

Recentered Influence Functions (RIF) in Stata

RIF-regression and RIF-decomposition

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- Interested in the commands. Download it from ssc: `ssc install rif`
- Latest Files: <https://bit.ly/2NFM3cH>

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Introductions: Why ?

- Once upon a time (2011), I was young(er), and came across a paper: Firpo, Fortin and Lemieux (2009): Unconditional Quantile Regressions (UQR).
- The premise was simple: A regression framework analysis to explore factors behind changes across the unconditional distributions (quantiles).
- Similar (Conditional) Quantile regression, but not quite the same.
- As many people. Sat down, read the paper and its companions many times. After understanding what it did, and apply it for my dissertation. (-rifreg-)

Introduction: Why?

- Few years later(2017), couple of papers with the method, decided to teach it in my econometrics class. There was a problem.
- Implementations of UQR in Stata were limited: `-rifreg-`, `-xtrifreg-`, `-rifireg-`. There was no "easy" applications for decompositions.
- I had programs that were too crude and clunky. Hard to share with students.
- So what to do: if the solution does not exist yet. Solve it yourself!

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How to compare distributional statistics?

- When comparing distributional statistics, one requires one of the following items:
 - Collection of data: $Y = [y_1, y_2, y_3, \dots, y_N]$
 - The Cumulative distribution function $F(Y)$ or F_Y
 - The probability density function $f(Y)$ or f_Y
- Once any one of these three pieces is obtained, any distributional statistic ($v()$) can be easily estimated. And differences across two groups can be obtained straight forward.

$$\Delta v = v(G_Y) - v(F_Y)$$

- Where Δv is the change in v when $F_Y \rightarrow G_Y$

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- Influence Functions (IF) can be thought as a generalization of the above experiment.
- It represents the re-scaled effect that a change in the distribution from $F_Y \rightarrow G_Y$ has on statistic v , when the change is infinitesimally small:

$$G_Y^{y_i} = (1 - \varepsilon)F_Y + \varepsilon 1_{y_i}$$

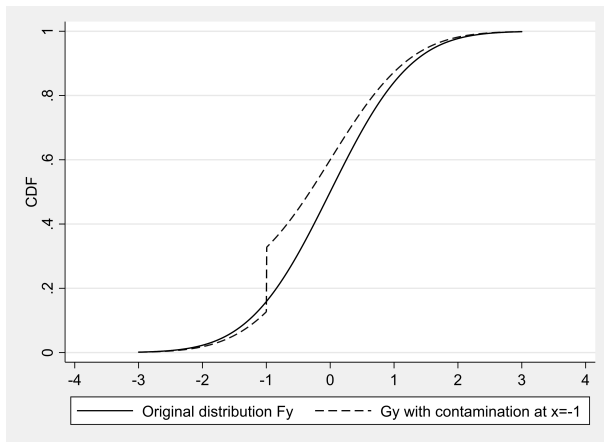
$$IF(y_i, v(F_Y)) = \lim_{\varepsilon \rightarrow 0} \frac{v(G_Y^{y_i}) - v(F_Y)}{\varepsilon}$$

- And, as introduced by FFL(2009)

$$RIF(y_i, v(F_Y)) = v(F_Y) + IF(y_i, v(F_Y))$$

- The contribution of y_i to the statistic $v()$

Visual Example of the change in F



- RIF has the following characteristics:

$$RIF(y_i, v(F_Y)) = v(F_Y) + IF(y_i, v(F_Y))$$

$$E(RIF(y_i, v(F_Y))) = v(F_Y)$$

$$E(IF(y_i, v(F_Y))) = 0$$

$$Var(v(F_Y)) = E(IF(y_i, v(F_Y))^2)$$

Why are they useful?

- Visual tool to inspect data, analyze statistics robustness to outliers (Cowel and Flatchaire, 2007)
- Simple estimation of standard errors of distributional statistic (Deville, 1999)
- Analysis of unconditional partial effects on distributional statistics based on regression and decomposition analysis (FFL, 2009,2018)

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How are RIF's Estimated?

