

# Influence Analysis with Panel Data using Stata

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# Motivation

- ▶ Short panel data sets (small  $N$  but  $N \gg T$ ) are common in many fields of Economics, e.g.
  - ▶ Macro-level panel data
  - ▶ Experimental panel data
- ▶ Observational data may contain “anomalous” observations (Rousseeuw and Van Zomeren, 1990; Silva, 2001)
  - ▶ Vertical outliers (VO), good leverage (GL) points, bad leverage (BL) points [▶ Example](#) [▶ DGP](#)
- ▶ Large influence on the Least Squares (LS) estimates  
⇒ Biased regression coefficients or standard errors (Donald and Maddala, 1993; Bramati and Croux, 2007; Verardi and Croux, 2009)

# Motivation

- ▶ **Diagnostic plots** (leverage-vs-residual plots)
  - ▶ for cross-sectional data: `lvr2plot/lvr2plot2`
  - ▶ Less handy for panel data
  
- ▶ **Measures of influence** (Cook (1979)'s distance)
  - ▶ for cross-sectional data: `predict c, cooks`
  - ▶ for panel data: `jackknife2, cooks(newvar)`  
`bpd(newvar): command`
  - ▶ These metrics may fail to flag multiple atypical cases (Atkinson and Mulira, 1993; Chatterjee and Hadi, 1988; Rousseeuw and Van Zomeren, 1990) unlike *pair-wise measures* (Lawrance, 1995)

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# In this presentation

- ▶ I present a method to
  1. Detect and identify the type of anomalous unit
  2. Show how these affect the LS estimates, and the influence of other units
- ▶ I follow a *unit-wise* approach (full history of a unit)
- ▶ I propose two commands in Stata
  - ▶ `xtlvr2plot` – Leverage-vs-residual plot for panel data
  - ▶ `xtinfluence` – Influence analysis with panel data

# Model and estimators

Static linear panel regression model with fixed effects

$$y_{it} = \mathbf{x}'_{it}\boldsymbol{\beta} + \alpha_i + u_{it}$$

Model after the *within-group* (WG) transformation

$$\tilde{y}_{it} = \tilde{\mathbf{x}}'_{it}\boldsymbol{\beta} + \tilde{u}_{it}$$

where  $\tilde{y}_{it} = y_{it} - T^{-1} \sum_t y_{it}$ , etc., and  $\boldsymbol{\beta}$  is a vector of parameters.

The WG Estimator

$$\hat{\boldsymbol{\beta}} = \left( \sum_{i=1}^N \sum_{t=1}^T \tilde{\mathbf{x}}_{it} \tilde{\mathbf{x}}'_{it} \right)^{-1} \sum_{i=1}^N \sum_{t=1}^T \tilde{\mathbf{x}}_{it} \tilde{y}_{it}$$

# Residual and Leverage

- ▶ The **average normalised residual** squared

$$\hat{u}_i^* = \frac{1}{T} \sum_{t=1}^T \left( \frac{\hat{u}_{it}}{\sqrt{\sum_i \hat{u}_{it}^2}} \right)^2$$

where  $\hat{u}_{it} = \tilde{y}_{it} - \tilde{\mathbf{x}}'_{it}\hat{\boldsymbol{\beta}}$  are LS Residuals.

Cut-off value:  $c_{\hat{u}_i^*} = \frac{2}{NT}$

- ▶ The **average individual leverage** of unit  $i$  at time  $t$  is

$$\bar{h}_i = \frac{1}{T} \sum_{t=1}^T h_{ii,tt}$$

where  $h_{ii,tt} = \tilde{\mathbf{x}}'_{it}(\tilde{\mathbf{X}}'\tilde{\mathbf{X}})^{-1}\tilde{\mathbf{x}}_{it}$ , and  $h_{ii,ts} = \tilde{\mathbf{x}}'_{it}(\tilde{\mathbf{X}}'\tilde{\mathbf{X}})^{-1}\tilde{\mathbf{x}}_{is}$  for  $t, s = 1, \dots, T$ .

