

2nd. STATA Users Group Meeting Mexico

Discussion of user-written Stata programs

Selection bias correction based on the multinomial logit: an application to the Mexican labour market.

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Goal.-

Application of the two step method ado-file *-selmlog-* explained in a robust manner by Bourguignon et al. (2004) and formally published by Bourguignon, Fourier and Gurgand (2007) *-JES-*.

Technical problem:

OLS becomes inefficient. Determination of wages generally causes a high correlation between the non-observable characteristics affecting wages and those that simultaneously determine the sector in which the individuals are currently /located/ functioning (working).

This will cause to obtain not only biased, but also inconsistent coefficients.

Evidence and facts

Bivariate selection bias
Heckman 1979

Earnings equations
Mincer 1974
Technique

Multilogit
Correction
(It works for
nested
especifications
as well)

Lee 1983 u_i and (η_j) are correlated, *iid?* $\rightarrow (\sigma \rho_1)$
Not true for the joint distribution.

Dubin and McFadden 1984 : No assump. on cov. u_1 and $(\eta_j - \eta_1)$ & multicollinearity exists.

Schmertmann 1994 u_1 and $(\eta_j - \eta_1)$ equal sign & *iid*

Dahl 2002

↓
Very strong hypothesis
in empirical studies

Bourguignon, et al. (2007)

Allows *corr* between choices $y_1 = x_1\beta_1 + \mu(P_1, \dots, P_M) + w_1$
 w_1 is mean independent of the regresors
 $(\sigma_1\rho_1), \dots, (\sigma_1\rho_s)$

Huesca (2005) and Zheren (2008). Recent applications BFG: Mexico and China.

Evidence and facts

As opposite to the bivariate case, when the number of events exceeds two categories, previous techniques (Lee, 1983 and so forth) present restrictions on the structure of the error terms and, generally, an inappropriate application – since those methods have been elaborated with the requirement of using an univariate transformation order–.



A correction for multivariate cases was developed in Dubin and McFadden (DM, 1984); this technique could not evaluate a model strong enough to admit maximum likelihood estimators, with complete information for the case were the number of choices were greater than two.

DM (1984) provides a model where the J sector must be required to establish a $J-1$ selection terms.



Bourguignon, et al. (2007) consider the case where the underlying selection process follows a polychotomous normal model, allowing correlations between alternatives.

Technique and ado selmlog

Must be understood by the self selection of individuals in the information and self-handling of the data that identical individual exists when using samples defined with a nonrandom criterion

two step generalized methodology proposed by BFG for polytomous cases is used, allowing OLS implementation in the calculations.

$$y_s = x_s \beta'_s + u_s$$

Let's assume the information follows a Gumbel distribution $G(.)$ *iid* for sake of normality. The following model is considered with a categorical variable $S = 1, \dots, M$ choices based on utilities y_s^* for the individuals as follows:

$$y_s^* = z_s \gamma_s + \eta_s,$$

Where Z and η_s compose a vector of independent variables and the disturbance term which confirms the usual conditions.

The impact on the dependent variable is observed just for the case where the alternative S is chosen which happens when:

$$y_s^* > \max_{j \neq s} (y_j^*)$$
$$\varepsilon_s = \max_{j \neq s} (y_j^* - \eta_s); \varepsilon_s < 0$$

the vector η_s is *iid* and Gumbel distributed; thus, their respective cumulative and density functions are

$$G(\eta) = \exp(-e^{-\eta})$$

$$g(\eta) = \exp(-\eta - e^{-\eta})$$

(See McFadden, 1973). It is in this part of the model where the multimominal logit specification applies in the traditional way:

$$P(z_s \gamma_s > \varepsilon_s) = \frac{\exp(z_s \gamma_s)}{\sum_j \exp(z_j \gamma_j)}$$

$$y_1 = x_1 \beta_1 - \sigma_1 [\rho_1 m(P_1) + \sum_{s>1} \rho_s \frac{P_s}{P_s - 1} m(P_s)] + v_s$$

β_{i1}' stands for coefficients and x_{i1} as attributes of the individual. The residual term displays the usual normality statistical conditions.

