# xtfeis.ado: Linear Fixed Effects Models with Individual Slopes

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> 8th German Stata Users Group Meeting June 25, 2010

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#### Outline

The Problem. Consider growth curves A straightforward and Simple Solution. FE with individual Slopes Implementation in Stata. xtfeis.ado Monte Carlo Simulation. Compare models. Real world example. The male marriage wage premium Conclusions

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Implementation in Stata. xtfeis.ado

Monte Carlo Simulation. Compare models.

Real world example. The male marriage wage premium

Conclusions

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 Fixed Effect framework: the ultimate method for causal analysis with non-experimental data



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- Problem recognized (Allison 1990, Heckman & Hotz 1989, Polachek & Kim 1994, Winship & Morgan 1999, Morgan & Winship 2007)

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- But many applications where conventional FE models fail because strict exogeneity is violated
  - here: time-constant unobserved factors correlate with observed factors
  - major example: unobserved effect changes over time
- Problem recognized (Allison 1990, Heckman & Hotz 1989, Polachek & Kim 1994, Winship & Morgan 1999, Morgan & Winship 2007)
- but seldomly solved in practice

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### The Problem. Consider growth curves

Allison (1990), Winship & Morgan (1999) consider 3 situations:

 Absent treatment, potential outcomes of treatment and control group develop along the same path



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  - solution ???
- ▶ 1., 2., 3.? Theories seldomly strong enough to decide

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### FE with individual Slopes, notation first

Let Y<sub>it</sub> : Outcome Y of individual i at time point t



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### Extending the FE framework

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- estimate

$$(Y_{it} - \overline{Y_i}) = \beta (D_{it} - \overline{D_i}) + \gamma (Z_{it} - \overline{Z_i}) + (\alpha_{1i} - \alpha_{1i}) + (\epsilon_{it} - \overline{\epsilon_{it}})$$

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Extended FE, including individual slopes

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- Extended FE, including individual slopes
  - substract not the mean, but time-varying estimate to eliminate  $\alpha_{1i}$  and  $\alpha_{2i}$

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- substract not the mean, but time-varying estimate to eliminate  $\alpha_{1i}$  and  $\alpha_{2i}$
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 $\widetilde{Y_{it}} = \beta \widetilde{D_{it}} + \widetilde{\epsilon_{it}}$ , where  $\widetilde{Y_{it}}$  is residual from individual time-series regression (OLS) of  $Y_{it}$  on  $Z_{it}$ , and analog for indep. variables

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transform "'by hand"' possible, but very very slow UNIVERSITÄT MANN

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### Extending the FE framework

• General approach to within transform (Wooldridge 2002):


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- General approach to within transform (Wooldridge 2002):
  - Premultiply all variables by matrix  $\Omega_i = I_T Z'_i (Z'_i Z_i)^{-1} Z'_i$



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  - Since Ω<sub>i</sub>Z<sub>i</sub> = 0, α<sub>1i</sub> and α<sub>2i</sub> are eliminated
- Conventional FE is a special case where Z<sub>i</sub> = (1) is (N × 1) vector of constants
- ▶ Random growth model is another special case where  $Z_i = (1, t)$  is  $(N \times 2)$  matrix of constants and time variable

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#### Stata ado xtfeis.ado

based on mata



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- ▶ appropriate when T<sub>i</sub> > J, where J is number of Z variables (including possibly individual constants)



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- ▶ appropriate when T<sub>i</sub> > J, where J is number of Z variables (including possibly individual constants)
- automatically selects estimation sample



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- fully robust s.e. on request
- Syntax: xtfeis varlist, [slope(varlist)]
  [noconstant] [cluster(clustvar)]

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### Simulation setup

▶ set up panel data set with 3000 cases (N = 1000, T = 3)



# Simulation setup

- ▶ set up panel data set with 3000 cases (N = 1000, T = 3)
- ▶ 500 receive treatment between t = 2 and t = 3



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- ▶ set up panel data set with 3000 cases (N = 1000, T = 3)
- ▶ 500 receive treatment between t = 2 and t = 3
- choose true parameters:
  - treatment effect β

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# Simulation setup

- ▶ set up panel data set with 3000 cases (N = 1000, T = 3)
- ▶ 500 receive treatment between t = 2 and t = 3
- choose true parameters:
  - treatment effect β
  - normally distributed individual constants  $\alpha_{1i}$

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• Generate outcome: 
$$Y_{it} = D_{it} + \alpha_{1i} + \alpha_{2i}T + \epsilon_{it}$$

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- Estimate treatment effect  $\hat{\beta}$  using 5 different models

