

Correlated Random Effects Methods for Panel Data Models with Heterogeneous Time Effects

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1. Introduction

- Microeconomic setting with small T , large N .
- Standard unobserved effects model for random draw from the population:

$$y_{it} = \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + c_i + u_{it}, t = 1, \dots, T$$

- c_i are unobserved random variables (heterogeneity).
- Time period dummies:

$$\mathbf{d}_t = (d2_t, \dots, dT_t)$$

- ▶ Used to flexibly control for aggregate factors.

$$y_{it} = \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + c_i + u_{it}, t = 1, \dots, T$$

- \mathbf{x}_{it} only includes variables that have variation across i and t .
- $\boldsymbol{\beta}$ is of interest.
- Use fixed effects estimation to remove c_i .
- Sometimes called “two-way fixed effects,” but $\boldsymbol{\gamma}$ are parameters, c_i are not.

- Limitation of the model: If \mathbf{d}_t represents flexible trends, all units i follow the same trend.
- Allow heterogeneous time effects:

$$\begin{aligned}y_{it} &= \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\mathbf{g}_i + c_i + u_{it} \\ &= \mathbf{x}_{it}\boldsymbol{\beta} + d2_t g_{i2} + d3_t g_{i3} + \cdots + dT_t g_{iT} + c_i + u_{it}\end{aligned}$$

- Each unit has its own intercept, and these also vary across t .

- Now write

$$\mathbf{g}_i \equiv \boldsymbol{\gamma} + \mathbf{h}_i, \quad E(\mathbf{h}_i) = \mathbf{0}$$

$$y_{it} = \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + c_i + \mathbf{d}_t\mathbf{h}_i + u_{it}$$

- Cannot treat (c_i, \mathbf{h}_i) as parameters to estimate without assuming $\{u_{it} : t = 1, \dots, T\}$ are IID. [Ahn, Lee, Schmidt (1993, Journal of Econometrics)].
- Should (but will not) carefully compare with small- T factor literature, especially ALS (2013, J of E).
 - ▶ Current approach allows more heterogeneity.
 - ▶ Current approach is computationally much simpler.

Remainder of Talk

- When are (generalized) FE estimators that ignore $\mathbf{d}_t\mathbf{h}_i$ consistent?
- Modeling $E(\mathbf{h}_i|\mathbf{x}_i)$ in

$$y_{it} = \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + c_i + \mathbf{d}_t\mathbf{h}_i + u_{it}$$

use CRE and equivalence of different estimators.

- Empirical example.
- Extensions to unit-specific trends.

2. Robustness of Standard Fixed Effects Estimators

- Suppose the true model is

$$y_{it} = \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + c_i + \mathbf{d}_t\mathbf{h}_i + u_{it}$$

$$E(u_{it}|\mathbf{x}_i, c_i, \mathbf{h}_i) = 0$$

but we ignore $\mathbf{d}_t\mathbf{h}_i$.

- Apply standard FE to remove c_i .
- Can apply an extension of Wooldridge (2005, REStat).

$$\bar{\mathbf{x}}_i = T^{-1} \sum_{t=1}^T \mathbf{x}_{it}$$

$$\ddot{\mathbf{x}}_{it} = \mathbf{x}_{it} - \bar{\mathbf{x}}_i$$

- Sufficient is

$$E(\ddot{\mathbf{x}}_{it} \otimes \mathbf{h}_i) = \mathbf{0}, \quad t = 1, \dots, T.$$

- Holds if

$$\mathbf{x}_{it} = \mathbf{q}_i + \mathbf{e}_{it}$$

$$E(\mathbf{e}_{it} \otimes \mathbf{h}_i) = \mathbf{0}, \quad t = 1, \dots, T$$

$$\text{Cov}(\mathbf{q}_i, \mathbf{h}_i) \text{ unrestricted}$$

- If $T > 2$, can remove more heterogeneity from \mathbf{x}_{it} , such as unit-specific linear trends.
- Now obtain $\ddot{\mathbf{x}}_{it}$ – detrended covariates – as residuals from

$$\mathbf{x}_{it} \text{ on } 1, t, \quad t = 1, \dots, T$$

- This allows for unit-specific trends:

$$\mathbf{x}_{it} = \mathbf{q}_i + \mathbf{m}_i t + \mathbf{e}_{it}$$

$$\text{Cov}(\mathbf{q}_i, \mathbf{h}_i), \text{Cov}(\mathbf{m}_i, \mathbf{h}_i) \text{ unrestricted}$$

- But this representation for \mathbf{x}_{it} is still special.

