

**Estimation of ordinal response models,
accounting for sample selection bias.**

ALFONSO MIRANDA

University of Keele

SOPHIA RABE-HESKETH

University of California, Berkeley

Correspondence to: A.Miranda@econ.keele.ac.uk

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ORDERED RESPONSES

- A limited number of H response categories $y_h, h = 1, 2, \dots, H$.
- Categories are ordered,

$$y_1 < y_2 < \dots < y_H$$

- Some examples:
 - Health condition status (excellent, good, regular, bad).
 - Opinions of a candidate in an election (strongly support, neutral, strongly opposed).
 - Job satisfaction (highly satisfied, satisfied, not satisfied).

LATENT REGRESSION MODELS

- The observed response for individual i , y_i is determined by a latent continuous variable process,

$$y_i^* = \mathbf{x}_i' \boldsymbol{\beta} + u_i \quad (1)$$

- A threshold model determines the observed response:

$$y_i = \begin{cases} 1 & \text{if } y_i^* \leq k_1 \\ 2 & \text{if } k_1 < y_i^* \leq k_2 \\ \cdot & \cdot \cdot \\ \cdot & \cdot \cdot \\ H & \text{if } k_{H-1} < y_i^* \end{cases}$$

- No constant is included in the covariate vector \mathbf{x}_i .

THE SAMPLE SELECTION PROBLEM

- the response variable is only observed if a particular condition ($sel = 1$) is met.
- A latent regression model for the selection variable is specified,

$$sel_i^* = \mathbf{z}_i' \boldsymbol{\gamma} + v_i, \quad (2)$$

where v_i is assumed to be normally distributed (z_i should include some variables not in x_i to secure identification).

- If $Cov(u_i, v_i) \neq 0$, using the observed sample of y and ordered Probit (ordered Logit) to estimate $\boldsymbol{\beta}$ will deliver biased estimators.
- This is known as the ‘sample selection bias’ problem (Heckman 1979).

THE SAMPLE SELECTION PROBLEM (CONT.)

- Notice that the correlation coefficient, ρ , is the only aspect of the covariance matrix that is identified. We impose therefore,

$$\Sigma = \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}.$$

Example

- Health condition only available for respondents who exercise at least twice a week. If healthier people exercise more, estimating a model for health condition on the basis of the observed sample will clearly deliver biased estimates.

GENERALISED LINEAR LATENT AND MIXED MODELS

- May use GLLAMMs to estimate the sample selection ordinal variable regression. We can write the ordinal variable model as,

$$\begin{aligned}
 y_i &\sim \text{multinomial}(1, \{\pi_{hi}, h = 1, \dots, H\}); \\
 g_1(\gamma_{hi}) &= \eta_{1hi} = \mathbf{x}_i' \boldsymbol{\beta} - \kappa_h + \lambda \varepsilon_i; \\
 &h = 1, \dots, H - 1.
 \end{aligned}$$

where

$$\gamma_{hi} = \sum_{s=h+1}^H \pi_{si} = Pr(y > h),$$

GLLAMMs (CONT. 1)

And the selection model as,

$$\begin{aligned} sel_i &\sim \text{binomial}(1, \pi_i) \\ g_2(\pi_i) &= \eta_{2i} = \mathbf{z}_i' \boldsymbol{\gamma} + \varepsilon_i \end{aligned}$$

where $\varepsilon_i \sim N(0, 1)$ is a latent variable representing unobserved heterogeneity and λ is a factor loading. This reparametrization reduces the dimensions of integration from 2 to 1.

A mixed response model

- Stack y_i and sel_i into a single variable q_{ji} , $j = 1, 2$.
- Viewing the ordinal variable $j = 1$ and the selection status $j = 2$ as clustered within individuals i , define the dummies $d_{1ji} = 1$ for the ordinal variable and $d_{2ji} = 1$ for the selection.

GLLAMMS (CONT. 2)

- Now we can define a mixed response model for q_{ji}

$$\begin{aligned}
 q_{ji} &\sim \begin{cases} \textit{multinomial} & \text{if } d_{1ji} = 1 \\ \textit{binomial} & \text{if } d_{2ji} = 1 \end{cases} \\
 \eta_{jhi} &= d_{1ji} [\mathbf{x}_i' \boldsymbol{\beta} - \kappa_h + \lambda \varepsilon_i] + d_{2ji} [\mathbf{z}_i' \boldsymbol{\gamma} + \varepsilon_i]; \\
 & \quad j = 1, 2, \quad h = 1, \dots, H - 1; \tag{3}
 \end{aligned}$$

- g_1 can be either the ordered Probit or the ordered Logit link. We use always a Probit link for g_2 .

GLLAMMs (CONT. 3)

- Due to the increase in the residual variance in (3) we expect β to increase by a factor of $\sqrt{1 + \lambda^2}$ if g_1 is oprobit or $\sqrt{\frac{\pi^2}{3} + \lambda^2}$ if g_1 is ologit.
- Similarly, we expect γ to increase by a factor of $\sqrt{2}$.
- Hence, after estimation β and γ must be rescaled!!
- Notice finally that,

$$\rho = \begin{cases} \frac{\lambda}{\sqrt{2(1+\lambda^2)}} & \text{if } g_1 \text{ is oprobit} \\ \frac{\lambda}{\sqrt{2\left(\frac{\pi^2}{3} + \lambda^2\right)}} & \text{if } g_1 \text{ is ologit} \end{cases}$$

THE **osm** COMMAND

- **osm** is a **gllamm** (Rabe-Hesketh, Skrondal & Pickles 2004) ‘wrapper’ program that fits endogenous switching and sample selection models for ordinal and count variables (endogenous switching is the default option).
- Accepts data in the usual wide format and then does the required changes to call **gllamm**.
- After estimation coefficients are rescaled and an output table that is easily interpretable is presented.
- **osm** exploits the adaptive quadrature capability of **gllamm**...one of the major **gllamm** strengths.