Cut-off value:  $c_{\bar{h}_i} = \frac{2(K+1)}{NT}$



## xtlvr2plot: Syntax

xtlvr2plot – Leverage-versus-normalised residual squared plot for panel data.

```
xtlvr2plot depvvar [indepvar] [if] [in] [, options]
```

*options*

---

*graph\_opts*

graph options allowed for twoway scatter

### Generated variables

*\_lev*

average individual leverage

*\_normres2*

average individual residual squared

## xtlvr2plot: Example of code

```
** Use of the 'xtlvr2plot' command
xtset id t

xtlvr2plot y x,                               ///
    xlabel(id)                                ///
    xlabel(, format(%9.3fc))                  ///
    ylabel(, angle(h) format(%9.3fc))         ///
    title("Unit-wise Evaluation", size(medsmall)) ///
    saving("xtlvr2plot_example.gph", replace)
```

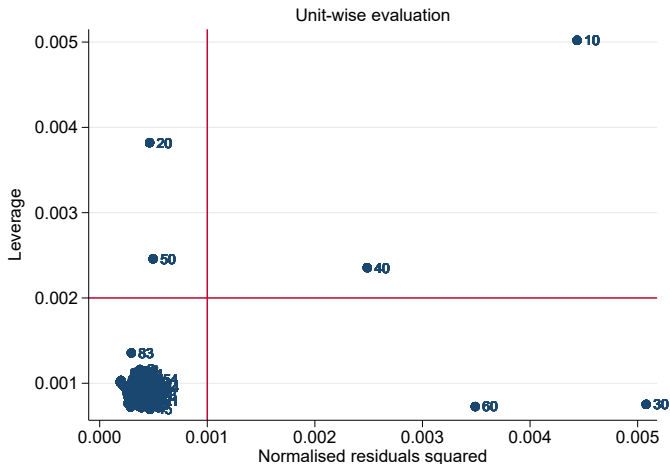
## xtlvr2plot: Summary Table

```
** Summary table w/detected anomalous units  
** generated by 'xtlvr2plot'
```

| Anomalous units     |       |
|---------------------|-------|
| x-cutoff =          | 0.001 |
| y-cutoff =          | 0.002 |
| Good leverage units |       |
| - Count :           | 2     |
| - List :            | 20 50 |
| Bad leverage units  |       |
| - Count :           | 2     |
| - List :            | 10 40 |
| Vertical outliers   |       |
| - Count :           | 2     |
| - List :            | 30 60 |

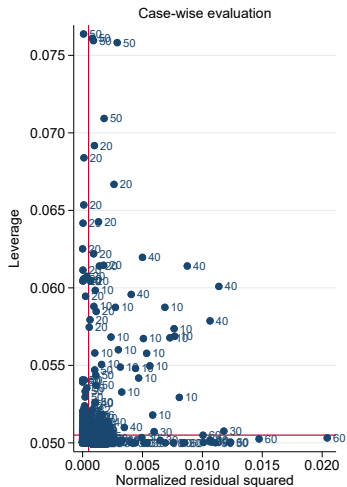
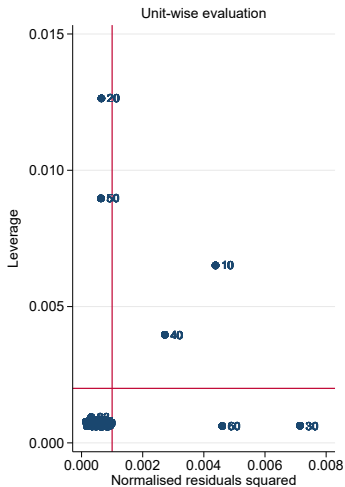
Note: Units 10 and 40 are set to be bad leverage units; units 20 and 50 good leverage units; units 30 and 60 vertical outliers.

# xtlvr2plot



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# xtlvr2plot vs lvr2plot



Note: Units 10 and 40 are set to be bad leverage units; units 20 and 50 good leverage units; units 30 and 60 vertical outliers.

