

# Influence functions at work

Philippe Van Kerm

Luxembourg Institute of Socio-Economic Research  
philippe.vankerm@liser.lu

2015 London Stata Users Group meeting  
September 10–11 2015, Cass Business School, London



# Introduction

- ▶ Illustration of practical uses of 'influence function' estimators with Stata
  1. Study structure of summary statistics (identification of 'influential observations')
  2. Variance estimation
  3. 'RIF regression'
- ▶ Application to income distribution analysis



## Definition

Let  $v(F)$  be a statistic of interest (a functional) calculated in distribution  $F$  (the mean, a percentile, the Gini coefficient of inequality, etc.)

The *influence function* of  $v$  is a function of  $y$  and  $F$  and is defined as (Hampel, 1974)

$$\text{IF}(y; v, F) = \lim_{\epsilon \downarrow 0} \frac{v((1 - \epsilon)F + \epsilon\Delta_y) - v(F)}{\epsilon}$$

The IF captures the effect on  $v(F)$  of an infinitesimal ‘contamination’ of  $F$  at point mass  $y$ .

(Note the expected value of the IF is zero:  $\int \text{IF}(y, v(F))dF(y) = 0$  )



## Definition (ctd.)

Expressions for  $\text{IF}(y; v, F)$  exist (or can be derived) for a wide range of statistics  $v^1$ :

... simple (linear) statistics, e.g., the mean

$$\text{IF}(y; \mu, F) = y - \mu(F)$$

... and more complex (non linear) statistics, e.g., a quantile

$$\text{IF}(y; Q_\theta, F) = \frac{-1}{f(Q_\theta(F))} (I(y \leq Q_\theta(F)) - \theta)$$

<sup>1</sup>See e.g., Essama-Nssah and Lambert (2012) for a catalogue of IFs relevant to income distribution analysis



# Practical use 1

## Practical use 1:

- ▶ visualising the 'structure' of a (possibly complex) index
- ▶ comparison of indices (think of the many inequality measures!)
- ▶ identification of influential observations (and robustness of the index)



## Income inequality indicators: the Atkinson index

The Atkinson inequality index (Atkinson, 1970):

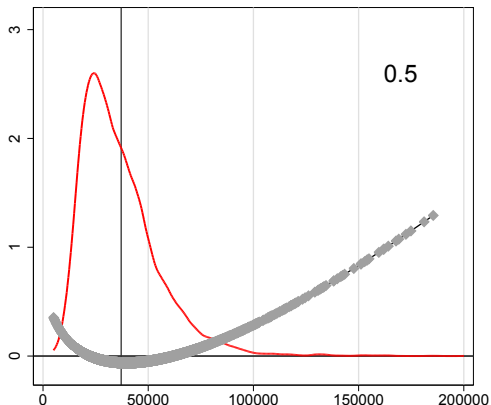
$$A(\epsilon) = 1 - \frac{1}{\mu} \left( \frac{1}{N} \sum_{i=1}^N y_i^{1-\epsilon} \right)^{\frac{1}{1-\epsilon}}$$

for  $\epsilon \geq 0$

The higher  $\epsilon$ , the higher 'inequality aversion'... Can we visualise that?



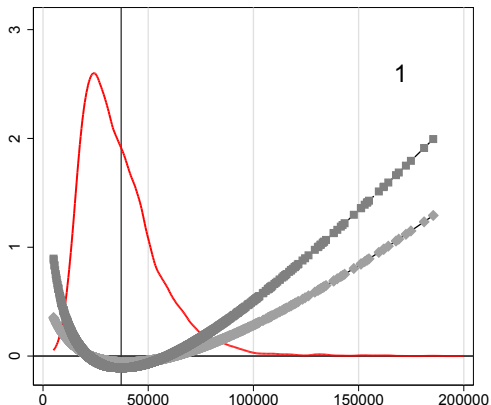
## Income inequality indicators: the Atkinson index IF



(annual household income data for Luxembourg 2012)