$m(P_s)$ are the probabilities and $(\sigma_1 \rho_1), \dots, (\sigma_1 \rho_s)$ the coefficient terms for the polychotomous correction of selectivity bias; v_s is an **orthogonal error** parameter towards the rest of terms, having a mean expectation equal to zero. This last property is what allows using directly the OLS procedure in the estimation.

$$1. P(z_s \gamma_s > \varepsilon_s) = \frac{\exp(z_s \gamma_s)}{\sum_j \exp(z_j \gamma_j)} \quad \text{Logit}$$

$$2. \ln y_s = \beta_s' x_s + \varepsilon_s - \sigma_{\eta u} \rho_s' \quad \text{Replacing terms, using a vector of Rhos}$$

One problem that arises from this occupational selection process technique is related to the IIA as stated by Hausman and McFadden (1984). Bourguignon, *et al.* (2004; 2007) can provide fairly good correction for the outcome equation, even when the IIA hypothesis is violated in nested models.

- | | | | |
|--------|---------------|---|---------------------------|
| 1. exp | Setting | $\rho_{12} = \rho_{13} = \rho_{23} = 0$ | misspecifies param. dist. |
| 2. exp | Small corr | $\rho_{12} = 0.1, \rho_{13} = 0.1, \rho_{23} = -0.2$ | |
| 3. exp | Violation IIA | $\rho_{12} = -0.1, \rho_{13} = 0.45, \rho_{23} = -0.35$ | |

$$\sigma^2 = 32, \text{corr}(u_1, \eta_1) = 0.64, \text{corr}(u_1, \eta_2) = -0.24, \text{corr}(u_1, \eta_3) = 0.14.$$

$$\begin{bmatrix} \sigma^2 & \sigma_{1u}^2 & \sigma_{2u}^2 & \sigma_{3u}^2 \\ \sigma_{1u}^2 & 1 & \rho_{12} & \rho_{13} \\ \sigma_{2u}^2 & \rho_{12} & 1 & \rho_{23} \\ \sigma_{3u}^2 & \rho_{13} & \rho_{23} & 1 \end{bmatrix}$$

Ensuring orthogonality so that $V[h(\eta_1, \eta_2, \eta_3)] \approx 0.16$

Empirical case

Answer the following questions: Will the differences in earnings between the formal and informal sectors of the labor market in Mexico be statistically significant? Which are the socioeconomic and occupational factors that mostly affect earnings amongst sectors?.

Logit has a practical advantage over probit when the sum of the predicted values equal to the sum of empirically observed values (Butcher and Dinardo, 1998.)

ENOE: Encuesta Nacional de Ocupación y Empleo: 2009-III.

Males and females aging from 16 to 65

Occupations = (1 ,..., 4)

Multinomial Logit

- 1: Formal self-employed
- 2: Informal self-employed
- 3: Formal wage-earners
- 4: Informal wage-earners

features for empirical application

To download it:

```
net from http://www.pse.ens.fr/gurgand/
```

To avoid endogeneity from the sample selection process we select for the objective earnings equation a vector of family background (highly recommended!).

Lee (1983), Dubin-McFadden (1984) and Dahl (2002) can be computed with `selmlog` as well. See `help selmlog`:

```
options [lee dmf(#) dhl(# [all])
```

`dhl` options include the order of the polynomials on the selection probabilities. With this number alone, the correction term includes only the probability to be selected on the observed outcome. If this number is followed by `all`, probabilities are included in polynomial form, with interactions, up to the specified order.