- The estimation of RIFs varies in complexity depending on the statistic of interest:

Mean:

$$RIF(y_i, \mu_Y) = y_i$$

Variance:

$$RIF(y_i, \sigma_Y^2) = (y_i - \mu_Y)^2$$

Quantile:

$$RIF(y_i, q_Y(p)) = q_Y(p) + \frac{p - 1(y \leq q_Y(p))}{f_Y(q_Y(p))}$$

- But complexity increases for other statistics.
- In Rios-Avila (2019) I provide a collection of RIFs for a large set of distribution statistics. They include the statistics from FFL(2018), Firpo and Pinto (2016), Chung and Vankerm (2018), Cowell and Flachaire (2007), Essama-Nssah and Lambert (2012) and Heckley et al (2016).

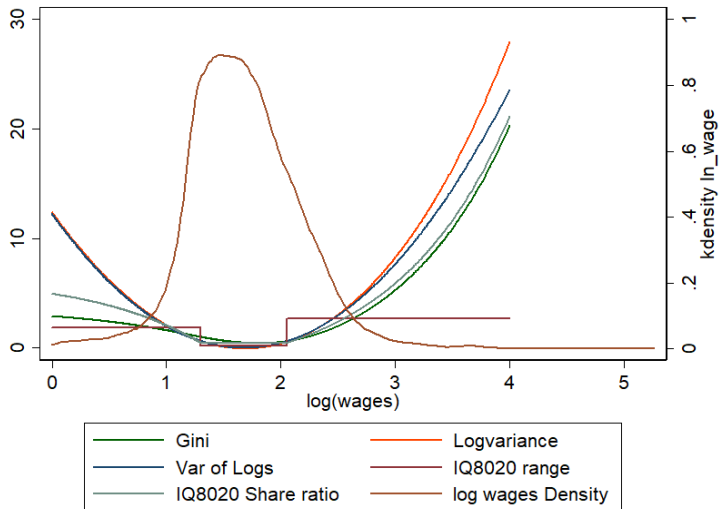
Using `_grifvar()`

- `_grifvar()` is an addon for `egen()`, that can be used to estimate all RIF's detailed in Rios-Avila(2019). It can be installed using (`ssc install rif`)
- The syntax is:
`egen new=rifvar(oldvar) [if/in], [by() weight()
rifoptions]`
rifoptions: Mean, variance, Coefficient of variation,
standard deviation, quantile, Interquantile range,
interquantile ratio, Gini, etc
For further detail `-help rifvar-`

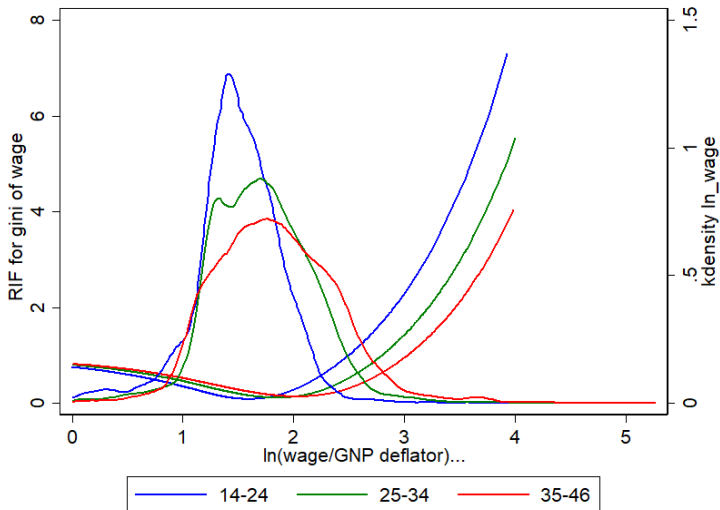
Example using `_rifvar`

```
webuse nlswork, clear
gen wage=exp(ln_wage)
egen rif_gini=rifvar(wage), gini
egen rif_log=rifvar(wage), logvar
egen rif_varlog=rifvar(ln_wage), var
egen rif_iqr=rifvar(ln_wage), iqr(20 80)
egen rif_iqsr=rifvar(wage), iqsr(20 80)
recode age (14/24=1 "14-24") (25/34=2 "25-34") (35/46=3 "35-46"),
gen(age_g)
egen rif_gini_age=rifvar(wage), gini by(age_g)
```

