

Quantile regression: Basics and recent advances

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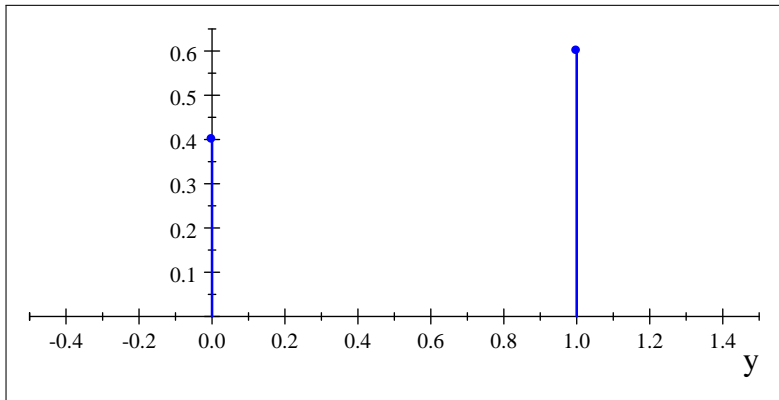
2019 UK Stata Conference
06/09/19

- Quantile regression (Koenker and Bassett, 1978) is increasingly used by practitioners but it is still not part of the standard econometric/statistics courses.
- Road map:
 - general introduction to quantile regression
 - two topics from recent research:
 - models with time-invariant individual (“fixed effects”) effects
 - structural quantile function.
- I will present the approach to these problems proposed by Machado and Santos Silva (2019), and illustrate the use of the corresponding Stata commands xtqreg and ivqreg2.

2. Conditional quantiles

- For $0 < \tau < 1$, the τ -th quantile of y given x is defined by

$$Q_y(\tau|x) = \min\{\eta | P(y \leq \eta|x) \geq \tau\}.$$



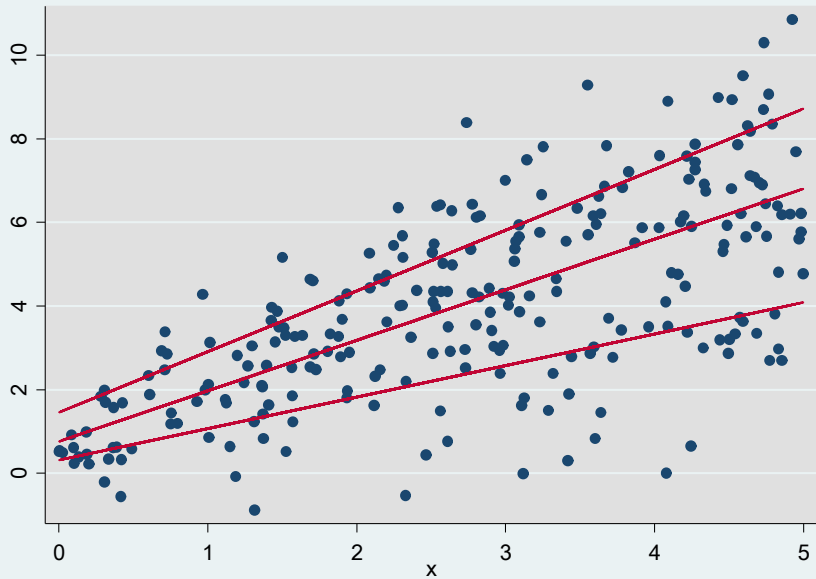
Bernoulli probability mass function with $\Pr(y = 1) = 0.6$

3. Basics of quantile regression

- Quantile regression estimates $Q_y(\tau|x)$.
- Throughout we assume linearity: $Q_y(\tau|x) = x'\beta(\tau)$.
- With linear quantiles, we can write

$$y = x'\beta(\tau) + u(\tau); \quad Q_{u(\tau)}(\tau|x) = 0.$$

- Note that the **errors** and the **parameters** depend on τ .
- For $\tau = 0.5$ we have the median regression.
- We need to restrict the **support** of x to ensure that quantiles do not cross.



- The estimator of $\beta(\tau)$ is defined by

$$\hat{\beta}(\tau) = \arg \min_b \frac{1}{n} \left\{ \sum_{y_i \geq x_i' b} \tau |y_i - x_i' b| + \sum_{y_i < x_i' b} (1 - \tau) |y_i - x_i' b| \right\}.$$

- The **F.O.C.** can be written as

$$\frac{1}{n} \sum_{i=1}^n ((\tau - \mathbf{1}((y_i - x_i' \hat{\beta}(\tau)) < 0))) x_i = 0.$$

- $\hat{\beta}(\tau)$ is **invariant** to perturbations of y_i that do not change the sign of $(y_i - x_i' \hat{\beta}(\tau))$.
- $\hat{\beta}(\tau)$ can be estimated by **linear programming** (see qreg).

