

Fitting interval-censored Cox model with time-varying covariates in Stata

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Outline

What are interval-censored event-time data?

Brief introduction to the algorithm

`stintcox`'s new features

- Using the `tvC()` option to create TVCs
- Testing the PH assumption using `tvC()`
- Fitting `stintcox` with multiple-record data
- Producing new postestimation graphs

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What are interval-censored event-time data?

- The event of interest is not always observed exactly but is known only to occur within some time interval. For example, cancer recurrence, time of COVID infection, etc.
- Interval-censored event-time data arise in many areas, including medical, epidemiological, economic, financial, and sociological studies.
- There are four types of censoring: left-censoring, right-censoring, interval-censoring, and no censoring.
- Data are usually stored in two formats.
- Ignoring interval-censoring may lead to biased estimates.

Types of censoring

For each subject i , event time T_i is not always exactly observed.
 $(L_i, R_i]$ denotes the interval in which T_i is observed.

No censoring

$$L_i = R_i = T_i$$

Right-censoring

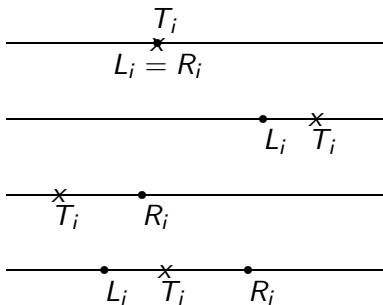
$$(L_i, R_i = +\infty)$$

Left-censoring

$$(L_i = 0, R_i]$$

Interval-censoring

$$(L_i, R_i]$$



Data formats

Single-record-per-subject (single-record) format:

- contains one record for a subject
- contains lower and upper endpoints of the event-time interval
- censoring type is determined by the event-time interval
- covariates are time-independent

	id	ltime	rtime	x1	x2	x3
1.	101	0	6	17	22	0
2.	102	4	9	12	22	1
3.	103	13	.	13	22	0

Data formats

Multiple-record-per-subject (multiple-record) format:

- typically contains multiple records for a subject
- contains an examination time and an event status for each record
- censoring type and the event-time interval can be determined by the examination time and event status
- easily records time-varying covariates

	id	time	status	x1	x2	x3
1.	101	6	1	17	22	0
2.	102	4	0	12	22	1
3.	102	6	0	12	22	0
4.	102	9	1	12	22	1
5.	103	13	0	13	22	0

Methods for analyzing interval-censored data

- Simple imputation methods
- Nonparametric maximum-likelihood estimation
- Parametric regression models – `stintreg`
- **Semiparametric Cox proportional hazards model** –
`stintcox`

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What is Cox proportional hazards model?

- The Cox proportional hazards model was first introduced by Cox in 1972 and was used routinely to analyze uncensored and right-censored event-time data.

$$h(t; \mathbf{x}) = h_0(t) \exp(\mathbf{x}'\boldsymbol{\beta})$$

- It does not require parameterization of the baseline hazard function.
- Also, under the proportional-hazard assumption, the hazard ratios are constant over time.

$$\frac{h(t; \mathbf{x}_i)}{h(t; \mathbf{x}_j)} = \frac{h_0(t) \exp(\mathbf{x}_i' \boldsymbol{\beta})}{h_0(t) \exp(\mathbf{x}_j' \boldsymbol{\beta})} = \exp(\mathbf{x}_i - \mathbf{x}_j)' \boldsymbol{\beta}$$

Cox model's challenge for interval-censored data

- Cox model is challenging for interval-censored event-time data because none of the event times are observed exactly. In particular, the traditional partial-likelihood approach is not applicable.
- Several authors have proposed spline methods to fit the Cox model to interval-censored data and those methods have their limitations.
- The direct maximum-likelihood optimization using the Newton-Raphson algorithm is highly unstable.
- Zeng et al. (2016) developed a genuine EM algorithm for efficient nonparametric maximum-likelihood estimation (NPMLE) method to fit the Cox model for interval-censored data.