3. CRE Approach to Heterogeneous Time Effects

- Now explicitly recognize the heterogeneity terms $\mathbf{d}_t \mathbf{h}_i$ in

$$y_{it} = \mathbf{x}_{it} \boldsymbol{\beta} + \mathbf{d}_t \boldsymbol{\gamma} + c_i + \mathbf{d}_t \mathbf{h}_i + u_{it}$$

$$E(u_{it} | \mathbf{x}_i, c_i, \mathbf{h}_i) = 0$$

- Remove c_i by within-unit demeaning:

$$\ddot{y}_{it} = \ddot{\mathbf{x}}_{it} \boldsymbol{\beta} + \ddot{\mathbf{d}}_t \boldsymbol{\gamma} + \ddot{\mathbf{d}}_t \mathbf{h}_i + \ddot{u}_{it}$$

- Make a Mundlak (1978) CRE assumption on \mathbf{h}_i :

$$E(h_{ir}|\mathbf{x}_i) = (\bar{\mathbf{x}}_i - \boldsymbol{\mu}_{\bar{\mathbf{x}}})\boldsymbol{\xi}_r, \quad r = 2, \dots, T$$

$$\boldsymbol{\mu}_{\bar{\mathbf{x}}} \equiv E(\bar{\mathbf{x}}_i)$$

- Leads to

$$\ddot{y}_{it} = \ddot{\mathbf{x}}_{it}\boldsymbol{\beta} + \ddot{\mathbf{d}}_t\boldsymbol{\gamma} + \ddot{d}_{2t}(\bar{\mathbf{x}}_i - \boldsymbol{\mu}_{\bar{\mathbf{x}}})\boldsymbol{\xi}_2 + \dots + \ddot{d}_{Tt}(\bar{\mathbf{x}}_i - \boldsymbol{\mu}_{\bar{\mathbf{x}}})\boldsymbol{\xi}_T + \ddot{e}_{it}$$

$$E(\ddot{e}_{it}|\mathbf{x}_i) = 0$$

- ▶ Replace $\boldsymbol{\mu}_{\bar{\mathbf{x}}}$ with the overall sample average, $\bar{\mathbf{x}}$.
- ▶ Can estimate all parameters consistently by POLS.

- Numerically identical to applying usual FE to the equation

$$y_{it} = \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + d2_t(\bar{\mathbf{x}}_i - \bar{\mathbf{x}})\boldsymbol{\xi}_2 + \cdots + dT_t(\bar{\mathbf{x}}_i - \bar{\mathbf{x}})\boldsymbol{\xi}_T + f_i + e_{it}$$

- Call the estimates $\hat{\boldsymbol{\beta}}_{FEH}, \hat{\boldsymbol{\gamma}}_{FEH}$.
- Use cluster-robust standard errors.
- Can use robust Wald test of

$$H_0 : \boldsymbol{\xi}_2 = \boldsymbol{\xi}_3 = \cdots = \boldsymbol{\xi}_T = \mathbf{0}$$

Comments

- It is possible to strongly reject

$$H_0 : \xi_2 = \xi_3 = \dots = \xi_T = \mathbf{0}$$

and have $\hat{\boldsymbol{\beta}}_{FEH} \approx \hat{\boldsymbol{\beta}}_{FE}$.

- ▶ Provides a robustness check on usual FE estimates.
- All estimates can be obtained from the Mundlak equation:

$$y_{it} = \alpha + \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + \bar{\mathbf{x}}_i\boldsymbol{\psi} + d2_t(\bar{\mathbf{x}}_i - \bar{\mathbf{x}})\xi_2 + \dots + dT_t(\bar{\mathbf{x}}_i - \bar{\mathbf{x}})\xi_T + r_{it}$$

- ▶ Pooled OLS or RE both produce $\hat{\boldsymbol{\beta}}_{FEH}, \hat{\boldsymbol{\gamma}}_{FEH}$.

- Can easily include time-constant covariates, say \mathbf{z}_i .
 - ▶ Just adding \mathbf{z}_i to the Mundlak equation changes none of the estimates (except intercept): Wooldridge (2019, Journal of Econometrics).
- Can add

$$ds_t(\mathbf{z}_i - \bar{\mathbf{z}})$$

to allow more heterogeneity in time effects.

- ▶ This will change the estimates of $\boldsymbol{\beta}$ (and $\boldsymbol{\gamma}$).

