

BOSTON COLLEGE  
 Department of Economics  
 EC 228 02 Econometric Methods  
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 Problem Set 3  
 Due at classtime, Tuesday 13 Oct 2009

C3.1

(i) (2 pts.) Probably  $\beta_2 > 0$ , as more income typically means better nutrition for the mother and better prenatal care.

(ii) (4 pts.) Yes, they are likely correlated and an argument can be made for both positive or negative correlation. On the one hand, an increase in income generally increases the consumption of a good, and *cigs* and *faminc* could be positively correlated. On the other hand, family incomes are also higher for families with more education, and more education and cigarette smoking tend to be negatively correlated. The sample correlation between *cigs* and *faminc* is about -0.173, indicating a negative correlation.

(iii) (4 pts.)

. regress bwght cigs

Source	SS	df	MS	Number of obs =	694
Model	10394.4794	1	10394.4794	F( 1, 692) =	25.33
Residual	283941.338	692	410.319852	Prob > F =	0.0000
Total	294335.817	693	424.727009	R-squared =	0.0353
				Adj R-squared =	0.0339
				Root MSE =	20.256

bwght	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
cigs	-.601789	.119565	-5.03	0.000	-.8365427	-.3670353
_cons	120.3839	.821228	146.59	0.000	118.7715	121.9963

. regress bwght cigs faminc

Source	SS	df	MS			
Model	11626.062	2	5813.03102	Number of obs = 694		
Residual	282709.755	691	409.131339	F( 2, 691) = 14.21		
				Prob > F = 0.0000		
				R-squared = 0.0395		
				Adj R-squared = 0.0367		
				Root MSE = 20.227		

  

bwght	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
cigs	-.5632265	.1214429	-4.64	0.000	-.801668	-.3247851
faminc	.073165	.0421699	1.74	0.083	-.0096316	.1559616
_cons	118.1664	1.518518	77.82	0.000	115.185	121.1479

For the regression without *faminc*:

$$\widehat{bwght} = 120.3839 - .601789cigs, n=694, R^2 = 0.0353$$

For the regression with *faminc*:

$$\widehat{bwght} = 118.1664 - .5632cigs + .0732faminc, n=694, R^2 = 0.0395$$

The effect of cigarette smoking is slightly smaller when *faminc* is added to the regression, but the difference is not great. This is due to the fact that *cigs* and *faminc* are not very correlated, and the coefficient on *faminc* is practically small. (The variable *faminc* is measured in thousands, so 10000 more dollars in 1988 income increases predicted weight by only 0.93 ounces.)

### C3.2

(i) (2 pts.)

```
. regress price sqrft bdrms
```

Source	SS	df	MS			
Model	580009.152	2	290004.576	Number of obs = 88		
Residual	337845.354	85	3974.65122	F( 2, 85) = 72.96		
				Prob > F = 0.0000		
				R-squared = 0.6319		
				Adj R-squared = 0.6233		
				Root MSE = 63.045		

  

price	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

<i>sqrft</i>		.1284362	.0138245	9.29	0.000	.1009495	.1559229
<i>bdrms</i>		15.19819	9.483517	1.60	0.113	-3.657582	34.05396
<i>_cons</i>		-19.315	31.04662	-0.62	0.536	-81.04399	42.414

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The estimated equation is

$$\widehat{price} = -19.32 + .128sqrft + 15.20bdrms$$

$$n = 88, R^2 = .632$$

- (ii) (2 pts.) Holding square footage constant,  $\Delta\widehat{price} = 15.20\Delta bdrms$ , and so  $\widehat{price}$  increases by 15.20, which means \$15,200.
- (iii) (2 pts.) Now  $\Delta\widehat{price} = .128\Delta sqrft + 15.20\Delta bdrms = .128(140) + 15.20 = 33.12$ , or \$33,120. Because the size of the house is increasing, this is a much larger effect than in(ii).
- (iv) (2 pts.) About 63.2% from  $R^2$ .
- (v) (2 pts.) The predicted price is  $-19.32 + .128(2,438) + 15.20(4) = 353.544$ , or \$353,544.
- (vi) (2 pts.) From part (v), the estimated value of the home based only on square footage and number of bedrooms is \$353,544. The actual selling price was \$300,000, which suggests the buyer underpaid by some margin. But, of course, there are many other features of a house (some that we cannot even measure) that affect price, and we have not controlled for these.

