

predcumi: A postestimation command for predicting and visualising cumulative incidence estimates after Cox regression models.



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Outline

- Background.
 - What motivated this work.
 - Competing risks: Survival vs. cumulative incidence function (CIF).
 - Cause-specific hazard formulation of competing risk models.
 - Competing risk models in Stata (what was missing).
- Implementation of **predcumi**
 - Syntax.
 - Programming aspects.
- Examples: CVD risk prediction
 - Cumulative incidence over observed times vs. within landmark times.
 - Graphical visualisation of the unadjusted/adjusted probabilities.
 - Comparison of covariate adjusted CIF vs. previous implementations.
- Remarks.

Motivation

- Involved in meta-analysis of individual participant data from multiple studies in **cardiovascular disease (CVD)**.
 - **CVD** is a composite endpoint, comprising **CHD** and **Stroke** (**each with further subtypes**).
 - Aetiological associations suggest some risk factors (e.g. lipids) have stronger associations with CHD and others (e.g. blood pressure) have stronger associations with stroke.
 - Yet it is somewhat traditional to analyse the composite CVD endpoint for risk prediction purposes.
- Question: Could disaggregation of prediction models for CVD to CHD and stroke components possibly lead to better predictions for the composite CVD outcome?
 - What about adjustment for other competing risks (e.g. **non-CVD death**)?

Competing risks (CR)

- Competing risks (e.g. death from other cause) prevent the event of interest from occurring altogether.
 - Hence, CRs should be handled differently from censoring when interest is on making **absolute risk predictions**.
 - Otherwise will overestimate the absolute risk of the event of interest.
- Kaplan-Meier failure estimates (**sts generate**) or absolute risk predictions from Cox models (**stcox**) do not account for competing risks ...
 - But the **stcox** model provides **cause-specific hazards (CSH)**, the fundamental quantities needed for calculating **cumulative incidence** to account for competing risks.

CSH, survival, and cumulative incidence

- **Cause-specific hazard (CSH) $h_k(t)$** : the hazard of failing from a specific cause (k) in the presence of competing risks.

$$h_k(t) = \lim_{\Delta t \rightarrow 0} \frac{\text{Prob}(t \leq T < t + \Delta t, D = k \mid T \geq t)}{\Delta t}$$

- **Cumulative hazard $H_k(t)$ and Survival $S_k(t)$, $S(t)$** :

$$H_k(t) = \sum_{j:t_j \leq t} h_k(t_j) \quad S_k(t) = \exp(-H_k(t)) \quad S(t) = \exp\left(-\sum_{k=1}^K H_k(t)\right)$$

- **Cumulative incidence function (CIF) $I_k(t)$** : Overall survival

$$I_k(t) = \sum_{j:t_j \leq t} h_k(t_j) S(t_{j-1})$$

- $h_k(t)$ is obtained after **stcox** using **--predict, basehc--** followed by careful sorting and summations by strata.

Back to motivating problem ...

- Planned to disaggregate the prediction of absolute risk of CVD to CHD and stroke components, treating both as competing events.
 - Calculate CIF of the composite CVD endpoint as sum of the predicted CIFs for CHD and stroke.
 - Additionally treat non-CVD deaths as competing risks.
 - Compare the predictions from the above approaches by calculating discrimination measures (e.g. C-index).
- Needed a program that could calculate the covariate-adjusted CIF over observed failure times, as well as the maximum within user-specified landmark times (e.g. 10-years) for each individual in the dataset.

Competing risk models in Stata (what was missing)

- User-written programs by Enzo Coviello (**stcompet** and **stcompadj**), could only provide CIFs over time adjusted to specific covariate values.
- **stcrreg** was formulated differently than the **stcox** model used for inference of hazard ratios (and was also difficult to converge in large datasets).
- Solution: Write a bespoke postestimation program that predicts CIFs after **stcox** model (**predcum**), including graphical visualisation and optional adjustment of covariates.

--predcumi-- Implementation

- Assumes fitted Cox model is for event of interest and gets the specification from **e(cmdline)**.
- User provides specification of the competing event(s) via options **compete(string)** or **compvars(varlist)** and **compcodes(numlist)**.
- **--predcumi--** then refits the model for each endpoint to obtain the cause-specific baseline hazards, cause-specific linear predictor, overall baseline survival, and overall linear predictor.
- It then summates $h_k(t_j) * \exp(\mathbf{x}\mathbf{b}_k) * S_0(t_{j-1})^{\exp(\mathbf{x}\mathbf{b})}$ product over time by strata to obtain the CIF for each distinct covariate pattern.
 - This summation part implemented in **mata**, hence faster than adofile.
- If **attimes(numlist)** option is specified, it calculates the maximum CIF over each landmark time in **attimes()**.

--predcumi-- Syntax

predcumi [*if exp*] [*in range*] [, **compvars**(*varlist*)
compcodes(*numlist*) **compete**(*string*) **attimes**(*numlist*
integer >0) **lifetimes**(*numlist integer* >0) **nograph**
nodots **adjust**(*string*) **xvar**(*varname*) **compare**(*varlist*)
nobaseline **nobytime** *graph_options*]

- Currently uses default output variable names xb^* , $s0^*$, hc^* , $cumhc^*$, pf^* , cif^* where $*$ = 1, 2, 3, ... endpoints, to save the predictions, but will make flexible in the future.
- **adjust**(*string*) options can be either **#**, **mean**, **p1**, **p5**, **p10**, **p25**, **p50**, **p75**, **p90**, **p95**, **p99**, or **round(#)**, in which case the adjusted CIF calculations are done after fixing the covariates at the values specified.

Example: CVD risk prediction

Contains data from data\predcumi_demo_data.dta

obs: 817

vars: 15

11 Sep 2012 15:28

variable name	storage type	display format	value label	variable label
cohort	str7	%9s		Cohort abbreviation
subjectid	str3	%9s		Subject ID
sex	byte	%8.0g	sex	Sex
duration	float	%9.0g		Time to event/censoring (yrs)
ep_chdmi	byte	%9.0g		CHD
ep_crbv	byte	%8.0g		Stroke
ep_cvd	byte	%9.0g		CVD (CHD or stroke)
ep_ncv_f	byte	%9.0g		Non-CVD death
ep_dead	byte	%9.0g		Any death
ages	float	%9.0g		Age at survey (yrs)
smallbin	byte	%9.0g	statbin	Smoking status
sbp	int	%8.0g		SBP (mmHg)
hxdiabbin	byte	%17.0g	hx	History of diabetes
tchol	float	%9.0g		Total cholesterol (mmol/l)
hdl	float	%9.0g		HDL-C (mmol/l)