4. Empirical Illustration

- Data on $N = 1,149$ U.S. air routes. $T = 4$.
- $lfare_{it}$ is log of average airfare. $concen_{it}$ is a market power measure.

```
. egen double concenbar = mean(concen), by(id)
```

```
. sum concenbar
```

Variable	Obs	Mean	Std. Dev.	Min	Max
concenbar	4,596	.6101149	.1888741	.1862	.9997

```
. gen double concenbar_dm = concenbar - r(mean)
```

```
. sum ldist
```

Variable	Obs	Mean	Std. Dev.	Min	Max
ldist	4,596	6.696482	.6593177	4.553877	7.909857

```
. gen double ldist_dm = ldist - r(mean)
```

```
. xtreg lfare concen y98 y99 y00, fe vce(cluster id)
```

```
Fixed-effects (within) regression      Number of obs   =      4,596
Group variable: id                    Number of groups =      1,149
```

```
R-sq:                                Obs per group:
    within = 0.1352                    min =          4
    between = 0.0576                    avg  =         4.0
    overall = 0.0083                    max  =          4
```

```
corr(u_i, Xb) = -0.2033                F(4,1148)      =      120.06
                                                Prob > F       =       0.0000
```

(Std. Err. adjusted for 1,149 clusters in id)

lfare	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
concen	.168859	.0494587	3.41	0.001	.0718194	.2658985
y98	.0228328	.004163	5.48	0.000	.0146649	.0310007
y99	.0363819	.0051275	7.10	0.000	.0263215	.0464422
y00	.0977717	.0055054	17.76	0.000	.0869698	.1085735
_cons	4.953331	.0296765	166.91	0.000	4.895104	5.011557
sigma_u	.43389176					
sigma_e	.10651186					
rho	.94316439	(fraction of variance due to u_i)				

. * Usual Mundlak regression, with time constant variables added:

. reg lfare concen concenbar c.ldist c.ldist_dm#c.ldist_dm y98 y99 y00 , vce(cluster id)

Linear regression

Number of obs = 4,596
 F(7, 1148) = 181.88
 Prob > F = 0.0000
 R-squared = 0.4068
 Root MSE = .33637

(Std. Err. adjusted for 1,149 clusters in id)

lfare	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
concen	.168859	.0494749	3.41	0.001	.0717877	.2659303
concenbar	.2136346	.0816403	2.62	0.009	.0534538	.3738155
ldist	.4818306	.0178697	26.96	0.000	.4467698	.5168914
c.ldist_dm#c.ldist_dm	.1038426	.0201911	5.14	0.000	.064227	.1434582
y98	.0228328	.0041643	5.48	0.000	.0146622	.0310033
y99	.0363819	.0051292	7.09	0.000	.0263183	.0464455
y00	.0977717	.0055072	17.75	0.000	.0869663	.108577
_cons	1.551289	.1473768	10.53	0.000	1.262131	1.840447

```

. * Heterogeneous time effects using Mundlak:
.
. reg lfare concen concenbar c.ldist c.ldist_dm#c.ldist_dm y98 y99 y00 ///
>     c.y98#c.concenbar_dm c.y99#c.concenbar_dm c.y00#c.concenbar_dm, ///
>     vce(cluster id)

```

```

Linear regression           Number of obs   =       4,596
                          F(10, 1148)         =       136.24
                          Prob > F           =       0.0000
                          R-squared          =       0.4078
                          Root MSE       =       .33619

```

(Std. Err. adjusted for 1,149 clusters in id)

lfare	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
concen	.168456	.0490432	3.43	0.001	.0722316	.2646805
concenbar	.116914	.083664	1.40	0.163	-.0472374	.2810655
ldist	.4818306	.0178755	26.95	0.000	.4467583	.5169029
c.ldist_dm#c.ldist_dm	.1038426	.0201977	5.14	0.000	.0642141	.1434711
y98	.0228364	.0041561	5.49	0.000	.0146819	.0309908
y99	.0363788	.0050715	7.17	0.000	.0264284	.0463291
y00	.0977672	.0053859	18.15	0.000	.0871999	.1083346

c.y98#c.concenbar_dm		.0616642	.0232143	2.66	0.008	.0161169	.1072114
c.y99#c.concenbar_dm		.1307868	.0285472	4.58	0.000	.0747762	.1867974
c.y00#c.concenbar_dm		.1960431	.0318187	6.16	0.000	.1336138	.2584724
_cons		1.610546	.1486973	10.83	0.000	1.318797	1.902295

. * FE gives same estimates:

```
.  
. xtreg lfare concen y98 y99 y00 ///  
>       c.y98#c.concenbar_dm c.y99#c.concenbar_dm c.y00#c.concenbar_dm, fe ///  
>       vce(cluster id)
```

(Std. Err. adjusted for 1,149 clusters in id)

lfare	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
concen	.168456	.0490272	3.44	0.001	.0722631	.264649
y98	.0228364	.0041548	5.50	0.000	.0146846	.0309881
y99	.0363788	.0050698	7.18	0.000	.0264317	.0463259
y00	.0977672	.0053842	18.16	0.000	.0872033	.1083312
c.y98#c.concenbar_dm	.0616642	.0232067	2.66	0.008	.0161318	.1071965
c.y99#c.concenbar_dm	.1307868	.0285379	4.58	0.000	.0747945	.1867791
c.y00#c.concenbar_dm	.1960431	.0318083	6.16	0.000	.1336342	.258452
_cons	4.953577	.0293317	168.88	0.000	4.896028	5.011127