A genuine model for `stintcox`

- Suppose that the observed data consist of $(t_{li}, t_{ui}, \mathbf{x}_i)$ for $i = 1, \dots, n$, where t_{li} and t_{ui} define the observed time interval and \mathbf{x}_i records covariate values for a subject i .
- Under the NPMLE approach, the baseline cumulative hazard function H_0 is regarded as a step function with nonnegative jumps h_1, \dots, h_m at t_1, \dots, t_m , respectively, where $t_1 < \dots < t_m$ are the distinct time points for all $t_{li} > 0$ and $t_{ui} < \infty$ for $i = 1, \dots, n$.
- The observed-data likelihood function is

$$\prod_{i=1}^n \exp \left\{ - \sum_{t_k \leq t_{li}} h_k \exp(\mathbf{x}_i \boldsymbol{\beta}) \right\} \left[1 - \exp \left\{ - \sum_{t_{li} < t_k \leq t_{ui}} h_k \exp(\mathbf{x}_i \boldsymbol{\beta}) \right\} \right]^{I(t_{ui} < \infty)} \quad (1)$$

A genuine model for `stintcox` (cont.)

- Let W_{ik} ($i = 1, \dots, n; k = 1, \dots, m$) be independent latent Poisson random variables with means $h_k \exp(\mathbf{x}_i \boldsymbol{\beta})$. Define $A_i = \sum_{t_k \leq t_{li}} W_{ik}$ and $B_i = I(t_{ui} < \infty) \sum_{t_{li} < t_k \leq t_{ui}} W_{ik}$. The likelihood for the observed data ($t_{li}, t_{ui}, \mathbf{x}_i, A_i = 0, B_i > 0$) is

$$\prod_{i=1}^n \prod_{t_k \leq t_{li}} \Pr(W_{ik} = 0) \left\{ 1 - \Pr\left(\sum_{t_{li} < t_k \leq t_{ui}} W_{ik} = 0 \right) \right\}^{I(t_{ui} < \infty)} \quad (2)$$

- (1) and (2) are exactly equal. The maximization of a weighted sum of Poisson log-likelihood functions is strictly concave and has a closed-form solution for h_k 's.

A genuine model for `stintcox` (cont.)

- We maximize (2) through an EM algorithm treating W_{ik} as missing data.
 - In the E-step, we evaluate the posterior means of W_{ik} .
 - In the M-step, we update β and h_k for $k = 1, \dots, m$.
- This method allows a completely arbitrary baseline hazard function, and the results are consistent, asymptotically normal, and asymptotically efficient.

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`stintcox`'s highlights

Stata 17 introduced the `stintcox` command for fitting a semiparametric Cox model to single-record interval-censored data.

- Provides four methods for standard-error computation.
- Provides standard-error computation on replay.
- Provides options to control the tradeoff between the execution speed and accuracy of the results.
- Supports two ways to choose the time intervals to be estimated for baseline hazard contributions.
- Supports stratification.
- Supports various postestimation features after fitting `stintcox`

`stintcox`'s new features

Stata 18 extended the functionality of `stintcox` command:

- Fits multiple-record formats
- Supports time-varying covariates (TVCs):
 - created automatically as deterministic functions of time using the `tvf()` option
 - use the `tvf()` option to test the proportional-hazards assumption
 - Supplied directly in a multiple-record data format
- Supports robust and cluster standard-error computation
- Produces goodness-of-fit plots
- Provides predictions with TVCs
- Plots functions with TVCs

Basic syntax

Single-record-per-subject data format

```
. stintcox [<indepvars>], interval(t_l t_u) ...
```

Multiple-record-per-subject data format

```
. stintcox [<indepvars>], id() time() status() ...
```

- `st` setting the data is not necessary and will be ignored.
- *indepvars* is optional. You can fit a Cox model without any covariates.

Motivating example background

Modified Bangkok IDU Preparatory Study

It is a cohort study of injecting drug users in Thailand.

- 1124 subjects were initially negative for HIV-1 virus.
- They were followed and tested for HIV approximately every four months.
- The event of interest was time to HIV-1 seropositivity.
- We want to identify the factors that influence time to HIV infection.
- Data are stored in both formats:
 - single-record dataset contains all baseline covariates;
 - multiple-record dataset contains both baseline covariates as well as time-varying covariates.