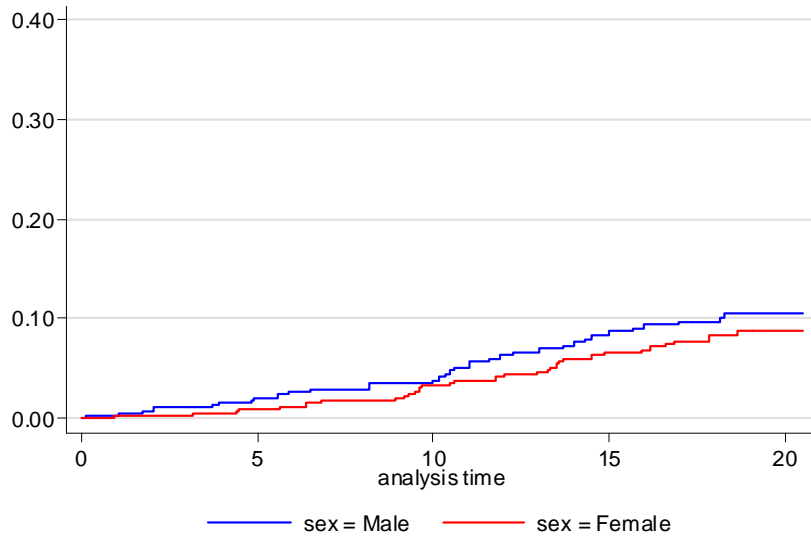
Sorted by: cohort subjectid

Data descriptive statistics

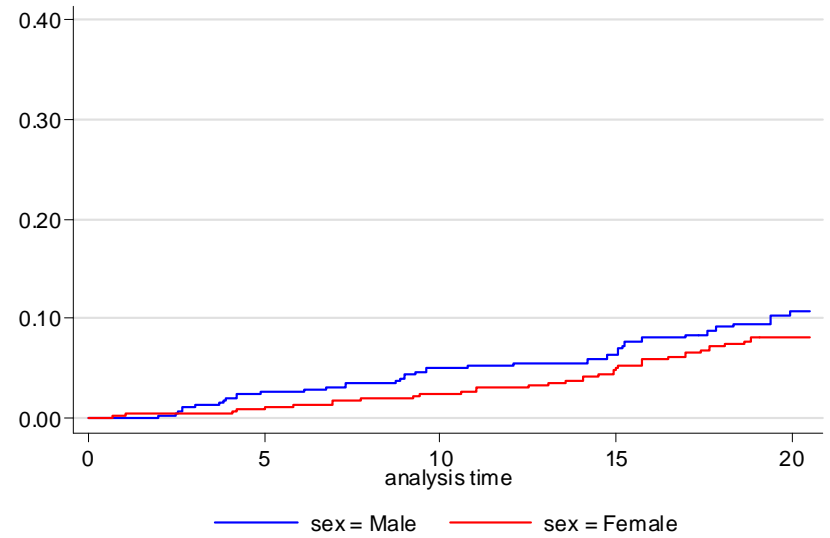
Variable	Obs	Mean	Std. Dev.	Min	Max
cohort	0				
subjectid	0				
sex	817	1.512852	.500141	1	2
sexbin	817	.5128519	.500141	0	1
duration	817	16.74673	5.693721	.0821355	20.4627
ages	817	57.85982	11.38797	40	80.096
smallbin	817	.2423501	.4287675	0	1
sbp	817	145.1004	21.50535	91	230
hxdiabbin	817	.0330477	.1788707	0	1
tchol	817	5.727442	1.028323	3.08	10.24
hdl	817	1.473035	.3640861	.7	2.97
variable	N	sum			
ep_chdmi	817	66	CHD		
ep_crbv	817	63	Stroke		
ep_cvd	817	129	CVD (CHD or stroke)		
ep_ncv_f	817	157	Non-CVD death		
ep_dead	817	240	Any death		

KM failure estimates

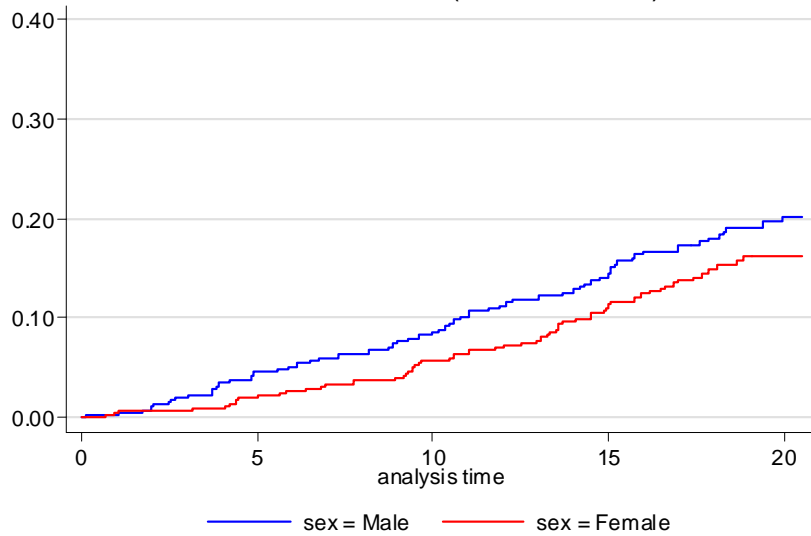
KM failure: CHD



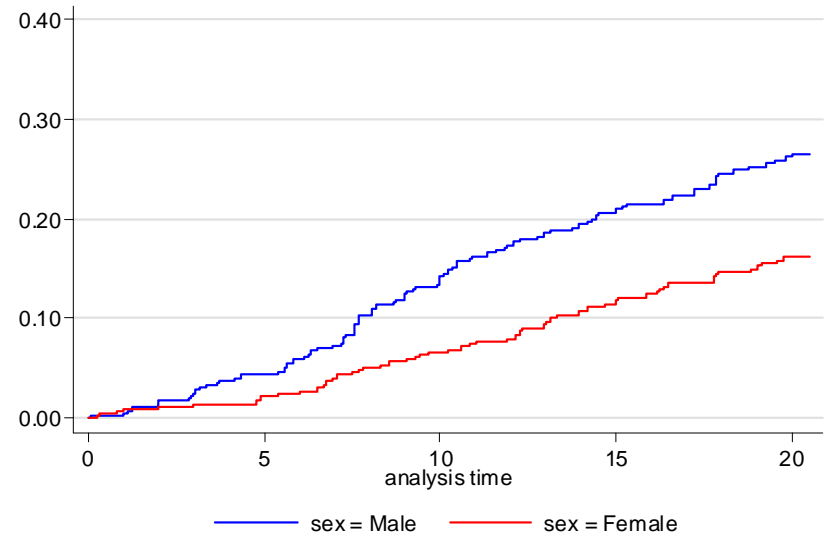
KM failure: Stroke



KM failure: CVD (CHD or stroke)



KM failure: Non-CVD death



User-fitted CHD model (event of interest)

```
. stset duration, failure("ep_chdmi == 1")
. xi: stcox ages smallbin sbp hxdiabbin tchol hdl, strata(sex) nolog
```

```
failure _d: ep_chdmi == 1
```

```
analysis time _t: duration
```

```
Stratified Cox regr. -- Breslow method for ties
```

```
No. of subjects =          817                Number of obs   =          817
```

```
No. of failures =           66
```

```
Time at risk    = 13682.08051
```

```
LR chi2(6)      =          89.58
```

```
Log likelihood  = -338.07286                Prob > chi2      =          0.0000
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
ages	1.085469	.0150106	5.93	0.000	1.056444 1.115292
smallbin	2.101201	.6167235	2.53	0.011	1.182038 3.735113
sbp	1.023793	.0060849	3.96	0.000	1.011936 1.035788
hxdiabbin	1.883857	.8553814	1.39	0.163	.773664 4.587153
tchol	1.041145	.128146	0.33	0.743	.8179821 1.325191
hdl	1.17283	.3919691	0.48	0.633	.6091954 2.257945

Stratified by sex

--predcumi-- Model 1 (CHD)

```
. predcumi, compvars("ep_crbv ep_ncv_f") compcodes(1) attimes(10)
```

```
failure _d: ep_chdmi == 1
```

```
analysis time _t: duration
```

```
Stratified Cox regr. -- Breslow method for ties
```

```
No. of subjects = 817 Number of obs = 817
```

```
No. of failures = 66
```

```
Time at risk = 13682.08051
```

```
LR chi2(6) = 89.58
```

```
Log likelihood = -338.07286 Prob > chi2 = 0.0000
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
ages	1.085469	.0150106	5.93	0.000	1.056444	1.115292
smallbin	2.101201	.6167235	2.53	0.011	1.182038	3.735113
sbp	1.023793	.0060849	3.96	0.000	1.011936	1.035788
hxdiabbin	1.883857	.8553814	1.39	0.163	.773664	4.587153
tchol	1.041145	.128146	0.33	0.743	.8179821	1.325191
hdl	1.17283	.3919691	0.48	0.633	.6091954	2.257945