### C3.4

- (i) (2 pts.) The minimum, maximum, and average values for these three variables are given in the table below. Use the command "summarize atndrte priGPA ACT".

Variable	Average	Minimum	Maximum
<i>atndrte</i>	81.71	6.25	100
<i>priGPA</i>	2.59	0.86	3.93
<i>ACT</i>	22.51	13	32

(ii) (4 pts.)

```
. regress atndrte priGPA ACT
```

Source	SS	df	MS			
Model	57336.7612	2	28668.3806	Number of obs =	680	
Residual	139980.564	677	206.765974	F( 2, 677) =	138.65	
Total	197317.325	679	290.59989	Prob > F =	0.0000	
				R-squared =	0.2906	
				Adj R-squared =	0.2885	
				Root MSE =	14.379	

  

atndrte	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
priGPA	17.26059	1.083103	15.94	0.000	15.13395	19.38724
ACT	-1.716553	.169012	-10.16	0.000	-2.048404	-1.384702
_cons	75.7004	3.884108	19.49	0.000	68.07406	83.32675

The estimated equation is

$$\widehat{atndrte} = 75.70 + 17.26priGPA - 1.72ACT$$

$$n = 680, R^2 = 0.291$$

The intercept means that, for a student whose prior GPA is zero and ACT score is zero, the predicted attendance rate is 75.7%. But this is clearly not an interesting segment of the population. (In fact, there are no students in the college population with  $priGPA = 0$  and  $ACT = 0$ , or with values even close to zero.)

- (iii) (2 pts) The coefficient on  $priGPA$  means that, if a student's prior GPA is one point higher (say, from 2.0 to 3.0), the attendance rate is about 17.3 percentage points higher. This holds  $ACT$  fixed. The negative coefficient on  $ACT$  is, perhaps initially a bit surprising. Five more points on the  $ACT$  is predicted to lower attendance by 8.6 percentage points at a given level of  $priGPA$ . As  $priGPA$  measures performance in college (and, at least partially, could reflect, past attendance rates), while  $ACT$  is a measure of potential in college, it appears that students that had more promise (which could mean more innate ability) think they can get by with missing lectures.

- (iv) (2 pts) We have  $\widehat{atndrte} = 75.70 + 17.267(3.65) - 1.72(20) \approx 104.3$ . Of course, a student cannot have higher than a 100% attendance rate. Getting predictions like this is always possible when using regression methods for dependent variables with natural upper or lower bounds. In practice, we would predict a 100% attendance rate for this student. (In fact, this student had an actual attendance rate of 87.5%.)
- (v) (2 pts) The difference in predicted attendance rates for A and B is  $17.26(3.1 - 2.1) - (21 - 26) = 25.86$ .

C3.8

- (i) (2 pts.)

```
. summarize prpblck income
```

Variable	Obs	Mean	Std. Dev.	Min	Max
prpblck	409	.1134864	.1824165	0	.9816579
income	409	47053.78	13179.29	15919	136529

The average of *prpblck* is .113 with standard deviation .182; the average of *income* is 47,053.78 with standard deviation 13,179.29. It is evident that *prpblck* is a proportion and that *income* is measured in dollars.