Single-record-per-subject data

```
. list id ltime rtime age_mean male needle inject jail ///  
> if id >= 271 & id <= 274, noobs
```

id	ltime	rtime	age_mean	male	needle	inject	jail
271	22.00	.	-6.46	Yes	Yes	No	No
272	3.80	9.41	8.54	No	No	No	Yes
273	20.66	.	-11.46	Yes	Yes	No	No
274	0.00	3.87	-4.46	Yes	Yes	Yes	Yes

Multiple-record-per-subject data

```
. list id time is_seropos age_mean male needle inject jail_vary ///
> if id >= 271 & id <=274, sepby(id) noobs abbreviate(10) compress
```

id	time	is_seropos	age_mean	male	needle	inject	jail_vary
271	4.89	No	-6.46	Yes	Yes	No	No
271	9.31	No	-6.46	Yes	Yes	No	No
271	13.38	No	-6.46	Yes	Yes	No	Yes
271	17.97	No	-6.46	Yes	Yes	No	Yes
271	22.00	No	-6.46	Yes	Yes	No	No
272	3.80	No	8.54	No	No	No	Yes
272	9.41	Yes	8.54	No	No	No	No
273	3.93	No	-11.46	Yes	Yes	No	No
273	8.00	No	-11.46	Yes	Yes	No	No
273	12.07	No	-11.46	Yes	Yes	No	Yes
273	15.97	No	-11.46	Yes	Yes	No	Yes
273	20.66	No	-11.46	Yes	Yes	No	Yes
274	3.87	Yes	-4.46	Yes	Yes	Yes	Yes

Fitting `stintcox` with single-record data

First, we fit a Cox model with time-independent covariates using the single-record data.

```
. stintcox age_mean i.male i.needle i.inject i.jail, interval(ltime rtime)
note: using adaptive step size to compute derivatives.
```

(iteration output omitted)

Interval-censored Cox regression	Number of obs	=	1,124
Baseline hazard: Reduced intervals	Uncensored	=	0
	Left-censored	=	41
Event-time interval:	Right-censored	=	991
Lower endpoint: ltime	Interval-cens.	=	92
Upper endpoint: rtime			
Log likelihood = -597.56443	Wald chi2(5)	=	17.10
	Prob > chi2	=	0.0043

--more--

Fitting `stintcox` with single-record data (cont.)

	Haz. ratio	OPG std. err.	z	P> z	[95% conf. interval]	
age_mean	.9684341	.0126552	-2.45	0.014	.9439452	.9935582
male						
Yes	.6846949	.1855907	-1.40	0.162	.4025073	1.164717
needle						
Yes	1.275912	.2279038	1.36	0.173	.8990401	1.810768
inject						
Yes	1.250154	.2414221	1.16	0.248	.8562184	1.825334
jail						
Yes	1.567244	.3473972	2.03	0.043	1.014982	2.419998

Note: Standard error estimates may be more variable for small datasets and datasets with low proportions of interval-censored observations.

Using the `tvC()` option

- `tvC()` specifies the variables to be included in the model as an interaction with a function of time to form time-varying covariates.
- It is a convenience tool to speed up calculations and avoid splitting the data over many analysis times.
- Option `texp()` is used in conjunction with `tvC()` to specify the function of time that multiplies covariates specified in the `tvC()` option, i.e., `texp(log(_t))`.
- Option `lrphtest` is used in conjunction with `tvC()` to perform the likelihood-ratio test between the full model and the model without specifying option `tvC()`.
- `tvC()` is also useful for testing the proportional-hazards (PH) assumption.