# Influence analysis: Overview

- ▶ How anomalous units may affect the LS estimates
  1. Joint influence
  2. Joint effects
  3. Conditional influence
  4. Conditional effects

# Influence analysis: Joint influence

- ▶ For  $i \neq j$ ,

$$C_{ij}(\hat{\beta}) = (\hat{\beta} - \hat{\beta}_{(i,j)})' (\tilde{\mathbf{X}}' \tilde{\mathbf{X}}) (\hat{\beta} - \hat{\beta}_{(i,j)}) (s^2 K)^{-1}$$

where  $\hat{\beta}_{(i,j)}$  is WG estimator w/t units  $i$  and  $j$ ,  $s$  is RMSE,  $K$  is #covariates

- ▶ Influence exerted by a pair  $(i,j)$  on LS estimates *jointly*
  - ▶ Comparison of LS estimates *with* and *without* the pair
  - ▶  $C_{ij}(\hat{\beta}) = C_{ji}(\hat{\beta})$
- 
- ▶  $C_{ij}(\hat{\beta}) \sim F(\nu_1, \nu_2)$   
where  $\nu_1 = k + 1$  and  $\nu_2 = NT - N - (k + 1)$

# Influence analysis: Joint influence

- ▶ For  $i = j$ ,

$$C_{ii}(\hat{\beta}) = (\hat{\beta} - \hat{\beta}_{(i)})' (\tilde{\mathbf{X}}' \tilde{\mathbf{X}}) (\hat{\beta} - \hat{\beta}_{(i)}) (s^2 K)^{-1}$$

where  $\hat{\beta}_{(i)}$  is WG estimator w/t unit  $i$

- ▶  $i$ 's influence on LS estimates (as in Belotti and Peracchi (2020))
- ▶  $C_{ii}(\hat{\beta}) \sim F(\nu_1, \nu_2)$   
where  $\nu_1 = k + 1$  and  $\nu_2 = NT - N - (k + 1)$



# Influence analysis: Joint effects

- ▶ For  $i \neq j$ ,

$$K_{j|i} = \frac{C_{ij}(\hat{\beta})}{C_{ii}(\hat{\beta})}$$

- ▶ How much the pair is influential wrt  $i$
- ▶ For  $i = j$ ,  $K_{j|i} = 1$
- ▶ For large values of  $K_{j|i}$ 
  - ▶ The most influential unit ( $j$ ) *alters* the effect of the least ( $i$ )
  - ▶  $j$  either *enhances* or *reduces* the effect of  $i$  on the LS estimates  
⇒ based on the conditional effect

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# Influence analysis: Conditional influence

For  $i \neq j$ ,

$$C_{i(j)}(\hat{\beta}) = (\hat{\beta}_{(i,j)} - \hat{\beta}_{(j)})' \left( \sum_{\substack{i=1 \\ i \neq j}}^N \tilde{\mathbf{X}}_{i(j)}' \tilde{\mathbf{X}}_{i(j)} \right) (\hat{\beta}_{(i,j)} - \hat{\beta}_{(j)}) (s^2 K)^{-1}$$

- ▶ Influence exerted by  $i$  on LS estimates without  $j$  in the sample
- ▶ How the absence of  $j$  affects the influence  $i$  on LS estimates
- ▶  $C_{i(j)}(\hat{\beta}) = 0$  for  $i = j$
- ▶  $C_{i(j)}(\hat{\beta}) \neq C_{j(i)}(\hat{\beta})$
- ▶  $C_{i(j)}(\hat{\beta}) \sim F(\nu_1, \nu_2)$

# Influence analysis: Conditional effects

- ▶ For  $i \neq j$

$$M_{i(j)} = \frac{C_{i(j)}(\hat{\beta})}{C_{ii}(\hat{\beta})}$$

- ▶ How influence of  $i$  changes before and after the deletion of  $j$
- ▶ If  $M_{i(j)} \geq 1$ 
  - ▶ influence of  $i$  **increases** without  $j$  in the sample
  - ▶  $j$  *masks*  $i$
- ▶ If  $M_{i(j)} < 1$ 
  - ▶ influence of  $i$  **decreases** without  $j$  in the sample
  - ▶  $j$  *boosts*  $i$

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## xtinfluence: Syntax

`xtinfluence` – Influence analysis for panel data displaying the measures and effects of unit  $j$  against unit  $i$ . The size of the symbols is proportional to the magnitude of the calculated measures.

```
xtinfluence depuar [indepvar] [if] [in] [, options]
```