. * Interact distance with time dummies:

```
. reg lfare concen concenbar ldist c.ldist_dm#c.ldist_dm y98 y99 y00 ///
>     c.y98#c.concenbar_dm c.y99#c.concenbar_dm c.y00#c.concenbar_dm ///
>     c.y98#c.ldist_dm c.y99#c.ldist_dm c.y00#c.ldist_dm, ///
>     vce(cluster id)
```

```
Linear regression                Number of obs    =      4,596
                                F(13, 1148)      =      106.35
                                Prob > F            =      0.0000
                                R-squared           =      0.4082
                                Root MSE        =      .33618
```

(Std. Err. adjusted for 1,149 clusters in id)

lfare	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]		

concen	.1681968	.0492655	3.41	0.001	.0715364	.2648573	
concenbar	.1497355	.0869212	1.72	0.085	-.0208067	.3202777	
ldist	.4986949	.0197181	25.29	0.000	.4600074	.5373823	
c.ldist_dm#c.ldist_dm	.1038426	.0202043	5.14	0.000	.0642011	.1434841	
y98	.0228387	.0041539	5.50	0.000	.0146886	.0309888	
y99	.0363768	.0050668	7.18	0.000	.0264355	.046318	
y00	.0977644	.0053295	18.34	0.000	.0873077	.108221	

c.y98#c.concenbar_dm	.0297475	.0267611	1.11	0.267	-.0227587	.0822536
c.y99#c.concenbar_dm	.1160657	.0343829	3.38	0.001	.0486053	.1835261
c.y00#c.concenbar_dm	.1124318	.036763	3.06	0.002	.0403016	.1845619
c.y98#c.ldist_dm	-.0165298	.0068546	-2.41	0.016	-.0299786	-.0030809
c.y99#c.ldist_dm	-.0076258	.009357	-0.81	0.415	-.0259845	.0107329
c.y00#c.ldist_dm	-.0433015	.0099287	-4.36	0.000	-.0627819	-.0238211
_cons	1.477749	.1633311	9.05	0.000	1.157288	1.79821

```
-----
. test c.y98#c.concenbar_dm c.y99#c.concenbar_dm c.y00#c.concenbar_dm ///
>      c.y98#c.ldist_dm c.y99#c.ldist_dm c.y00#c.ldist_dm
```

```
( 1) c.y98#c.concenbar_dm = 0
( 2) c.y99#c.concenbar_dm = 0
( 3) c.y00#c.concenbar_dm = 0
( 4) c.y98#c.ldist_dm = 0
( 5) c.y99#c.ldist_dm = 0
( 6) c.y00#c.ldist_dm = 0
```

```
F( 6, 1148) = 12.72
Prob > F = 0.0000
```

5. Allowing for Heterogeneous Slopes

- A model where everything is heterogeneous:

$$y_{it} = \mathbf{x}_{it}\mathbf{b}_i + c_i + \mathbf{d}_t\mathbf{g}_i + u_{it}$$

$$\mathbf{b}_i = \boldsymbol{\beta} + \mathbf{a}_i$$

$$\mathbf{g}_i = \boldsymbol{\gamma} + \mathbf{h}_i$$

$$y_{it} = \mathbf{x}_{it}\boldsymbol{\beta} + c_i + \mathbf{d}_t\boldsymbol{\gamma} + \mathbf{x}_{it}\mathbf{a}_i + \mathbf{d}_t\mathbf{h}_i + u_{it}$$

$$\mathbf{a}_i = \boldsymbol{\Lambda}(\bar{\mathbf{x}}_i - \boldsymbol{\mu}_{\bar{\mathbf{x}}})' + \mathbf{q}_i$$

$$\mathbf{h}_i = \boldsymbol{\Xi}(\bar{\mathbf{x}}_i - \boldsymbol{\mu}_{\bar{\mathbf{x}}})' + \mathbf{f}_i$$

$$\begin{aligned}
y_{it} &= \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + \mathbf{x}_{it}[\boldsymbol{\Lambda}(\bar{\mathbf{x}}_i - \boldsymbol{\mu}_{\bar{\mathbf{x}}})'] + \mathbf{d}_t[\boldsymbol{\Lambda}(\bar{\mathbf{x}}_i - \boldsymbol{\mu}_{\bar{\mathbf{x}}})'] + c_i + v_{it} \\
&= \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + [(\bar{\mathbf{x}}_i - \boldsymbol{\mu}_{\bar{\mathbf{x}}}) \otimes \mathbf{x}_{it}]\boldsymbol{\lambda} + [(\bar{\mathbf{x}}_i - \boldsymbol{\mu}_{\bar{\mathbf{x}}}) \otimes \mathbf{d}_t]\boldsymbol{\xi} + c_i + v_{it}
\end{aligned}$$

$$E(v_{it}|\mathbf{x}_i) = 0, t = 1, \dots, T$$

- Remove c_i using fixed effects.
- Or apply Mundlak, adding

$$\bar{\mathbf{x}}_i, (\bar{\mathbf{x}}_i - \bar{\mathbf{x}}) \otimes \bar{\mathbf{x}}_i$$

