061	13.68	0.000	1.591053	2.123012

country						
var(_cons)	.2623196	.0849424			.1390597	.494835

LR test vs. ologit model: chibar2(01) = 222.09

Prob >= chibar2 = 0.0000

● Fixed effects

● Thresholds

● Random effect

What does Stata offer to assess the fit?

- Akaike (AIC) und Schwarz Bayesian Information Criterion (BIC)
 - ▶ Decision rule: Choose the model with the lowest AIC or BIC

```
. estimates stats fiom riom rim
```

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
fiom	11,168	.	-9326.802	3	18659.6	18681.57
riom	11,168	.	-9033.234	4	18074.47	18103.75
rim	11,168	.	-7410.712	18	14857.42	14989.2

Note: N=Obs used in calculating BIC; see [R] BIC note.

- ▶ Looking at AIC and BIC, the rim fits best of all bad models
- ▶ But we do not know how well the rim fits

Output of my fit_meologit_2lev.ado

1

- Assessing the fit by the McKelvey & Zavoina-Pseudo R^2 s and the Intra-Class-Correlation

```
. fit_meologit_2lev
```

```
Fit-measures for the MELOGIT/MEOLOGIT-model:
```

```
McKelvey&Zavoina-Pseudo-R2 (fixed&random effects)= 0.5137
```

```
McKelvey&Zavoina-Pseudo-R2 (fixed effects only)= 0.4774
```

```
Just estimating the Random-/Fixed Intercept Only Logit-Model
```

```
Intra-Class-Correlation (Level 2) = 0.1507
```

Output of my fit_melogit_2lev.ado

2

● McFadden Pseudo R²s and corresponding Likelihood-Ratio- χ^2 tests

McFadden Pseudo-R2 (M_A vs. Random-Intercept-Only-Logit Model) = 0.1796

McFadden Pseudo-R2 (M_A vs. Fixed-Intercept-Only-Logit Model) = 0.2054

Likelihood-Ratio-chi2-Test (H0: All fixed effects = 0)

Likelihood-ratio test	LR chi2(14) =	3245.04
(Assumption: <u>riom</u> nested in <u>ma</u>)	Prob > chi2 =	0.0000

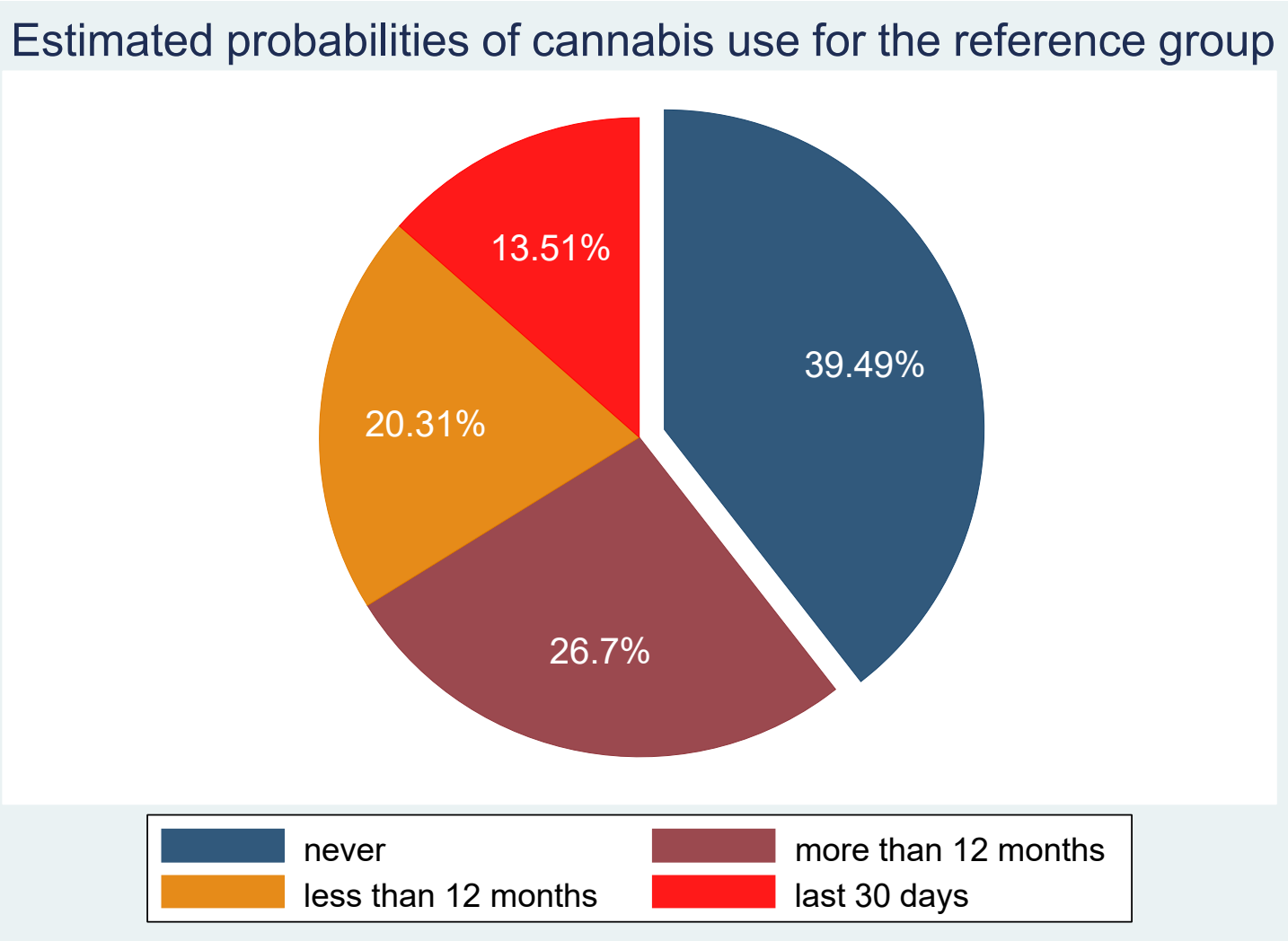
Likelihood-Ratio-chi2-Test (H0: All fixed & random effects = 0)