Stratified by sex

```
Duplicates in terms of sex _t0 _t hcl
```

dupl1	Freq.	Percent	Cum.
0	60	90.91	90.91
1	6	9.09	100.00
Total	66	100.00	

--predcumi-- Model 2 (Stroke)

```
-> stset duration, failure(ep_crbv==1)
```

```
failure _d: ep_crbv == 1
```

```
analysis time _t: duration
```

```
Stratified Cox regr. -- Breslow method for ties
```

```
No. of subjects = 817 Number of obs = 817
```

```
No. of failures = 63
```

```
Time at risk = 13682.08051
```

```
LR chi2(6) = 45.67
```

```
Log likelihood = -341.26046 Prob > chi2 = 0.0000
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
ages	1.080394	.0146957	5.68	0.000	1.051971	1.109584
smallbin	1.239853	.3939733	0.68	0.499	.6651106	2.311247
sbp	1.004778	.0065243	0.73	0.463	.9920712	1.017647
hxdiabbin	1.123873	.6831708	0.19	0.848	.341425	3.699464
tchol	.9629966	.1232594	-0.29	0.768	.7493327	1.237584
hdl	.6935387	.2667539	-0.95	0.341	.3263438	1.473893

Stratified by sex

```
Duplicates in terms of sex _t0 _t hc2
```

dupl2	Freq.	Percent	Cum.
0	57	90.48	90.48
1	6	9.52	100.00
Total	63	100.00	

--predcumi-- Model 3 (Non-CVD death)

```
-> stset duration, failure(ep_ncv_f==1)
```

```
failure _d: ep_ncv_f == 1
```

```
analysis time _t: duration
```

```
Stratified Cox regr. -- Breslow method for ties
```

```
No. of subjects = 817 Number of obs = 817
```

```
No. of failures = 157
```

```
Time at risk = 13682.08051
```

```
LR chi2(6) = 175.09
```

```
Log likelihood = -822.29807
```

```
Prob > chi2 = 0.0000
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]	
ages	1.105477	.0099111	11.18	0.000	1.086221	1.125074
smallbin	1.661562	.3161485	2.67	0.008	1.144344	2.412549
sbp	.9982142	.0041068	-0.43	0.664	.9901974	1.006296
hxdiabbin	1.981136	.6249526	2.17	0.030	1.067583	3.676438
tchol	.8962753	.0760461	-1.29	0.197	.7589614	1.058432
hdl	1.454097	.3178565	1.71	0.087	.9473837	2.231828

Stratified by sex

```
Duplicates in terms of sex _t0 _t hc3
```

dupl3	Freq.	Percent	Cum.
0	109	69.43	69.43
1	32	20.38	89.81
2	12	7.64	97.45
3	4	2.55	100.00
Total	157	100.00	

--predcumi-- Model 4 (Any event)

```
-> stset duration, failure(__000002)
```

```
failure _d: __000002
```

```
analysis time _t: duration
```

```
Stratified Cox regr. -- Breslow method for ties
```

```
No. of subjects = 817 Number of obs = 817
```

```
No. of failures = 286
```

```
Time at risk = 13682.08051
```

```
LR chi2(6) = 290.60
```

```
Log likelihood = -1511.4978
```

```
Prob > chi2 = 0.0000
```

```
-----
```

_t	Haz. Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
ages	1.094803	.0071982	13.78	0.000	1.080785 1.109002
smallbin	1.636837	.2329696	3.46	0.001	1.238382 2.163496
sbp	1.006025	.0030151	2.00	0.045	1.000133 1.011952
hxdiabbin	1.789991	.4242442	2.46	0.014	1.124885 2.84835
tchol	.9462233	.0580374	-0.90	0.367	.8390435 1.067094
hdl	1.187661	.1965384	1.04	0.299	.8586825 1.642678

```
-----
```

Stratified by sex

```
Duplicates in terms of sex _t0 _t hc
```

```
-----
```

dupl	Freq.	Percent	Cum.
0	144	50.35	50.35
1	94	32.87	83.22
2	36	12.59	95.80
3	12	4.20	100.00
Total	286	100.00	

```
-----
```

--predcumi-- Model 1 CIF over time

Calculating cumulative incidence of ep_chdmi over time among 817 subjects
in 2 strata at 817 linear predictor values

Stratum = 1 (Male), n = 398, and 398 linear predictors

```
..... 50
..... 100
..... 150
..... 200
..... 250
..... 300
..... 350
.....
```

Stratum = 2 (Female), n = 419, and 419 linear predictors

```
..... 50
..... 100
..... 150
..... 200
..... 250
..... 300
..... 350
..... 400
.....
```

--predcumi-- Model 1 CIF up to time 10

Calculating cumulative incidence of ep_chdmi at time = 10

in 2 strata

Stratum = 1 (Male), n = 398, and 398 linear predictors

```
..... 50
..... 100
..... 150
..... 200
..... 250
..... 300
..... 350
.....
```

Stratum = 2 (Female), n = 419, and 419 linear predictors

```
..... 50
..... 100
..... 150
..... 200
..... 250
..... 300
..... 350
..... 400
.....
```

... and similarly for each of Model 2 (Stroke) and Model 3 (Non-CVD death)

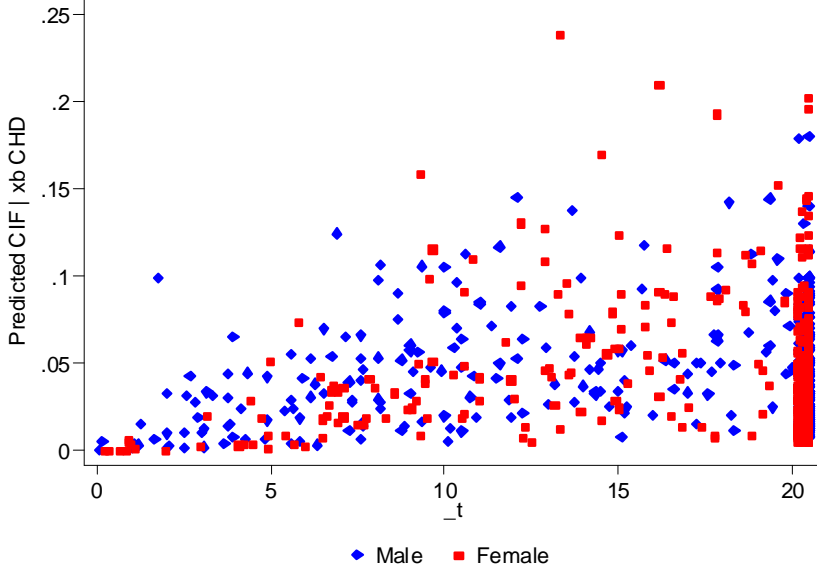
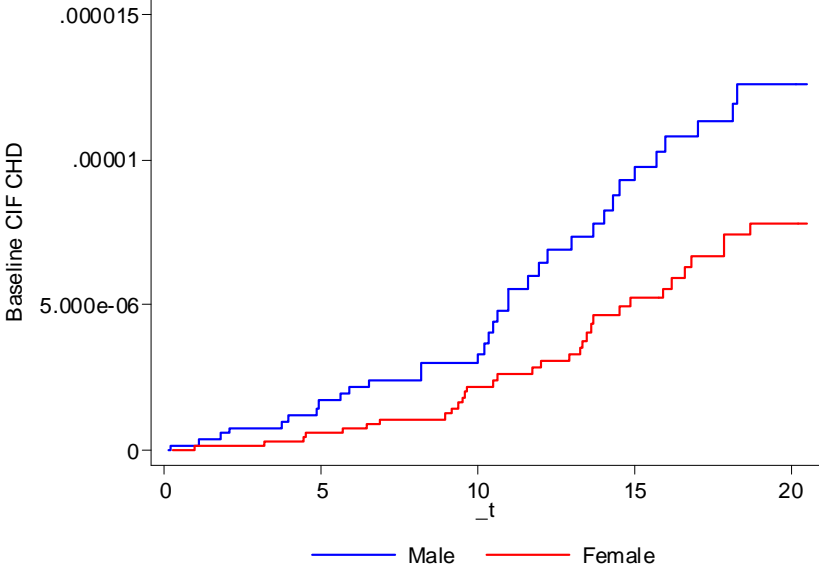
Execution time = 4 sec for calculations + 12 sec for graphs