Syntax

1. Compute the earnings distribution using `selmlog` command.

```
selmlog depvar varlist [ifexp][inrange],select(depvar_m=varlist_m)
      [lee dmf(#) dh1(# [all]) showmlogit wls
      bootstrap(number_of_replications[sample_size])
      mloptions(mlogit options) gen(variable generic name)]
```

2. Computing the empirical case (Weighted Least Squares `-wls-` to account for heteroskedasticity present in the model due to selectivity).

****Formal Self-employed:

```
selmlog logw1 anios_esc eda eda2 rama2 rama4 rama5 rama6 rama8 ///
if logw>0, select(logitp = eda hijos jefe ur conyugal) ///
dmf(2) wls bootstrap(100) mloptions(rrr level (95)) gen(rho_1)
```

****Informal Self-employed:

```
selmlog logw2 anios_esc eda eda2 rama2 rama4 rama5 rama6 rama8 ///
if logw>0, select(logitp = eda hijos jefe ur conyugal) ///
dmf(2) wls bootstrap(100) mloptions(rrr level (95)) gen(rho_2)
```

****Formal wage-earner:

```
selmlog logw3 anios_esc eda eda2 rama2 rama4 rama5 rama6 rama8 ///
if logw>0, select(logitp = eda hijos jefe ur conyugal) ///
dmf(2) wls bootstrap(100) mloptions(rrr level (95)) gen(rho_3)
```

****Informal wage-earner:

```
selmlog logw4 anios_esc eda eda2 rama2 rama4 rama5 rama6 rama8 ///
if logw>0, select(logitp = eda hijos jefe ur conyugal) ///
dmf(2) wls bootstrap(100) mloptions(rrr level (95)) gen(rho_4)
```

Multi-Logit

STATA (R)
 Statistical Software
 Copyright 2009 StataCorp LP
 4405 Lakeview Drive
 College Station, Texas 77845 USA
 800-STATA-PC http://www.stata.com
 979-896-1000 Stata@stata.com
 979-896-1000 (fax)

Single-user, perpetual license:
 Serial number: 4003042568
 Licensed to: Luis Rivera
 Centro de Investigacion en Alimentacion y Desarrollo

Notes:
 1. *(/no option or -set memory-)* 50.00 MB allocated to data
 2. *(/no option or -set memory-)* 5000 maximum variables
 3. New update available: type update all

running c:\ado\personal\profiles.do ...
 - use "C:\Users\LUISRIV\AppData\Local\Temp\ST0000000000.tmp"
 - set memory 100M

Current memory allocation

variable	current value	description	memory usage (MB * 1024K)
set maxvar	5000	max. variables allowed	1.417M
set memory	100M	max. data space	100.000M
set maxsize	400	max. 805 vars. in models	1.25M
			01.700M

- use "C:\congreso\FNCE 2009.dta"
 - end of do-file

Command

```
logit y1 y2 y3 y4
```

Viewer [R] [new] [C:\congreso\FNCE 2009.dta]

Multi-nomial logit regression
 Log likelihood = -11981.247
 Number of obs = 10685
 Likelihood chi2(21) = 1330.37
 Prob > chi2 = 0.0000
 Pseudo R2 = 0.0610