Testing the PH assumption using `tvvc()`

- One way of testing the PH assumption for a covariate (say, x_1) is to test whether the coefficient associated with that covariate is time invariant.
- This can be accomplished by including an interaction between this covariate and a function of time ($g(t)$) in the model and testing whether the corresponding coefficient equals zero ($\gamma_1 = 0$).

$$\begin{aligned}h(t) &= h_0(t) \exp\{\beta_1 x_1 + \gamma_1 g(t) x_1\} \\ &= h_0(t) \exp[\{\beta_1 + \gamma_1 g(t)\} x_1]\end{aligned}$$

Example: testing the PH assumption

We now include all covariates in option `tvc()` to additionally include their interactions with the analysis time in the model. Thus we can test the PH assumption individually and globally:

```
. stintcox age_mean i.male i.needle i.inject i.jail, interval(ltime rtime) ///
> tvc(age_mean i.male i.needle i.inject i.jail) nohr
note: using adaptive step size to compute derivatives.
```

(iteration output omitted)

Interval-censored Cox regression	Number of obs	=	1,124
Baseline hazard: Reduced intervals	Uncensored	=	0
	Left-censored	=	41
Event-time interval:	Right-censored	=	991
Lower endpoint: ltime	Interval-cens.	=	92
Upper endpoint: rtime			
Log likelihood = -590.43386	Wald chi2(10)	=	31.99
	Prob > chi2	=	0.0004

--more--

Example: testing the PH assumption (cont.)

	Coefficient	OPG std. err.	z	P> z	[95% conf. interval]	
main						
age_mean	-.0310177	.0233817	-1.33	0.185	-.076845	.0148097
male						
Yes	-1.271583	.4604788	-2.76	0.006	-2.174105	-.3690615
needle						
Yes	-.1819587	.3297493	-0.55	0.581	-.8282554	.464338
inject						
Yes	.6852961	.3431924	2.00	0.046	.0126513	1.357941
jail						
Yes	-.529615	.4021087	-1.32	0.188	-1.317734	.2585036
--more--						

Example: testing the PH assumption (cont.)

<code>tv</code>						
age_mean	-0.000129	.0017099	-0.08	0.940	-.0034804	.0032224
male						
Yes	.0884102	.042994	2.06	0.040	.0041434	.1726769
needle						
Yes	.0358545	.0238562	1.50	0.133	-.0109027	.0826118
inject						
Yes	-.0361192	.0228754	-1.58	0.114	-.0809541	.0087157
jail						
Yes	.0916036	.0348915	2.63	0.009	.0232176	.1599896

Notes: Standard error estimates may be more variable for small datasets and datasets with low proportions of interval-censored observations.

Variables in `tv` equation interacted with `_t`.

Wald test that `[tv] = 0`: $\chi^2(5) = 13.3282$

Prob > $\chi^2 = 0.0205$

Fitting `stintcox` with multiple-record data

Fit a Cox model using multiple-record data, including the time-varying covariate `jail_vary`

```
. stintcox age_mean i.male i.needle i.inject i.jail_vary, id(id) time(time) ///
> status(is_seropos)
```

note: time-varying covariates detected in the data; using method `nearleft` to impute their values between examination times.

note: using adaptive step size to compute derivatives.

(iteration output omitted)

Interval-censored Cox regression
Baseline hazard: Reduced intervals

Number of obs	=	6,453
Number of subjects	=	1,124
Uncensored	=	0
Left-censored	=	41
Right-censored	=	991
Interval-cens.	=	92

ID variable: `id`
Examination time: `time`
Status indicator: `is_seropos`

Wald chi2(5)	=	17.03
Prob > chi2	=	0.0044

Log likelihood = -598.34887

--more--

Fitting `stintcox` with multiple-record data (cont.)

time	OPG					
	Haz. ratio	std. err.	z	P> z	[95% conf. interval]	
age_mean	.9714605	.012757	-2.20	0.027	.9467762	.9967884
male						
Yes	.6678044	.1816576	-1.48	0.138	.3918353	1.138138
needle						
Yes	1.271409	.2275426	1.34	0.180	.8952546	1.805609
inject						
Yes	1.370672	.2575405	1.68	0.093	.9484142	1.980928
jail_vary						
Yes	1.440966	.2916178	1.81	0.071	.9691488	2.142481

Time varying: `jail_vary`

Note: Standard error estimates may be more variable for small datasets and datasets with low proportions of interval-censored observations.