--predcumi-- Output variables (default)

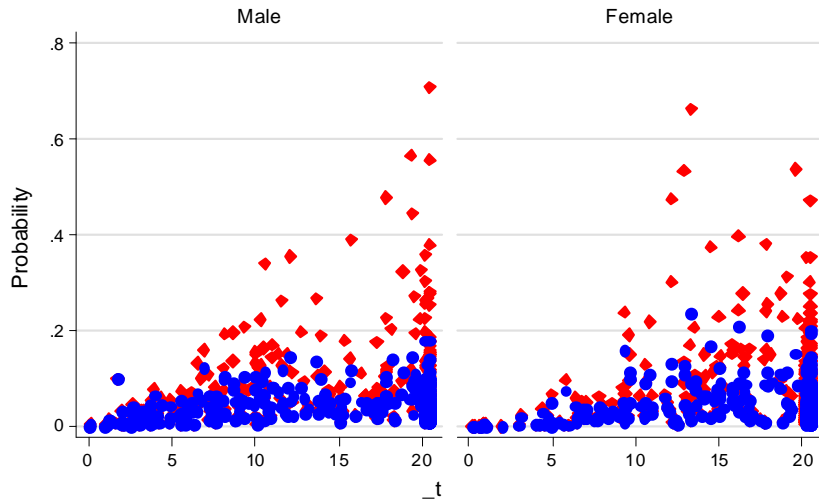
Generated variables

variable name	storage type	display format	value label	variable label
xb1	double	%10.0g		Linear prediction CHD
xb2	double	%10.0g		Linear prediction Stroke
xb3	double	%10.0g		Linear prediction Non-CVD death
xb	double	%10.0g		Linear prediction overall
cif01	double	%10.0g		Baseline CIF CHD
cif1	double	%10.0g		Predicted CIF xb CHD
cif01t10	double	%10.0g		Baseline CIF (T <= 10) CHD
cif1t10	double	%10.0g		Predicted CIF (T <= 10) CHD
cif02	double	%10.0g		Baseline CIF Stroke
cif2	double	%10.0g		Predicted CIF xb Stroke
cif02t10	double	%10.0g		Baseline CIF (T <= 10) Stroke
cif2t10	double	%10.0g		Predicted CIF (T <= 10) Stroke
cif03	double	%10.0g		Baseline CIF Non-CVD death
cif3	double	%10.0g		Predicted CIF xb Non-CVD death
cif03t10	double	%10.0g		Baseline CIF (T <= 10) Non-CVD death
cif3t10	double	%10.0g		Predicted CIF (T <= 10) Non-CVD death
pf1	double	%10.0g		Cox Pr(failure) CHD
pf1t10	double	%10.0g		Cox Pr(T <= 10) CHD
pf2	double	%10.0g		Cox Pr(failure) Stroke
pf2t10	double	%10.0g		Cox Pr(T <= 10) Stroke
pf3	double	%10.0g		Cox Pr(failure) Non-CVD death
pf3t10	double	%10.0g		Cox Pr(T <= 10) Non-CVD death
pf	double	%10.0g		Cox Pr(failure) overall
pft10	double	%10.0g		Cox Pr(T <= 10) overall

Baseline CIF and predicted CIF vs. time

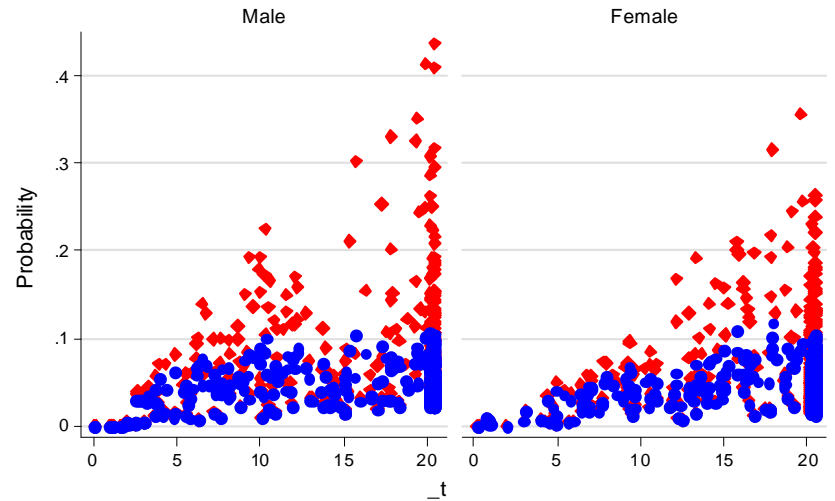


Cox Pr(failure) and predicted CIF vs. time



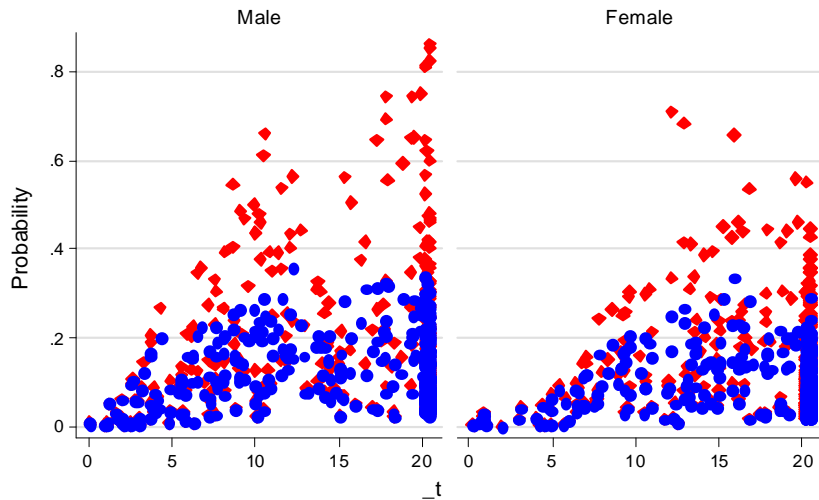
◆ Cox Pr(failure) CHD ● Predicted CIF | xb CHD

Graphs by group(sex)



◆ Cox Pr(failure) Stroke ● Predicted CIF | xb Stroke

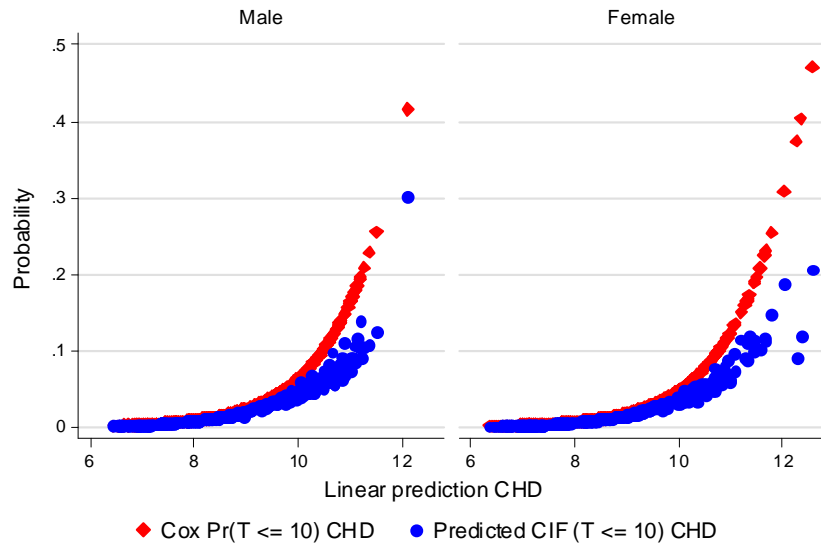
Graphs by group(sex)



