

Bivariate probit

$$y_1^* = \beta_1 x + \epsilon_1$$

$$y_2^* = \beta_2 x + \epsilon_2$$

$$\mathbf{y} = (\mathbf{1}\{y_1^* > 0\}, \mathbf{1}\{y_2^* > 0\})'$$

$$\boldsymbol{\epsilon} = (\epsilon_1, \epsilon_2)' \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$$

$$\boldsymbol{\Sigma} = \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}$$

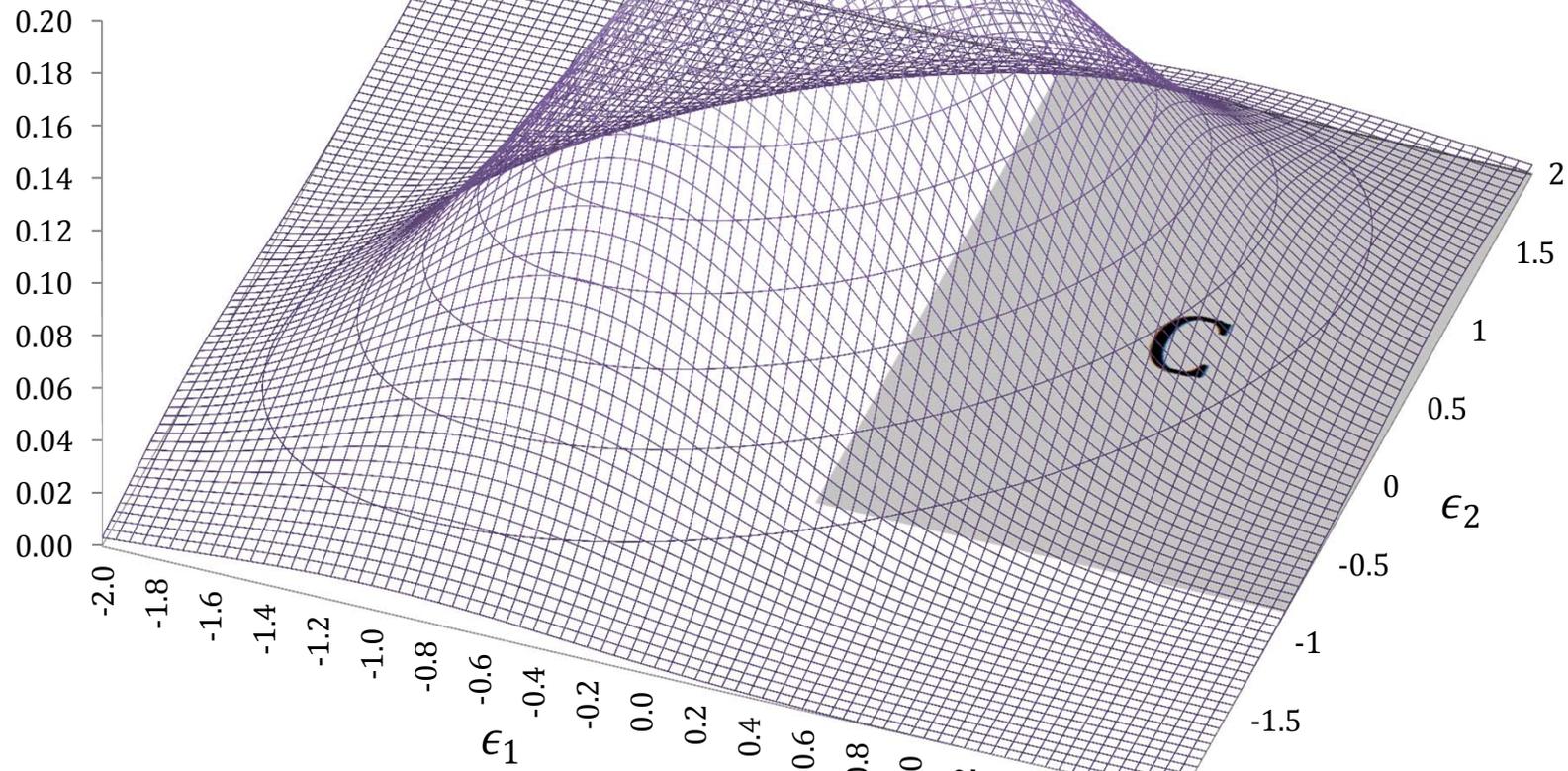
Case $\mathbf{y} = (1,1)'$

$$\epsilon_1 > -\beta_1 x, \epsilon_2 > -\beta_2 x$$

$$C = \{(\epsilon_1, \epsilon_2) : \epsilon_1 > -\beta_1 x, \epsilon_2 > -\beta_2 x\}$$

$$\mathcal{L}(\beta_1, \beta_2, \boldsymbol{\Sigma}; \mathbf{y}_i | x_i) = \int_C \phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$$