# Income inequality indicators: the Atkinson index IF

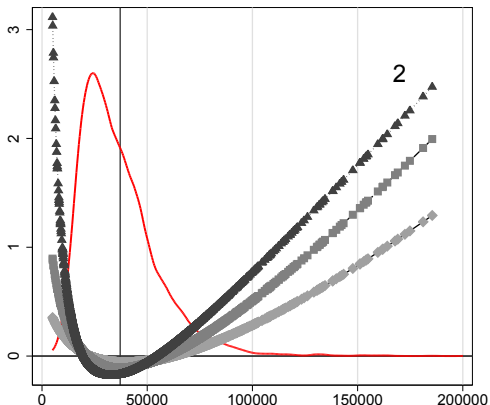


(annual household income data for Luxembourg 2012)





## Income inequality indicators: the Atkinson index IF

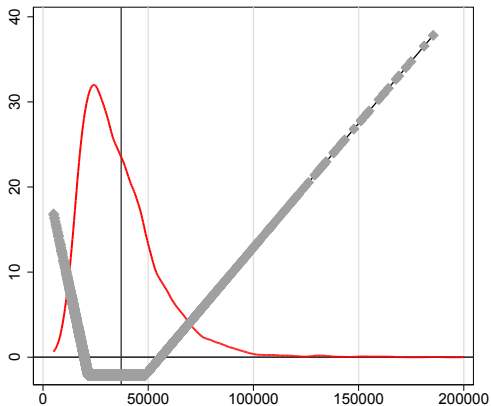


(annual household income data for Luxembourg 2012)

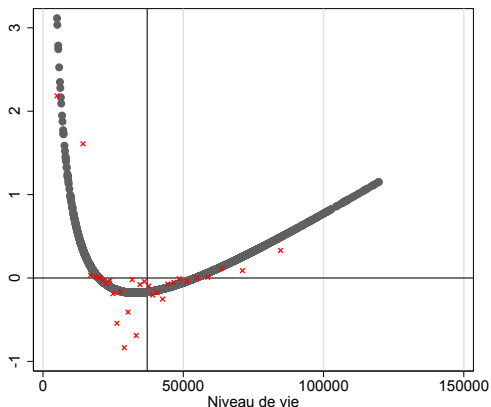


## Income inequality: the Quintile Share Ratio

How does it compare with the Quintile (Group) Share Ratio?



## An aside: comparison with delete-one jackknife influence measure



Not the same when observations are *weighted*!



## Practical use 2

### Practical use 2:

- ▶ estimation of the sampling variance of the index
- ▶ asymptotic approximation that works with complex non-linear statistics
- ▶ valid with complex survey design!
- ▶ (... it is all in the Stata manuals already)



## Variance estimation

An asymptotic approximation of the variance of  $v$  is given by (Hampel, 1974)

$$V(v, F) \approx \int \text{IF}(y; v, F)^2 dF(y)$$

Practically boils down to estimation of a total (Deville, 1999):

$$V(\hat{v}, F) \approx V \left( \sum_{i=1}^N w_i \text{IF}(y_i; v, \hat{F}) \right)$$

... and formula well-known for the variance of a total even with complex survey design: implemented in Stata!



## Variance estimation

### Code template

```
svyset ...  
generate rif= ... // note: point estimate is added to  
IF  
svy: mean rif
```

Silly example with the mean:

```
svyset [pw=W] , ...  
su y [aw=W]  
gen rifmean = r(mean) + (y - r(mean))  
svy: mean rifmean  
svy: mean y // yes, it works!
```



# Variance estimation

Built in some user-written commands

```
.          svy : inequaly nivie , atkinson(0.5 1 2) s80s20
(running inequaly on estimation sample)
```

Survey data analysis

Number of strata	=	4	Number of obs	=	5,988
Number of PSUs	=	5,988	Population size	=	515,407.14
			Design df	=	5,984
			F( 0, 5984)	=	.
			Prob > F	=	.