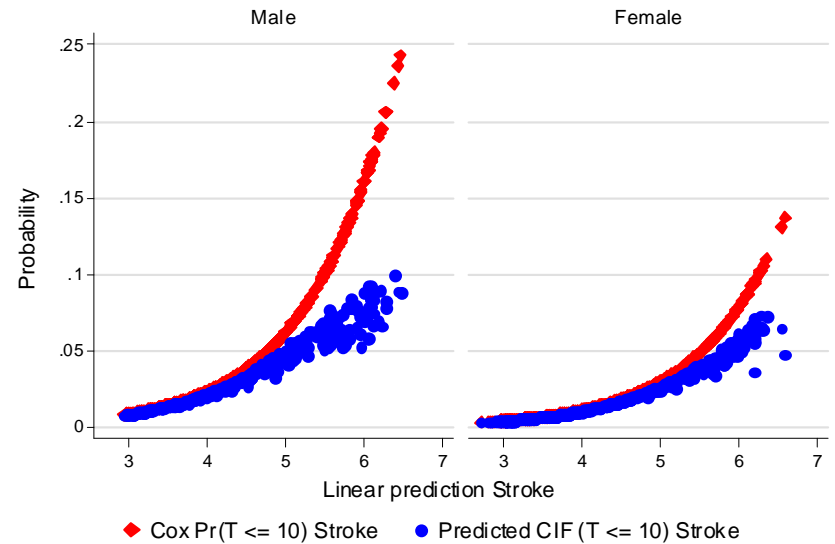
◆ Cox Pr(failure) Non-CVD death ● Predicted CIF | xb Non-CVD death

Graphs by group(sex)

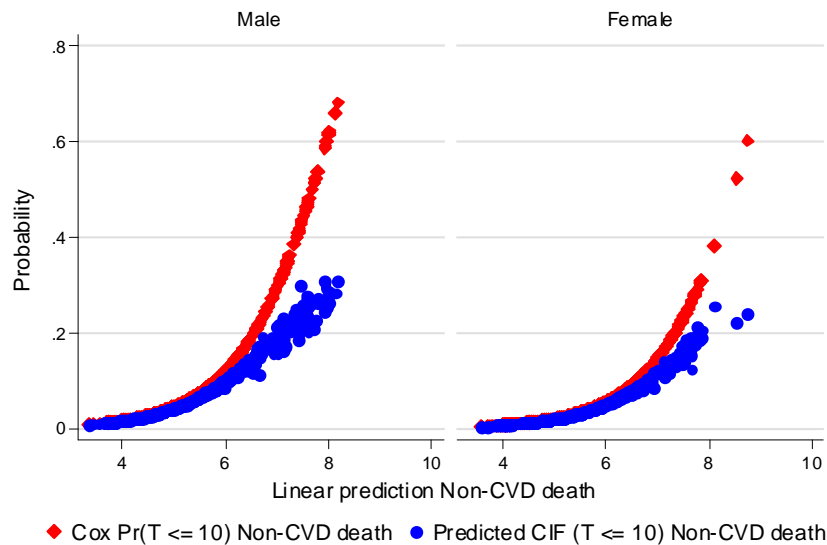
Cox Pr(failure) and predicted CIF vs. linear predictor



Graphs by group(sex)

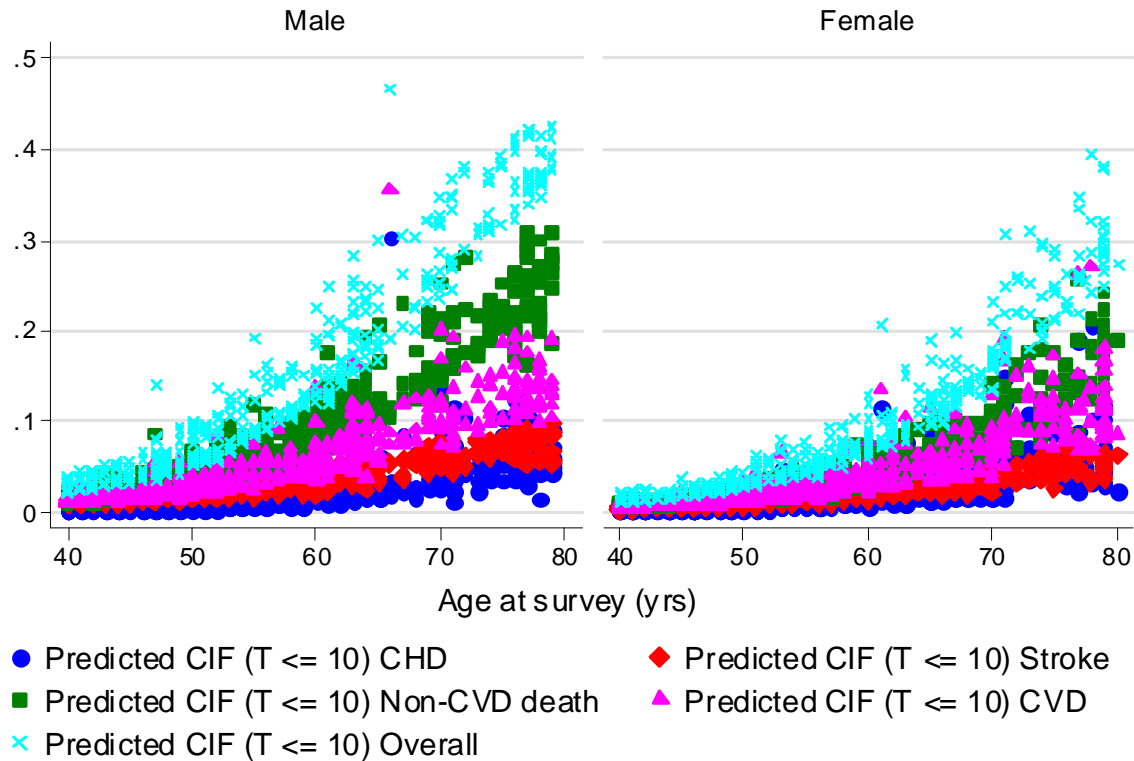


Graphs by group(sex)



Graphs by group(sex)

DIY combined summary of the predicted CIFs vs. age



Graphs by Sex

$$\text{Where: } \text{CIF}_{\text{CVD}} = \text{CIF}_{\text{CHD}} + \text{CIF}_{\text{Stroke}}$$

$$\text{CIF}_{\text{Overall}} = \text{CIF}_{\text{CHD}} + \text{CIF}_{\text{Stroke}} + \text{CIF}_{\text{Non-CVD}}$$

--predcumi-- Covariate adjusted CIF (1)

```
. predcumi, compvars("ep_crbv ep_ncv_f") compcodes(1) attimes(10) ///  
  adjust(smallbin = 0 sbp = mean hxdiabbin = 0 tchol = mean hdl = mean) ///  
  xvar(ages) nobaseline nobytime
```

< output omitted, same models as described before >

Adjusting covariates to specified values ...

```
Adjusting smallbin      = 0  
Adjusting sbp          = 145.1003671970624      (mean)  
Adjusting hxdiabbin    = 0  
Adjusting tchol        = 5.727441867678956      (mean)  
Adjusting hdl          = 1.473035494437854      (mean)
```

Unadjusted variables ages

Variable	Obs	Mean	Std. Dev.	Min	Max
ages	817	57.85982	11.38797	40	80.096

Calculating cumulative incidence of ep_chdmi at time = 10
in 2 strata

Stratum = 1 (Male), n = 398, and 40 linear predictors

.....

