Simulating simple and complex survival data Stata UK User Group Meeting Cass Business School

11th September 2014

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Background		Discussion	

Outline

- 1. Background
- 2. Motivating dataset
- 3. Simulating survival times from standard distributions
- 4. A general algorithm for generating survival times
- 5. Discussion

Background		Discussion	
Backgroi	und		

- Simulation studies are conducted to assess the performance of current and novel statistical models in pre-defined scenarios
- Guidelines for the reporting of simulation studies in medical research have been published (Burton et al., 2006)
- Many simulation studies involving survival data use the exponential or Weibull models
- Often in clinical trials and population based studies, at least one turning point in the baseline hazard function is observed

Motivating dataset

- ▶ webuse brcancer
- ▶ 686 women diagnosed with breast cancer in Germany
- 246 were randomised to receive hormonal therapy and 440 to receive a placebo
- Outcome of interest is recurrence-free survival, with 299 patients experiencing the event

Analysis

- Weibull proportional hazards model
- Flexible parametric model with 5 degrees of freedom
- Treatment included in both models

Fitted survival functions



Fitted hazard functions



Simulating survival times

Bender et al. (2005) provided a simple and efficient method to simulate survival times from standard parametric distributions

$$\begin{split} h(t|X) &= h_0(t) \exp(X\beta), \qquad H(t|X) = H_0(t) \exp(X\beta) \\ S(t|X) &= \exp[-H(t|X)], \qquad F(t|X) = 1 - \exp[-H(t|X)] \end{split}$$

.

	Survival simulation	Discussion	

If we let T be the simulated survival time

$$F(T|X) = 1 - \exp[-H(T|X)] = u,$$
 where $u \sim U(0,1)$

and

$$S(T|X) = 1 - u$$
 (or equivalently $= u$)

This can then simply be re-arranged and solved for T

$$T = H_0^{-1}[-\log(u)\exp(-X\beta)]$$

For example in Stata

```
. //simulate 1000 survival times
```

```
. set obs 1000
```

```
obs was 0, now 1000
```

- . //set seed for reproducibility
- . set seed 398894
- . //get uniform draws, representing centiles
- . gen u = runiform()
- . //geenrated a binary treatment group indicator
- . gen treatment = runiform()>0.5
- . //Weibull baseline parameters
- . local lambda = 0.1
- . local gamma = 1.2
- . //treatment effect
- . local loghr = 0.7
- . //simulate survival times from Weibull PH model
- . gen stimes = (-log(u)/(`lambda´*exp(`loghr`*treatment)))^(1/`gamma´)

survsim (from SSC)

survsim newvarname1 [newvarname2] [, options]

- distribution(exp|gomp|weib)
- lambda(#), gamma(#)
- > covariates(varname # [varname #] ...)
- ▶ tde(varname # [varname #] ...)
- maxtime(#)
- . survsim stime event, dist(weib) lambda(0.1)
- > gamma(1.2) cov(treatment 0.7)

Recent use of survival simulation

- Paul Lambert and I recently proposed a general parametric framework for survival analysis, implemented in stgenreg (Crowther and Lambert, 2013b, 2014)
- Reviews raised questions about benefits/pitfalls compared to the Cox model
- We set out to compare the efficiency of the Kaplan-Meier estimate of survival with a parametric function using splines, when data is sparse in the right tail

Core of simulation program

- . //simulate from a Weibull distribution
- . survsim stime died, lambda(0.2) gamma(1.3) maxt(5)
- . //censoring times
- . gen cens = runiform()*6
- . replace died = 0 if cens<stime
- . replace stime = cens if cens<stime
- . stset stime, f(died=1)
- . //KM estimate
- . sts gen s1 = s sells = se(lls) lb = lb(s) ub = ub(s)
- . //Fit parametric model
- . stgenreg, loghaz([xb]) xb(#rcs(df(3)))
- . //Get predicted survival at 4 and 5 years
- . range t45 4 5 2
- . predict surv, survival timevar(t45) ci

	Survival simulation	Discussion	
Results			

Table : Bias and mean squared error of log(-log(S(t))) at 4 and 5 years.

Time		Kaplan-Meier	Parametric model	
Aveare	Bias	-0.0019	-0.0038	
4 years	MSE	0.1251	0.1100	
Evente	Bias	0.0066	0.0063	
5 years	MSE	0.1565	0.1481	
Median $\#$ events = 101				
Median $\#$ events in final year $= 5$				

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Benefits of the Bender et al. (2005) approach

- Extremely easy to implement
- Quite often we simulate survival times and then apply Cox models -> baseline hazard from which we simulate is irrelevent
- What if we wish to simulate from a more complex and biologically plausible underlying hazard function?
- There is a growing interest in parametric survival models (Royston and Lambert, 2011; Crowther and Lambert, 2014)

Limitations with simulating survival times from standard distributions with proportional hazards

$$T = H_0^{-1}[-\log(u)\exp(-X\beta)]$$

- Must be able to integrate the hazard function in order to calculate the cumulative hazard function
- We then must be able to invert the cumulative hazard function to obtain the simulated survival time

Simulating from a more complex baseline hazard function

We can use a mixture of parametric distributions

$$S_0(t) = pS_{01}(t) + (1-p)S_{02}(t)$$
 (1)

For example a 2-component mixture Weibull

$$S_0(t) = p \exp(-\lambda_1 t^{\gamma_1}) + (1-p) \exp(-\lambda_2 t^{\gamma_2})$$
(2)
with $0 \le p \le 1$, and $\lambda_1, \lambda_2, \gamma_1, \gamma_2 > 0$



		Survivar Sintalación		
Incorpor	ating proporti	onal hazards give	s us a survival func	tion
S(t)	$= [p \exp(-\lambda_1)]$	$(1+t^{\gamma_1})+(1-p)\exp(t^{\gamma_1})$ ex	$p(-\lambda_2 t^{\gamma_2})]^{\exp(X_{eta})}$	(3)
This mo Attempt	del is impleme ing to apply t	ented in the stmi he inversion meth	.x command from S 10d, gives	SSC.
	S(t) =	u, where u ~	- <i>U</i> (0, 1)	(4)
which c	annot be re-ar	ranged to directly	i solve for t .	