Example using `_grifvar`



Example using `_grifvar`



Example using _grifvar

Bootstrap with INEQDECO vs Mean RIF

	Observed Coef.	Bootstrap Std. Err.	Mean	Std. Err.
e_1	.1332862	.0018928	.1332862	.0019675
e_0	.1248232	.0021064	.1248232	.0021383
a05	.0635771	.0013159	.0635771	.0013272
a2	.2104676	.0023613	.2104676	.0024529
p9010	3.033107	.0167076	3.033602	.017874
p7525	1.826839	.0101527	1.825998	.0077394
gini	.2731835	.0020243	.2731836	.0020572

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- FFL(2009) Introduced the a new type of quantile regression that they call unconditional quantile regression. This was a special case of RIF regressions.
- The core of the idea was:
 - In a linear regression $y = b_0 + b_1 * x_1 + b_2 * x_2 + e$ we are modeling how changes in x 's may cause a change in y .
 - $RIF(y_i, v(F_Y))$ is the contribution of an observation y_i has on the construction of statistic v .
 - then, if we model $RIF(y_i, v(F_Y)) = a_0 + a_1 * x_1 + a_2 * x_2 + e$, we are modeling how changes in X 's relate to the contributions of observation i to the statistic of interest.
- FFL(2009) proposed using the RIF instead of IF. (No impact on regressions)

- So now that we are modeling RIF's as functions of X's. The interpretation requires some care. why?

$$RIF(y_i, v(F_Y)) = a_0 + a_1 * x_1 + a_2 * x_2 + e$$

- The simple partial effect tell us...nothing, except for few exceptions (for example Mean, FGT and Watts poverty indices).

$$\frac{\partial RIF(.)}{\partial x_1} = a_1$$

$$\frac{\partial E(RIF(.)|x_1, x_2)}{\partial x_1} = a_1$$

why? if x_1 changes for person i , that persons influence on the outcome will change in a_1 . But, in a population of millions, one person won't make a difference on v .

- Alternatively, if we take unconditional expectations:

$$E\left(\text{RIF}(y_i, v(F_Y))\right) = E(a_0 + a_1 * x_1 + a_2 * x_2 + e)$$

$$v(F_Y) = a_0 + a_1 * E(x_1) + a_2 * E(x_2)$$

- So we can derive correct partial effect

$$\frac{\partial v(F_Y)}{\partial E(x_1)} = a_1$$

- a_1 is the effect that a unit change in the average value of $E(x_1)$ will have on statistic v , assuming everything else constant.
- For most $v()$, one needs to assume everyone's x change in 1 unit.
- For Dummy variables, one needs to assume the change is in the Proportion of people in a particular group.

- Up until now, 3 other options were available for the estimation of RIF regressions: `-rifreg-` (FFL2009); `-xtrifreg-` (Borgen2016); `-rifireg-` Heckley et al (2016).
- the command `-rifhdreg-` does everything this other commands do with additional capabilities.
 - Can estimate all RIFs using `_grifvar()`
 - It is a wrapper around `regress` and `reghdfe` (Correia 2017). So most of their capabilities are used.
 - Different weight options, robust standard errors, fixed effects and allows for factor variables.
- It has a simple syntax:
`rifhdreg depvar indepvar [aw pw iw] [if in],
rif(rifoptions) regress_options reghdfe_options`

rifhdreg: Example

RIF regression with rescaled RIF:

```
rifhdreg wage age grade union tenure hours wks_work,  
rif(gini) scale(1000) robust
```

```
rifhdreg wage age grade union tenure hours wks_work,  
rif(ucs(80)) scale(100) robust
```