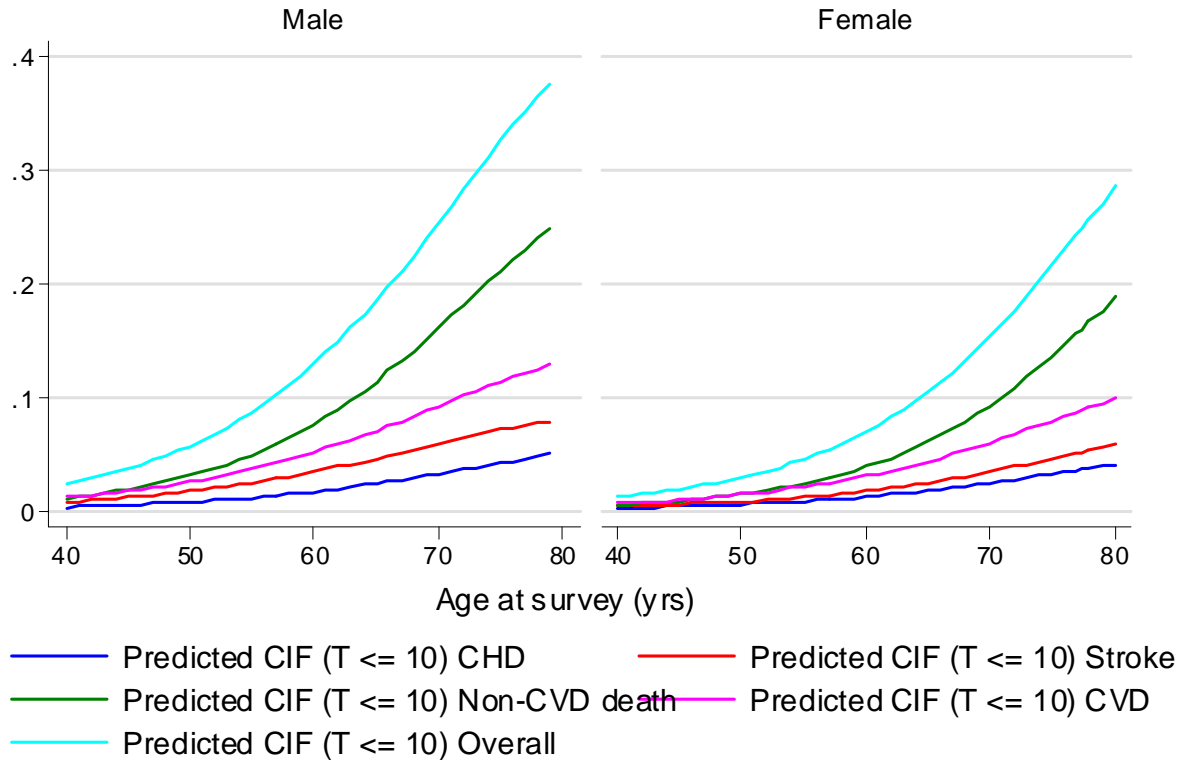
Stratum = 2 (Female), n = 419, and 42 linear predictors

.....

... and similarly for each of Model 2 (Stroke) and Model 3 (Non-CVD death)

Execution time = 1 sec for calculations + 6 sec for graphs

DIY combined summary of the predicted CIFs vs. age



Graphs by Sex

$$\text{Where: } \text{CIF}_{\text{CVD}} = \text{CIF}_{\text{CHD}} + \text{CIF}_{\text{Stroke}}$$

$$\text{CIF}_{\text{Overall}} = \text{CIF}_{\text{CHD}} + \text{CIF}_{\text{Stroke}} + \text{CIF}_{\text{Non-CVD}}$$

--predcumi-- Covariate adjusted CIF (2)

```
. predcumi, compvars("ep_crbv ep_ncv_f") compcodes(1) attimes(10) ///
  adjust(sbp = p95 hxdiabbin = 0 tchol = p95 hdl = p5) ///
  xvar(ages) nobaseline nobytime compare(smallbin)
```

< output omitted, same models as described before >

Adjusting covariates to specified values ...

```
Adjusting sbp          = 183                (p95)
Adjusting hxdiabbin   = 0
Adjusting tchol        = 7.46999979019165    (p95)
Adjusting hdl          = .9599999785423279    (p5)
Unadjusted variables  ages smallbin
```

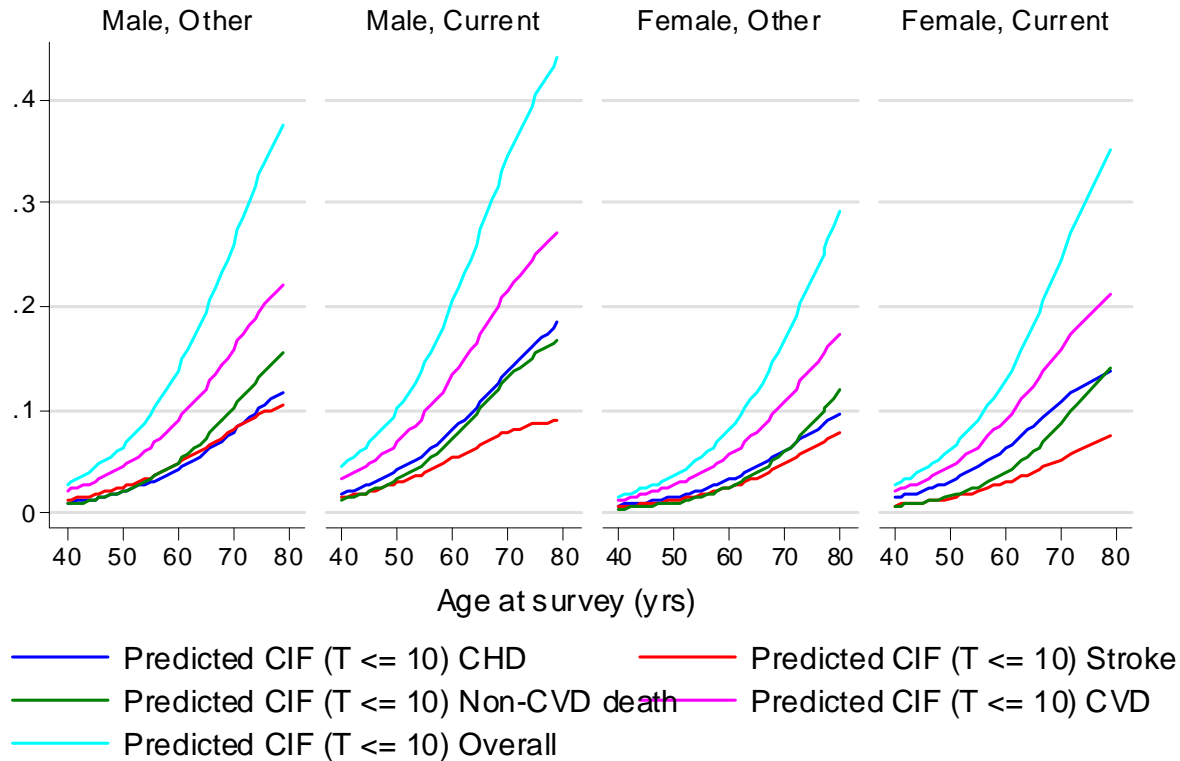
Variable	Obs	Mean	Std. Dev.	Min	Max
ages	817	57.85982	11.38797	40	80.096
smallbin	817	.2423501	.4287675	0	1

Calculating cumulative incidence of ep_chdmi at time = 10
in 2 strata

```
Stratum = 1 (Male), n = 398, and 78 linear predictors
..... 50
.....
Stratum = 2 (Female), n = 419, and 71 linear predictors
..... 50
.....
```

... and similarly for each of Model 2 (Stroke) and Model 3 (Non-CVD death)

