

nwxtregress: Network regressions in Stata

Italian Stata Conference 2022

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May 19, 2022

Economic and social agents are not independent

- Empirical analysis in social sciences (nearly) invariably relies on the assumption of cross-sectional independence.
 - ▶ E.g., the Gauss-Markov theorem assumes independence of disturbances.
- Most real-world applications involve interactions between units of observation.
 - ▶ E.g., companies buy and sell from one another, individuals share information with family and friends, etc.

Many applications of interactions are best represented using networks

- The inherent interactions between social entities has spawned a wide literature on social networks (Borgatti et al., 2009).
- Networks (or graphs) parsimoniously capture many economic settings:
 - ▶ Economic entities are the vertices or nodes of the network.
 - ▶ E.g., the US economy is made up of a collection of industries.
 - ▶ Relationships between entities are the edges, links, or ties of the network.
 - ▶ E.g., the dollar value of transactions between two industries.
- A key question remains: how do we analyze outcomes in a regression framework in the context of networks?
 - ▶ cross-sectional independence cannot be assumed!

Literature

- Network description and mathematical models (Newman, 2010; Jackson, 2010)
- Spatial econometrics
 - ▶ Initially used in regional science to model neighbouring regions
 - ▶ Empirical models and estimation techniques with a priori knowledge of relationship between units (LeSage and Pace, 2009; Kelejian and Piras, 2017)
- Using spatial models to identify peer effects in social networks (Bramoullé et al., 2009; Grieser et al., 2021)

Interactions pose identification challenges

- Consider a traditional panel model with 2 units:

$$y_{1t} = X_{1t}\beta + \epsilon_{1t}$$

$$y_{2t} = X_{2t}\beta + \epsilon_{2t}$$

- The independence assumption implies: $E[\epsilon_1\epsilon_2] = E[\epsilon_1]E[\epsilon_2]$
- This rules out the possibility that units 1 and 2 interact.
- Thus, for many applications, a more appropriate model is:

$$y_{1t} = \rho y_{2t} + X_{1t}\beta + \epsilon_{1t}$$

$$y_{2t} = \rho y_{1t} + X_{2t}\beta + \epsilon_{2t}$$

- This clearly violates independence (endogenous outcome y on RHS)
- Simultaneity invalidates inferences based on direct estimation

A parsimonious model of interactions

- Generalizing the panel model gives:

$$y_{it} = \sum_{j \neq i} \rho_{ij} y_{jt} + X_{it} \beta + \epsilon_{it}$$

- Considering all interactions ($\approx N^2$) is impractical
- Ord (1975) proposed the parsimonious parameterization:

$$y_{it} = \rho \sum_{j \neq i} w_{ij,t} y_{jt} + X_{it} \beta + \epsilon_{it}$$

- w_{ij} represents a priori link between i and j

We must invert the model to solve it

- It is more convenient to use matrix notation
- If we stack all elements in conforming vectors/matrices:

$$y = \rho W y + X\beta + \epsilon$$

- This is known as the Spatial Autoregressive (SAR) model
- Estimating the model “as is” poses various challenges (Manski, 1993; Angrist, 2014)
- Solving for a reduced-form data generating process is more useful:

$$y = (I - \rho W)^{-1}(X\beta + \epsilon)$$

- Note y s only appear on LHS, but model is nonlinear in parameters

The Model implies geometrically-decaying propagation

- Given mathematical restrictions on ρ and W :

$$(I - \rho W)^{-1} = I + \rho W + \rho^2 W^2 + \dots$$

- Interpret outcome as geometric sum of:
 - ▶ Own effect (I term)
 - ▶ Immediate peers' effect (W term)
 - ▶ Peers of peers effect (W^2 term)
 - ▶ etc

Partial derivatives are no longer β s

- In traditional model:

$$\frac{\partial y_i}{\partial x_i} = \beta, \text{ and } \frac{\partial y_i}{\partial x_j} = 0, i \neq j$$