RIF regression with rescaled RIF and Fixed effects:

```
rifhdreg wage age grade union tenure hours wks_work,  
rif(gini) scale(1000) vce(robust) abs(idcode) keepsingleton  
rifhdreg wage age grade union tenure hours wks_work,  
rif(ucs(80)) scale(100) vce(robust) abs(idcode)  
keepsingleton
```

RIF Regression: rifhdreg

```
. rifhdreg wage age grade union tenure hours wks_work, rif(lor(20)) scale(100) robust
```

```
Linear regression                               Number of obs   =    18,601
                                                F(6, 18594)    =    107.15
                                                Prob > F       =    0.0000
                                                R-squared     =    0.0402
                                                Root MSE     =    6.539
```

RIF(wage)	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
age	-.1328956	.0109562	-12.13	0.000	-.1543707	-.1114205
grade	-.2921282	.025608	-11.41	0.000	-.3423223	-.2419341
union	-.1795715	.1245977	-1.44	0.150	-.4237943	.0646514
tenure	-.08538	.0124141	-6.88	0.000	-.1097128	-.0610472
hours	.0643307	.0085744	7.50	0.000	.047524	.0811373
wks_work	.0084812	.0022489	3.77	0.000	.0040731	.0128892
_cons	15.27706	.4965183	30.77	0.000	14.30384	16.25028

Distributional Statistic: lor(20)

Sample Mean RIF lor(20) : 9.8941

RIF Regression: rifhdreg

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	gini	ucs80	lor20	mean X	FE gini	FE ucs80	FE lor20
age	4.905 (0.548)	0.350 (0.0470)	-0.133 (0.0110)	31.39 (0.0454)	5.424 (1.079)	0.398 (0.0936)	-0.134 (0.0183)
grade	4.604 (1.240)	0.0708 (0.108)	-0.292 (0.0256)				
union	0.746 (7.176)	-0.501 (0.630)	-0.180 (0.125)	0.236 (0.00311)	-13.91 (7.065)	-1.101 (0.615)	0.474 (0.153)
tenure	-0.102 (0.588)	-0.129 (0.0525)	-0.0854 (0.0124)	4.003 (0.0303)	0.644 (1.006)	-0.0584 (0.0887)	-0.105 (0.0189)
hours	-3.142 (0.327)	-0.251 (0.0274)	0.0643 (0.00857)	36.82 (0.0698)	-3.037 (0.552)	-0.263 (0.0473)	0.0478 (0.0115)
wks_work	-0.288 (0.103)	-0.0168 (0.00885)	0.00848 (0.00225)	63.29 (0.208)	-0.204 (0.105)	-0.0128 (0.00913)	0.00510 (0.00218)
._cons	185.2 (18.44)	35.34 (1.567)	15.28 (0.497)		218.9 (40.58)	34.80 (3.511)	12.31 (0.714)
<i>N</i>	18601	18601	18601	18601	18601	18601	18601
rifmean	263.8	36.31	9.894		263.8	36.31	9.894

Standard errors in parentheses

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- -rifhdreg- provides a simple framework for analyzing the impact of changes in the distribution of X 's on distributional statistics, at the margin. (RIF's are Local linear approximations)
- Large changes in distributions require other methods, for example decomposition methods: Oaxaca Blinder
- The premise, allow for all distributions of X 's to change between two groups.
- As long as the conditional independence assumptions holds, we can apply OB to decompose differences in statistics as functions of differences in characteristics, and differences in returns to those characteristics.

- The simple OB framework:

$$\Delta v = v(H_y) - v(F_y)$$

$$\Delta v = \beta_h \bar{X}_h - \beta_f \bar{X}_f$$

$$\Delta v = \beta_h (\bar{X}_h - \bar{X}_f) - (\beta_h - \beta_f) \bar{X}_f$$

This assumes a linear counterfactual $v(CF_y) = \beta_h \bar{X}_f$

- A better counterfactual can be obtained using IPW ($\omega(x)$).