*options*

---

|                                    |  |
|------------------------------------|--|
| <u>figure</u> ( <i>graphtype</i> ) | display diagnostic plots like <i>graphtype</i> allows for the choice between scatter plot or heat plot; default is scatter |
| <i>graph_opts</i>                  | graph options allowed for scatter and heatmap  |
| <u>saving</u> ( <i>filename</i> )  | save .dta and .pdf file with the specified name and location   |

### Saved data sets

*filename\_adj\_mtx.dta* Automatically saves a data set with the influence measures and effects generated by the command

## xtinfluence: Example

**\*\*Use of the 'xtinfluence' command**

```
xtset id t
```

**\*\* Heat plot**

```
xtinfluence y x, figure(heat) ///  
    keylabels(all, interval) color(RdBu, reverse) ///  
    lev(30) statistic(max) ///  
    xlabel(5(10)100, angle(h) labsize(small)) ///  
    xmtick(##10) xlabel(##2, angle(h)) ///  
    ylabel(5(10)100, angle(h)) ///  
    ymtick(##10) ylabel(##2, angle(h)) ///  
    saving("xtinfluence_heat")
```

**\*\* Scatter plot**

```
xtinfluence y x, figure(scatter) ///  
    xlabel(5(10)100, angle(h) labsize(small)) ///  
    xmtick(##10) xlabel(##2, angle(h)) ///  
    ylabel(5(10)100, angle(h)) ///  
    ymtick(##10) ylabel(##2, angle(h)) ///  
    saving("xtinfluence_scatter")
```



# Influence analysis: Summary table

| Variable | Obs    | Mean     | Std. dev. | Min      | Max      |
|----------|--------|----------|-----------|----------|----------|
| C        | 10,000 | .3811386 | 2.200585  | 2.35e-11 | 33.58732 |
| K        | 10,000 | 16156.08 | 1242556   | 4.42e-08 | 1.23e+08 |
| cC       | 10,000 | .0038312 | .0353837  | 0        | .6169614 |
| M        | 9,900  | .0305928 | .6922132  | 4.39e-06 | 65.47916 |

---

## Influence analysis

---

v1 = k+1 = 2

v2 = NT-N-k-1 = 1898

c1 = 4/N = .04

c2 = F(v1,v2,.5) = 0.6934

---

Cii >= c1

- Count : 8

- List : 8 10 20 34 40 43 50 65

Cii >= c2

- Count : 2

- List : 10 40

i with K >= p99

- Count : 30

- List : 3 4 6 9 11 13 14 19 24 27 47 49 55 57 62 64 67 68 69 71 72 74 76 77 79 84 86 89 93 95

j with K >= p99

- Count :

- List :

i with M >= 1

- Count : 2

- List : 9 74

j with M >= 1

- Count : 2

- List : 10 40

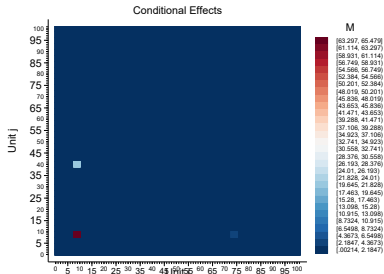
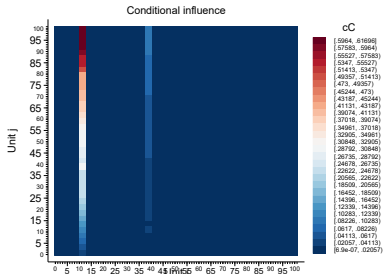
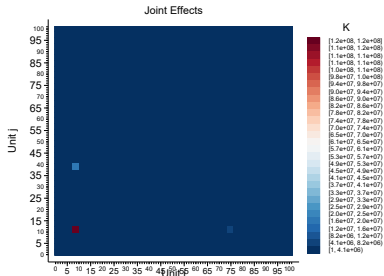
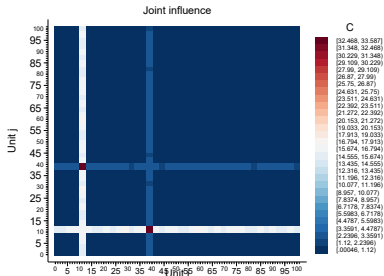
---