SYNTAX

`osm depvar [varlist] [if exp] [in range], i(varname)`

`switch(varname= varlist) switch(varlist) Family(familyname)`

`selection quadrature(#) Link(linkname) From(initial values)`

`Trace nolog Trace Eval Commands`

Table 1: Available Families and links

Family	Link
Poisson	log
Binomial	ordinal Probit
	ordinal Logit

SAMPLE SELECTION ORDERED PROBIT: AN EXAMPLE

```
. osm ordvar x1 x2, id(id) s(sel = x1 x2 x3 x4) q(15) adapt family(bin) link(oprobit) sel
```

Running adaptive quadrature

Iteration 0: log likelihood = -5444.942

(output omitted)

Iteration 4: log likelihood = -5175.5835

Adaptive quadrature has converged, running Newton-Raphson

Iteration 0: log likelihood = -5175.5835

(output omitted)

Iteration 3: log likelihood = -5175.5765

Sample Selection Ordered Probit Regression

(Adaptive quadrature -- 15 points)

	Number of obs = 3500
	Wald chi2(6) = 1114.42
Log likelihood = -5175.5765	Prob > chi2 = 0.0000

		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
ordvar							
	x1	.3154251	.0408026	7.73	0.000	.2354535	.3953968
	x2	.1470225	.026887	5.47	0.000	.0943249	.1997202

```

selection |
      x1 |  .9573865  .0356374  26.86  0.000  .8875385  1.027234
      x2 |  .4217439  .0286755  14.71  0.000  .365541  .4779468
      x3 | - .5968153  .0303954 -19.64  0.000 - .6563893 - .5372414
      x4 |  .6372245  .0308598  20.65  0.000  .5767403  .6977087
      _cons |  .5448698  .0288654  18.88  0.000  .4882947  .601445
-----+-----
aux_ordvar |
      _cut1 | - .4012284  .0325979 -12.31  0.000 - .4651192 - .3373376
      _cut2 |  .1583416  .048699  3.25  0.001  .0628932  .2537899
      _cut3 |  .4265045  .0598836  7.12  0.000  .3091348  .5438743
      _cut4 |  .7873888  .0763759  10.31  0.000  .6376948  .9370827
      _cut5 |  1.229029  .0981156  12.53  0.000  1.036726  1.421332
-----+-----
      rho |  .2901614  .0654419  4.43  0.000  .1458488  .4012554
-----+-----
Likelihood ratio test for rho=0: chi2(1)= 21.32 Prob>=chi2 = 0.000

```

SAMPLE SELECTION ORDERED LOGIT: AN EXAMPLE

```
. osm ordvar x1 x2, id(id) s(sel = x1 x2 x3 x4) q(15) adapt family(bin) link(ologit) sel
```

Running adaptive quadrature

Iteration 0: log likelihood = -5468.3146

(output omitted)

Iteration 6: log likelihood = -5180.7342

Adaptive quadrature has converged, running Newton-Raphson

Iteration 0: log likelihood = -5180.7342

(output omitted)

Iteration 3: log likelihood = -5180.7303

Sample Selection Ordered Logit Regression

(Adaptive quadrature -- 15 points)

```

                                Number of obs =      3500
                                Wald chi2(6)   =    1123.24
                                Prob > chi2    =      0.0000
Log likelihood = -5180.7303

```

```

-----+-----
                |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
ordvar         |
  x1           |   .3267404   .0369461     8.84  0.000     .2543274   .3991535
  x2           |   .1510283   .0248903     6.07  0.000     .1022441   .1998125
-----+-----

```

```

selection |
      x1 |  .9563906  .0356895  26.80  0.000  .8864404  1.026341
      x2 |  .4199398  .0287014  14.63  0.000  .3636862  .4761935
      x3 | - .5964406  .0305075 -19.55  0.000  -.6562341  -.536647
      x4 |  .6383798  .0309312  20.64  0.000  .5777557  .6990038
      _cons | .5446456  .0288797  18.86  0.000  .4880425  .6012486
-----+-----
aux_ordvar |
      _cut1 |  -.447559  .0312788 -14.31  0.000  -.5088643  -.3862536
      _cut2 |  .0840034  .0416869   2.02  0.044  .0022986  .1657082
      _cut3 |  .4016877  .0529198   7.59  0.000  .2979668  .5054085
      _cut4 |  .7749008  .0673403  11.51  0.000  .6429162  .9068854
      _cut5 |  1.123638  .0803356  13.99  0.000  .9661834  1.281093
-----+-----
      rho |  .1214521  .0392716   3.09  0.002  .0426992  .1958209
-----+-----
Likelihood ratio test for rho=0: chi2(1)= 11.43 Prob>=chi2 = 0.001

```

FINAL REMARKS

- Besides estimating sample selection models **osm** fits endogenous switching models (i.e., when an endogenous dummy is present in the main equation).
- Using the Poisson Family and the Log link **osm** can fit models for count data.
- In the near future **osm** will be extended to allow for:
 - Probit/Logit links.
 - Weights.