Imperial College London

Somers' D: A common currency for associations

Roger B. Newson r.newson@imperial.ac.uk http://www.imperial.ac.uk/nhli/r.newson/

Department of Primary Care and Public Health, Imperial College London

21st UK Stata Users' Group Meeting, 10–11 September, 2015 Downloadable from the conference website at http://ideas.repec.org/s/boc/usug15.html

- ▶ We assume that pairs of (*X*, *Y*)-pairs (*X_i*, *Y_i*) and (*X_j*, *Y_j*) are sampled from a population of (*X*, *Y*)-pairs, under a specified sampling scheme.
- Kendall's τ_a is defined (symmetrically) as the expectation

$$\tau_{XY} = \mathrm{E}[\mathrm{sign}(X_i - X_j)\mathrm{sign}(Y_i - Y_j)],$$

or as the difference between the probabilities of **concordance** and **discordance** between the two (X, Y)-pairs.

► **Somers'** *D* is defined (asymmetrically) as the ratio

$$D(Y|X) = \tau_{XY}/\tau_{XX},$$

or as the difference between the two corresponding *conditional* probabilities, given that one *X*–value is known to be larger than the other *X*–value.

These definitions can be extended to cases where the X-values and/or the Y-values may be left-censored and/or right-censored.

- ► We assume that pairs of (X, Y)-pairs (X_i, Y_i) and (X_j, Y_j) are sampled from a population of (X, Y)-pairs, under a specified sampling scheme.
- Kendall's τ_a is defined (symmetrically) as the expectation

$$\tau_{XY} = \mathrm{E}[\mathrm{sign}(X_i - X_j)\mathrm{sign}(Y_i - Y_j)],$$

or as the difference between the probabilities of **concordance** and **discordance** between the two (X, Y)-pairs.

▶ **Somers'** *D* is defined (asymmetrically) as the ratio

$$D(Y|X) = \tau_{XY}/\tau_{XX},$$

or as the difference between the two corresponding *conditional* probabilities, given that one *X*–value is known to be larger than the other *X*–value.

These definitions can be extended to cases where the X-values and/or the Y-values may be left-censored and/or right-censored.

- ► We assume that pairs of (X, Y)-pairs (X_i, Y_i) and (X_j, Y_j) are sampled from a population of (X, Y)-pairs, under a specified sampling scheme.
- Kendall's τ_a is defined (symmetrically) as the expectation

$$\tau_{XY} = \mathrm{E}[\mathrm{sign}(X_i - X_j)\mathrm{sign}(Y_i - Y_j)],$$

or as the difference between the probabilities of **concordance** and **discordance** between the two (X, Y)-pairs.

► **Somers'** *D* is defined (asymmetrically) as the ratio

$$D(Y|X) = \tau_{XY}/\tau_{XX},$$

or as the difference between the two corresponding *conditional* probabilities, given that one *X*–value is known to be larger than the other *X*–value.

These definitions can be extended to cases where the X-values and/or the Y-values may be left-censored and/or right-censored.

- ► We assume that pairs of (X, Y)-pairs (X_i, Y_i) and (X_j, Y_j) are sampled from a population of (X, Y)-pairs, under a specified sampling scheme.
- Kendall's τ_a is defined (symmetrically) as the expectation

$$\tau_{XY} = \mathrm{E}[\mathrm{sign}(X_i - X_j)\mathrm{sign}(Y_i - Y_j)],$$

or as the difference between the probabilities of **concordance** and **discordance** between the two (X, Y)-pairs.

► **Somers'** *D* is defined (asymmetrically) as the ratio

$$D(Y|X) = \tau_{XY}/\tau_{XX},$$

or as the difference between the two corresponding *conditional* probabilities, given that one *X*-value is known to be larger than the other *X*-value.

These definitions can be extended to cases where the X-values and/or the Y-values may be left-censored and/or right-censored

- ► We assume that pairs of (X, Y)-pairs (X_i, Y_i) and (X_j, Y_j) are sampled from a population of (X, Y)-pairs, under a specified sampling scheme.
- Kendall's τ_a is defined (symmetrically) as the expectation

$$\tau_{XY} = \mathrm{E}[\mathrm{sign}(X_i - X_j)\mathrm{sign}(Y_i - Y_j)],$$

or as the difference between the probabilities of **concordance** and **discordance** between the two (X, Y)-pairs.

► **Somers'** *D* is defined (asymmetrically) as the ratio

$$D(Y|X) = \tau_{XY}/\tau_{XX},$$

or as the difference between the two corresponding *conditional* probabilities, given that one *X*-value is known to be larger than the other *X*-value.

► These definitions can be extended to cases where the X-values and/or the Y-values may be left-censored and/or right-censored.

- The package somersd[4] can be downloaded from SSC, and estimates Somers' D or Kendall's τ_a .
- ► These **rank parameters** can be interval–estimated under a wide range of sampling schemes, with or without censorship.
- The cluster() option allows for sampling clusters from a population of clusters.
- Sampling-probability weights are enabled by pweights, allowing us to estimate target-population parameters from a sampled population by direct standardization.
- ► The wstrata() and bstrata() options allow for restriction to comparisons within or between strata.
- And the parameter estimates are saved as Stata estimation results, using delta–jackknife variances and a choice of Normalizing and/or variance–stabilizing transformations.

- The package somersd[4] can be downloaded from SSC, and estimates Somers' D or Kendall's τ_a .
- ► These **rank parameters** can be interval–estimated under a wide range of sampling schemes, with or without censorship.
- The cluster() option allows for sampling clusters from a population of clusters.
- Sampling-probability weights are enabled by pweights, allowing us to estimate target-population parameters from a sampled population by direct standardization.
- The wstrata() and bstrata() options allow for restriction to comparisons within or between strata.
- And the parameter estimates are saved as Stata estimation results, using delta–jackknife variances and a choice of Normalizing and/or variance–stabilizing transformations.

- The package somersd[4] can be downloaded from SSC, and estimates Somers' D or Kendall's τ_a .
- ► These **rank parameters** can be interval–estimated under a wide range of sampling schemes, with or without censorship.
- The cluster() option allows for sampling clusters from a population of clusters.
- Sampling-probability weights are enabled by pweights, allowing us to estimate target-population parameters from a sampled population by direct standardization.
- ► The wstrata() and bstrata() options allow for restriction to comparisons within or between strata.
- And the parameter estimates are saved as Stata estimation results, using delta–jackknife variances and a choice of Normalizing and/or variance–stabilizing transformations.

- The package somersd[4] can be downloaded from SSC, and estimates Somers' D or Kendall's τ_a .
- ► These **rank parameters** can be interval–estimated under a wide range of sampling schemes, with or without censorship.
- The cluster() option allows for sampling clusters from a population of clusters.
- Sampling-probability weights are enabled by pweights, allowing us to estimate target-population parameters from a sampled population by direct standardization.
- The wstrata() and bstrata() options allow for restriction to comparisons within or between strata.
- And the parameter estimates are saved as Stata estimation results, using delta–jackknife variances and a choice of Normalizing and/or variance–stabilizing transformations.