# filename\_adj\_mtx.dta

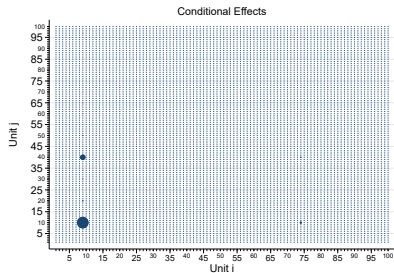
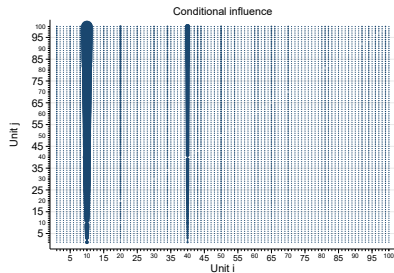
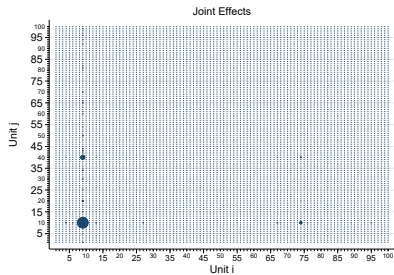
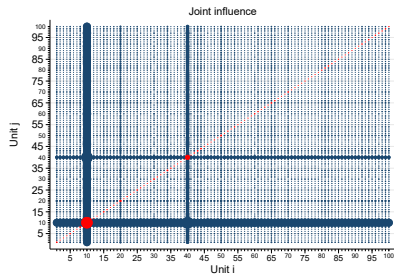
The saved data set resembles a directed and weighted adjacency list

|    | i | j  | C        | K        | cC       | M        |
|----|---|----|----------|----------|----------|----------|
| 1  | 1 | 1  | .0318985 | 1        | 0        | 0        |
| 2  | 1 | 2  | .0779802 | 2.444638 | 8.05e-06 | .0002523 |
| 3  | 1 | 3  | .0379366 | 1.189292 | .000065  | .0020391 |
| 4  | 1 | 4  | .0812006 | 2.545595 | .0000804 | .0025191 |
| 5  | 1 | 5  | .0384888 | 1.206603 | .0000916 | .0028703 |
| 6  | 1 | 6  | .0619195 | 1.941144 | .000091  | .0028528 |
| 7  | 1 | 7  | .0802803 | 2.516744 | .0001116 | .0034988 |
| 8  | 1 | 8  | .0322271 | 1.010302 | .0001236 | .003874  |
| 9  | 1 | 9  | .0102966 | .3227937 | .0001144 | .0035852 |
| 10 | 1 | 10 | 34.86443 | 1092.981 | .0001167 | .0036569 |
| 11 | 1 | 11 | .0380862 | 1.193983 | .0001264 | .0039615 |
| 12 | 1 | 12 | .0524164 | 1.643225 | .0001519 | .0047621 |
| 13 | 1 | 13 | .0510088 | 1.599099 | .0001667 | .005226  |
| 14 | 1 | 14 | .0550416 | 1.725525 | .0001834 | .0057488 |
| 15 | 1 | 15 | .0617752 | 1.936618 | .0001679 | .0052648 |
| 16 | 1 | 16 | .0591808 | 1.855285 | .000202  | .0063336 |
| 17 | 1 | 17 | .0512263 | 1.605917 | .0001969 | .0061739 |
| 18 | 1 | 18 | .067513  | 2.116496 | .0002049 | .006424  |
| 19 | 1 | 19 | .0904264 | 2.834818 | .000237  | .0074296 |
| 20 | 1 | 20 | 11.59427 | 363.474  | .0005592 | .0175295 |
| 21 | 1 | 21 | .0564583 | 1.769938 | .0002562 | .0080332 |
| 22 | 1 | 22 | .0020566 | .0644732 | .0002375 | .0074454 |
| 23 | 1 | 23 | .091529  | 2.869384 | .0002585 | .0081049 |
| 24 | 1 | 24 | .026083  | .8176892 | .0002669 | .0083674 |
| 25 | 1 | 25 | .0945991 | 2.965631 | .0003046 | .0095503 |