	nivie	Coef.	Linearized Std. Err.	t	P> t	[95% Conf. Interval]	
atkp5							
	_cons	.0621767	.0039672	15.67	0.000	.0543995	.069954
atk1							
	_cons	.1174055	.0061389	19.12	0.000	.1053711	.12944
atk2							
	_cons	.2146223	.0083138	25.82	0.000	.1983242	.2309204
s80s20							
	_cons	3.953448	.1226658	32.23	0.000	3.712978	4.193917



## Practical use 3

### Practical use 3:

- ▶ 'Recentered IF regression' (Firpo et al., 2007, 2009)
- ▶ evaluate impact of covariates on distribution statistics





## RIF regression

The effect of interest

For example, how do foreign households affect  $v(F)$ ?

$$F(y) = \sum_{x \in \Omega_X} s_x F_x(y)$$

Consider an infinitesimal variation: swap native for foreign workers

$$G_r^{F,t,k}(y) = (s_k + t) F_k(y) + (s_r - t) F_r(y) + \sum_{x \in \Omega_X \setminus \{k,r\}} s_x F_x(y).$$

(Choe and Van Kerm, 2014)

What is the impact of this swap on the statistic of interest?



## Methods (ctd.)

### Recentered influence function estimator

Firpo et al. (2009) show that effect of interest is given by:

$$E[\text{RIF}(y; v, F)|X = k] - E[\text{RIF}(y; v, F)|X = r]$$

where  $\text{RIF}(y; v, F) = v(F) + \text{IF}(y; v, F)$

Regression-based estimator,  $\beta$  in :

$$E[\text{RIF}(y; v, F)|X = x] = \alpha + x\beta$$

(Note: N. Fortin provides the Stata package `rifreg` for regressions on quantile, variance and Gini functionals

(<http://faculty.arts.ubc.ca/nfortin/datahead.html>).



## Interpretation of RIF regression coefficients

- ▶ The RIF at  $y$  gives the influence on  $v(F)$  of an infinitesimal increase in the density of the data at  $y$
- ▶ Regression coefficients reveal how much the average influence of observations vary with  $X$  (holding other covariates constant)
- ▶ It also reveals how much  $v(F)$  would respond to a change in the distribution of  $X$  in the population holding distribution of other covariates constant
  - ▶ linear approximation valid only for *marginal* changes in  $X$ !



# Illustrative example 1

## Data and problem

- ▶ Panel Study Liewen zu Letzebuerg 2011 (official source for poverty and inequality statistics in Luxembourg)
- ▶ Permanent panel (started in 2003) gradually converted into rotating panel from 2010:
  - ▶ new samples ('rotation groups') now added for four years (so dataset for 2010, 2011, 2012 is mix of old and new samples)
  - ▶ Concern about how much 'new samples' differ from 'old sample' (due, e.g., to attrition (and slightly different sampling frame))
  - ▶ Impact on trends in inequality and poverty indicators? Break in series when 'old sample' abandoned?

⇒ Use RIF regression to check if rotation group impacts on inequality 

## Illustrative example 1

### Code

```
svy: inequaly nivie , atkinson(0.5 1 2)
predict rif* , rif // predict after -inequaly- gives (R)IF
svy: regress rif1 ib9.(rot)
svy: regress rif2 ib9.(rot)
svy: regress rif3 ib9.(rot)
svy: inequaly nivie , s80s20
predict rifs80s20 , rif
svy: regress rifs80s20 ib9.(rot)
svy: newpoverty nivie , fracmedian(.6)
predict rifh , rif
svy: regress rifh ib9.(rot)
```



## Results: Atkinson(0.5)

```
.          suy: regress rif1 ib9.(rot)
(running regress on estimation sample)
```

Survey: Linear regression

```
Number of strata =      4
Number of PSUs  =    5,988
```

```
Number of obs   =    5,988
Population size = 515,407.14
Design df       =    5,984
F( 3, 5982)     =    2.96
Prob > F        =    0.0312
R-squared       =    0.0023
```

rif1	Coef.	Linearized Std. Err.	t	P> t	[95% Conf. Interval]	
rot						
1	.0157444	.0186107	0.85	0.398	-.0207392	.052228
2	-.0048659	.0043721	-1.11	0.266	-.0134369	.003705
3	.0111881	.0059868	1.87	0.062	-.0005481	.0229243
_cons	.0575857	.0033763	17.06	0.000	.0509669	.0642044