Likelihood-ratio test	LR chi2(15) =	3832.18
(Assumption: <u>fiom</u> nested in <u>ma</u>)	Prob > chi2 =	0.0000

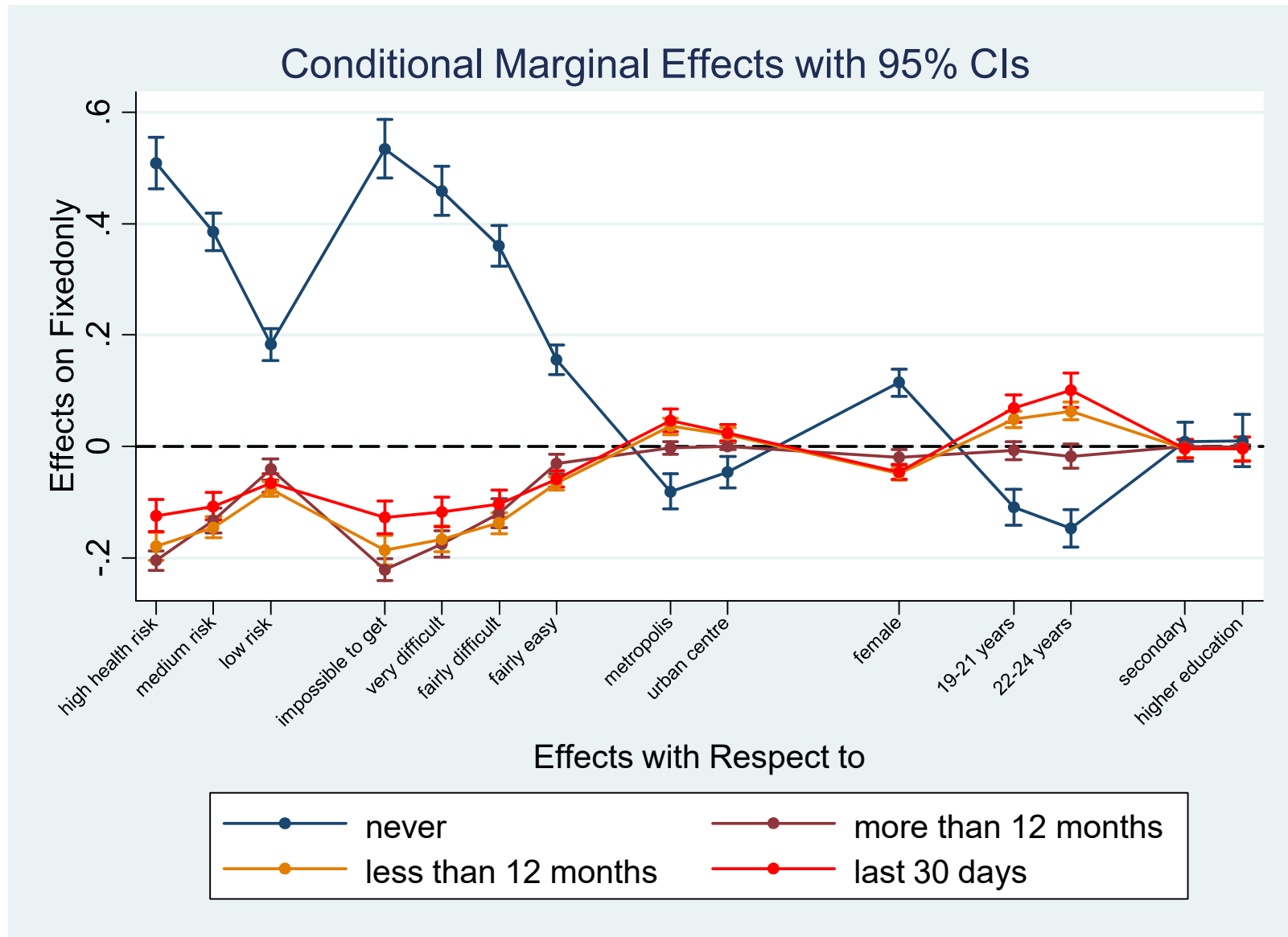
Note: The reported degrees of freedom assumes the null hypothesis is not on the boundary of the parameter space. If this is not true, then the reported test is conservative.