# Influence analysis: Heat plot



# Influence analysis: Scatter plot



# Conclusion

- ▶ The proposed *STATA* commands allow to
  1. Identify anomalous units and their type (unit-wise leverage-vs-residual plot)
  2. Investigate how anomalous units may affect the LS estimates (joint and conditional influence and effects)
  
- ▶ Once identified the type of anomaly in the sample
  1. Methods for measurement error if error in the data entry
  2. Robust estimation techniques if VO and BL (Bramati and Croux, 2007; Verardi and Croux, 2009; Aquaro and Čížek, 2013, 2014; Jiao, 2022)
  3. Jackknife-type standard errors if GL (MacKinnon and White, 1985; Davidson et al., 1993; MacKinnon, 2013; Belotti and Peracchi, 2020; Polselli, 2022)

Thank you for your attention!

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🐙 <https://github.com/POLSEAN/Influence-Analysis>

🐦 @AnnalivPolselli

# References I

- Aquaro, M. and Čížek, P. (2013). One-step robust estimation of fixed-effects panel data models. *Computational Statistics & Data Analysis*, 57(1):536–548.
- Aquaro, M. and Čížek, P. (2014). Robust estimation of dynamic fixed-effects panel data models. *Statistical Papers*, 55(1):169–186.
- Atkinson, A. and Mulira, H.-M. (1993). The stalactite plot for the detection of multivariate outliers. *Statistics and Computing*, 3(1):27–35.
- Belotti, F. and Peracchi, F. (2020). Fast leave-one-out methods for inference, model selection, and diagnostic checking. *The Stata Journal*, 20(4):785–804.
- Berka, M., Devereux, M. B., and Engel, C. (2018). Real exchange rates and sectoral productivity in the eurozone. *American Economic Review*, 108(6):1543–81.
- Bramati, M. C. and Croux, C. (2007). Robust estimators for the fixed effects panel data model. *The Econometrics Journal*, 10(3):521–540.
- Chatterjee, S. and Hadi, A. S. (1988). Impact of simultaneous omission of a variable and an observation on a linear regression equation. *Computational Statistics & Data Analysis*, 6(2):129–144.
- Cook, R. D. (1979). Influential observations in linear regression. *Journal of the American Statistical Association*, 74(365):169–174.
- Davidson, R., MacKinnon, J. G., et al. (1993). Estimation and inference in econometrics. *OUP Catalogue*.

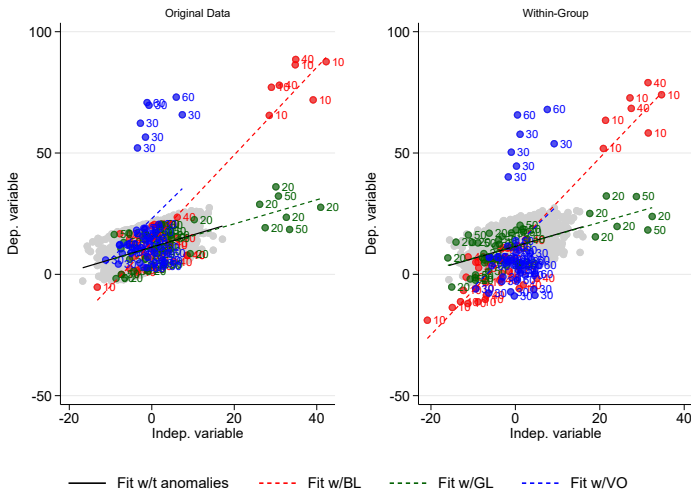
## References II

- Donald, S. G. and Maddala, G. (1993). 24 identifying outliers and influential observations in econometric models. In *Econometrics*, volume 11 of *Handbook of Statistics*, pages 663 – 701. Elsevier.
- Jiao, X. (2022). A simple robust procedure in instrumental variables regression. Unpublished, Last accessed: 07/02/2023.
- Lawrance, A. (1995). Deletion influence and masking in regression. *Journal of the Royal Statistical Society: Series B (Methodological)*, 57(1):181–189.
- MacKinnon, J. G. (2013). Thirty years of heteroskedasticity-robust inference. In *Recent advances and future directions in causality, prediction, and specification analysis*, pages 437–461. Springer.
- MacKinnon, J. G. and White, H. (1985). Some heteroskedasticity-consistent covariance matrix estimators with improved finite sample properties. *Journal of econometrics*, 29(3):305–325.
- PolSELLI, A. (2022). *Essays on Econometric Methods*. PhD thesis, University of Essex.
- Rousseeuw, P. J. and Van Zomeren, B. C. (1990). Unmasking multivariate outliers and leverage points. *Journal of the American Statistical Association*, 85(411):633–639.
- Silva, J. S. (2001). Influence diagnostics and estimation algorithms for powell's scl. *Journal of Business & Economic Statistics*, 19(1):55–62.
- Verardi, V. and Croux, C. (2009). Robust regression in stata. *The Stata Journal*, 9(3):439–453.