- The package somersd[4] can be downloaded from SSC, and estimates Somers' D or Kendall's τ_a .
- ► These **rank parameters** can be interval–estimated under a wide range of sampling schemes, with or without censorship.
- The cluster() option allows for sampling clusters from a population of clusters.
- Sampling-probability weights are enabled by pweights, allowing us to estimate target-population parameters from a sampled population by direct standardization.
- The wstrata() and bstrata() options allow for restriction to comparisons within or between strata.
- And the parameter estimates are saved as Stata estimation results, using delta–jackknife variances and a choice of Normalizing and/or variance–stabilizing transformations.

- The package somersd[4] can be downloaded from SSC, and estimates Somers' D or Kendall's τ_a .
- ► These **rank parameters** can be interval–estimated under a wide range of sampling schemes, with or without censorship.
- The cluster() option allows for sampling clusters from a population of clusters.
- Sampling-probability weights are enabled by pweights, allowing us to estimate target-population parameters from a sampled population by direct standardization.
- ► The wstrata() and bstrata() options allow for restriction to comparisons within or between strata.
- And the parameter estimates are saved as Stata estimation results, using delta–jackknife variances and a choice of Normalizing and/or variance–stabilizing transformations.

- The package somersd[4] can be downloaded from SSC, and estimates Somers' D or Kendall's τ_a .
- ► These **rank parameters** can be interval–estimated under a wide range of sampling schemes, with or without censorship.
- The cluster() option allows for sampling clusters from a population of clusters.
- Sampling-probability weights are enabled by pweights, allowing us to estimate target-population parameters from a sampled population by direct standardization.
- ► The wstrata() and bstrata() options allow for restriction to comparisons within or between strata.
- And the parameter estimates are saved as Stata estimation results, using delta–jackknife variances and a choice of Normalizing and/or variance–stabilizing transformations.

- ► Somers' *D* has the useful feature that a larger Somers' *D* cannot be secondary to a smaller Somers' *D* in the same direction.
- ▶ That is to say, if a positive Somers' *D* between *X* and *Y* is secondary to a positive Somers' *D* between *X* and *W*, then, from the argument of Newson (1987)[5], we must have the inequality

- ► So, if a confidence interval for D(Y|X) D(W|X) contains only positive values, then a positive association of X with Y cannot be caused by a positive association of X with W.
- Such a confidence interval can be produced using lincom or nlcom after somersd.
- ▶ This is especially useful if *Y* is an outcome, *X* is an exposure, and *W* is a positive predictor of *X*.

- ► Somers' *D* has the useful feature that a larger Somers' *D* cannot be secondary to a smaller Somers' *D* in the same direction.
- ▶ That is to say, if a positive Somers' *D* between *X* and *Y* is secondary to a positive Somers' *D* between *X* and *W*, then, from the argument of Newson (1987)[5], we must have the inequality

- ► So, if a confidence interval for D(Y|X) D(W|X) contains only positive values, then a positive association of X with Y cannot be caused by a positive association of X with W.
- Such a confidence interval can be produced using lincom or nlcom after somersd.
- ► This is especially useful if *Y* is an outcome, *X* is an exposure, and *W* is a positive predictor of *X*.

- ► Somers' *D* has the useful feature that a larger Somers' *D* cannot be secondary to a smaller Somers' *D* in the same direction.
- ► That is to say, if a positive Somers' D between X and Y is secondary to a positive Somers' D between X and W, then, from the argument of Newson (1987)[5], we must have the inequality

- ► So, if a confidence interval for D(Y|X) D(W|X) contains only positive values, then a positive association of X with Y cannot be caused by a positive association of X with W.
- Such a confidence interval can be produced using lincom or nlcom after somersd.
- ► This is especially useful if *Y* is an outcome, *X* is an exposure, and *W* is a positive predictor of *X*.

- ► Somers' *D* has the useful feature that a larger Somers' *D* cannot be secondary to a smaller Somers' *D* in the same direction.
- ► That is to say, if a positive Somers' D between X and Y is secondary to a positive Somers' D between X and W, then, from the argument of Newson (1987)[5], we must have the inequality

- ► So, if a confidence interval for D(Y|X) D(W|X) contains only positive values, then a positive association of X with Y cannot be caused by a positive association of X with W.
- Such a confidence interval can be produced using lincom or nlcom after somersd.
- ► This is especially useful if *Y* is an outcome, *X* is an exposure, and *W* is a positive predictor of *X*.

- ► Somers' *D* has the useful feature that a larger Somers' *D* cannot be secondary to a smaller Somers' *D* in the same direction.
- ► That is to say, if a positive Somers' D between X and Y is secondary to a positive Somers' D between X and W, then, from the argument of Newson (1987)[5], we must have the inequality

- ► So, if a confidence interval for D(Y|X) D(W|X) contains only positive values, then a positive association of X with Y cannot be caused by a positive association of X with W.
- Such a confidence interval can be produced using lincom or nlcom after somersd.
- ► This is especially useful if *Y* is an outcome, *X* is an exposure, and *W* is a positive predictor of *X*.

- ► Somers' *D* has the useful feature that a larger Somers' *D* cannot be secondary to a smaller Somers' *D* in the same direction.
- ► That is to say, if a positive Somers' D between X and Y is secondary to a positive Somers' D between X and W, then, from the argument of Newson (1987)[5], we must have the inequality

- ► So, if a confidence interval for D(Y|X) D(W|X) contains only positive values, then a positive association of X with Y cannot be caused by a positive association of X with W.
- Such a confidence interval can be produced using lincom or nlcom after somersd.
- ► This is especially useful if *Y* is an outcome, *X* is an exposure, and *W* is a positive predictor of *X*.

- ▶ In particular, W may be a **propensity score**[1], predicting X from a list of multiple confounders V_1, \ldots, V_K .
- ▶ In the 21st-century Rubin method[6], the propensity score is a predictor from a regression model in the exposure, which we find in the joint distribution of the exposure and the confounders.
- ► We then add in the outcome data, and use the propensity score to estimate a propensity-adjusted exposure effect on the outcome.
- This adjusted exposure effect is typically interpreted as an exposure effect in a fantasy target population, with real-world marginal exposure and propensity distributions, but no exposure-propensity association.
- ► This is usually done using propensity weighting, propensity matching, or propensity stratification, in regression models of *Y* with respect to *X*. *However*...

- ► In particular, W may be a propensity score[1], predicting X from a list of multiple confounders V₁,..., V_K.
- ▶ In the 21st-century Rubin method[6], the propensity score is a predictor from a regression model in the exposure, which we find in the joint distribution of the exposure and the confounders.
- ► We then add in the outcome data, and use the propensity score to estimate a propensity-adjusted exposure effect on the outcome.
- This adjusted exposure effect is typically interpreted as an exposure effect in a fantasy target population, with real-world marginal exposure and propensity distributions, but no exposure-propensity association.
- ► This is usually done using propensity weighting, propensity matching, or propensity stratification, in regression models of *Y* with respect to *X*. *However*...