- In the model with interactions:

$$\frac{\partial y_i}{\partial x_j} = (I - \rho W)_{ij}^{-1} \beta, \forall i, j$$

- Listing all partial derivatives is impractical.
- LeSage and Pace (2009) propose summarizing partial derivative estimates into direct and indirect effect averages:
 - ▶ Direct: $\frac{1}{N} \sum_i \frac{\partial y_i}{\partial x_i}$
 - ▶ Indirect: $\frac{1}{N} \sum_i \sum_{j \neq i} \frac{\partial y_i}{\partial x_j}$

The SDM adds contextual effects to SAR

- The Spatial Durbin Model (SDM) is given by:

$$y = \rho Wy + X\beta + WX\theta + \epsilon$$

- The values in WX represent the covariates of peers
- The effect of these covariates is often referred to a contextual effect
- These values are assumed exogenous and do not materially change the estimation

A short primer on estimation

- Focusing on one cross-section (for notational convenience), the likelihood function of the model is:

$$f(Y, X; \rho, \beta, \sigma^2) = |I_N - \rho W| (2\pi\sigma^2)^{-N/2} \exp\left(-\frac{e'e}{2\sigma^2}\right)$$
$$e = (I - \rho W)Y - X\beta$$

- If ρ is known (say ρ_0), then β (and σ^2) can be integrated out in a maximum likelihood estimation (MLE)
- The problem becomes an optimization w.r.t. ρ only
- The estimation proceeds with an MCMC sampler using the above likelihood

nwxtregress¹

Syntax

Spatial Autocorrelation Model (SAR)

```
nwxtregress depvar indepvars [if] , dvarlag(W1[,options1])  
[ mcmc_options nosparse ]
```

Spatial Durbin Model (SDM)

```
nwxtregress depvar indepvars [if] , dvarlag(W1[,options1])  
ivarlag(W2[,options1]) [ mcmc_options nosparse ]
```

- W1 and W2 define spatial weight matrices, default is Sp object.
- Note: nwxtregress allows for unbalanced panels and time varying W1 and W2 (unlike spxtreg)

¹This command is work in progress. Options, functions and results might change.

nwxtregress

Spatial Weight Options

```
nwxtregress depvar indepvars [if] , dvarlag(W1[,options1]) ]  
[  ivarlag(W2[,options1]) mcmc_options nosparse ]
```

- options1 controls the spatial weight matrices:

- ▶ mata declares weight matrix is mata matrix. [Details](#)
- ▶ sparse if weight matrix is sparse. [Details](#)
- ▶ timesparse weight matrix is sparse and varying over time. [Details](#)
- ▶ id(string) vector of IDs if W is a non sparse mata matrix.

nwxtregress

Further Options

```
nwxtregress depvar indepvars [if] , dvarlag(W1[options1]) ]  
[ ivarlag(W2[options1]) mcmc_options nosparse ]
```

- `nospars`e do not convert weight matrix internally to a sparse matrix.
- `mcmc_options` control the Markov Chain Monte Carlo:
 - ▶ `draws`(integer 2000) number of griddy gibbs draws.
 - ▶ `gridlength`(integer 1000) grid length
 - ▶ `nomit`(integer 500) number of omitted draws
 - ▶ `barrypace`(numlist) settings for BarryPace Trick, iterations, maxorder default: 50 100
 - ▶ `usebp` use BarryPace trick instead of LUD for inverse of $I - \rho W$.
 - ▶ `seed`(#) sets the seed.

Example: BEA I/O Tabela I

Data

- We collect USE/MAKE table data from the BEA's website
- These data represent the goods that were used (USE) and made (MAKE) by each industry in the US
- To construct links between industries, we convert into flows between industries
- Loaded data as S_p matrix using `spmatrix fromdata W = sam* , replace`, but only for year 1998.
- We also collect key variables about each industry: capital consumption, compensation, and net surplus.