How does the effects look like?

- The baseline



The joint marginsplot for the 4 categories



- Known
 - ▶ The Monte-Carlo-simulation studies show that the McKelvey & Zavoina Pseudo R^2 is the best fit measure for binary and ordinal logit models
- New
 - ▶ Generalization of the M & Z-Pseudo R^2 to binary and ordinal multilevel logit models. The prediction of estimated logits bases upon the fixed effects only or upon fixed and random effects of exogenous variables
 - ▶ The McFadden-Pseudo R^2 bases upon the fixed effects only or upon fixed and random-effects of the exogenous variables using a context-independent zero model

- New
 - ▶ Simultaneous Likelihood-Ratio- χ^2 test for the estimated fixed effects using the random-intercept-only (RIOM) as the zero model
 - ▶ Simultaneous Likelihood-Ratio- χ^2 test for the estimated fixed and random effects using the fixed-intercept-only (FIOM) as the zero model
- That's why
 - ▶ I suggest to use my `fit_meologit_2lev.ado` and `fit_meologit_3lev.ado` to assess the fit of 2- and 3-level logit models with binary and ordinal outcome

Closing words

- Thank you for your attention
- Do you have some questions?

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Stata code for fit_meologit_2lev.ado

1

```
program fit_meologit_2lev, rclass
version 14

tempvar plgt1

quietly estimates store ma

quietly predict `plgt1' if e(sample), eta

quietly sum `plgt1'

display as text "Fit-measures for the MELOGIT/MEOLOGIT-model:"
display as text " "
display as text "McKelvey&Zavoina-Pseudo-R2 (fixed&random effects)= " as result %6.4f ///
  abs(r(Var)*r(N)-1) / ((r(N)*(_pi^2 / 3)+ (r(Var)*r(N)-1)))
display as text " "
drop `plgt1'

tempvar plgt2
quietly predict `plgt2' if e(sample), xb

quietly sum `plgt2'

display as text "McKelvey&Zavoina-Pseudo-R2 (fixed effects only)= " as result %6.4f ///
  abs(r(Var)*r(N)-1) / ((r(N)*(_pi^2 / 3)+ (r(Var)*r(N)-1)))

drop `plgt2'

dis " "

capture drop llma

tempvar llma

gen llma=`e(ll)'

dis as text " "
dis as text "Just estimating the Random-/Fixed Intercept Only Logit-Model"
dis as text " "
```

Stata code for fit_meologit_2lev.ado

2

```
* Schaetzung des RIOM
quietly: `e(cmd2)' `e(depvar)' if e(sample), || `e(ivars)':

quietly: estimates store riom

* Berechnung der Intra-Class-Correlation (ICC)
display as text "Intra-Class-Correlation (Level 2) = " as result %6.4f ///
  (_b[var(_cons[`e(ivars)']):_cons]) / (_b[var(_cons[`e(ivars)']):_cons] + (_pi^2 / 3))

dis as text " "
dis as text "McFadden Pseudo-R2 (M_A vs. Random-Intercept-Only-Logit Model) = " ///
  as result %6.4f abs(1- (llma / `e(ll)'))
dis as text " "

* Schätzung des FIOM
quietly: `e(cmd2)' `e(depvar)' if e(sample)

quietly: estimates store fiom

dis as text "McFadden Pseudo-R2 (M_A vs. Fixed-Intercept-Only-Logit Model) = " ///
  as result %6.4f abs(1- (llma / `e(ll)'))
dis as text " "

drop llma

dis as text " "
dis as text "Likelihood-Ratio-chi2-Test (H0: All fixed effects = 0) "
lrtest riom ma

dis as text " "
dis as text "Likelihood-Ratio-chi2-Test (H0: All fixed & random effects = 0) "
lrtest fiom ma

exit
```