DIY combined summary of the predicted CIFs vs. age



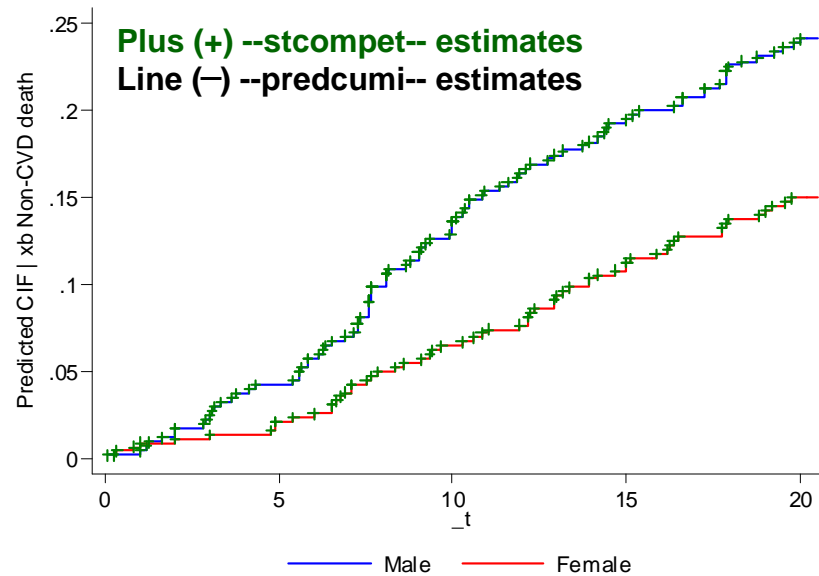
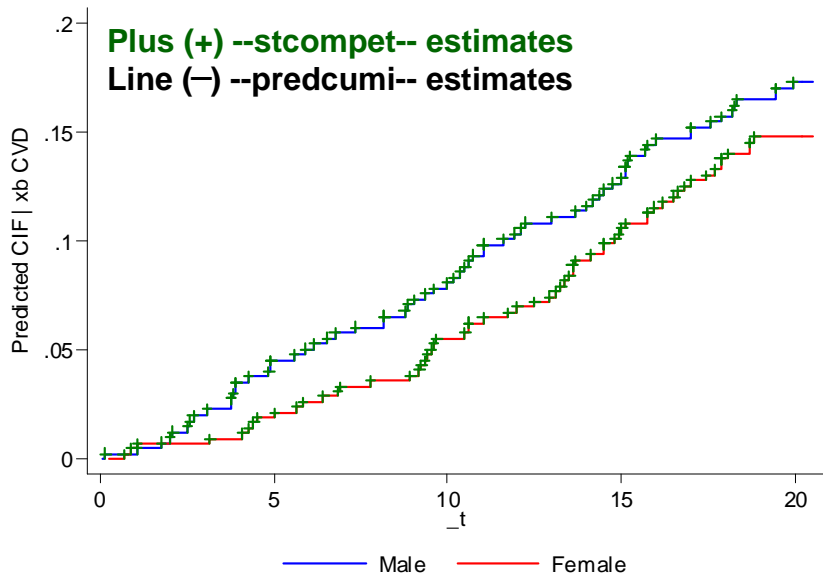
Graphs by Sex and Smoking status

$$\text{Where: } \text{CIF}_{\text{CVD}} = \text{CIF}_{\text{CHD}} + \text{CIF}_{\text{Stroke}}$$

$$\text{CIF}_{\text{Overall}} = \text{CIF}_{\text{CHD}} + \text{CIF}_{\text{Stroke}} + \text{CIF}_{\text{Non-CVD}}$$

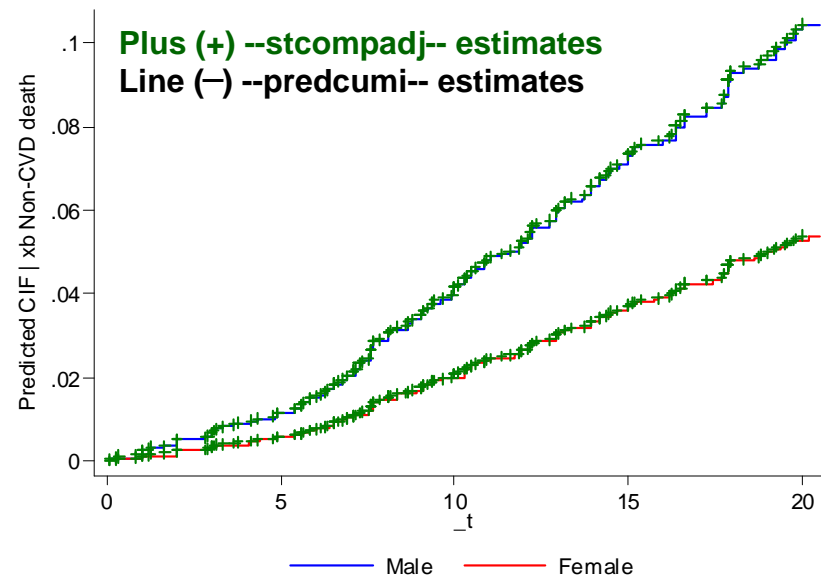
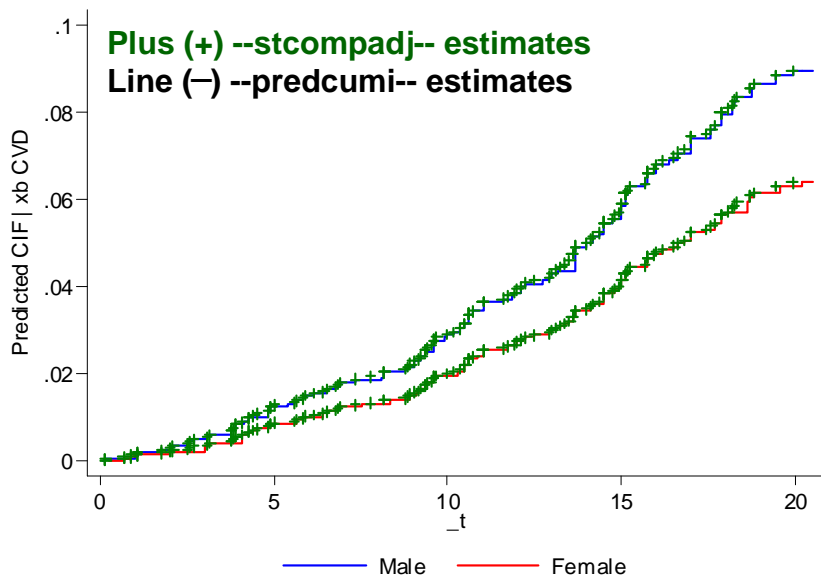
Check against --stcompet-- (unadjusted)

```
. stset duration, failure("cause == 1")
. xi: stcox, strata(sex) estimate nolog
. predcumi, compete("cause == 2")
. stcompet cumi = ci, compet1(2) by(sex)
foreach m of numlist 1 2 {
  overlay cif`m' _t, over(sex) ms(i i) c(J) ylabel(#5, angle(horiz)) ///
  scheme("slcolor_sk") addplot(scatter cumi _t if cause == `m', ms(+))
  name(gr1_cause`m', replace)
}
```



Check against --stcompadj-- (adjusted)

```
. stset duration, failure("cause == 1")
. xi: stcox sexbin ages, nolog
. predcumi, compete("cause==2") adjust(ages = 50)
. local fitopts = "maineffect(sexbin ages) competeffect(sexbin ages) showmod"
. stcompadj sexbin=0 ages = 50, compet(2) gen(cifmain0 cifcomp0) `fitopts'
. stcompadj sexbin=1 ages = 50, compet(2) gen(cifmain1 cifcomp1) `fitopts'
foreach m of numlist 1 2 {
  local compvars = cond(`m' == 1, "cifmain*", "cifcomp*")
  overlay cif`m' _t, over(sexbin) ms(i i) c(J) ylabel(#5, angle(horiz)) ///
  scheme("slcolor_sk") addplot(scatter `compvars' _t if cause == `m', ms(+ +))
  name (gr3_cause`m', replace)
}
```