- (ii) (2 pts)

```
. regress psoda prpblck income
```

Source	SS	df	MS	Number of obs =	401
Model	.202552215	2	.101276107	F( 2, 398) =	13.66
Residual	2.95146493	398	.007415741	Prob > F =	0.0000
Total	3.15401715	400	.007885043	R-squared =	0.0642
				Adj R-squared =	0.0595
				Root MSE =	.08611

  

psoda	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
prpblck	.1149882	.0260006	4.42	0.000	.0638724 .1661039
income	1.60e-06	3.62e-07	4.43	0.000	8.91e-07 2.31e-06
_cons	.9563196	.018992	50.35	0.000	.9189824 .9936568

The results from the OLS regression are

$$\widehat{psoda} = .956 + .115prpbck + .0000016income$$

$$n = 401, R^2 = .064$$

If say *prpbck* increases by .10 (ten percentage point), the price of soda is estimated to increase by .0115 dollars, or about 1.2 cents. While this does not seem large, there are communities with no black population and others that are almost all black, in which case the difference in *psoda* is estimated to be almost 11.5 cents.

(iii) (2 pts.)

```
. regress psoda prpbck
```

Source	SS	df	MS	Number of obs =	401
Model	.057010466	1	.057010466	F( 1, 399) =	7.34
Residual	3.09700668	399	.007761922	Prob > F =	0.0070
Total	3.15401715	400	.007885043	R-squared =	0.0181
				Adj R-squared =	0.0156
				Root MSE =	.0881

  

psoda	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
prpbck	.0649269	.023957	2.71	0.007	.0178292 .1120245
_cons	1.037399	.0051905	199.87	0.000	1.027195 1.047603

The simple regression estimate on *prpbck* is .065, so the simple regression estimate is actually lower. This is because *prpbck* and *income* are negatively correlated (-.43) and *income* has a positive coefficient in the multiple regression. You can see the negative correlation by using the command "corr prpbck income".

(iv) (2 pts.)

```
. regress lpsoda prpblck lincome
```

Source	SS	df	MS			
Model	.196020672	2	.098010336	Number of obs =	401	
Residual	2.68272938	398	.006740526	F( 2, 398) =	14.54	
Total	2.87875005	400	.007196875	Prob > F =	0.0000	
				R-squared =	0.0681	
				Adj R-squared =	0.0634	
				Root MSE =	.0821	

  

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
lpsoda						
prpblck	.1215803	.0257457	4.72	0.000	.0709657	.1721948
lincome	.0765114	.0165969	4.61	0.000	.0438829	.1091399
_cons	-.793768	.1794337	-4.42	0.000	-1.146524	-.4410117

$$\log(\widehat{psoda}) = -.794 + .122prpblck + .077lincome$$

$$n = 401, R^2 = .068$$

If *prpblck* increases by .20,  $\log(\widehat{psoda})$  is estimated to increase by  $.20(.122) = .0244$ , or about 2.44 percent.

(v) (2 pts.)

```
. regress lpsoda prpblck lincome prppov
```

Source	SS	df	MS			
Model	.250340622	3	.083446874	Number of obs =	401	
Residual	2.62840943	397	.006620679	F( 3, 397) =	12.60	
Total	2.87875005	400	.007196875	Prob > F =	0.0000	
				R-squared =	0.0870	
				Adj R-squared =	0.0801	
				Root MSE =	.08137	

  

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
lpsoda						
prpblck	.0728072	.0306756	2.37	0.018	.0125003	.1331141
lincome	.1369553	.0267554	5.12	0.000	.0843552	.1895553
prppov	.38036	.1327903	2.86	0.004	.1192999	.6414201
_cons	-1.463333	.2937111	-4.98	0.000	-2.040756	-.8859092

$\hat{\beta}_{prpblck}$  falls to about .073 when *prppov* is added to the regression.

(vi) (2 pts.)

```
. corr lincome prppov
(obs=409)

          | lincome  prppov
-----+-----
lincome |   1.0000
prppov  |  -0.8385   1.0000
```

The correlation is about -.84, which makes sense because poverty rates are determined by income (but not directly in terms of median income).

(vii) (2 pts.) There is no argument that they are highly correlated, but we are using them simply as controls to determine if there is price discrimination against blacks. In order to isolate the pure discrimination effect, we need to control for as many measures of income as we can; therefore, including both variables makes sense.