	Logit#	RRR	Std. Err.	Z	P> z	[95% Conf. Interval]
1	cca	.947802	.0161164	3.13	0.002	.916733 .9799218
	edu2	1.00109	.0002061	5.28	0.000	1.000685 1.001490
	hi_jos	2.324742	.1078182	11.82	0.000	2.103320 2.543813
	jeife	1.258188	.0860988	2.50	0.001	1.079008 1.467010
	ur	2.328507	.1732320	11.36	0.000	2.01268 2.6941
	conyuga1	1.817908	.1217216	8.00	0.000	1.610710 2.027121
	r3	1.370490	.1684802	4.26	0.000	1.278337 1.543803
2	cca	.9818878	.0255618	-0.66	0.510	.931727 1.031000
	edu2	1.00007	.0002050	2.27	0.023	1.000091 1.000123
	hi_jos	2.502068	.2588818	8.88	0.000	2.001801 3.001817
	jeife	1.431183	.1339900	3.30	0.001	1.136837 1.770389
	ur	1.251958	.1407508	5.17	0.000	1.015188 1.516807
	conyuga1	2.002881	.1932401	7.20	0.000	1.637783 2.419817
	r3	1.679807	.2581787	4.15	0.000	1.181167 2.009912
3		(base outcome)				
4	cca	.8324302	.0110230	12.33	0.000	.8111017 .8743183
	edu2	1.001788	.0001676	10.71	0.000	1.001066 1.002511
	hi_jos	2.730418	.2032133	13.43	0.000	2.303341 3.172633
	jeife	1.188086	.0776875	2.61	0.009	1.0012 1.408567
	ur	1.841070	.1170009	9.36	0.000	1.623201 2.087003
	conyuga1	.9307181	.0508125	-1.68	0.094	.818195 1.071958
	r3	1.244302	.1147427	2.37	0.018	1.038810 1.461034

**** Hausman tests of IIA assumption
 Ho: cccs(outcome j vs outcome k) are independent of other alternatives.

unfitcd	chi2	df	p>chi2	evidence
1	4.438	14	1.000	for Ho
2	0.688	10	1.000	for Ho
3	1022.363	13	0.000	against Ho
4	-58.709	10	1.000	for Ho

Multi-Logit

Viewer (#1) [view "H:\Congreso\BFG.log"]

view "H:\Congreso\BFG.log"

Advice Contents What's New News

```

Multinomial logistic regression
Log likelihood = -11981.247
Number of obs   = 10685
LR chi2(21)    = 1556.57
Prob > chi2    = 0.0000
Pseudo R2     = 0.0610
  
```

	logitp	RRR	Std. Err.	z	P> z	[95% Conf. Interval]	

1							
	eda	.947802	.0161164	-3.15	0.002	.916735	.9799218
	eda2	1.00109	.0002063	5.28	0.000	1.000685	1.001494
	hijos	2.524742	.1978182	11.82	0.000	2.165329	2.943813
	jefe	1.258384	.0984989	2.94	0.003	1.079409	1.467034
	ur	2.328567	.1732329	11.36	0.000	2.01263	2.6941
	conyugal	1.837908	.1237216	9.04	0.000	1.610734	2.097121
	r5	1.576496	.1684892	4.26	0.000	1.278557	1.943863

2							
	eda	.9836979	.0245639	-0.66	0.510	.9367127	1.03304
	eda2	1.00067	.0002956	2.27	0.023	1.000091	1.00125
	hijos	2.502469	.2584438	8.88	0.000	2.043903	3.063917
	jefe	1.431183	.1553966	3.30	0.001	1.156837	1.770589
	ur	1.753294	.1902546	5.17	0.000	1.417388	2.168807
	conyugal	2.002881	.1932461	7.20	0.000	1.657783	2.419817
	r5	1.829801	.2543787	4.35	0.000	1.393382	2.402912

3	(base outcome)						

4							
	eda	.8524362	.0110236	-12.35	0.000	.8311017	.8743183
	eda2	1.001794	.0001676	10.71	0.000	1.001466	1.002123
	hijos	2.739418	.2052133	13.45	0.000	2.365341	3.172655
	jefe	1.186096	.0776875	2.61	0.009	1.0432	1.348567
	ur	1.841976	.1176699	9.56	0.000	1.625201	2.087665
	conyugal	.9107181	.0508125	-1.68	0.094	.8163795	1.015958
	r5	1.244562	.1147427	2.37	0.018	1.038819	1.491054

```

. mlogtest, hausman
**** Hausman tests of IIA assumption
Ho: odds(Outcome-J vs Outcome-K) are independent of other alternatives.

Omitted |      chi2   df   P>chi2   evidence
-----|-----
1 |      -4.438   14   1.000   for Ho
2 |       0.699   14   1.000   for Ho
3 |    1922.563   15   0.000   against Ho
4 |     -58.749   14   1.000   for Ho
  