## Results: Atkinson(2)

```
.          suy: regress rif3 ib9.(rot)
(running regress on estimation sample)
```

Survey: Linear regression

Number of strata	=	4	Number of obs	=	5,988
Number of PSUs	=	5,988	Population size	=	515,407.14
			Design df	=	5,984
			F( 3, 5982)	=	2.04
			Prob > F	=	0.1056
			R-squared	=	0.0023

	rif3	Coef.	Linearized Std. Err.	t	P> t	[95% Conf. Interval]	
	rot						
	1	.0246292	.0347917	0.71	0.479	-.0435751	.0928336
	2	-.006993	.0135387	-0.52	0.606	-.0335338	.0195478
	3	.0339101	.0171911	1.97	0.049	.0002094	.0676108
	_cons	.2036115	.009367	21.74	0.000	.1852488	.2219741



## Results: S80/S20

```
.          svy: regress rifs80s20 ib9.(rot)
(running regress on estimation sample)
```

```
Survey: Linear regression
```

```
Number of strata =      4
Number of PSUs  =    5,988
```

```
Number of obs   =    5,988
Population size = 515,407.14
Design df       =    5,984
F( 3, 5982)     =    2.44
Prob > F        =    0.0622
R-squared       =    0.0027
```

rifs80s20	Coef.	Linearized Std. Err.	t	P> t	[95% Conf. Interval]	
rot						
1	.4599947	.5221966	0.88	0.378	-.5636989	1.483688
2	-.0630366	.2020467	-0.31	0.755	-.4591209	.3330477
3	.5306184	.2375919	2.23	0.026	.0648526	.9963842
_cons	3.753404	.1405622	26.70	0.000	3.477851	4.028956





## Results: Poverty rate

```
.          suy: regress r1fh ib9.(rot)
(running regress on estimation sample)
```

Survey: Linear regression

Number of strata = 4  
 Number of PSUs = 5,988

Number of obs = 5,988  
 Population size = 515,407.14  
 Design df = 5,984  
 F( 3, 5982) = 1.32  
 Prob > F = 0.2659  
 R-squared = 0.0020

r1fh	Coef.	Linearized Std. Err.	t	P> t	[95% Conf. Interval]	
rot						
1	.0146267	.0233548	0.63	0.531	-.0311571	.0604105
2	.0069663	.0211085	0.33	0.741	-.034414	.0483466
3	.043081	.0221201	1.95	0.052	-.0002823	.0864443
_cons	-.1342025	.0129755	10.34	0.000	-.1087658	-.1596393



## Illustrative example 2

Effect of foreign households on inequality and poverty?

- ▶ Effect of a marginal increase in share of foreign-headed households on indicators
  - ▶ assuming no change in income structure otherwise
- ▶ Condition on age of foreign households



## Results: Atkinson(0.5)

```
.          svy: regress rif1 ib9.(rot) i.(chme11)
(running regress on estimation sample)
```

Survey: Linear regression

Number of strata	=	4	Number of obs	=	5,987
Number of PSUs	=	5,987	Population size	=	515,119.97
			Design df	=	5,983
			F( 6, 5978)	=	7.42
			Prob > F	=	0.0000
			R-squared	=	0.0099

		Linearized				[95% Conf. Interval]	
	rif1	Coef.	Std. Err.	t	P> t		
	rot						
	1	.0142397	.0183501	0.78	0.438	-.0217331	.0502126
	2	-.0062364	.0043078	-1.45	0.148	-.0146812	.0022083
	3	.0080582	.0059128	1.36	0.173	-.003533	.0196495
	chme11						
	Portugais	.0135398	.0039052	3.47	0.001	.0058841	.0211955
	Autres UE-15	.0262902	.0178432	1.47	0.141	-.008689	.0612694
	Non UE-15	.0418941	.0086454	4.85	0.000	.0249461	.0588421
	_cons	.0476424	.0050467	9.44	0.000	.0377491	.0575358