- ► In particular, W may be a propensity score[1], predicting X from a list of multiple confounders V₁,..., V_K.
- ► In the 21st-century Rubin method[6], the propensity score is a predictor from a regression model in the exposure, which we find in the joint distribution of the exposure and the confounders.
- ► We then add in the outcome data, and use the propensity score to estimate a propensity-adjusted exposure effect on the outcome.
- This adjusted exposure effect is typically interpreted as an exposure effect in a fantasy target population, with real-world marginal exposure and propensity distributions, but no exposure-propensity association.
- ► This is usually done using propensity weighting, propensity matching, or propensity stratification, in regression models of *Y* with respect to *X*. *However*...

- ► In particular, W may be a propensity score[1], predicting X from a list of multiple confounders V₁,..., V_K.
- ► In the 21st-century Rubin method[6], the propensity score is a predictor from a regression model in the exposure, which we find in the joint distribution of the exposure and the confounders.
- ► We then add in the outcome data, and use the propensity score to estimate a propensity-adjusted exposure effect on the outcome.
- This adjusted exposure effect is typically interpreted as an exposure effect in a fantasy **target population**, with real-world marginal exposure and propensity distributions, but no exposure-propensity association.
- ► This is usually done using propensity weighting, propensity matching, or propensity stratification, in regression models of *Y* with respect to *X*. *However*...

- ► In particular, W may be a propensity score[1], predicting X from a list of multiple confounders V₁,..., V_K.
- ► In the 21st-century Rubin method[6], the propensity score is a predictor from a regression model in the exposure, which we find in the joint distribution of the exposure and the confounders.
- ► We then add in the outcome data, and use the propensity score to estimate a propensity-adjusted exposure effect on the outcome.
- This adjusted exposure effect is typically interpreted as an exposure effect in a fantasy target population, with real-world marginal exposure and propensity distributions, but no exposure-propensity association.
- ► This is usually done using propensity weighting, propensity matching, or propensity stratification, in regression models of *Y* with respect to *X*. *However*...

- ► In particular, W may be a propensity score[1], predicting X from a list of multiple confounders V₁,..., V_K.
- ► In the 21st-century Rubin method[6], the propensity score is a predictor from a regression model in the exposure, which we find in the joint distribution of the exposure and the confounders.
- ► We then add in the outcome data, and use the propensity score to estimate a propensity-adjusted exposure effect on the outcome.
- This adjusted exposure effect is typically interpreted as an exposure effect in a fantasy target population, with real-world marginal exposure and propensity distributions, but no exposure-propensity association.
- ► This is usually done using propensity weighting, propensity matching, or propensity stratification, in regression models of *Y* with respect to *X*. *However*...

- ... a *possible* alternative candidate for the adjusted exposure effect is a **propensity–weighted Somers'** D(Y|X).
- ► This uses **sampling-probability weights**, generated from the exposure *X* and the propensity score *W*, so that subjects with high exposure and low propensity, or low exposure and high propensity, are weighted upwards to remove the association between *W* and *X*.
- Before we add in the outcome data, we might want to be sure that our propensity weights are indeed removing this association.
- ► If the propensity-weighted D(W|X) is close to zero, then we might be confident that a larger propensity-weighted D(Y|X) cannot be secondary to it.
- ► And, if the propensity-weighted D(W|X) is not close to zero, then we have diagnosed a problem with non-overlap, because there are not enough high-exposure low-propensity and low-exposure high-propensity subjects for us to weight upwards.

Somers' D: A common currency for associations

- ... a *possible* alternative candidate for the adjusted exposure effect is a **propensity-weighted Somers'** D(Y|X).
- ► This uses **sampling-probability weights**, generated from the exposure *X* and the propensity score *W*, so that subjects with high exposure and low propensity, or low exposure and high propensity, are weighted upwards to remove the association between *W* and *X*.
- Before we add in the outcome data, we might want to be sure that our propensity weights are indeed removing this association.
- ► If the propensity-weighted D(W|X) is close to zero, then we might be confident that a larger propensity-weighted D(Y|X) cannot be secondary to it.
- ► And, if the propensity-weighted D(W|X) is not close to zero, then we have diagnosed a problem with non-overlap, because there are not enough high-exposure low-propensity and low-exposure high-propensity subjects for us to weight upwards.

- ... a *possible* alternative candidate for the adjusted exposure effect is a **propensity-weighted Somers'** D(Y|X).
- ► This uses **sampling-probability weights**, generated from the exposure *X* and the propensity score *W*, so that subjects with high exposure and low propensity, or low exposure and high propensity, are weighted upwards to remove the association between *W* and *X*.
- Before we add in the outcome data, we might want to be sure that our propensity weights are indeed removing this association.
- ► If the propensity-weighted D(W|X) is close to zero, then we might be confident that a larger propensity-weighted D(Y|X) cannot be secondary to it.
- ► And, if the propensity-weighted D(W|X) is not close to zero, then we have diagnosed a problem with non-overlap, because there are not enough high-exposure low-propensity and low-exposure high-propensity subjects for us to weight upwards.

- ... a *possible* alternative candidate for the adjusted exposure effect is a **propensity-weighted Somers'** D(Y|X).
- ► This uses **sampling-probability weights**, generated from the exposure *X* and the propensity score *W*, so that subjects with high exposure and low propensity, or low exposure and high propensity, are weighted upwards to remove the association between *W* and *X*.
- Before we add in the outcome data, we might want to be sure that our propensity weights are indeed removing this association.
- ► If the propensity-weighted D(W|X) is close to zero, then we might be confident that a larger propensity-weighted D(Y|X) cannot be secondary to it.
- ► And, if the propensity-weighted D(W|X) is not close to zero, then we have diagnosed a problem with non-overlap, because there are not enough high-exposure low-propensity and low-exposure high-propensity subjects for us to weight upwards.

- ... a *possible* alternative candidate for the adjusted exposure effect is a **propensity-weighted Somers'** D(Y|X).
- ► This uses **sampling-probability weights**, generated from the exposure *X* and the propensity score *W*, so that subjects with high exposure and low propensity, or low exposure and high propensity, are weighted upwards to remove the association between *W* and *X*.
- Before we add in the outcome data, we might want to be sure that our propensity weights are indeed removing this association.
- ► If the propensity-weighted D(W|X) is close to zero, then we might be confident that a larger propensity-weighted D(Y|X) cannot be secondary to it.
- ► And, if the propensity-weighted D(W|X) is not close to zero, then we have diagnosed a problem with non-overlap, because there are not enough high-exposure low-propensity and low-exposure high-propensity subjects for us to weight upwards.

- ... a *possible* alternative candidate for the adjusted exposure effect is a **propensity-weighted Somers'** D(Y|X).
- ► This uses **sampling-probability weights**, generated from the exposure *X* and the propensity score *W*, so that subjects with high exposure and low propensity, or low exposure and high propensity, are weighted upwards to remove the association between *W* and *X*.
- Before we add in the outcome data, we might want to be sure that our propensity weights are indeed removing this association.
- ► If the propensity-weighted D(W|X) is close to zero, then we might be confident that a larger propensity-weighted D(Y|X) cannot be secondary to it.
- ► And, if the propensity-weighted D(W|X) is not close to zero, then we have diagnosed a problem with non-overlap, because there are not enough high-exposure low-propensity and low-exposure high-propensity subjects for us to weight upwards.