Example: BEA I/O Tabela II

Data

- We are estimating:
 - ▶ SAR:

$$\begin{aligned} cap_cons &= \beta_0 + \rho W_1 cap_cons \\ &+ \beta_1 compensation + \beta_3 net_surplus + \epsilon \end{aligned}$$

- ▶ SDM:

$$\begin{aligned} cap_cons &= \beta_0 + \rho W_1 cap_cons + \gamma_1 W_2 compensation \\ &+ \beta_1 compensation + \beta_3 net_surplus + \epsilon \end{aligned}$$

SAR

Time constant spatial weights

```
. nwxtregress cap_cons compensation net_surplus , ///
> dvarlag(W) seed(1234)
```

Griddy Gibbs (2000)

```
-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|
      10      20      30      40      50      %
.....
..... 50
..... 100
```

```
Spatial SAR                Number of obs    =    1358
Panel Variable (i): ID      Number of groups =     62
Time Variable (t): Year     Obs. of group:   22
                               min =    19
                               avg =    22
                               max =    22
                               R-squared =    0.73
                               Adj. R-squared =    0.73
```

	cap_cons	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	compensation	-1.303834	.0230025	-56.68	0.000	-1.386128	-1.232569
	net_surplus	-1.187692	.0236269	-50.27	0.000	-1.268585	-1.100094
W	cap_cons	.1119915	.0272865	4.10	0.000	.007	.199
	\sigma_u	.2659483	.0103939			.2328073	.3016006

Example

Time varying spatial weight

- Saved network data in *timesparse* format in mata as *W*.
- The first column identifies the year, second and third the IDs and the last one the value of the weight.
- Non standardized timesparse *W*:

```
. mata W[1..10,.]
```

	1	2	3	4
1	1997	1	1	120.445105
2	1997	1	2	2646.806067
4	1997	1	4	1594.653373
5	1997	1	5	93.56892452
9	1997	1	9	444.9500985
10	1997	1	10	1884.318874

SAR

```
. nwxregress cap_cons compensation net_surplus , ///
> dvarlag(W,mata timesparse) seed(1234)
```

```
Griddy Gibbs (2000)
```

```
———|——— 10 |——— 20 |——— 30 |——— 40 |——— 50 %
..... 50
..... 100
```

```
Spatial SAR                      Number of obs      =       1358
Panel Variable (i): ID             Number of groups   =        62
Time Variable (t): Year            Obs. of group:     =        22
                                   min =              19
                                   avg =              22
                                   max =              22
                                   R-squared          =        0.73
                                   Adj. R-squared     =        0.73
```

	cap_cons	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	compensation	-1.310997	.0226358	-57.92	0.000	-1.391276	-1.24096
	net_surplus	-1.195375	.0232523	-51.41	0.000	-1.274683	-1.109078
W	cap_cons	.094567	.025236	3.75	0.000	-.003	.175
	\sigma_u	.2665574	.0104167			.2333413	.3022325

SAR

Direct Indirect Effects

```
. estat impact
```

Average Impacts

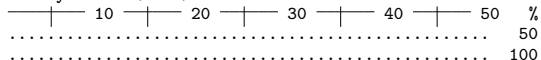
Number of obs = 1358

cap_cons	dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]	
direct						
compensation	-1.316631	.0227273	-57.93	0.000	-1.393996	-1.245314
net_surplus	-1.200513	.0233336	-51.45	0.000	-1.279492	-1.111448
indirect						
compensation	-.118322	.0315852	-3.75	0.000	-.2292999	.0038375
net_surplus	-.1078809	.0287743	-3.75	0.000	-.2066906	.0034616
total						
compensation	-1.434953	.0406058	-35.34	0.000	-1.612898	-1.297665
net_surplus	-1.308393	.0386178	-33.88	0.000	-1.453865	-1.161205

SDM

```
. nwxregress cap_cons compensation net_surplus , ///  
> dvarlag(W,mata timesparse) ///  
> ivarlag(W: compensation,mata timesparse ) seed(1234)
```