```

Selmlog command using BFG (Lee)

```
selectivity correction based on multinomial logit
Second step regression
Bootstrapped standard errors (100 replications)
```

logw1	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
logw1						
anios_esc	-.0015537	.0045909	-0.34	0.735	-.0105516	.0074442
eda	-.0070126	.0110103	-0.64	0.524	-.0285924	.0145672
eda2	.0000231	.0001264	0.18	0.855	-.0002245	.0002708
rama2	.2603791	.1421182	1.83	0.067	-.0181675	.5389256
rama4	-.0673728	.0629758	-1.07	0.285	-.1908031	.0560575
rama5	-.6115123	.281147	-2.18	0.030	-1.16255	-.0604743
rama6	.013802	.0369808	0.37	0.709	-.058679	.086283
rama8	-.222056	.2259293	-0.98	0.326	-.6648694	.2207573
_m1	.1546043	.1118948	1.38	0.167	-.0647056	.3739141
_cons	6.215879	.334825	18.56	0.000	5.559634	6.872124
Anciliary						
sigma2	.5954783	.1259639	4.73	0.000	.3485936	.8423631
rho	.2003496	.1206472	1.66	0.097	-.0361146	.4368138

```
.
end of do-file
```

Command



Selmlog command using (dmf(0)) Dubin-McFadden [1]

```

selectivity correction based on multinomial logit
1 Second step regression
1 Bootstrapped standard errors (100 replications)

```

logw1	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
logw1						
anios_esc	-.0017757	.0039499	-0.45	0.653	-.0095173	.005966
eda	-.0090108	.0114329	-0.79	0.431	-.0314188	.0133972
eda2	.0000463	.0001292	0.36	0.720	-.000207	.0002995
rama2	.2633093	.1395896	1.89	0.059	-.0102812	.5368998
rama4	-.0640174	.0540833	-1.18	0.237	-.1700188	.041984
rama5	-.6069665	.3037591	-2.00	0.046	-1.202323	-.0116096
rama6	.0148896	.0354548	0.42	0.675	-.0546005	.0843797
rama8	-.2273555	.2071633	-1.10	0.272	-.6333882	.1786772
_m2	.0868744	.5257237	0.17	0.869	-.9435251	1.117274
_m3	.0084358	.2283435	0.04	0.971	-.4391092	.4559808
_m4	.0013133	.2748971	0.00	0.996	-.5374752	.5401018
_cons	6.225136	.35089	17.74	0.000	5.537404	6.912868
Anciliary						
Sigma2	.5230322	.3951511	1.32	0.186	-.2514498	1.297514
rho2	.1540642	.6441352	0.24	0.811	-1.108418	1.416546
rho3	.0149602	.3021221	0.05	0.961	-.5771881	.6071086
rho4	.002329	.355908	0.01	0.995	-.6952379	.6998959

```

.
end of do-file
.

```

Command



Selmlog command using (dmf(1)) Dubin-McFadden [2]

-all correlation coefficients sum-up to zero-

selectivity correction based on multinomial logit
 Second step regression
 Bootstrapped standard errors (100 replications)