- ► Somers' *D* is expressed on a sensible scale from -1 to 1, as it is a difference between probabilities.
- ► *However*, an audience accustomed to other association measures, arising from specific models, may be culture–shocked when presented with arguments using Somers' *D*.
- ► *Fortunately*, under a range of familiar models, there is a one-to-one mapping between Somers' *D* and the association measure defined by the model.
- ► And, even better, these mappings are *nearly* linear (or at least log–linear), at least for Somers' *D* values between -0.5 and 0.5.
- ► This makes Somers' *D* a **common currency** for comparing associations measured using different models.
- We will now examine the currency conversions involved under four familiar example models.

- ► Somers' *D* is expressed on a sensible scale from -1 to 1, as it is a difference between probabilities.
- ► *However*, an audience accustomed to other association measures, arising from specific models, may be culture–shocked when presented with arguments using Somers' *D*.
- ► *Fortunately*, under a range of familiar models, there is a one-to-one mapping between Somers' *D* and the association measure defined by the model.
- ► And, even better, these mappings are *nearly* linear (or at least log–linear), at least for Somers' *D* values between -0.5 and 0.5.
- ► This makes Somers' *D* a **common currency** for comparing associations measured using different models.
- We will now examine the currency conversions involved under four familiar example models.

- ➤ Somers' D is expressed on a sensible scale from -1 to 1, as it is a difference between probabilities.
- ► *However*, an audience accustomed to other association measures, arising from specific models, may be culture–shocked when presented with arguments using Somers' *D*.
- ► *Fortunately*, under a range of familiar models, there is a one-to-one mapping between Somers' *D* and the association measure defined by the model.
- ► And, even better, these mappings are *nearly* linear (or at least log–linear), at least for Somers' *D* values between -0.5 and 0.5.
- ► This makes Somers' *D* a **common currency** for comparing associations measured using different models.
- We will now examine the currency conversions involved under four familiar example models.

- ► Somers' *D* is expressed on a sensible scale from -1 to 1, as it is a difference between probabilities.
- ► *However*, an audience accustomed to other association measures, arising from specific models, may be culture–shocked when presented with arguments using Somers' *D*.
- ► *Fortunately*, under a range of familiar models, there is a one-to-one mapping between Somers' *D* and the association measure defined by the model.
- ► And, even better, these mappings are *nearly* linear (or at least log–linear), at least for Somers' *D* values between -0.5 and 0.5.
- ► This makes Somers' *D* a **common currency** for comparing associations measured using different models.
- We will now examine the currency conversions involved under four familiar example models.

- ► Somers' *D* is expressed on a sensible scale from -1 to 1, as it is a difference between probabilities.
- ► *However*, an audience accustomed to other association measures, arising from specific models, may be culture–shocked when presented with arguments using Somers' *D*.
- ► *Fortunately*, under a range of familiar models, there is a one-to-one mapping between Somers' *D* and the association measure defined by the model.
- ► And, even better, these mappings are *nearly* linear (or at least log–linear), at least for Somers' *D* values between -0.5 and 0.5.
- ► This makes Somers' *D* a **common currency** for comparing associations measured using different models.
- ► We will now examine the currency conversions involved under four familiar example models.
Problem: Not everybody understands Somers' D

- ► Somers' *D* is expressed on a sensible scale from -1 to 1, as it is a difference between probabilities.
- ► *However*, an audience accustomed to other association measures, arising from specific models, may be culture–shocked when presented with arguments using Somers' *D*.
- ► *Fortunately*, under a range of familiar models, there is a one-to-one mapping between Somers' *D* and the association measure defined by the model.
- ► And, even better, these mappings are *nearly* linear (or at least log–linear), at least for Somers' *D* values between -0.5 and 0.5.
- ► This makes Somers' *D* a **common currency** for comparing associations measured using different models.
- We will now examine the currency conversions involved under four familiar example models.

Problem: Not everybody understands Somers' D

- ► Somers' *D* is expressed on a sensible scale from -1 to 1, as it is a difference between probabilities.
- ► *However*, an audience accustomed to other association measures, arising from specific models, may be culture–shocked when presented with arguments using Somers' *D*.
- ► *Fortunately*, under a range of familiar models, there is a one-to-one mapping between Somers' *D* and the association measure defined by the model.
- ► And, even better, these mappings are *nearly* linear (or at least log–linear), at least for Somers' *D* values between -0.5 and 0.5.
- ► This makes Somers' *D* a **common currency** for comparing associations measured using different models.
- We will now examine the currency conversions involved under four familiar example models.

- ▶ In our first (trivial) example, we assume that the variables *X* and *Y* are both binary, with possible values 0 and 1.
- Somers' D(Y|X) is then simply the difference between proportions

 $D(Y|X) = \Pr(Y = 1|X = 1) - \Pr(Y = 1|X = 0).$

So, the "scientific" rank parameter, measuring the strength of associations, is also the "practically–useful" regression parameter, measuring how much good we can do by intervening to change X.

- ► In our first (trivial) example, we assume that the variables *X* and *Y* are both binary, with possible values 0 and 1.
- Somers' D(Y|X) is then simply the difference between proportions

 $D(Y|X) = \Pr(Y = 1|X = 1) - \Pr(Y = 1|X = 0).$

So, the "scientific" rank parameter, measuring the strength of associations, is also the "practically–useful" regression parameter, measuring how much good we can do by intervening to change X.

- ► In our first (trivial) example, we assume that the variables *X* and *Y* are both binary, with possible values 0 and 1.
- Somers' D(Y|X) is then simply the difference between proportions

 $D(Y|X) = \Pr(Y = 1|X = 1) - \Pr(Y = 1|X = 0).$

So, the "scientific" rank parameter, measuring the strength of associations, is also the "practically–useful" regression parameter, measuring how much good we can do by intervening to change X.

- ► In our first (trivial) example, we assume that the variables *X* and *Y* are both binary, with possible values 0 and 1.
- Somers' D(Y|X) is then simply the difference between proportions

$$D(Y|X) = \Pr(Y = 1|X = 1) - \Pr(Y = 1|X = 0).$$

► So, the "scientific" rank parameter, measuring the strength of associations, is also the "practically–useful" regression parameter, measuring how much good we can do by intervening to change X.

- We plot Somers' D on the vertical axis against the difference between proportions on the horizontal axis.
- ► Note that the axes are labelled at multiples of 1/12, including ±1/2, ±1/3 and ±1/4.
- The reason for this will become clear in subsequent examples.



- We plot Somers' D on the vertical axis against the difference between proportions on the horizontal axis.
- Note that the axes are labelled at multiples of 1/12, including ±1/2, ±1/3 and ±1/4.
- The reason for this will become clear in subsequent examples.



- We plot Somers' D on the vertical axis against the difference between proportions on the horizontal axis.
- ► Note that the axes are labelled at multiples of 1/12, including ±1/2, ±1/3 and ±1/4.
- The reason for this will become clear in subsequent examples.



- We plot Somers' D on the vertical axis against the difference between proportions on the horizontal axis.
- ► Note that the axes are labelled at multiples of 1/12, including ±1/2, ±1/3 and ±1/4.
- ► The reason for this will become clear in subsequent examples.