- Can also add

$$(\mathbf{z}_i - \bar{\mathbf{z}}) \otimes \mathbf{x}_{it}$$


```

. * Now heterogeneous slope on concen, too.
.
. reg lfare concen c.concen#c.concenbar_dm c.concen#c.ldist_dm ///
>   concenbar ldist c.ldist_dm#c.ldist_dm y98 y99 y00 ///
>   c.concenbar#c.concenbar_dm c.concenbar#c.ldist_dm ///
>   c.y98#c.concenbar_dm c.y99#c.concenbar_dm c.y00#c.concenbar_dm ///
>   c.y98#c.ldist_dm c.y99#c.ldist_dm c.y00#c.ldist_dm, ///
>   vce(cluster id)

```

(Std. Err. adjusted for 1,149 clusters in

lfare	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval
concen	.1690247	.0500444	3.38	0.001	.0708359 .2672134
c.concen#c.concenbar_dm	.3101472	.4075416	0.76	0.447	-.4894626 1.109757
c.concen#c.ldist_dm	-.2097112	.0907065	-2.31	0.021	-.3876804 -.0317421
concenbar	.1396369	.2389655	0.58	0.559	-.3292212 .6084949
ldist	.3880281	.0856013	4.53	0.000	.2200756 .5559807
c.ldist_dm#c.ldist_dm	.1286053	.0225076	5.71	0.000	.0844446 .172766
y98	.0230028	.0041427	5.55	0.000	.0148747 .0311308
y99	.0355884	.0050872	7.00	0.000	.0256071 .0455698
y00	.0976618	.005291	18.46	0.000	.0872806 .108043

c.concenbar#c.concenbar_dm	-.2588744	.5716612	-0.45	0.651	-1.380492	.8627436
c.concenbar#c.ldist_dm	.4014202	.1687987	2.38	0.018	.0702316	.7326088
c.y98#c.concenbar_dm	.0262151	.0272121	0.96	0.336	-.027176	.0796061
c.y99#c.concenbar_dm	.1131609	.0340452	3.32	0.001	.0463631	.1799587
c.y00#c.concenbar_dm	.1065938	.0363228	2.93	0.003	.0353272	.1778604
c.y98#c.ldist_dm	-.0144794	.0069094	-2.10	0.036	-.0280359	-.000923
c.y99#c.ldist_dm	-.0080089	.0093604	-0.86	0.392	-.0263743	.0103564
c.y00#c.ldist_dm	-.0442223	.0096118	-4.60	0.000	-.0630809	-.0253637
_cons	2.225276	.6831325	3.26	0.001	.8849479	3.565604

6. Allowing for Unit-Specific Trends

- Now allow the base model to have heterogeneous linear trends:

$$y_{it} = \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + c_i + g_{it} + \mathbf{d}_t\mathbf{h}_i + u_{it}$$

- In removing the linear trends, lose one element of \mathbf{d}_t .
- In running the regressions (for each i),

$$\mathbf{x}_{it} \text{ on } 1, t, \quad t = 1, \dots, T,$$

let

$$\hat{\mathbf{x}}_{it} = \hat{\mathbf{a}}_{i0} + \hat{\mathbf{a}}_{i1}t$$

- Detrended values:

$$\ddot{\mathbf{x}}_{it} = \mathbf{x}_{it} - \hat{\mathbf{x}}_{it}$$

$$\ddot{y}_{it} = \ddot{\mathbf{x}}_{it}\boldsymbol{\beta} + \ddot{\mathbf{d}}_t\boldsymbol{\gamma} + \ddot{\mathbf{d}}_t\mathbf{h}_i + \ddot{u}_{it}$$

- Model \mathbf{h}_i as a (linear) function of

$$\hat{\mathbf{a}}_i = (\hat{\mathbf{a}}_{i0}, \hat{\mathbf{a}}_{i1})$$

- Use pooled OLS on

$$\ddot{y}_{it} = \ddot{\mathbf{x}}_{it}\boldsymbol{\beta} + \ddot{\mathbf{d}}_t\boldsymbol{\gamma} + (\ddot{\mathbf{d}}_t \otimes \hat{\mathbf{a}}_i)\boldsymbol{\xi} + \ddot{e}_{it}$$