- Asymptotic theory is **non-standard** because the objective function is not differentiable.
- However, under certain regularity conditions, $\hat{\beta}(\tau)$ has standard properties:

$$\sqrt{n}(\hat{\beta}(\tau) - \beta(\tau)) \xrightarrow{d} \mathcal{N}(0, D^{-1}AD^{-1}),$$

$$D = E \left[f_{u(\tau)}(0|x_i) x_i x_i' \right], \quad A = E \left[(\tau - \mathbf{1}(u(\tau)_i \leq 0))^2 x_i x_i' \right].$$

- It is possible to estimate A and D under different assumptions (see qreg and qreg2).

- The main advantage of quantile regression is the **informational gains** they provide.
- Quantiles are “**robust**” measures of location and are estimated using a “**robust**” estimator.
- Quantiles and means have very **different** properties.
 - Quantiles are not **additive**; the quantile of the sum is not the sum of the quantiles.
 - Quantiles are **equivariant** to non-decreasing transformations; for example, if y_i is non-negative with

$$Q_{y_i}(\tau|x_i) = \exp(x_i'\beta(\tau)),$$

then,

$$Q_{\ln(y_i)}(\tau|x_i) = x_i'\beta(\tau).$$

- The plain-vanilla quantile regression estimator has been extended to different settings:
 - Censored regression; Powell (1984)
 - Binary data; Manski (1975, 1985), Horowitz (1992)
 - Ordered data; M.-j. Lee (1992)
 - Count data; Machado and Santos Silva (2005)
 - Corner-solutions data; Machado, Santos Silva, and Wei (2016)
 - Clustering; Parente and Santos Silva (2016)
- Two areas of active research are:
 - quantile regressions with time-invariant individual ("fixed") effects, and
 - structural quantile function.

7. Quantiles via moments

- Consider a location-scale model

$$y_i = x_i' \beta + (x_i' \gamma) u_i,$$

where x_i and u_i are independent and $\Pr(x_i' \gamma > 0) = 1$.

- In this case the mean and all conditional quantiles are linear

$$\begin{aligned} Q_y(\tau|x) &= x_i' \beta + (x_i' \gamma) Q_u(\tau|x_i) \\ &= x_i' \beta(\tau) \end{aligned}$$

$$\beta(\tau) = \beta + \gamma Q_u(\tau).$$

- In this model, the information provided by β , γ , and $Q_u(\tau)$ is equivalent to the information provided by regression quantiles.

- Machado and Santos Silva (2019) noted that, assuming $E(U) = 0$ and using the normalization $E(|U|) = 1$, β and γ are identified by conditional expectations:

$$E[y_i | x_i] = \beta_0 + \beta_1 x_i$$

$$E[|y_i - \beta_0 - \beta_1 x_i| | x_i] = \gamma_0 + \gamma_1 x_i$$

- $Q_u(\tau | x_i)$ can be estimated from the scaled errors

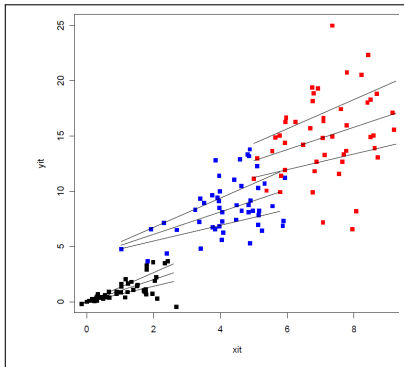
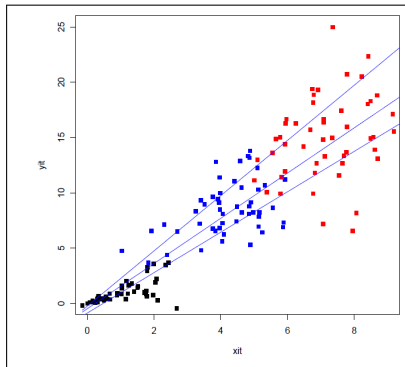
$$\frac{y_i - \beta_0 - \beta_1 x_i}{\gamma_0 + \gamma_1 x_i}$$

- This provides a way to estimate quantile regression using two OLS regressions and the computation of a univariate quantile.

- Suppose now that we are interested in estimating

$$Q_{y_{it}}(\tau|x_{it}, \eta_i) = x'_{it}\beta(\tau) + \eta(\tau)_i, \text{ with } i = 1, \dots, n; t = 1, \dots, T.$$

- As in mean regression, “**fixed effects**” can be important.



- Estimation of quantile regression with fixed effects is difficult because there is **no transformation** that can be used to eliminate the incidental parameters.
- Therefore, due to the **incidental parameter problem**, consistency requires that both $n \rightarrow \infty$ and $T \rightarrow \infty$.
- For fixed T , the only realistic option is the "**correlated random effects**" (Mundlak) estimator; see Abrevaya and Dahl (2008).
- Roger Koenker (2004) and Canay (2011) proposed estimators based on the assumption that $\eta(\tau)_i = \eta_i$ but this goes against the spirit of quantile regression.