Appendix

Multilevel ordered logit model

1

Equations of the 2-level-ordered logit model

Level 2: Between-Context Regression

2a) Logistic Intercept-as-Outcome-Model:

$$\beta_{0j} = 0 + \gamma_{01} \times Z_{.j} + u_{0j}$$

2b) Logistic Slope-as-Outcome-Model:

$$\beta_{1j} = \gamma_{10} + \gamma_{11} \times Z_{.j} + u_{1j}$$

Level 1: Within-Context Regression

$$1) \ln \left[\frac{P(Y > k)}{P(Y \leq k)} \right] = \beta_{0j} + \beta_{1j} \times X_{ij} - \sum_{K=1}^{k-1} \delta_k \{+r_{ij}\}$$

Single equation notation: 2a) and 2b) in 1)

$$\ln \left[\frac{P(Y > k)}{P(Y \leq k)} \right] = \left(0 + \gamma_{01} \times Z_{.j} + u_{0j} \right) + \left(\gamma_{10} \times X_{ij} + \gamma_{11} \times X_{ij} \times Z_{.j} + u_{1j} \times X_{ij} \right) - \sum_{k=1}^{K-1} \delta_k \{+r_{ij}\}$$

Notation of Raudenbush&Bryk (2002):

γ : fixed-effect estimator

Z : exogenous level 2 variable

β : random-effect estimator

X : exogenous level 1 variable

u_{0j} : residuum random-intercept

u_{1j} : residuum random-slope

r_{ij} : residuum of within-context-logistic regression

δ_k : threshold for category k of Y

- Interpretation of the residua of the Between- Context-Regression

$$3a) u_{0j} = \beta_{0j} - [\gamma_{00} + \gamma_{01} \times Z_{.j}] = \beta_{0j} - \widehat{\beta}_{0j}$$

$$3b) u_{1j} = \beta_{1j} - [\gamma_{10} + \gamma_{11} \times Z_{.j}] = \beta_{1j} - \widehat{\beta}_{1j}$$

- Assumptions for the residua of the logistic 2-level logit model

Level 1:

1.1) r_{ij} is binomial distributed with an expected value of zero

and a variance $\sigma_{r_{ij}}^2 = \widehat{P}_{ij}(Y=1) \times (1 - \widehat{P}_{ij}(Y=1))$

1.2) Heteroscedasticity of r_{ij} in all contextual units j

- Implication for the level 1 residuum r_{ij}
 - ▶ Because of its own heteroscedasticity the variance $\sigma^2(r_{ij})$ can not be estimated. It is replaced by the variance of the logistic density function ($\pi^2 / 3$)

● Residua of level 2

2.1) u_{kj} is normal distributed with an expected value of zero and a covariance matrix T of the residua

$$E \begin{bmatrix} u_{0j} \\ u_{1j} \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad T = \begin{bmatrix} \tau_{00} & \tau_{01} \\ \tau_{10} & \tau_{11} \end{bmatrix} \quad \sigma_{u_{0j}}^2 = \tau_{00} \quad \sigma_{u_{1j}}^2 = \tau_{11}$$

$$\sigma_{u_{0j}, u_{1j}} = \tau_{10} = \tau_{01}$$

2.2) The residua of level 1 and level 2 are not correlated:

$$\sigma_{u_{0j}, r_{ij}} = \sigma_{u_{1j}, r_{ij}} = 0$$

Alternative in Stata: Information criteria

- Calculation of Akaike- (AIC) and Schwarz Bayesian-Information-Criteria (BIC)

$$AIC = -2 \times \log L_{M_A} + 2 \times k$$

$$BIC = -2 \times \log L_{M_A} + \log N \times k$$

↑
deviance

↑
complexity of the model

Range: $0 < AIC \leq +\infty$

$$0 < BIC \leq +\infty$$

Legend:

log: *Logarithmus naturalis*

k: *Number of estimated parameters*

N: *Sample size*

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