# Scatter Plot DGP

▶ Back



Note: Units 10 and 40 are bad leverage units; units 20 and 50 are good leverage units; units 30 and 60 are vertical outliers.

```
loc numobs 100
set obs 100
gen id = _n
expand 20

bys id: generate t = _n
bys id: gen z = rnormal(0,5)
**GL
bys id: replace z = z + rnormal(30,1) if id==20 & t<=5
bys id: replace z = z + rnormal(30,1) if id==50 & t<=2
**for BL
bys id: replace z = z + rnormal(30,1) if id==10 & t<=5
bys id: replace z = z + rnormal(30,1) if id==40 & t<=2
**line
bys id: gen a = runiform(0,20)
bys id: gen y = 1 + .5*z + a + runiform()
**BL
bys id: replace y = y + rnormal(50,1) if id==10 & t<=5
bys id: replace y = y + rnormal(50,1) if id==40 & t<=2
**V0
bys id: replace y = y + rnormal(50,1) if id==30 & t<=5
bys id: replace y = y + rnormal(50,1) if id==60 & t<=2
```

## Example: Berka et al. (2018)

- ▶ They study relationship between real exchange rate and sectoral productivity in the Eurozone
- ▶ Regression model:

$$RER_{it} = \beta TFP_{it} + \mathbf{x}'_{it}\boldsymbol{\gamma} + \alpha_i + u_{it}$$

$RER_{it}$ : real exchange rate in log

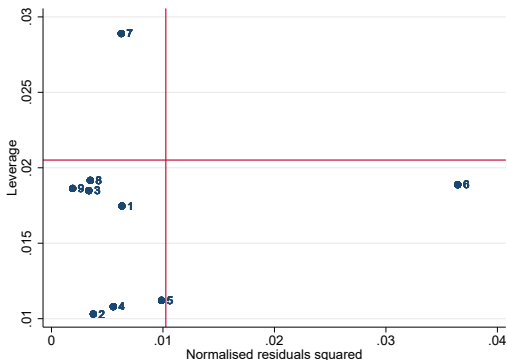
$TFP_{it}$ : total factor productivity in log

$\mathbf{x}_{it}$ : other controls

$\alpha_i$ : country fixed effects

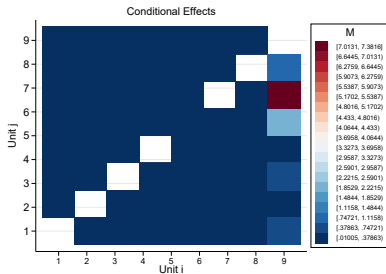
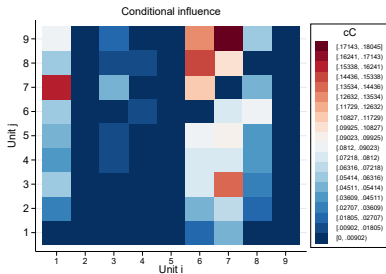
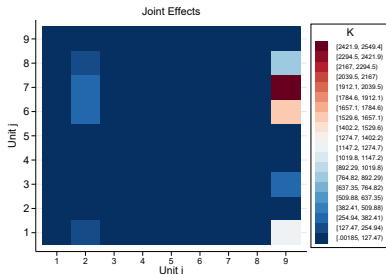
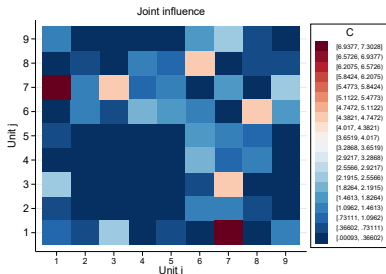
- ▶ Finding strong correlation between TFP and RER among high-income countries with floating nominal exchange rates
- ▶ Sample: 9 countries
- ▶ Time Period: 1995–2007
- ▶ Table 4, specification (2a)

## Example: Leverage-vs-residual plot ▶ Scatter



Note: 1-Austria, 2-Belgium, 3-Finland, 4-France, 5-Germany, 6-Ireland, 7-Italy, 8-Netherlands, 9-Spain.