- In this example, we assume that the variable X is binary, and that Y has Normal distributions conditionally on each X-value, with respective means μ_0 and μ_1 , and a common standard deviation (SD) σ .
- Somers' D(Y|X) is then given by

$$D(Y|X) = 2\Phi\left(\frac{\delta}{\sqrt{2}}\right) - 1,$$

- ▶ This relationship is not linear, but sigmoid.
- However, it implies that δ is approximately ±0.954 SDs when Somers' D is ±0.5, and approximately linear in between.

- In this example, we assume that the variable X is binary, and that Y has Normal distributions conditionally on each X-value, with respective means μ_0 and μ_1 , and a common standard deviation (SD) σ .
- Somers' D(Y|X) is then given by

$$D(Y|X) = 2\Phi\left(\frac{\delta}{\sqrt{2}}\right) - 1,$$

- ▶ This relationship is not linear, but sigmoid.
- However, it implies that δ is approximately ±0.954 SDs when Somers' D is ±0.5, and approximately linear in between.

- In this example, we assume that the variable X is binary, and that Y has Normal distributions conditionally on each X-value, with respective means μ_0 and μ_1 , and a common standard deviation (SD) σ .
- Somers' D(Y|X) is then given by

$$D(Y|X) = 2\Phi\left(\frac{\delta}{\sqrt{2}}\right) - 1,$$

- ► This relationship is not linear, but sigmoid.
- However, it implies that δ is approximately ±0.954 SDs when Somers' D is ±0.5, and approximately linear in between.

- In this example, we assume that the variable X is binary, and that Y has Normal distributions conditionally on each X-value, with respective means μ_0 and μ_1 , and a common standard deviation (SD) σ .
- Somers' D(Y|X) is then given by

$$D(Y|X) = 2\Phi\left(\frac{\delta}{\sqrt{2}}\right) - 1,$$

- This relationship is not linear, but sigmoid.
- However, it implies that δ is approximately ±0.954 SDs when Somers' D is ±0.5, and approximately linear in between.

- In this example, we assume that the variable X is binary, and that Y has Normal distributions conditionally on each X-value, with respective means μ_0 and μ_1 , and a common standard deviation (SD) σ .
- Somers' D(Y|X) is then given by

$$D(Y|X) = 2\Phi\left(\frac{\delta}{\sqrt{2}}\right) - 1,$$

- This relationship is not linear, but sigmoid.
- ► However, it implies that δ is approximately ±0.954 SDs when Somers' D is ±0.5, and approximately linear in between.

- This time, the curve is sigmoid, as Somers' D is bounded between -1 and 1.
- However, there is a range of near–linearity between Somers' D values of -0.5 and 0.5.
- At these values, the mean differences are -0.954 and +0.954 SDs, respectively.



- ► This time, the curve is sigmoid, as Somers' *D* is bounded between -1 and 1.
- ► *However*, there is a range of near–linearity between Somers' *D* values of -0.5 and 0.5.
- At these values, the mean differences are -0.954 and +0.954 SDs, respectively.



- ► This time, the curve is sigmoid, as Somers' *D* is bounded between -1 and 1.
- ► However, there is a range of near-linearity between Somers' D values of -0.5 and 0.5.
- At these values, the mean differences are -0.954 and +0.954 SDs, respectively.



- ► This time, the curve is sigmoid, as Somers' *D* is bounded between -1 and 1.
- ► However, there is a range of near-linearity between Somers' D values of -0.5 and 0.5.
- At these values, the mean differences are -0.954 and +0.954 SDs, respectively.



- ► Alternatively, under the same model, we might reverse the roles of *X* and *Y*, making *X* a binary disease–case status variable, *Y* a continuous logistic disease predictor, and [D(Y|X) + 1]/2 the area under the receiver–operating characteristic (ROC) curve[2].
- The interesting *model* parameter may then be the **interquartile** odds ratio (IQOR) between the 75th and 25th percentiles of the *control* distribution of Y, given by

IQOR = exp { $\delta [\Phi^{-1}(0.75) - \Phi^{-1}(0.25)]$ },

where δ is still the mean case–control difference (in SDs).

- ► This time, the curve is nearly log–linear between Somers' *D* values of -0.5 and 0.5.
- And, at these values, the IQORs are approximately 0.276 and 3.621, respectively.
- ► *So*, the IQORs corresponding to intermediate Somers' *D* values can be approximated well by weighted geometric means, interpolated log–linearly between 0.276 and 3.621.

- ► Alternatively, under the same model, we might reverse the roles of X and Y, making X a binary disease-case status variable, Y a continuous logistic disease predictor, and [D(Y|X) + 1]/2 the area under the receiver-operating characteristic (ROC) curve[2].
- The interesting *model* parameter may then be the **interquartile** odds ratio (IQOR) between the 75th and 25th percentiles of the *control* distribution of Y, given by

IQOR = exp { $\delta \left[\Phi^{-1}(0.75) - \Phi^{-1}(0.25) \right]$ },

where δ is still the mean case–control difference (in SDs).

- ► This time, the curve is nearly log–linear between Somers' *D* values of -0.5 and 0.5.
- And, at these values, the IQORs are approximately 0.276 and 3.621, respectively.
- ► *So*, the IQORs corresponding to intermediate Somers' *D* values can be approximated well by weighted geometric means, interpolated log–linearly between 0.276 and 3.621.

- ► Alternatively, under the same model, we might reverse the roles of X and Y, making X a binary disease-case status variable, Y a continuous logistic disease predictor, and [D(Y|X) + 1]/2 the area under the receiver-operating characteristic (ROC) curve[2].
- ► The interesting *model* parameter may then be the **interquartile** odds ratio (IQOR) between the 75th and 25th percentiles of the *control* distribution of *Y*, given by

$$\text{IQOR} = \exp\left\{\delta\left[\Phi^{-1}(0.75) - \Phi^{-1}(0.25)\right]\right\},$$

where δ is still the mean case–control difference (in SDs).

- ► This time, the curve is nearly log–linear between Somers' *D* values of -0.5 and 0.5.
- And, at these values, the IQORs are approximately 0.276 and 3.621, respectively.
- ► *So*, the IQORs corresponding to intermediate Somers' *D* values can be approximated well by weighted geometric means, interpolated log–linearly between 0.276 and 3.621.

- ► Alternatively, under the same model, we might reverse the roles of X and Y, making X a binary disease-case status variable, Y a continuous logistic disease predictor, and [D(Y|X) + 1]/2 the area under the receiver-operating characteristic (ROC) curve[2].
- ► The interesting *model* parameter may then be the **interquartile** odds ratio (IQOR) between the 75th and 25th percentiles of the *control* distribution of *Y*, given by

$$IQOR = \exp \left\{ \delta \left[\Phi^{-1}(0.75) - \Phi^{-1}(0.25) \right] \right\},\label{eq:IQOR}$$

where δ is still the mean case–control difference (in SDs).