Using `tvcovimpute()` option

- Use `tvcovimpute()` to specify how to impute unobserved covariate values between two examination times for time-varying covariates.
- The imputation methods include `nearleft` (default), `nearright`, `nearest`, or `first`.

```
. stintcox age_mean i.male i.needle i.inject i.jail_vary, id(id) time(time) ///
> status(is_seropos) tvcovimpute(nearright)
note: time-varying covariates detected in the data; using method nearright to
      impute their values between examination times.
note: using adaptive step size to compute derivatives.
```

(iteration output omitted)

Interval-censored Cox regression	Number of obs	=	6,453
Baseline hazard: Reduced intervals	Number of subjects	=	1,124
	Uncensored	=	0
ID variable: id	Left-censored	=	41
Examination time: time	Right-censored	=	991
Status indicator: is_seropos	Interval-cens.	=	92
	Wald chi2(5)	=	18.41
Log likelihood = -597.00103	Prob > chi2	=	0.0025

Using `tvcovimpute()` option (cont.)

time	Haz. ratio	OPG		z	P> z	[95% conf. interval]	
		std. err.					
age_mean	.9726438	.0126471		-2.13	0.033	.9481692	.9977502
male							
Yes	.6561992	.1780502		-1.55	0.121	.3855444	1.116856
needle							
Yes	1.267405	.228118		1.32	0.188	.890654	1.803523
inject							
Yes	1.367475	.252569		1.69	0.090	.9521488	1.963966
jail_vary							
Yes	1.640746	.3346384		2.43	0.015	1.100106	2.44708

Time varying: `jail_vary`

Note: Standard error estimates may be more variable for small datasets and datasets with low proportions of interval-censored observations.

Postestimation features after `stintcox`

`stintcox` provides several postestimation features after estimation:

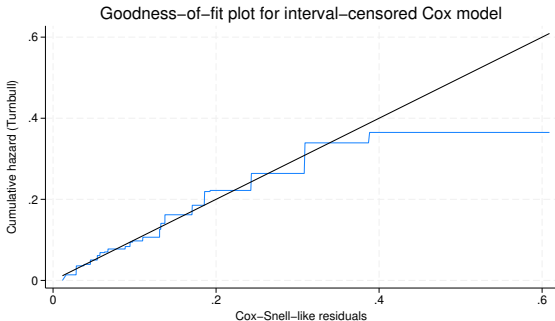
- Predictions of hazard ratios, linear predictions, and standard errors with support for TVCs
- Predictions of baseline survivor, baseline cumulative hazard, and baseline hazard contribution functions
- Prediction of martingale-like residuals and Cox–Snell-like residuals
- goodness-of-fit plot
- Plots for survivor, hazard, and cumulative hazard functions

Producing Goodness-of-fit (GOF) plot

- `estat gofplot` is used to assess the goodness of fit of the model visually.
- It plots the Cox–Snell-like residuals versus the estimated cumulative hazard function corresponding to these residuals.
- The estimated cumulative hazards are calculated using the self-consistency algorithm proposed by Turnbull (1976).
- The Cox–Snell-like residuals form the 45° reference line. If the model fits the data well, the plotted estimated cumulative hazards should be close to the reference line.

Goodness-of-fit (GOF) plot

```
. estat gofplot
```

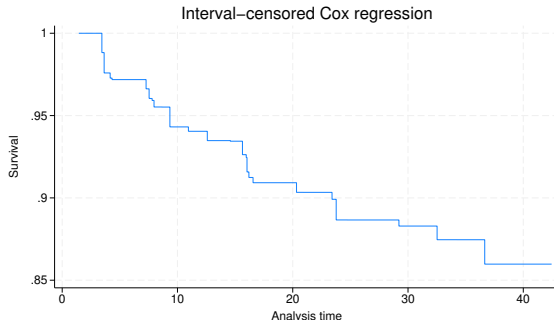


Graph survivor function

- Use `stcurve` to plot the estimated survivor function.
- By default, `stcurve` evaluates the functions at the overall means of covariates.

```
. stcurve, survival
```