logw1	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
logw1						
anios_esc	-.001937	.0066486	-0.29	0.771	-.014968	.011094
eda	-.0047785	.0148917	-0.32	0.748	-.0339657	.0244087
eda2	-1.55e-06	.0001717	-0.01	0.993	-.0003381	.000335
rama2	.250801	.1822725	1.38	0.169	-.1064466	.6080486
rama4	-.0605817	.0633696	-0.96	0.339	-.1847838	.0636205
rama5	-.6005671	.2961673	-2.03	0.043	-1.181044	-.0200898
rama6	.0184748	.0405116	0.46	0.648	-.0609265	.0978762
rama8	-.2428746	.2314719	-1.05	0.294	-.6965512	.210802
_m1	-.0623649	.1937439	-0.32	0.748	-.442096	.3173662
_m2	.772115	2.668067	0.29	0.772	-4.4572	6.00143
_m3	.4767232	1.60533	0.30	0.766	-2.669665	3.623111
_m4	.4728991	1.584504	0.30	0.765	-2.632671	3.578469
_cons	6.750295	2.153209	3.13	0.002	2.530084	10.97051
Anciliary						
sigma2	1.25447	3.443743	0.36	0.716	-5.495144	8.004083
rho1	-.0714142	.1494277	-0.48	0.633	-.3642872	.2214588
rho2	.8841504	1.330742	0.66	0.506	-1.724055	3.492356
rho3	.5458967	.7634907	0.72	0.475	-.9505175	2.042311
rho4	.5415177	.7639026	0.71	0.478	-.9557039	2.038739

end of do-file

Command

Selmlog command using BFG (dmf(2))

selectivity correction based on multinomial logit
 Second step regression
 Bootstrapped standard errors (100 replications)

logw1	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
logw1						
anios_esc	.0491248	.0048183	10.20	0.001	.0106928	.0819471
rama2	.2426938	.0806902	3.01	0.029	.1811145	.2968432
rama4	-.0559203	.0610801	-0.92	0.361	-.1756351	.0637945
rama5	-.6100826	.3005651	-2.03	0.042	-1.199179	-.0209857
rama6	.0225234	.0045593	4.94	0.022	-.0668373	.0311884
rama8	-.2639496	.2446007	-1.08	0.081	-.7433581	.0154593
r1	.1154463	.0096481	11.97	0.000	.0736529	.1545455
r2	.0558856	.0088745	6.30	0.019	.0280532	.0824401
r3	.0481392	.0892543	0.54	0.596	-.1267961	.2230745
r4	.1249463	.0517734	2.41	0.039	.0949263	.1848189
r6	.1829422	.0986251	1.85	0.064	.1035946	.2262438
_m1	-.0545378	.3358121	-0.16	0.871	-.1271664	.6036427
_m2	1.162403	1.005957	1.16	0.068	.6983051	2.023111
_m3	.3831932	.0771588	4.97	0.032	.2909283	.4954781
_m4	.2251297	.0696599	3.23	0.047	.1901813	.2904432
_cons	6.162259	.5968858	10.32	0.000	4.9923841	7.332133

>

Anciliary	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
Anciliary						
sigma2	.8620013	.4056217	2.13	0.054	-1.892966	3.61696
rho1	-.0587404	.0250426	-2.35	0.051	-0.549567	.432086
rho2	1.251995	.0948975	13.19	0.000	1.227962	1.28111
rho3	.4127279	.0538643	7.66	0.004	.3993362	.468449
rho4	.2424817	.0531211	4.56	0.031	.2198673	.263537

```
. ****Cuenta propia informal:
. selmlog logw2 anios_esc rama2 rama4 rama5 rama6 rama8 r1 r2 r3 r4 r6 ///
> if logw>0, select(logitp = eda eda2 hijos jefe ur conyugal r5) ///
> dmf(2) wls bootstrap(100) mloptions(rrr level (95)) gen(rho_2)
```

selectivity correction based on multinomial logit
 Second step regression
 Bootstrapped standard errors (100 replications)

logw2	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
logw2						
anios_esc	.0186409	.0087613	2.13	0.033	.0014691	.0358127

Conclusions :

`selmlog` command is a useful tool to correct selection bias in polytomous cases (From Lee to BFG).

The empirical application confirms for the Mexican case, that choices are selected in a non-randomly process: Individuals decide where to work!

An advantage is not to depend on the IIA-Hausman-Mc Fadden's test for nested models.

Our suggestion is not to specify models with a great number of covariates when computing the ado.

In earnings equations use familiar background as variables for selection.

The inference with a great number of reps is time consuming, 100 reps is recommended.

References

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