# Example: Network-like plots ▶ Summary



# Example: Summary [▶ Back](#)

| Variable | Obs | Mean     | Std. dev. | Min      | Max      |
|----------|-----|----------|-----------|----------|----------|
| C        | 81  | 1.0233   | 1.472976  | .0009253 | 7.30281  |
| K        | 81  | 97.87085 | 368.2484  | .0018538 | 2549.404 |
| cC       | 81  | .032125  | .0439157  | 0        | .1804506 |
| M        | 72  | .2303033 | .8915019  | .0046645 | 7.381636 |

---

## Influence analysis

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$v1 = k+1 = 2$

$v2 = NT-N-k-1 = 184$

$c1 = 4/N = .4444444444444444$

$c2 = F(v1,v2,.5) = 0.6958$

---

Cii >= c1

- Count : 4

- List : 1 6 7 8

Cii >= c2

- Count : 3

- List : 1 6 7

i with K >= p99

- Count : 1

- List : 9

j with K >= p99

- Count :

- List :

i with M >= 1

- Count : 1

- List : 9

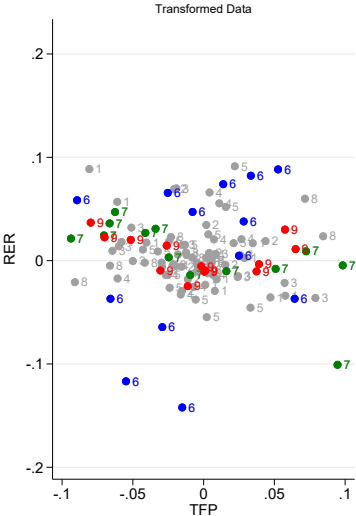
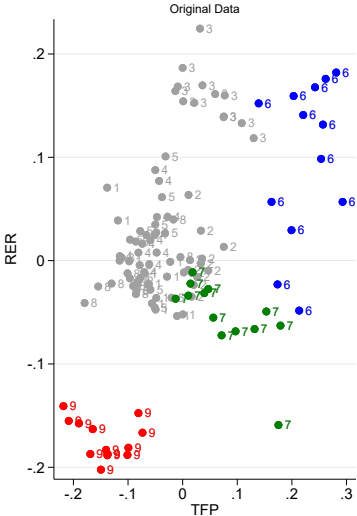
j with M >= 1

- Count : 2

- List : 6 7

---

# Example: Scatter [▶ Back](#)



Note: 1-Austria, 2-Belgium, 3-Finland, 4-France, 5-Germany, 6-Ireland, 7-Italy, 8-Netherlands, 9-Spain.

# Summary of Method

1. Identify anomalous units and their type with `xtlvr2plot`
2. Conduct the influence analysis with `xtinfluence`
  - 2.1 **Joint Influence Plot**
    - Identify units with high individual influence (main diagonal)
    - Identify pairs with high joint influence (off-diagonal)
    - Highly influential units swamp all other units
  - 2.2 **Joint Effect Plot**
    - Identify pairs with largest effect
    - $j$  swamps the effect of  $i$
    - $j$  must be detected in (1) and (2.1)
  - 2.3 **Conditional Influence Plot**
    - Identify influential  $i$  conditional to removing  $j$
    - Check if same units as (1) and (2.1)
  - 2.4 **Conditional Effect Plot**
    - Identify pairs with largest effect
    - $j$  masks the effect of  $i$
    - Compare identified pairs with (2.2)
3. Units detected in (1), (2.1) and (2.3) are anomalous; (2.2) and (2.4) explain how they affect the influence of other units and, hence, LS estimates