- ► This time, the curve is nearly log-linear between Somers' *D* values of -0.5 and 0.5.
- And, at these values, the IQORs are approximately 0.276 and 3.621, respectively.
- ► *So*, the IQORs corresponding to intermediate Somers' *D* values can be approximated well by weighted geometric means, interpolated log–linearly between 0.276 and 3.621.

- ► Alternatively, under the same model, we might reverse the roles of X and Y, making X a binary disease-case status variable, Y a continuous logistic disease predictor, and [D(Y|X) + 1]/2 the area under the receiver-operating characteristic (ROC) curve[2].
- ► The interesting *model* parameter may then be the **interquartile** odds ratio (IQOR) between the 75th and 25th percentiles of the *control* distribution of *Y*, given by

$$IQOR = \exp \left\{ \delta \left[\Phi^{-1}(0.75) - \Phi^{-1}(0.25) \right] \right\},\label{eq:IQOR}$$

where δ is still the mean case–control difference (in SDs).

- ► This time, the curve is nearly log-linear between Somers' *D* values of -0.5 and 0.5.
- And, at these values, the IQORs are approximately 0.276 and 3.621, respectively.
- ► *So*, the IQORs corresponding to intermediate Somers' *D* values can be approximated well by weighted geometric means, interpolated log–linearly between 0.276 and 3.621.

- ► Alternatively, under the same model, we might reverse the roles of X and Y, making X a binary disease-case status variable, Y a continuous logistic disease predictor, and [D(Y|X) + 1]/2 the area under the receiver-operating characteristic (ROC) curve[2].
- ► The interesting *model* parameter may then be the **interquartile** odds ratio (IQOR) between the 75th and 25th percentiles of the *control* distribution of *Y*, given by

$$IQOR = \exp \left\{ \delta \left[\Phi^{-1}(0.75) - \Phi^{-1}(0.25) \right] \right\},\label{eq:IQOR}$$

where δ is still the mean case–control difference (in SDs).

- ► This time, the curve is nearly log-linear between Somers' *D* values of -0.5 and 0.5.
- ► And, at these values, the IQORs are approximately 0.276 and 3.621, respectively.
- ► So, the IQORs corresponding to intermediate Somers' D values can be approximated well by weighted geometric means, interpolated log–linearly between 0.276 and 3.621.

- Note that the IQOR (on the horizontal axis) is plotted on a binary log scale.
- However, after this transformation, there is again a range of near-linearity between Somers' D values of -0.5 and 0.5.
- And, at these values, the IQORs are 0.276 and 3.621, respectively.



- Note that the IQOR (on the horizontal axis) is plotted on a binary log scale.
- ► *However*, after this transformation, there is again a range of near–linearity between Somers' *D* values of -0.5 and 0.5.
- And, at these values, the IQORs are 0.276 and 3.621, respectively.



- Note that the IQOR (on the horizontal axis) is plotted on a binary log scale.
- ► *However*, after this transformation, there is again a range of near-linearity between Somers' *D* values of -0.5 and 0.5.
- And, at these values, the IQORs are 0.276 and 3.621, respectively.



- Note that the IQOR (on the horizontal axis) is plotted on a binary log scale.
- ► *However*, after this transformation, there is again a range of near-linearity between Somers' *D* values of -0.5 and 0.5.
- And, at these values, the IQORs are 0.276 and 3.621, respectively.



- ▶ This time, *Y* is a positive–valued continuous lifetime variable, with a constant hazard ratio *R* between the subpopulations identified by X = 1 and X = 0 (as in a Cox regression).
- ▶ In the absence of censorship, Somers' D(Y|X) is then given by

D(Y|X) = (1-R)/(1+R).

- So, the Somers' D values corresponding to hazard ratios of 3, 2, 1, 1/2 and 1/3 are -1/2, -1/3, 0, 1/3 and 1/2, respectively.
- ► This relationship is decreasing, and nearly log–linear for Somers' *D* between -0.5 and 0.5.
- ▶ Note that the hazard ratio (unlike Somers' *D*) is still the same in the *presence* of censorship.

- ► This time, *Y* is a positive–valued continuous lifetime variable, with a constant hazard ratio *R* between the subpopulations identified by X = 1 and X = 0 (as in a Cox regression).
- ▶ In the absence of censorship, Somers' D(Y|X) is then given by

$$D(Y|X) = (1-R)/(1+R).$$

- So, the Somers' D values corresponding to hazard ratios of 3, 2, 1, 1/2 and 1/3 are -1/2, -1/3, 0, 1/3 and 1/2, respectively.
- ► This relationship is decreasing, and nearly log–linear for Somers' *D* between -0.5 and 0.5.
- ▶ Note that the hazard ratio (unlike Somers' *D*) is still the same in the *presence* of censorship.

- ► This time, Y is a positive-valued continuous lifetime variable, with a constant hazard ratio R between the subpopulations identified by X = 1 and X = 0 (as in a Cox regression).
- ► In the absence of censorship, Somers' D(Y|X) is then given by

$$D(Y|X) = (1-R)/(1+R).$$

- So, the Somers' D values corresponding to hazard ratios of 3, 2, 1, 1/2 and 1/3 are -1/2, -1/3, 0, 1/3 and 1/2, respectively.
- ► This relationship is decreasing, and nearly log–linear for Somers' *D* between -0.5 and 0.5.
- ▶ Note that the hazard ratio (unlike Somers' *D*) is still the same in the *presence* of censorship.

- ► This time, Y is a positive-valued continuous lifetime variable, with a constant hazard ratio R between the subpopulations identified by X = 1 and X = 0 (as in a Cox regression).
- ► In the absence of censorship, Somers' D(Y|X) is then given by

$$D(Y|X) = (1-R)/(1+R).$$

- So, the Somers' D values corresponding to hazard ratios of 3, 2, 1, 1/2 and 1/3 are -1/2, -1/3, 0, 1/3 and 1/2, respectively.
- ► This relationship is decreasing, and nearly log–linear for Somers' *D* between -0.5 and 0.5.
- ▶ Note that the hazard ratio (unlike Somers' *D*) is still the same in the *presence* of censorship.

- ► This time, Y is a positive-valued continuous lifetime variable, with a constant hazard ratio R between the subpopulations identified by X = 1 and X = 0 (as in a Cox regression).
- ► In the absence of censorship, Somers' D(Y|X) is then given by

$$D(Y|X) = (1 - R)/(1 + R).$$

- So, the Somers' D values corresponding to hazard ratios of 3, 2, 1, 1/2 and 1/3 are -1/2, -1/3, 0, 1/3 and 1/2, respectively.
- ► This relationship is decreasing, and nearly log-linear for Somers' *D* between -0.5 and 0.5.
- ▶ Note that the hazard ratio (unlike Somers' *D*) is still the same in the *presence* of censorship.

- ► This time, Y is a positive-valued continuous lifetime variable, with a constant hazard ratio R between the subpopulations identified by X = 1 and X = 0 (as in a Cox regression).
- ► In the absence of censorship, Somers' D(Y|X) is then given by

$$D(Y|X) = (1 - R)/(1 + R).$$

- So, the Somers' D values corresponding to hazard ratios of 3, 2, 1, 1/2 and 1/3 are -1/2, -1/3, 0, 1/3 and 1/2, respectively.
- ► This relationship is decreasing, and nearly log-linear for Somers' *D* between -0.5 and 0.5.
- Note that the hazard ratio (unlike Somers' D) is still the same in the presence of censorship.