- Kato, Galvão, and Montes-Rojas (2012) studied the properties of quantile regression in a model where the fixed effects are explicitly included as **dummies**.
- The estimator is consistent and asymptotically normal when both $n \rightarrow \infty$ and $T \rightarrow \infty$ with $n^2 [\ln(n)]^3 / T \rightarrow 0$.
- This is an issue because in many applications n is much larger than T (e.g. for $T = 40$, $n = 100$, $n^2 [\ln(n)]^3 / T = 24,416$).
- An alternative is to use the quantiles-via-moments estimator.

- Consider the location-scale model for panel data

$$y_{it} = \alpha_i + x'_{it}\beta + (\delta_i + x'_{it}\gamma)u_{it}$$

$$\eta(\tau)_i = \alpha_i + \delta_i Q_u(\tau), \quad \beta(\tau) = \beta + \gamma Q_u(\tau),$$

where x_i and u_i are independent and $\Pr((\delta_i + x'_{it}\gamma) > 0) = 1$.

- Estimation is performed using two fixed effects regressions (`xtreg`) and computing a univariate quantile.
- Consistency requires $(n, T) \rightarrow \infty$ with $n = o(T)$.
- For fixed T the estimator will have a bias but:
 - simulations suggest that the bias is negligible for $n/T \leq 10$;
 - the bias can be removed using **jackknife**.
- The estimator is implemented in the `xtqreg` command (available from SSC)

```
xtqreg depvar [indepvars] [if] [in] [, options]
```

quantile(#[#[# ...]]): estimates # quantile; default is
quantile(.5)

id: specifies the variable defining the panel

ls: displays the estimates of the location and scale
parameters

- Suppose that we have a structural relationship defined by

$$\begin{aligned}y &= d\alpha + x'\beta + u, \\d &= \delta(x, z, v)\end{aligned}$$

where v may not be independent of u

- We are interested in

$$S_y(\tau|d, x) = d\alpha(\tau) + x'\beta(\tau),$$

the structural quantile function such that:

- $\Pr[y < S_y(\tau|d, x) | z, x] = \tau,$
- $S_y(\tau|d, x) = Q_y(\tau|z, x) \neq Q_y(\tau|d, x).$

- **Chernozhukov and Hansen (2008)** propose an estimator of $S_Y(\tau|d, x)$ based on the observation that

$$Q_{y-d\alpha(\tau)}(\tau|z, x) = x'\beta(\tau) + z\gamma(\tau)$$

with $\gamma(\tau) = 0$.

- We can implement the estimator by:
 - estimating $\beta(\tau)$ and $\gamma(\tau)$ for a range of values of $\alpha(\tau)$
 - and choosing as estimates the ones corresponding to the value of $\alpha(\tau)$ for which $\gamma(\tau)$ is in some sense closer to zero.
- Chernozhukov and Hansen (2008) prove the consistency and asymptotic normality of the estimator.
- The estimator is difficult to implement when there are multiple endogenous variables, but there have been a number of recent **developments** on this.

- Again, the quantile-via-moments estimator can be useful.
- Consider a location-scale structural relationship

$$y = d\alpha + x'\beta + (d\delta + x'\gamma) u, \quad d = \delta(x, z, v),$$

where v may not be independent of u but u is independent of x and z .

- Because $S_y(\tau|d, x)$ is such that $\Pr[y < S_y(\tau|d, x)|z, x] = \tau$,

$$\begin{aligned} S_y(\tau|d, x) &= d\alpha + x'\beta + (d\delta + x'\gamma) Q_u(\tau) \\ &= d(\alpha + \delta Q_u(\tau)) + x(\beta + \gamma Q_u(\tau)). \end{aligned}$$

- GMM can be used to estimate the structural parameters:

$$E \left[\left(\frac{y_i - d\alpha - x'\beta}{d\delta + x'\gamma} \right) \middle| z_i \right] = 0,$$
$$E \left[\left(\frac{|y_i - d\alpha - x'\beta|}{d\delta + x'\gamma} - 1 \right) \middle| z_i \right] = 0.$$

- $Q_u(\tau)$ can be estimated from the standardized errors

$$(y_i - d\hat{\alpha} - x'\hat{\beta}) / (d\hat{\delta} + x'\hat{\gamma}).$$

- The estimator has the usual properties.
- The estimator is implemented in the ivqreg2 command (available from SSC)

```
ivqreg2 depvar [indepvars] [if] [in] [, options]
```

quantile(#[#[# ...]]): estimates # quantile; default is
quantile(.5)

instruments(varlist): list of instruments, including control
variables; by default no instruments are used and
restricted quantile regression is performed

ls: displays the estimates of the location and scale
parameters

- Quantile regression can be very useful and it is now easy to implement in a variety of cases.
- In some contexts, however, quantile regression can be challenging.
- The Method of Moments-Quantile Regression estimator can be useful in some of these cases.
- xtqreg and ivqreg2 make it easy to estimate quantile regressions with “fixed effects” or endogenous variables.

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