#### 4.1

(i) (2 pts.) Heteroskedasticity generally causes the t statistics not to have a t distribution under  $H_0$ . Homoskedasticity is one of the CLM assumptions.

(ii) (2 pts.) The CLM assumptions contain no mention of the sample correlations among independent variables, except to rule out the case where the correlation is one. If two independent variables are perfectly correlated, then the X matrix is not of full rank and we have a problem. Otherwise, partial correlations are acceptable (and likely). (iii) (2 pts.) An important omitted variable violates Assumption MLR.4 (zero conditional mean), so then the t statistics don't have a t distribution under  $H_0$ . For example, suppose we are trying to predict consumption of cigarettes. On the right hand side, we include income but we do not include education. Since income and education are almost surely positively correlated, then the errors would not have zero conditional mean. This would lead to biased estimates of  $\beta$ .



### 4.3

(i) (4 pts.) Holding *profmarg* fixed,  $\widehat{\Delta rdintens} = .321\Delta\log(sales) = (.321/100)[100\Delta\log(sales)] \approx .00321(\%\Delta sales)$ . Therefore, if  $\%\Delta sales = 10$ ,  $\widehat{\Delta rdintens} \approx .032$ , or only about 3/100 of a percentage point. For such a large percentage increase in sales, this seems like a very small effect.

(ii) (4 pts.)  $H_0 : \beta_1 = 0$  versus  $H_1 : \beta_1 > 0$ , where  $\beta_1$  is the population slope on  $\log(sales)$ . The t statistic is  $.321/.216 \approx 1.486$ . The 5% critical value for a one-tailed test, with  $df = 32 - 3 = 29$ , is obtained from Table G.2 as 1.699; so we cannot reject  $H_0$  at the 5% level. But the 10% critical value is 1.311; since the t statistic is above this value, we reject  $H_0$  in favor of  $H_1$  at the 10% level.

(iii) (2 pts.) With an increase of profit margin by 1 percentage point, expenditures on R&D rise by 0.05 percentage points. Economically that is quite significant, as given a 10 % increase in profit margin then they will increase expenditures on R& D by 0.5 percentage point.

(iv) 2 pts.) Not really. Its t statistic is only  $0.05/0.046=1.087$ , so we are not able to reject at even the 10% level.

### 4.5

(i) (2 pts.)  $.412 \pm 1.96(.094)$ , or about  $[.228 , .596]$ .

(ii) (2 pts.) No, because the value .4 is well inside the 95% CI.

(iii)(2 pts.) Yes, because 1 is well outside the 95% CI.

### C4.1

(i) (2 pts.) Holding other factors fixed,

$$\Delta voteA = \beta_1\Delta\log(expendA) = (\beta_1/100)[100\Delta\log(expendA)] \approx (\beta_1/100)(\%\Delta expendA) \quad (1)$$

So a .01 increase in expenditure will result in a  $(\beta_1/100) * (100 * .01) = .01\beta_1$  change in the vote for A.

(ii) (2 pts.) The null hypothesis is  $H_0 : \beta_2 = -\beta_1$ , which means a  $z\%$  increase in expenditure by A and a  $z\%$  increase in expenditure by B leaves voteA unchanged. We can equivalently write  $H_0 : \beta_1 + \beta_2 = 0$ .

(iii) (4 pts.)

```
. reg voteA lexpendA lexpendB prtystrA
```

Source	SS	df	MS			
Model	38405.1089	3	12801.703	Number of obs =	173	
Residual	10052.1396	169	59.4801161	F( 3, 169) =	215.23	
Total	48457.2486	172	281.728189	Prob > F =	0.0000	
				R-squared =	0.7926	
				Adj R-squared =	0.7889	
				Root MSE =	7.7123	

  

voteA	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
lexpendA	6.083316	.38215	15.92	0.000	5.328914	6.837719
lexpendB	-6.615417	.3788203	-17.46	0.000	-7.363247	-5.867588
prtystrA	.1519574	.0620181	2.45	0.015	.0295274	.2743873
_cons	45.07893	3.926305	11.48	0.000	37.32801	52.82985