Somers' D plotted against the hazard ratio

- Again, the horizontal axis is on a binary log scale.
- The curve is decreasing, but nearly log–linear for Somers' D between -0.5 and 0.5.
- The respective hazard ratios, corresponding to these limits, are 3 and 1/3.
- This log-linearity range includes the typical smoking-related hazard ratio of 2, corresponding to an *uncensored* Somers' D of -1/3.


- ► Again, the horizontal axis is on a binary log scale.
- The curve is decreasing, but nearly log–linear for Somers' D between -0.5 and 0.5.
- The respective hazard ratios, corresponding to these limits, are 3 and 1/3.
- This log–linearity range includes the typical smoking–related hazard ratio of 2, corresponding to an *uncensored* Somers' D of -1/3.



- ► Again, the horizontal axis is on a binary log scale.
- ► The curve is decreasing, but nearly log-linear for Somers' D between -0.5 and 0.5.
- The respective hazard ratios, corresponding to these limits, are 3 and 1/3.
- This log–linearity range includes the typical smoking–related hazard ratio of 2, corresponding to an *uncensored* Somers' D of -1/3.



Somers' D: A common currency for associations

- ► Again, the horizontal axis is on a binary log scale.
- ► The curve is decreasing, but nearly log-linear for Somers' D between -0.5 and 0.5.
- ► The respective hazard ratios, corresponding to these limits, are 3 and 1/3.
- This log-linearity range includes the typical smoking-related hazard ratio of 2, corresponding to an *uncensored* Somers' D of -1/3.



- ► Again, the horizontal axis is on a binary log scale.
- ► The curve is decreasing, but nearly log-linear for Somers' D between -0.5 and 0.5.
- ► The respective hazard ratios, corresponding to these limits, are 3 and 1/3.
- This log-linearity range includes the typical smoking-related hazard ratio of 2, corresponding to an *uncensored* Somers' D of -1/3.



- ► This time, *X* and *Y* have a bivariate Normal joint distribution, with means μ_X and μ_Y , SDs σ_X and σ_Y , and a Pearson correlation coefficient ρ .
- ► Somers' *D*(*Y*|*X*) is then equal to Kendall's tau-a, and is related to the Pearson correlation by **Greiner's relation**,

$$D(Y|X) = \tau_{XY} = \frac{2}{\pi} \arcsin(\rho).$$

- So, the Somers' D values of 0, ±1/3, ±1/2, and ±1 correspond to Pearson correlations of 0, ±1/2, ±√1/2, and ±1, respectively.
- ► This relationship, again, is nearly linear (with slope 2/π) for Somers' D between -0.5 and 0.5.
- ▶ Note that Greiner's relation still holds if *X* and/or *Y* is derived using a Normalizing transformation.

- ► This time, *X* and *Y* have a bivariate Normal joint distribution, with means μ_X and μ_Y , SDs σ_X and σ_Y , and a Pearson correlation coefficient ρ .
- ► Somers' *D*(*Y*|*X*) is then equal to Kendall's tau-a, and is related to the Pearson correlation by **Greiner's relation**,

$$D(Y|X) = \tau_{XY} = \frac{2}{\pi} \arcsin(\rho).$$

- So, the Somers' D values of 0, ±1/3, ±1/2, and ±1 correspond to Pearson correlations of 0, ±1/2, ±√1/2, and ±1, respectively.
- ► This relationship, again, is nearly linear (with slope 2/π) for Somers' D between -0.5 and 0.5.
- ▶ Note that Greiner's relation still holds if *X* and/or *Y* is derived using a Normalizing transformation.

- ► This time, *X* and *Y* have a bivariate Normal joint distribution, with means μ_X and μ_Y , SDs σ_X and σ_Y , and a Pearson correlation coefficient ρ .
- ► Somers' D(Y|X) is then equal to Kendall's tau-a, and is related to the Pearson correlation by Greiner's relation,

$$D(Y|X) = \tau_{XY} = \frac{2}{\pi} \arcsin(\rho).$$

- So, the Somers' D values of 0, ±1/3, ±1/2, and ±1 correspond to Pearson correlations of 0, ±1/2, ±√1/2, and ±1, respectively.
- ► This relationship, again, is nearly linear (with slope 2/π) for Somers' D between -0.5 and 0.5.
- ▶ Note that Greiner's relation still holds if *X* and/or *Y* is derived using a Normalizing transformation.

- ► This time, *X* and *Y* have a bivariate Normal joint distribution, with means μ_X and μ_Y , SDs σ_X and σ_Y , and a Pearson correlation coefficient ρ .
- ► Somers' D(Y|X) is then equal to Kendall's tau-a, and is related to the Pearson correlation by Greiner's relation,

$$D(Y|X) = \tau_{XY} = \frac{2}{\pi} \arcsin(\rho).$$

- So, the Somers' D values of 0, ±1/3, ±1/2, and ±1 correspond to Pearson correlations of 0, ±1/2, ±√1/2, and ±1, respectively.
- ► This relationship, again, is nearly linear (with slope 2/π) for Somers' D between -0.5 and 0.5.
- ▶ Note that Greiner's relation still holds if *X* and/or *Y* is derived using a Normalizing transformation.

- ► This time, *X* and *Y* have a bivariate Normal joint distribution, with means μ_X and μ_Y , SDs σ_X and σ_Y , and a Pearson correlation coefficient ρ .
- ► Somers' D(Y|X) is then equal to Kendall's tau-a, and is related to the Pearson correlation by Greiner's relation,

$$D(Y|X) = \tau_{XY} = \frac{2}{\pi} \arcsin(\rho).$$

- So, the Somers' D values of 0, ±1/3, ±1/2, and ±1 correspond to Pearson correlations of 0, ±1/2, ±√1/2, and ±1, respectively.
- ► This relationship, again, is nearly linear (with slope 2/π) for Somers' D between -0.5 and 0.5.
- ▶ Note that Greiner's relation still holds if *X* and/or *Y* is derived using a Normalizing transformation.

- ► This time, *X* and *Y* have a bivariate Normal joint distribution, with means μ_X and μ_Y , SDs σ_X and σ_Y , and a Pearson correlation coefficient ρ .
- ► Somers' D(Y|X) is then equal to Kendall's tau-a, and is related to the Pearson correlation by Greiner's relation,

$$D(Y|X) = \tau_{XY} = \frac{2}{\pi} \arcsin(\rho).$$

- So, the Somers' D values of 0, ±1/3, ±1/2, and ±1 correspond to Pearson correlations of 0, ±1/2, ±√1/2, and ±1, respectively.
- ► This relationship, again, is nearly linear (with slope 2/π) for Somers' D between -0.5 and 0.5.
- ► Note that Greiner's relation still holds if *X* and/or *Y* is derived using a Normalizing transformation.