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- repeat 1000 times

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  - treatment effect  $\beta$
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- Estimate treatment effect  $\hat{\beta}$  using 5 different models
- repeat 1000 times
- ▶ get mean of  $\hat{\beta}$ , s.e., % coefs. diff. from true  $\beta$

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### Simulation: 5 different models



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### Simulation: 5 different models

#### Model Equation



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### Simulation: 5 different models

Model	Equation
Pooled OLS	$Y_{it} = \beta D_{it} + \gamma T + \epsilon_{it}$



### Simulation: 5 different models

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Pooled OLS	$Y_{it} = \beta D_{it} + \gamma T + \epsilon_{it}$
ANCOVA	$Y_{it+1} = \delta Y_{it} + \beta D_i + \gamma T + \epsilon_{it}$

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ANCOVA	$Y_{it+1} = \delta Y_{it} + \beta D_i + \gamma T + \epsilon_{it}$
Change score	$Y_{it+1} - Y_{it} = \beta D_i + \gamma T + \epsilon_{it}$

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Change score	$Y_{it+1} - Y_{it} = \beta D_i + \gamma T + \epsilon_{it}$
Fixed Effects	$\ddot{Y}_{it} = \beta \ddot{D}_{it} + \gamma \ddot{T} + \ddot{\epsilon}_{it}$

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Fixed Effects IS	$\widetilde{Y_{it}} = \beta \widetilde{D_{it}} + \widetilde{\epsilon_{it}}$

V. Ludwig Linear FE Models with Ind. Slopes

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### Simulation results

Scenario 1: Absent treatment, potential outcomes follow same path



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# Simulation results

Scenario 1: Absent treatment, potential outcomes follow same path

• 
$$\beta = 0, \ \alpha_{1i} = N(.25D_i, .1), \ \alpha_{2i} = 0, \ \epsilon_{it} = N(0, .1)$$



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Model	$\hat{eta}$	s.e.	<b>%</b> p > .05
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Model	$\hat{eta}$	s.e.	<b>%</b> <i>p</i> > .05
Pooled OLS	.214	.010	100

V. Ludwig Linear FE Models with Ind. Slopes

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Model	$\hat{eta}$	s.e.	<b>%</b> <i>p</i> > .05
Pooled OLS	.214	.010	100
ANCOVA	.125	.007	100



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Model	$\hat{eta}$	s.e.	<b>%</b> <i>p</i> > .05
Pooled OLS	.214	.010	100
ANCOVA	.125	.007	100
Change score	<.001	.006	1.1

V. Ludwig Linear FE Models with Ind. Slopes

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Model	$\hat{eta}$	s.e.	<b>%</b> <i>p</i> > .05
Pooled OLS	.214	.010	100
ANCOVA	.125	.007	100
Change score	<.001	.006	1.1
Fixed Effects	<.001	.007	4.7

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Pooled OLS	.214	.010	100	-
ANCOVA	.125	.007	100	_
Change score	<.001	.006	1.1	_
Fixed Effects	<.001	.007	4.7	-
Fixed Effects IS	<.001	.011	5.1	_
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### Simulation results

Scenario 2: Absent treatment, potential outcomes converge



## Simulation results

Scenario 2: Absent treatment, potential outcomes converge

•  $\beta = 0, \ \alpha_{1i} = N(.25D_i, .1), \ \alpha_{2i} = N(-.05D_i, .1), \ \epsilon_{it} = N(0, .1)$ 



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## Simulation results

Scenario 2: Absent treatment, potential outcomes converge

•  $\beta = 0, \ \alpha_{1i} = N(.25D_i, .1), \ \alpha_{2i} = N(-.05D_i, .1), \ \epsilon_{it} = N(0, .1)$ 

**Model** 
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FE model: fine when α<sub>1i</sub> = 0, but fails when potential outcomes w/o treatment do not follow same path (α<sub>2i</sub> = 0 required)



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- FE-IS is the only model which performs nicely regardles of α<sub>1i</sub> and α<sub>2i</sub>
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#### The story

 Married men earn higher (hourly) wages than never-married, cohabiting, divorced

V. Ludwig



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- Married men earn higher (hourly) wages than never-married, cohabiting, divorced
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  - not clear whether the effect is causal
  - possible bias because of α<sub>2i</sub>

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Two country studies: West Germany and US



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- controls: divorce, remarriage, number of children, yrs.
  education, tenure, year dummies

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### Results: Effect of marriage on male wages

GSOEP			NLSY		
Model	$\hat{eta}$	robust s.e.	Model	$\hat{eta}$	robust s.e.
POLS	.078**	(.014)	POLS	.146**	(.010)
FE	.036**	(.013)	FE	.082**	(.008)
FE-IS	.015	(.010)	FE-IS	.021*	(.008)

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### Conclusions

- FE-IS valid under more general conditions than alternative models
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