- Using an extension of the Mundlak result, can show it is algebraically equivalent to using POLS on

$$y_{it} = \alpha + \mathbf{x}_{it}\boldsymbol{\beta} + \mathbf{d}_t\boldsymbol{\gamma} + \bar{\mathbf{x}}_i\boldsymbol{\eta} + \hat{\mathbf{x}}_{it}\boldsymbol{\theta} + (\mathbf{d}_t \otimes \hat{\mathbf{a}}_i)\boldsymbol{\xi} + r_{it}$$

$$\hat{\mathbf{x}}_{it} = \hat{\mathbf{a}}_{i0} + \hat{\mathbf{a}}_{i1}t$$

- Apply to a passenger “demand” model.

```
. * First apply the previous model where only time averages are used.

. reg lpassen concen concenbar ldist ldistsq y98 y99 y00 ///
> c.y98#c.concenbar_dm c.y99#c.concenbar_dm c.y00#c.concenbar_dm ///
> lfare lfarebar c.y98#c.lfarebar_dm c.y99#c.lfarebar_dm ///
> c.y00#c.lfarebar_dm, vce(cluster id)
```

```
Linear regression                               Number of obs   =       4,596
                                                F(15, 1148)    =       39.97
                                                Prob > F       =       0.0000
                                                R-squared      =       0.0770
                                                Root MSE      =       .85065
```

(Std. Err. adjusted for 1,149 clusters in id)

lpassen	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
concen	.1390506	.0899681	1.55	0.122	-.0374697	.3155709
concenbar	-.6315775	.2026034	-3.12	0.002	-1.029092	-.2340631
ldist	-1.303931	.7013855	-1.86	0.063	-2.680072	.0722106
ldistsq	.0949819	.05304	1.79	0.074	-.0090842	.199048
y98	.045365	.004841	9.37	0.000	.0358667	.0548632
y99	.1035984	.006301	16.44	0.000	.0912357	.1159612
y00	.1966456	.0101262	19.42	0.000	.1767777	.2165135

c.y98#c.concenbar_dm	-.0011952	.026645	-0.04	0.964	-.0534735	.0510832
c.y99#c.concenbar_dm	-.035792	.0343854	-1.04	0.298	-.1032573	.0316732
c.y00#c.concenbar_dm	-.1049551	.0379231	-2.77	0.006	-.1793614	-.0305489
lfare	-1.15989	.1091893	-10.62	0.000	-1.374123	-.9456567
lfarebar	.7215856	.1377878	5.24	0.000	.4512414	.9919297
c.y98#c.lfarebar_dm	-.053968	.0112635	-4.79	0.000	-.0760674	-.0318685
c.y99#c.lfarebar_dm	-.054394	.0133306	-4.08	0.000	-.0805491	-.028239
c.y00#c.lfarebar_dm	-.0275061	.0166339	-1.65	0.098	-.0601424	.0051302
_cons	12.89575	2.323339	5.55	0.000	8.337285	17.45422

```
. * Now use the unit-specific linear trend heterogeneity:
```

```
. gen double lpassen_h = .  
(4,596 missing values generated)
```

```
. gen double lpassen_dt = .  
(4,596 missing values generated)
```

```
. gen double lfare_h = .  
(4,596 missing values generated)
```

```
. gen double lfare_dt = .  
(4,596 missing values generated)
```

```
. gen double lfare_a0 = .  
(4,596 missing values generated)
```

```
. gen double lfare_a1 = .  
(4,596 missing values generated)
```

```
. gen double concen_h = .  
(4,596 missing values generated)
```

```
. gen double concen_dt = .  
(4,596 missing values generated)
```

```
. gen double concen_a0 = .  
(4,596 missing values generated)
```



```
. gen double concen_a1 = .  
(4,596 missing values generated)  
  
. gen double y98_h = .  
(4,596 missing values generated)  
  
. gen double y99_h = .  
(4,596 missing values generated)  
  
. gen double y00_h = .  
(4,596 missing values generated)  
  
. gen double y98_dt = .  
(4,596 missing values generated)  
  
. gen double y99_dt = .  
(4,596 missing values generated)  
  
. gen double y00_dt = .  
(4,596 missing values generated)  
  
. gen double y98_a1_t = .  
(4,596 missing values generated)  
  
. gen double y99_a1_t = .  
(4,596 missing values generated)  
  
. gen double y00_a1_t = .  
(4,596 missing values generated)
```

```

. gen t = year - 1997

. local i = 1

. while `i' <= 1149 {
2.
.     qui reg lpassen t if id == `i'
3.         predict lpassen_t, xb
4.         qui replace lpassen_h = lpassen_t if id == `i'
5.         qui replace lpassen_dt = lpassen - lpassen_t if id == `i'
6.
.     qui reg lfare t if id == `i'
7.         qui replace lfare_a0 = _b[_cons] if id == `i'
8.         qui replace lfare_a1 = _b[t] if id == `i'
9.         predict lfare_t, xb
10.        qui replace lfare_h = lfare_t if id == `i'
11.        qui replace lfare_dt = lfare - lfare_t if id == `i'
12.
.     qui reg concen t if id == `i'
13.        qui replace concen_a0 = _b[_cons] if id == `i'
14.        qui replace concen_a1 = _b[t] if id == `i'
15.        predict concen_t, xb
16.        qui replace concen_h = concen_t if id == `i'
17.        qui replace concen_dt = concen - concen_t if id == `i'
18.