- The curve is nearly linear (with slope 2/π) for Somers' D values between -0.5 and 0.5.
- ► The dashed lines on the vertical axis denote the Somers' D values of 0, ±1/3, ±1/2, and ±1.
- The dashed lines on the horizontal axis give the corresponding Pearson ρ values of 0, ±1/2, ±√1/2, and ±1.



- The curve is nearly linear (with slope 2/π) for Somers' D values between -0.5 and 0.5.
- ► The dashed lines on the vertical axis denote the Somers' D values of 0, ±1/3, ±1/2, and ±1.
- The dashed lines on the horizontal axis give the corresponding Pearson ρ values of 0, ±1/2, ±√1/2, and ±1.



- The curve is nearly linear (with slope 2/π) for Somers' D values between -0.5 and 0.5.
- ► The dashed lines on the vertical axis denote the Somers' D values of 0, ±1/3, ±1/2, and ±1.
- The dashed lines on the horizontal axis give the corresponding Pearson ρ values of 0, ±1/2, ±√1/2, and ±1.



- The curve is nearly linear (with slope 2/π) for Somers' D values between -0.5 and 0.5.
- ► The dashed lines on the vertical axis denote the Somers' D values of 0, ±1/3, ±1/2, and ±1.
- ► The dashed lines on the horizontal axis give the corresponding Pearson *ρ* values of 0, ±1/2, ±√1/2, and ±1.



- Somers' D defines a hierarchy between monotonic associations, stating (scientifically) which ones may *not* be secondary to which others, without committing ourselves to a particular model.
- However, a lot of statistically-minded scientists prefer to measure associations using model parameters, and answer practical questions about how much good they can do.
- ► *Fortunately*, under a range of familiar models, Somers' *D* is monotonic in the more familiar model parameter, and nearly linear (or log–linear) for Somers' *D* values between -0.5 and 0.5.
- ► So, given a Somers' D between -0.5 and 0.5, we can interpolate linearly (or log–linearly) to limit the range of proportion/mean differences, odds/hazard ratios, or Pearson correlations, which *may* be secondary to an association with that Somers' D.
- A more equation-intensive version of this story can be found in the "Miscellaneous documents" section of the author's website[3].

- ► Somers' *D* defines a hierarchy between monotonic associations, stating (scientifically) which ones may *not* be secondary to which others, without committing ourselves to a particular model.
- However, a lot of statistically-minded scientists prefer to measure associations using model parameters, and answer practical questions about how much good they can do.
- Fortunately, under a range of familiar models, Somers' D is monotonic in the more familiar model parameter, and nearly linear (or log–linear) for Somers' D values between -0.5 and 0.5.
- ► So, given a Somers' D between -0.5 and 0.5, we can interpolate linearly (or log–linearly) to limit the range of proportion/mean differences, odds/hazard ratios, or Pearson correlations, which may be secondary to an association with that Somers' D.
- A more equation-intensive version of this story can be found in the "Miscellaneous documents" section of the author's website[3].

- Somers' D defines a hierarchy between monotonic associations, stating (scientifically) which ones may *not* be secondary to which others, without committing ourselves to a particular model.
- However, a lot of statistically-minded scientists prefer to measure associations using model parameters, and answer practical questions about how much good they can do.
- Fortunately, under a range of familiar models, Somers' D is monotonic in the more familiar model parameter, and nearly linear (or log–linear) for Somers' D values between -0.5 and 0.5.
- ► So, given a Somers' D between -0.5 and 0.5, we can interpolate linearly (or log–linearly) to limit the range of proportion/mean differences, odds/hazard ratios, or Pearson correlations, which *may* be secondary to an association with that Somers' D.
- A more equation-intensive version of this story can be found in the "Miscellaneous documents" section of the author's website[3].

- Somers' D defines a hierarchy between monotonic associations, stating (scientifically) which ones may *not* be secondary to which others, without committing ourselves to a particular model.
- However, a lot of statistically-minded scientists prefer to measure associations using model parameters, and answer practical questions about how much good they can do.
- ► *Fortunately*, under a range of familiar models, Somers' *D* is monotonic in the more familiar model parameter, and nearly linear (or log–linear) for Somers' *D* values between -0.5 and 0.5.
- ► So, given a Somers' D between -0.5 and 0.5, we can interpolate linearly (or log–linearly) to limit the range of proportion/mean differences, odds/hazard ratios, or Pearson correlations, which *may* be secondary to an association with that Somers' D.
- A more equation—intensive version of this story can be found in the "Miscellaneous documents" section of the author's website[3].

- Somers' D defines a hierarchy between monotonic associations, stating (scientifically) which ones may *not* be secondary to which others, without committing ourselves to a particular model.
- However, a lot of statistically-minded scientists prefer to measure associations using model parameters, and answer practical questions about how much good they can do.
- ► *Fortunately*, under a range of familiar models, Somers' *D* is monotonic in the more familiar model parameter, and nearly linear (or log–linear) for Somers' *D* values between -0.5 and 0.5.
- So, given a Somers' D between -0.5 and 0.5, we can interpolate linearly (or log-linearly) to limit the range of proportion/mean differences, odds/hazard ratios, or Pearson correlations, which may be secondary to an association with that Somers' D.
- A more equation—intensive version of this story can be found in the "Miscellaneous documents" section of the author's website[3].

- ► Somers' *D* defines a hierarchy between monotonic associations, stating (scientifically) which ones may *not* be secondary to which others, without committing ourselves to a particular model.
- However, a lot of statistically-minded scientists prefer to measure associations using model parameters, and answer practical questions about how much good they can do.
- ► *Fortunately*, under a range of familiar models, Somers' *D* is monotonic in the more familiar model parameter, and nearly linear (or log–linear) for Somers' *D* values between -0.5 and 0.5.
- ► So, given a Somers' D between -0.5 and 0.5, we can interpolate linearly (or log-linearly) to limit the range of proportion/mean differences, odds/hazard ratios, or Pearson correlations, which may be secondary to an association with that Somers' D.
- A more equation-intensive version of this story can be found in the "Miscellaneous documents" section of the author's website[3].

References

- [1] Guo, S. and Fraser, M. W. 2014. *Propensity score analysis. Second edition.* Los Angeles, CA: Sage.
- [2] Hanley, J. A. and B. J. McNeil. 1982. The meaning and use of the area under a receiver operating characteristic (ROC) curve. *Radiology* 143: 29–36.
- [3] Newson, R. B. 2014. Interpretation of Somers' D under four simple models. Downloaded on 25 February 2015 from the author's website at http://www.imperial.ac.uk/nhli/r.newson/papers.htm#miscellaneous_documents
- [4] Newson, R. 2006. Confidence intervals for rank statistics: Somers' *D* and extensions. *The Stata Journal* **6(3)**: 309–334.
- [5] Newson, R. B. 1987. An analysis of cinematographic cell division data using U-statistics. DPhil thesis, Sussex University, Brighton, UK.
- [6] Rubin, D. B. 2008. For objective causal inference, design trumps analysis. The Annals of Applied Statistics 2(3): 808–840.

This presentation, and the do-file producing the examples, can be downloaded from the conference website at *http://ideas.repec.org/s/boc/usug15.html*

The packages described and used in this presentation can be downloaded from SSC, using the ssc command.

Somers' D: A common currency for associations