## Results: Atkinson(0.5)

```
.          svy: regress rif1 ib9.(rot) i.(chme11) ib6.(chme09)
(running regress on estimation sample)
```

Survey: Linear regression

Number of strata	=	4	Number of obs	=	5,987
Number of PSUs	=	5,987	Population size	=	515,119.97
			Design df	=	5,983
			F( 10, 5974)	=	5.90
			Prob > F	=	0.0000
			R-squared	=	0.0108

	Coef.	Linearized Std. Err.	t	P> t	[95% Conf. Interval]	
-----						
rif1						
-----						
rot						
1	.0141385	.0178513	0.79	0.428	-.0208565	.0491335
2	-.0061202	.0043087	-1.42	0.156	-.0145667	.0023264
3	.0082485	.0059444	1.39	0.165	-.0034047	.0199016
chme11						
Portugais	.0159073	.0039841	3.99	0.000	.008097	.0237177
Autres UE-15	.0269482	.0181449	1.49	0.138	-.0086224	.0625188
Non UE-15	.0442646	.0085935	5.15	0.000	.0274182	.061111
chme09						
[16-24]	.0032088	.0138918	0.23	0.817	-.0240242	.0304418
[25-34]	-.0125041	.0078015	-1.60	0.109	-.0277978	.0027897
[35-49]	-.0050532	.007715	-0.65	0.512	-.0201773	.0100708
[50-64]	.0021521	.0129697	0.17	0.868	-.0232733	.0275774
_cons	.0499364	.0094564	5.28	0.000	.0313985	.0684744
-----						



## Results: Atkinson(2)

```
.          svy: regress rif3 ib9.(rot) i.(chme11) ib6.(chme09)
(running regress on estimation sample)
```

Survey: Linear regression

```
Number of strata   =          4
Number of PSUs    =       5,987
Number of obs     =       5,987
Population size   = 515,119.97
Design df        =       5,983
F( 10, 5974)     =         2.68
Prob > F         =       0.0029
R-squared        =       0.0173
```

	Coef.	Linearized Std. Err.	t	P> t	[95% Conf. Interval]	
-----						
rif3						
rot						
1	.0197437	.033453	0.59	0.555	-.0458363	.0853236
2	-.0112723	.0133852	-0.84	0.400	-.0375122	.0149676
3	.0245481	.0159634	1.54	0.124	-.0067459	.0558422
chme11						
Portugais	.0076757	.0114413	0.67	0.502	-.0147533	.0301047
Autres UE-15	.0634587	.0338566	1.87	0.061	-.0029124	.1298298
Non UE-15	.1332594	.0373124	3.57	0.000	.0601136	.2064052
chme09						
[16-24]	.046107	.0539939	0.85	0.393	-.0597405	.1519544
[25-34]	-.0031152	.0206771	-0.15	0.880	-.0436499	.0374194
[35-49]	.0021401	.0159684	0.13	0.893	-.0291637	.0334439
[50-64]	.0165315	.0241836	0.68	0.494	-.0308771	.06394
_cons	.1764981	.0180461	9.78	0.000	.1411213	.2118749
-----						



## Results: Poverty rate

```
.          svy: regress rifh ib9.(rot) i.(chme11) ib6.(chme09)
(running regress on estimation sample)
```

Survey: Linear regression

```
Number of strata   =          4
Number of PSUs    =       5,987
Number of obs     =       5,987
Population size   = 515,119.97
Design df        =       5,983
F( 10, 5974)     =         2.89
Prob > F         =       0.0013
R-squared        =       0.0142
```