Check against --stcrreg-- (adjusted)

```
. stset duration, failure("cause == 1")
. xi: stcox sexbin ages, nolog
      failure _d:  cause == 1
      analysis time _t:  duration
Cox regression -- Breslow method for ties
No. of subjects =          817                Number of obs   =          817
No. of failures =          129
Time at risk    = 13682.08051
Log likelihood  = -780.89034
LR chi2(2)      =          111.93
Prob > chi2     =           0.0000
```

```
-----
      _t | Haz. Ratio   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      sexbin |   .681508   .1204439   -2.17   0.030   .4819888   .9636182
      ages   |   1.092081   .0096398    9.98   0.000   1.07335   1.111139
-----
```

```
. predcumi, compete("cause==2") adjust(ages = 50)
```

```
<output, omitted >
```

```
. foreach m of numlist 1 2 {
  levelsof cause if !inlist(cause, 0, `m'), local(compevent)
  stset duration, failure("cause == `m'")
  xi: stcrreg sexbin ages, compete(cause == `compevent')
  local atvall = "at1(sexbin = 0 ages = 50) at2(sexbin = 1 ages = 50)"
  forvalues k = 1/1 {
    local gropts = `"'title("") ylabel(#5, angle(horiz))"'
    local atvals "`atval`k'"
    local addplot = "scatter cif`m' _t, ms(+)"
    stcurve, cif `atvals' `gropts' addplot(`addplot') ///
    legend(order(1 2)) name(gr4at`k'_cause`m', replace)
  }
}
```

Check against --stcrreg-- Model 1

```
failure _d: cause == 1
analysis time _t: duration
```

```
Iteration 0: log pseudolikelihood = -821.17236
Iteration 1: log pseudolikelihood = -817.50996
Iteration 2: log pseudolikelihood = -817.48958
Iteration 3: log pseudolikelihood = -817.48958
```

```
Competing-risks regression                No. of obs      =      817
                                           No. of subjects =      817
Failure event   : cause == 1             No. failed      =      129
Competing event: cause == 2             No. competing   =      157
                                           No. censored    =      531
                                           Wald chi2(2)    =      65.94
Log pseudolikelihood = -817.48958        Prob > chi2     =      0.0000
```

_t	SHR	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
sexbin	.8124038	.143884	-1.17	0.241	.5741383	1.149549
ages	1.068916	.0090733	7.85	0.000	1.05128	1.086848

Check against --stcrreg-- Model 2

```
failure _d: cause == 2
analysis time _t: duration
```

```
Iteration 0: log pseudolikelihood = -971.95067
Iteration 1: log pseudolikelihood = -969.36457
Iteration 2: log pseudolikelihood = -969.3526
Iteration 3: log pseudolikelihood = -969.3526
```

Competing-risks regression

```
Failure event : cause == 2
Competing event: cause == 1
```

Log pseudolikelihood = -969.3526

```
No. of obs      =      817
No. of subjects =      817
No. failed      =      157
No. competing   =      129
No. censored    =      531
Wald chi2(2)    =     117.48
Prob > chi2     =      0.0000
```

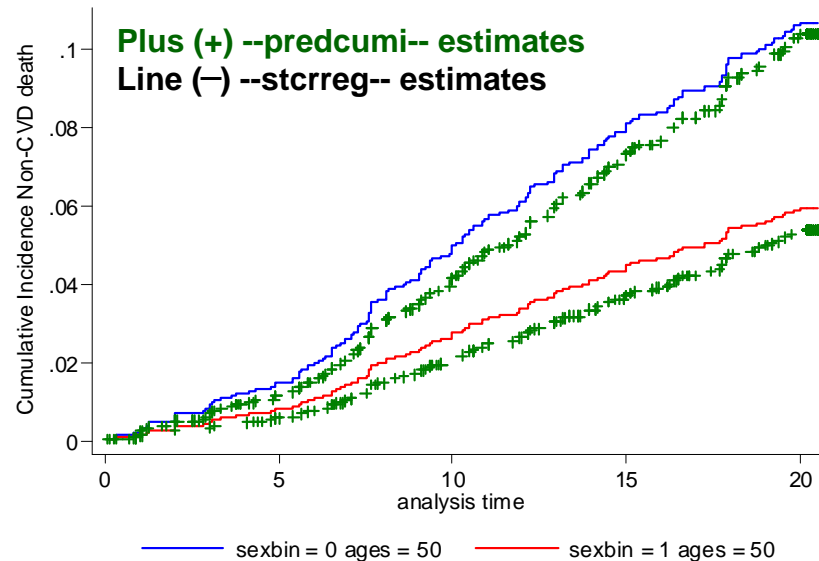
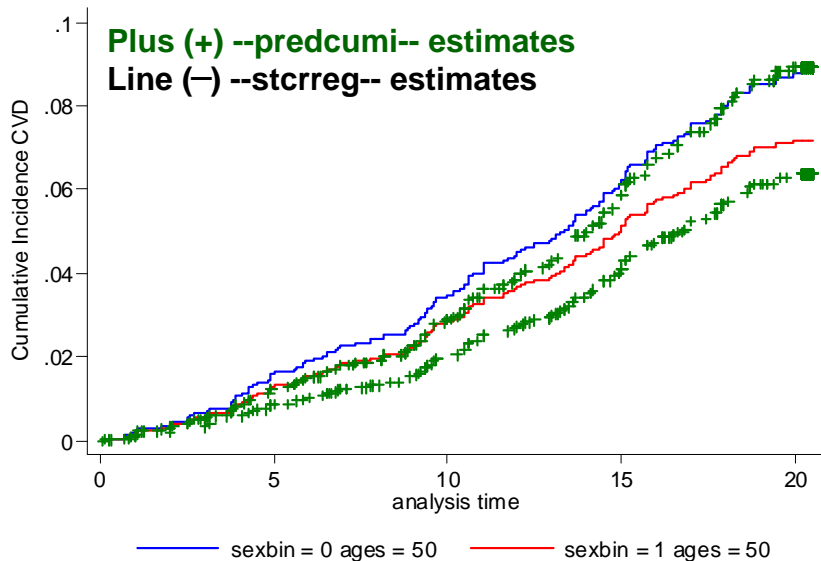
```
-----
```

_t	SHR	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
sexbin	.5422991	.0886548	-3.74	0.000	.393627	.7471242
ages	1.08355	.0084414	10.30	0.000	1.067131	1.100222

```
-----
```

Check against `--stcrreg--`

- Not so good agreement ..., but not a surprise considering the model formulation in `--stcrreg--` is different from `--stcox--`.



Remarks

- **--predcumi--** provides some added functionality not addressed in previous user-written programs, i.e. individual predictions and CIF within landmark times.
 - Ongoing considerations include extension to lifetime-risk calculation with age-as-timescale models.
- **--stcompet--** and **--stcompadj--**, however, do a lot more with respect to covariate adjusted CIF estimates, including confidence intervals, hypothesis testing, and CIF inferences based on flexible parametric models.
 - Also discovered during this meeting that **--stpm2--** has a wrapper function **--stpm2cif--** for adjusted CIF calculations.
- For large datasets, competing risk models based on cause-specific hazards could be more tractable than currently possible with **--stcrreg--** (??)

Acknowledgements

- Mentors
 - S Thompson, J Danesh
- Statistical collaborators
 - S Thompson, I White, L Pennells, A Wood
- ERFC collaborators
 - <http://www.phpc.cam.ac.uk/ceu/research/erfc/studies/>