To solve we can apply iterative root finding techniques, such as Newton-Raphson iterations, or Brent's univariate root finder. I favour the latter, using mm_root() from Ben Jann's moremata (Jann, 2005)

	Survival simulation	Discussion	
survsim			

survsim newvarname1 [newvarname2] [, options]

- ▶ mixture
- distribution(exp|gomp|weib)
- lambdas(#), gammas(#)
- > covariates(varname # [varname #] ...)
- maxtime(#)
- . survsim stime event, mixture dist(weib)
- > lambdas(0.1 0.2) gammas(1.2 0.5) p(0.3)

Simulating survival times when the cumulative hazard doesn't have a closed form expression - joint model data

$$h(t) = h_0(t) \exp \left[X \beta + \alpha m(t) \right]$$

where

$$m(t) = \beta_{0i} + \beta_{1i}t$$

- To obtain the cumulative hazard function we require numerical integration
- We then require root finding techniques to solve for the simulated survival time, t

Numerical integration

$$\int_{-1}^1 g(x) \mathrm{d}x = \int_{-1}^1 W(x) g(x) \mathrm{d}x \approx \sum_{i=1}^m w_i g(x_i)$$

where W(x) is a known weighting function and g(x) can be approximated by a polynomial function.

$$\int_{t_{0i}}^{t_i} h(x) dx = \frac{t_i - t_{0i}}{2} \int_{-1}^{1} h\left(\frac{t_i - t_{0i}}{2}x + \frac{t_{0i} + t_i}{2}\right) dx$$
$$\approx \frac{t_i - t_{0i}}{2} \sum_{i=1}^{m} w_i h\left(\frac{t_i - t_{0i}}{2}x_i + \frac{t_{0i} + t_i}{2}\right)$$

	Survival simulation	Discussion	
survsim			

survsim newvarname1 [newvarname2] [, options]

- [log]hazard()
- [log]cumhazard()
- > nodes(#)
- > covariates(varname # [varname #] ...)
- ▶ tde(varname # [varname #] ...)
- tdefunction()
- centol(#)
- maxtime(#)
- . survsim stime event, hazard(0.1:*1.2:*t:(1.1:-1))

Simulating survival data - recap



General survival simulation

Given a well-defined hazard function, h(t), this two-stage algorithm involving

- 1. Numerical integration
- 2. Root-finding

provides a framework for general survival simulation which can incorporate:

- Practically any user-defined baseline hazard function
- Time-varying covariates
- Time-dependent effects
- Delayed entry
- Extends to competing risks, frailty etc.

	Survival simulation	Discussion	

Examples

Fractional polynomial baseline
survsim stime event, logh(-18 :+
7.3:*log(#t):-11.5:*#t:^(0.5):*log(#t))

	Survival simulation	Discussion	

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Fractional polynomial baseline survsim stime event, logh(-18 :+ 7.3:*log(#t):-11.5:*#t:^(0.5):*log(#t))

Non-proportional hazards

survsim stime event, logh(-18 :+
7.3:*log(#t):-11.5:*#t:^(0.5):*log(#t)) cov(trt -0.7)
tde(trt 1) tdefunc(0.01:*t :+ 0.4:*log(t))

		Survival simulation	Discussion	
Examp	les			
▶ .	loınt model data (t	ime-varying cova	riate)	
	//Simulate 1000 survi set obs 1000	val times		
	<pre>//Define the associat local alpha = 0.25</pre>	ion between the bio	marker and survival	
	<pre>//Generate the random gen b0 = rnormal(0,1)</pre>	n intercept and rand	om slopes	
	<pre>gen b1 = rnormal(1,0.</pre>	5)		
	<pre>survsim stime event, :+ `alpha´ :* (b0 :+</pre>	loghazard(-2.3:+2:* b1 :* #t)) maxt(5)	#t:-#t:^(2):+0.12:*#t	:^3
	<pre>//Generate observed b gen id = _n</pre>	piomarker values at	times 0, 1, 2, 3 , 4	years
	expand 5			
	bys id: gen meastime	= _n-1		
	<pre>//Remove observations bys id: drop if meast</pre>	s after event or cen cime>=stime	soring time	

. //Generate observed biomarker values incorporating measurement error $% \left({{\left({{{{\bf{G}}_{{\bf{G}}}}} \right)}_{{\bf{G}}}} \right)$

. gen response = b0 + b1 + meastime + rnormal(0, 0.5)

Practical advice

 Although computation time is often minimal, it may be of use to simulate your 1000 datasets, say, before applying any model fits

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- With the numerical integration, it is important to assess the approximation by setting a seed and using an increasing number of quadrature points

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 We have described a general framework for the generation of survival data, incorporating any combination of complex hazard functions, time-dependent effects, time-varying covariates, delayed entry, random effects and covariates measured with error (Crowther and Lambert, 2013a)

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 - As the procedure relies on numerical integration, it is important to establish the consistency of the simulated survival times by setting a seed and using an increasing number of quadrature nodes
 - You can also specify a user-defined [log] cumulative hazard function (Royston, 2012) (stsurvsim)
 - Simulating from a fitted model (or observed censoring) distribution) can be particularly useful (Royston, 2012)

		Discussion	References
References			

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