	Coef.	Linearized Std. Err.	t	P> t	[95% Conf. Interval]	
-----						
rifh						
-----						
rot						
1	.0114206	.0234601	0.49	0.626	-.0345697	.0574109
2	.0040279	.0209573	0.19	0.848	-.037056	.0451117
3	.0380812	.0219422	1.74	0.083	-.0049334	.0810959
chme11						
Portugais	-.0770121	.0238951	-3.22	0.001	-.1238552	-.030169
Autres UE-15	.0329958	.019137	1.72	0.085	-.0045195	.0705112
Non UE-15	.0465431	.0511668	0.91	0.363	-.0537624	.1468486
chme09						
[16-24]	.0411687	.0877057	0.47	0.639	-.130766	.2131034
[25-34]	.0480566	.0257379	1.87	0.062	-.002399	.0985123
[35-49]	.0379237	.0192831	1.97	0.049	.0001218	.0757256
[50-64]	.0562376	.0190103	2.96	0.003	.0189707	.0935046
_cons	.1004148	.0159395	6.30	0.000	.0691676	.131662
-----						



## Results: Poverty rate (fixed poverty line)

```
.          svy: regress rifhbis1 ib9.(rot) i.(chme11) ib6.(chme09)
(running regress on estimation sample)
```

Survey: Linear regression

Number of strata	=	4	Number of obs	=	5,987
Number of PSUs	=	5,987	Population size	=	515,119.97
			Design df	=	5,983
			F( 10, 5974)	=	16.31
			Prob > F	=	0.0000
			R-squared	=	0.1074

	Coef.	Linearized Std. Err.	t	P> t	[95% Conf. Interval]	
-----						
rifhbis1						
rot						
1	.0178995	.0213374	0.84	0.402	-.0239295	.0597284
2	.0059625	.0192948	0.31	0.757	-.0318624	.0437873
3	.038296	.0218492	1.75	0.080	-.0045364	.0811283
chme11						
Portugais	.2040203	.0231579	8.81	0.000	.1586225	.2494181
Autres UE-15	.052723	.0189874	2.78	0.006	.0155009	.0899452
Non UE-15	.3275548	.0547087	5.99	0.000	.2203059	.4348036
chme09						
[16-24]	.1271475	.0982018	1.29	0.195	-.0653634	.3196585
[25-34]	.047582	.0231179	2.06	0.040	.0022626	.0929015
[35-49]	.0561238	.015962	3.52	0.000	.0248325	.0874151
[50-64]	.0382346	.014837	2.58	0.010	.0091487	.0673204
_cons	.0181208	.0122119	1.48	0.138	-.0058189	.0420606
-----						



- Atkinson, A. B. (1970), 'On the measurement of inequality', *Journal of Economic Theory* **2**(3), 244–263.
- Choe, C. and Van Kerm, P. (2014), Foreign workers and the wage distribution: Where do they fit in?, CEPS/INSTEAD Working Paper 2014-02, Luxembourg Institute of Socio-Economic Research, Esch-sur-Alzette, Luxembourg.
- Deville, J.-C. (1999), 'Variance estimation for complex statistics and estimators: linearization and residual techniques', *Survey Methodology* **25**, 193–204.





- Essama-Nssah, B. and Lambert, P. J. (2012), Influence functions for policy impact analysis, *in* J. A. Bishop and R. Salas, eds, 'Inequality, Mobility and Segregation: Essays in Honor of Jacques Silber', Vol. 20 of *Research on Economic Inequality*, Emerald Group Publishing, chapter 6, pp. 135–159.
- Firpo, S., Fortin, N. M. and Lemieux, T. (2007), Unconditional quantile regressions, Technical Working Paper 339, National Bureau of Economic Research, Cambridge MA, USA.
- Firpo, S., Fortin, N. M. and Lemieux, T. (2009), 'Unconditional quantile regressions', *Econometrica* **77**(3), 953–973.
- Hampel, F. R. (1974), 'The influence curve and its role in robust estimation', *Journal of the American Statistical Association* **69**(346), pp. 383–393.



Support from the Luxembourg 'Fonds National de la Recherche' is gratefully acknowledged (project 'Tax-benefit systems, employment structures and cross-country differences in income inequality in Europe: a micro-simulation approach–SIMDECO').