The estimated equation (with standard errors in parentheses below estimates) is

$$\widehat{voteA} = 45.08(3.93) + 6.08(0.38)\log(expendA) - 6.62(0.39)\log(expendB) + .15(0.06)prtystrA$$

$$n = 173, R^2 = .793$$

The coefficient on  $\log(expendA)$  is very significant (t statistic  $\approx 15.92$ ), as is the coefficient on  $\log(expendB)$  (t statistic  $\approx -17.45$ ). The estimates imply that a 10% ceteris paribus increase in spending by candidate A increases the predicted share of the vote going to A by about .61 percentage points. [Recall that, holding other factors fixed,  $\Delta \widehat{voteA} \approx (6.083/100)\% \Delta \log(expendA)$ ] Similarly, a 10% ceteris paribus increase in spending by B reduces A's vote by about .66 percentage points. These effects certainly cannot be ignored. While the coefficients on  $\log(expendA)$  and  $\log(expendB)$  are of similar magnitudes (and opposite in sign, as we expect), we do not have the standard error of  $\hat{\beta}_1 + \hat{\beta}_2$ , which is what we would need to test the hypothesis from part (ii).

(iv) (2 pts.)

```
. test lexpendA=-lexpendB
```

```
( 1) lexpendA + lexpendB = 0
```

```
F( 1, 169) = 1.00
Prob > F = 0.3196
```

So we fail to reject  $\beta_1 + \beta_2 = 0$ .

### C4.3

(i) (2 pts.) The estimated model is

```
. regress lprice sqrft bdrms
```

Source	SS	df	MS			
Model	4.71671468	2	2.35835734	Number of obs =	88	
Residual	3.30088884	85	.038833986	F( 2, 85) =	60.73	
Total	8.01760352	87	.092156362	Prob > F =	0.0000	
				R-squared =	0.5883	
				Adj R-squared =	0.5786	
				Root MSE =	.19706	

  

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
sqrft	.0003794	.0000432	8.78	0.000	.0002935	.0004654
bdrms	.0288844	.0296433	0.97	0.333	-.0300543	.0878232
_cons	4.766027	.0970445	49.11	0.000	4.573077	4.958978

$$\log(\widehat{price}) = 4.766(0.10) + .000379(.000043)sqrft + .0289(.0296)bdrms$$

$$n = 88, R^2 = .588$$

Therefore,  $\hat{\theta}_1 = 150(.000379) + .0289 = .858$ , which means that an additional 150 square foot bedroom increases the predicted price by about 8.6 %.

(ii) (2 pts.)  $\beta_2 = \theta_1 - 150\beta_1$ , and so  $\log(price) = \beta_0 + \beta_1sqrft + (\theta_1 - 150\beta_1)bdrms + u = \beta_0 + \beta_1(sqrft - 150bdrms) + \theta_1bdrms + u$ .

(iii) (2 pts.) From part (ii) we run the regression

```
. gen sqrft150=sqrft-150*bdrms
```

```
. regress lprice sqrft150 bdrms
```

Source	SS	df	MS			
Model	4.71671468	2	2.35835734	Number of obs =	88	
Residual	3.30088884	85	.038833986	F( 2, 85) =	60.73	
Total	8.01760352	87	.092156362	Prob > F =	0.0000	
				R-squared =	0.5883	
				Adj R-squared =	0.5786	
				Root MSE =	.19706	

lprice	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
sqrtft150	.0003794	.0000432	8.78	0.000	.0002935	.0004654
bdrms	.0858013	.0267675	3.21	0.002	.0325804	.1390223
_cons	4.766027	.0970445	49.11	0.000	4.573077	4.958978

Really,  $\hat{\theta}_1 = .0858$ ; note we also get  $se(\hat{\theta}_1) = .0268$ . The 95% confidence interval is .0326 to .1390 (or about 3.3% to 13.9%).