```

```

.      qui reg y98 t if id == `i'
19.    predict y98_t, xb
20.      qui replace y98_h = y98_t if id == `i'
21.    qui replace y98_dt = y98 - y98_t if id == `i'
22.      qui replace y98_a1_t = _b[t]*t if id == `i'
23.
.      qui reg y99 t if id == `i'
24.    predict y99_t, xb
25.      qui replace y99_h = y99_t if id == `i'
26.    qui replace y99_dt = y99 - y99_t if id == `i'
27.      qui replace y99_a1_t = _b[t]*t if id == `i'
28.
.    qui reg y00 t if id == `i'
29.      predict y00_t, xb
30.        qui replace y00_h = y00_t if id == `i'
31.      qui replace y00_dt = y00 - y00_t if id == `i'
32.        qui replace y00_a1_t = _b[t]*t if id == `i'
33.
.    drop lfare_t concen_t lpassen_t y98_t y99_t y00_t
34.
.    local i = `i' + 1
35. }

```

```
. sum concen_h
```

Variable	Obs	Mean	Std. Dev.	Min	Max
concen_h	4,596	.6101149	.1929988	.15482	1.0747

```
. gen double concen_h_dm = concen_h - r(mean)
```

```
. sum lfare_h
```

Variable	Obs	Mean	Std. Dev.	Min	Max
lfare_h	4,596	5.095601	.4322098	3.836819	6.26246

```
. gen double lfare_h_dm = lfare_h - r(mean)
```

```
. sum concen_a0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
concen_a0	4,596	.6175428	.1959924	.15482	1.01686

```
. gen double concen_a0_dm = concen_a0 - r(mean)
```

```
. sum concen_a1
```

Variable	Obs	Mean	Std. Dev.	Min	Max
concen_a1	4,596	-.0049519	.0351504	-.16346	.1463

```
. gen double concen_a1_dm = concen_a1 - r(mean)
```

```
. sum lfare_a0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
lfare_a0	4,596	5.050825	.4526443	3.836819	6.069202

```
. gen double lfare_a0_dm = lfare_a0 - r(mean)
```

```
. sum lfare_a1
```

Variable	Obs	Mean	Std. Dev.	Min	Max
lfare_a1	4,596	.0298502	.0637264	-.3344203	.5016729

```
. gen double lfare_a1_dm = lfare_a1 - r(mean)
```

```

. reg lpassen_dt concen_dt lfare_dt y98_dt y99_dt y00_dt ///
> c.y98_dt#c.concen_a0_dm c.y99_dt#c.concen_a0_dm c.y00_dt#c.concen_a0_dm ///
> c.y98_dt#c.lfare_a0_dm c.y99_dt#c.lfare_a0_dm c.y00_dt#c.lfare_a0_dm ///
> c.y98_dt#c.concen_a1_dm c.y99_dt#c.concen_a1_dm c.y00_dt#c.concen_a1_dm ///
> c.y98_dt#c.lfare_a1_dm c.y99_dt#c.lfare_a1_dm c.y00_dt#c.lfare_a1_dm,
> nocons vce(cluster id)

```

```

note: y00_dt omitted because of collinearity
note: c.y00_dt#c.concen_a0_dm omitted because of collinearity
note: c.y00_dt#c.lfare_a0_dm omitted because of collinearity
note: c.y00_dt#c.concen_a1_dm omitted because of collinearity
note: c.y00_dt#c.lfare_a1_dm omitted because of collinearity

```

```

Linear regression                Number of obs    =      4,596
                                F(12, 1148)      =      34.15
                                Prob > F              =      0.0000
                                R-squared              =      0.4038
                                Root MSE           =      .07314

```

(Std. Err. adjusted for 1,149 clusters in id)

lpassen_dt	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
concen_dt	.0049276	.0951786	0.05	0.959	-.181816	.1916712
lfare_dt	-1.006965	.0660814	-15.24	0.000	-1.136619	-.8773108
y98_dt	-.017328	.004024	-4.31	0.000	-.0252232	-.0094328
y99_dt	-.0231335	.0044971	-5.14	0.000	-.0319568	-.0143101
y00_dt	0	(omitted)				