Griddy Gibbs (2000)



```
Spatial SDM                      Number of obs      =       1358  
Panel Variable (i): ID           Number of groups   =        62  
Time Variable (t): Year         Obs. of group:    =        22  
                                min =             19  
                                avg =             22  
                                max =             22  
                                R-squared          =       0.73  
                                Adj. R-squared     =       0.73
```

	cap_cons	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	compensation	-1.311007	.023423	-55.97	0.000	-1.399435	-1.235371
	net_surplus	-1.195046	.0239872	-49.82	0.000	-1.267235	-1.108135
W	cap_cons	.1033565	.0270195	3.83	0.000	.013	.19
	compensation	.0177968	.0267131	0.67	0.505	-.0679343	.107463
	\sigma_u	.2669513	.0101299			.2355723	.3033771

SDM

Direct Indirect Effects

```
. estat impact
```

Average Impacts

Number of obs = 1358

cap_cons	dy/dx	Std. Err.	z	P> z	[95% Conf. Interval]	
direct						
compensation	-1.311734	.0234539	-55.93	0.000	-1.40014	-1.23592
net_surplus	-1.195803	.0240079	-49.81	0.000	-1.26789	-1.108757
indirect						
compensation	-.1318438	.0523719	-2.52	0.012	-.3261685	.0389637
net_surplus	-.1382094	.0400208	-3.45	0.001	-.2875287	-.0158213
total						
compensation	-1.443578	.059526	-24.25	0.000	-1.656334	-1.250907
net_surplus	-1.334012	.0484298	-27.55	0.000	-1.526842	-1.179319

Future steps

- Most social networks contain very few edges relative to possible edges
 - ▶ Such networks are best represented via sparse matrices
 - ▶ Current support for sparse matrices in Mata is limited
 - ▶ Future improvements will rely on additional sparse matrix functions
- Streamlined support for fixed effects/intercept
- Implementation of convex combination of multiple networks (Debarsy and LeSage, 2020)
- Available on GitHub (<https://janditzen.github.io/nwxtregress/>) or directly in Stata:

```
net install nwxtregress ,  
from(https://janditzen.github.io/nwxtregress/)
```
- Please, help us by providing feedback

References I

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Newman, M. 2010. Networks: An Introduction. Oxford University Press.

Weight Matrices back

Square

Square matrix format

- The spatial weights are a matrix with dimension $N_g \times N_g$. It is time constant. An Example for a 5×5 matrix is:

	1	2	3	4		
	+-----+					
1		0	.1	.2	0	
2		0	0	.1	.2	
3		.3	.1	0	0	
4		.2	0	.2	0	
	+-----+					

Weight Matrices [back](#)

Sparse format

- The sparse matrix format is a $v \times 3$ matrix, where v is the number of non-zero elements in the spatial weight matrix.
- The weight matrix is time constant. The first column indicates the destination, the second the origin of the flow. A sparse matrix of the matrix from above is:

Destination	Origin	Flow
1	2	0.1
1	3	0.2
2	3	0.1
2	4	0.2
3	1	0.3
3	2	0.1
4	1	0.2
4	3	0.2

Weight Matrices back

Time-Sparse format

- The time sparse format can handle time varying spatial weights.
- The first column indicates the time period, the remaining are the same as for the sparse matrix. For example, if there are two time periods and we have the matrix from above for the first and the square for the second period:

Time	Destination	Origin	Flow
1	1	2	0.1
1	1	3	0.2
1	2	3	0.1
1	2	4	0.2
1	3	1	0.3
1	3	2	0.1
1	4	1	0.2
1	4	3	0.2
	<i>(next time period)</i>		
2	1	2	0.1
2	1	3	0.4
2	2	3	0.1
2	2	4	0.4
2	3	1	0.9
2	3	2	0.1
2	4	1	0.4
2	4	3	0.4