$$\Delta v = v\left(\int (H_{Y|X} * dH_X)\right) - v\left(\int (F_{Y|X} * dF_X)\right)$$

$$v(CF_y) = v\left(\int (H_{Y|X} * dF_X)\right) = v\left(\int (H_{Y|X} * \omega(x) * dH_X)\right) = \beta_c \bar{X}_c$$

- `-oaxaca_rif-` is a wrapper around `-oaxaca-` (Jann 2008) that implements these two types of decompositions.
- It basically estimates the appropriate RIFs, uses them as dependent variables, and re-arranges the results.
- the syntax
`oaxaca_rif depvar indepvar [aw pw iw] [if in], by(var)
rif(rifoptions) IPW_options oaxaca_options`
- Many features of `-oaxaca-` are kept.

Oaxaca_rif example

```
bootstrap:oaxaca_rif wage age grade tenure hours wks_work,  
rif(mean) by(union) swap w(1)
```

```
bootstrap: oaxaca_rif wage age grade tenure hours wks_work,  
rif(mean) by(union) rwlogit(age grade tenure hours  
wks_work) swap w(1)
```

```
bootstrap: oaxaca_rif wage age grade tenure hours wks_work,  
rif(gini) by(union) rwlogit(age grade tenure hours  
wks_work) scale(1000) swap w(1)
```

```
bootstrap: oaxaca_rif wage age grade tenure hours wks_work,  
rif(ucs(80)) by(union) rwlogit(age grade tenure hours  
wks_work) scale(100) swap w(1)
```

Oaxaca_rif example

Model : Blinder-Oaxaca RIF-decomposition

Type : Standard

RIF : mean

Scale : 1

Group 1: union = 1

N of obs 1 = 4382

Group c: x2*b1

N of obs C = .

Group 2: union = 0

N of obs 2 = 14219

wage	Observed Coef.	Bootstrap Std. Err.	z	P> z	Normal-based [95% Conf. Interval]	
overall						
group_1	7.611619	.0794412	95.81	0.000	7.455917	7.767321
group_2	6.174817	.0243625	253.46	0.000	6.127067	6.222566
difference	1.436802	.0808799	17.76	0.000	1.278281	1.595324
explained	.2467854	.0639477	3.86	0.000	.1214503	.3721205
unexplained	1.190017	.1289997	9.22	0.000	.9371822	1.442852

Oaxaca_rif example

```

Group 1: union = 1                N of obs 1    = 4382
Group c: X1~>rw~>X2              N of obs C    = 4382
Group 2: union = 0                N of obs 2    = 14219
    
```

wage	Observed Coef.	Bootstrap Std. Err.	z	P> z	Normal-based [95% Conf. Interval]	
Overall						
Group_1	7.611619	.0684663	111.17	0.000	7.477427	7.74581
Group_c	7.296463	.1015878	71.82	0.000	7.097355	7.495572
Group_2	6.174817	.0243648	253.43	0.000	6.127062	6.222571
Tdifference	1.436802	.0766688	18.74	0.000	1.286534	1.587071
ToT_Explained	.3151556	.0540289	5.83	0.000	.2092609	.4210503
ToT_Unexplained	1.121647	.1086084	10.33	0.000	.9087782	1.334515
Explained						
Total	.3151556	.0540289	5.83	0.000	.2092609	.4210503
Pure_explained	.285553	.0510947	5.59	0.000	.1854092	.3856968
Specif_err	.0296026	.0075979	3.90	0.000	.014711	.0444942
Unexplained						
Total	1.121647	.1086084	10.33	0.000	.9087782	1.334515
Reweight_err	-.0235812	.0124732	-1.89	0.059	-.0480282	.0008659
Pure_Unexplained	1.145228	.1133489	10.10	0.000	.9230683	1.367388