c.y98_dt#c.concen_a0_dm	.0247107	.0192804	1.28	0.200	-.013118	.0625394
c.y99_dt#c.concen_a0_dm	.0308164	.0193918	1.59	0.112	-.007231	.0688639
c.y00_dt#c.concen_a0_dm	0	(omitted)				
c.y98_dt#c.lfare_a0_dm	-.0329983	.0083758	-3.94	0.000	-.0494319	-.0165648
c.y99_dt#c.lfare_a0_dm	-.0375524	.0079381	-4.73	0.000	-.0531273	-.0219776
c.y00_dt#c.lfare_a0_dm	0	(omitted)				
c.y98_dt#c.concen_a1_dm	.0450727	.1310509	0.34	0.731	-.2120534	.3021988
c.y99_dt#c.concen_a1_dm	.0196799	.1975331	0.10	0.921	-.3678863	.4072462
c.y00_dt#c.concen_a1_dm	0	(omitted)				
c.y98_dt#c.lfare_a1_dm	.198575	.1346018	1.48	0.140	-.0655181	.4626681
c.y99_dt#c.lfare_a1_dm	.0075498	.2000518	0.04	0.970	-.3849584	.400058
c.y00_dt#c.lfare_a1_dm	0	(omitted)				

```

. reg lpassen concen lfare c.ldist c.ldist_dm#c.ldist_dm y98 y99 y00 ///
>   concenbar lfarebar concen_h lfare_h ///
>   c.y98#c.concen_a0_dm c.y99#c.concen_a0_dm c.y00#c.concen_a0_dm ///
>   c.y98#c.lfare_a0_dm c.y99#c.lfare_a0_dm c.y00#c.lfare_a0_dm ///
>   c.y98#c.concen_a1_dm c.y99#c.concen_a1_dm c.y00#c.concen_a1_dm ///
>   c.y98#c.lfare_a1_dm c.y99#c.lfare_a1_dm c.y00#c.lfare_a1_dm, vce(cluster id)

```

```

Linear regression                               Number of obs   =       4,596
                                                F(23, 1148)     =       33.03
                                                Prob > F        =       0.0000
                                                R-squared       =       0.1057
                                                Root MSE       =       .83806

```

(Std. Err. adjusted for 1,149 clusters in id)

	lpassen	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	

	concen	.0049276	.0953035	0.05	0.959	-.1820609	.1919161
	lfare	-1.006965	.0661681	-15.22	0.000	-1.136789	-.877141
	ldist	-.01635	.0592676	-0.28	0.783	-.132635	.099935
	c.ldist_dm#c.ldist_dm	.121346	.0530347	2.29	0.022	.0172902	.2254018
	y98	.0242774	.0112826	2.15	0.032	.0021406	.0464141
	y99	.0600772	.0212456	2.83	0.005	.0183926	.1017618
	y00	.124816	.0300917	4.15	0.000	.0657752	.1838568

concenbar	-2.458434	.4773988	-5.15	0.000	-3.395106	-1.521762
lfarebar	-.475291	.3082788	-1.54	0.123	-1.080144	.1295621
concen_h	2.055851	.4549934	4.52	0.000	1.163139	2.948563
lfare_h	.962927	.3047852	3.16	0.002	.3649286	1.560925
c.y98#c.concen_a0_dm	-.0040696	.024358	-0.17	0.867	-.0518608	.0437216
c.y99#c.concen_a0_dm	-.026744	.0318316	-0.84	0.401	-.0891988	.0357107
c.y00#c.concen_a0_dm	-.0863407	.0366607	-2.36	0.019	-.1582702	-.0144112
c.y98#c.lfare_a0_dm	-.0423505	.0109446	-3.87	0.000	-.0638241	-.0208769
c.y99#c.lfare_a0_dm	-.0562567	.0141338	-3.98	0.000	-.0839876	-.0285258
c.y00#c.lfare_a0_dm	-.0280564	.0171804	-1.63	0.103	-.061765	.0056521
c.y98#c.concen_a1_dm	-1.814706	.474012	-3.83	0.000	-2.744733	-.8846795
c.y99#c.concen_a1_dm	-3.699879	.9978239	-3.71	0.000	-5.657642	-1.742116
c.y00#c.concen_a1_dm	-5.579338	1.415692	-3.94	0.000	-8.356972	-2.801704
c.y98#c.lfare_a1_dm	-1.028896	.3812075	-2.70	0.007	-1.776838	-.2809546
c.y99#c.lfare_a1_dm	-2.447393	.7684671	-3.18	0.001	-3.955151	-.9396358
c.y00#c.lfare_a1_dm	-3.682415	1.09094	-3.38	0.001	-5.822874	-1.541956
_cons	8.910393	.4332756	20.57	0.000	8.060292	9.760494

7. Extension to Unbalanced Panels

- Unbalanced panels, using the complete cases.
- Use functions of

$$\{(s_{it}, s_{it}\mathbf{X}_{it}) : t = 1, \dots, T\}$$

- Now more care is needed with the aggregate time dummies [Wooldridge (2019)].