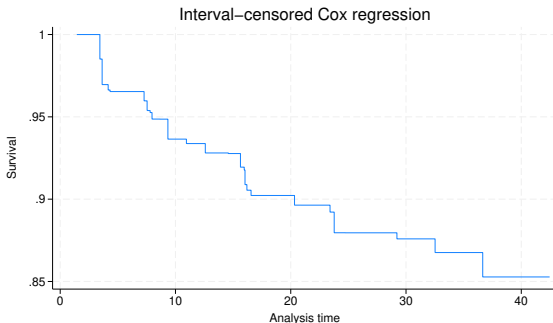
```
note: function evaluated at overall means of covariates.
```



Graph survivor function with TVCs

- Use option `attmeans` to evaluate the function at time-specific means.

```
. stcurve, survival attmeans  
note: function evaluated at time-specific means of covariates.
```



Graph survivor function using frame

We can also use option `atframe()` to specify your own TVC values to be used to evaluate the survivor function.

- Suppose we want to plot the survivor curve for an individual with the same covariate pattern as subject 2.
- We create a new frame called `id2` and use `frame put` to copy the relevant information to the new frame.
- We list the data in frame `id2`.

```
. frame put time age_mean male needle inject jail_vary if id==2, into(id2)
. frame id2: list
```

	time	age_mean	male	needle	inject	jail_vary
1.	4.13	-6.46	Yes	No	Yes	Yes
2.	8.26	-6.46	Yes	No	Yes	No
3.	12.30	-6.46	Yes	No	Yes	No
4.	16.07	-6.46	Yes	No	Yes	No
5.	20.10	-6.46	Yes	No	Yes	No
6.	24.26	-6.46	Yes	No	Yes	No

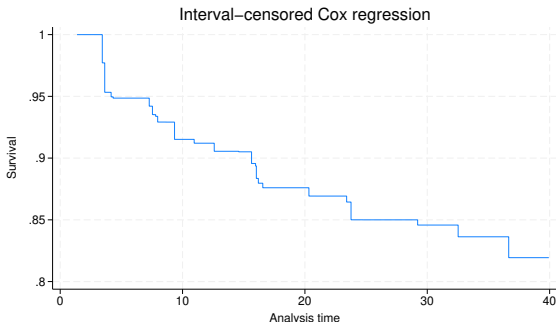
Graph survivor function using frame (cont.)

- Use option `atframe()` to graph the survivor curve for this particular profile,

```
. stcurve, survival atframe(id2)
```

```
note: function evaluated at specified values of selected covariates and  
      overall means of other covariates (if any).
```

```
note: covariate values from frame id2 used to evaluate function.
```



Conclusions for `stintcox`

- Fits a genuine semiparametric Cox proportional hazards model with two formats of interval-censored data.
- Supports different methods for standard error computation; also support VCE computation on replay.
- Supports creating TVCs automatically and testing the PH assumption.
- Provides diagnostic measures, predictions, and much more after fitting the model.
- Provides convenient graphical tools for assessing the goodness of fit of the model, and for plotting the survivor, cumulative hazards, and hazard functions.
- Supports TVCs with predictions and graphs.

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References

- Farrington, C. P. (2000). Residuals for proportional hazards models with interval-censored survival data. *Biometrics* 56, 473–482.
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- Zeng, D., F. Gao, and D. Lin (2017). Maximum likelihood estimation for semiparametric regression models with multivariate interval-censored data. *Biometrika* 104, 505–525.
- Zeng, D., L. Mao, and D. Lin (2016). Maximum likelihood estimation for semiparametric transformation models with interval-censored data. *Biometrika* 103, 253–271.

More resources

<https://www.stata.com/manuals/ststintcox.pdf>

<https://www.stata.com/manuals/ststintcoxpostestimation.pdf>

<https://www.stata.com/manuals/ststintcoxph-assumptionplots.pdf>

<https://www.stata.com/manuals/stestatgofplot.pdf>

<https://www.stata.com/manuals/ststcurve.pdf>

Thank you!