Oaxaca_rif example

	(1)	(2)	(3)	(4)	(5)	(6)
	gini	ucs80	lor20	q10	q50	q90
Overall						
Group_1	246.1 (5.946)	34.72 (0.536)	10.35 (0.116)	1.394 (0.00860)	1.919 (0.00922)	2.446 (0.00749)
Group_c	261.5 (9.477)	36.17 (0.920)	10.19 (0.168)	1.337 (0.0107)	1.848 (0.00987)	2.407 (0.0107)
Group_2	262.8 (2.029)	36.44 (0.167)	10.03 (0.0575)	1.197 (0.00545)	1.683 (0.00359)	2.309 (0.00480)
Explained						
Total	-15.43 (4.287)	-1.446 (0.448)	0.161 (0.0758)	0.0565 (0.00642)	0.0710 (0.00825)	0.0389 (0.00756)
Pure_explained	-16.27 (3.761)	-1.451 (0.395)	0.227 (0.0708)	0.0587 (0.00474)	0.0634 (0.00621)	0.0296 (0.00603)
Specif_err	0.840 (0.686)	0.00470 (0.0655)	-0.0659 (0.0174)	-0.00220 (0.00659)	0.00765 (0.00413)	0.00924 (0.00516)
Unexplained						
Total	-1.318 (9.336)	-0.273 (0.937)	0.161 (0.169)	0.140 (0.0122)	0.166 (0.00957)	0.0978 (0.0114)
Reweight_err	-2.119 (0.705)	-0.174 (0.0647)	0.0451 (0.0139)	0.00106 (0.00119)	-0.000364 (0.00157)	-0.00454 (0.00160)
Pure_Unexplained	0.801 (9.394)	-0.0997 (0.931)	0.115 (0.168)	0.139 (0.0123)	0.166 (0.00959)	0.102 (0.0119)

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- Since its inception, early this year, a few other expansions have been added to this program. Some very recent.
- `-rifhdreg-` It allows for SVY. Specially useful for the estimation of standard errors of distributional statistics.
- `-rifhdreg-` adds "over". This may be used as a partial conditional RIF. Useful for standard errors across multiple groups.
- `-rifhdreg-` can now estimate effects similar IPWRA treatment effects, using `rwlogit` or `rwprobit`. This is similar to Firpo and Pinto (2016). Allows for `att`, `ate` and `atu`. Useful for analyzing Inequality treatment effects

- `-rifsureg-`. This would be the equivalent to `sqreg`, but unconditional quantile regressions.
- Handy for making plots across quantiles.
- `-rifsureg2-` is similar to `rifsureg`, but allows to simultaneously estimate RIF regressions for non colinear models.
- `-uqreg-` Stand alone command to estimate Unconditional Partial effects for UQR with alternative model specifications (logit/probit/other)

rifsureg: Example

```
rifsureg ln_wage age grade union tenure hours wks_work,  
qnts(10(10)90)
```

```
margins, dydx(grade) nose
```

```
marginsplot
```

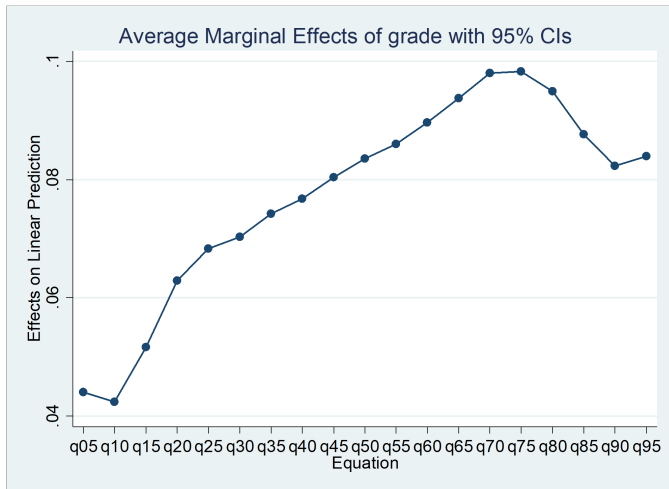


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Conclusions

- RIF and IF are powerful tools for analyzing and visualizing distributional statistics.
- The three main commands presented today (`_grifvar`, `rifhdreg`, `oaxaca_rif`) aim to facilitate the use of RIF's for this type of analysis
- Questions, comments and suggestions are welcome.
- Thank you
- Latest Files: <https://bit.ly/2NFM3cH>

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