$\phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$



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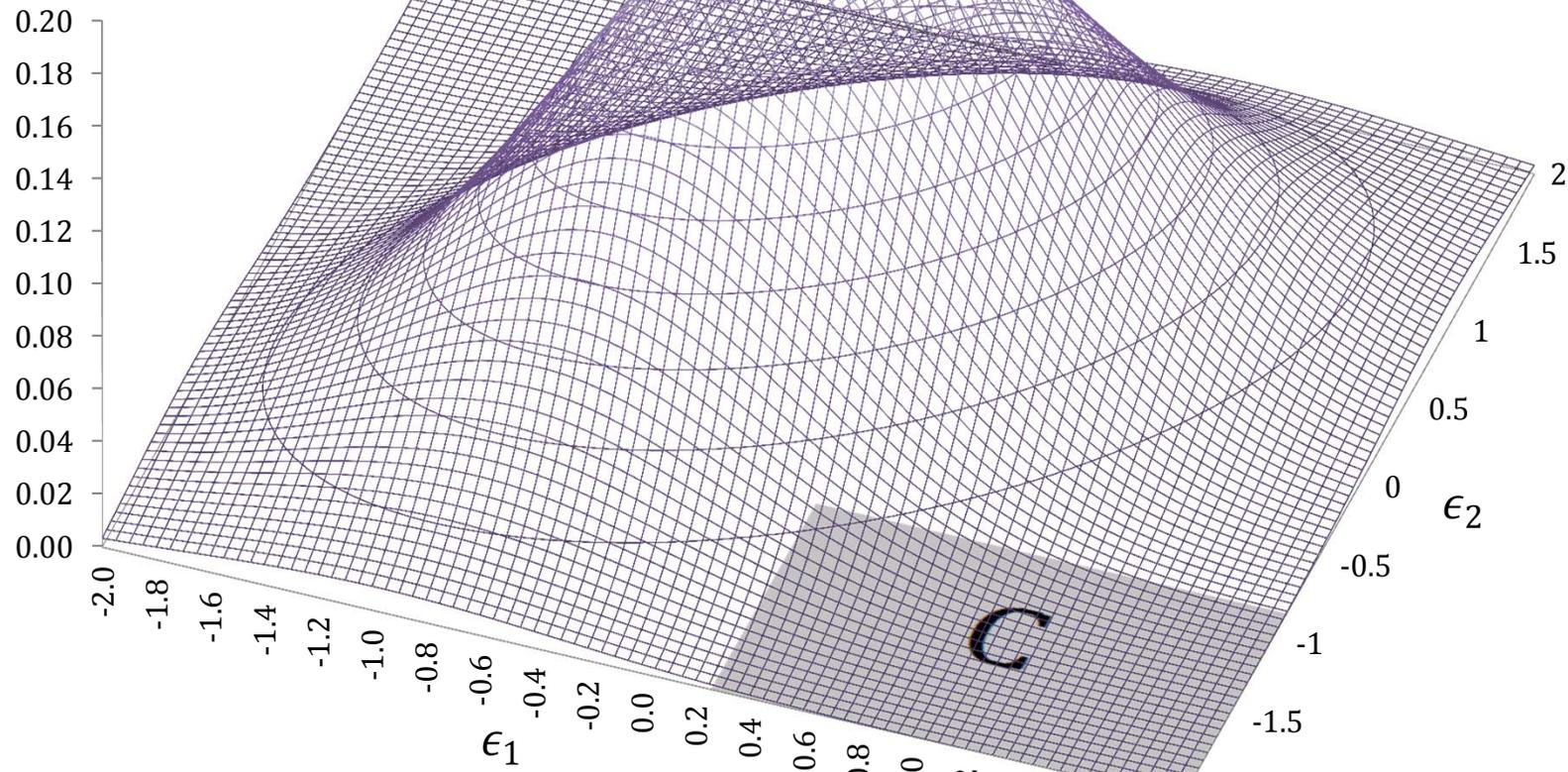
Case $\mathbf{y} = (1, 0)'$

$$\epsilon_1 > -\beta_1 x, \epsilon_2 < -\beta_2 x$$

$$C = \{(\epsilon_1, \epsilon_2) : \epsilon_1 > -\beta_1 x, \epsilon_2 < -\beta_2 x\}$$

$$\mathcal{L}(\beta_1, \beta_2, \boldsymbol{\Sigma}; \mathbf{y}_i | x_i) = \int_C \phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$$

$\phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$



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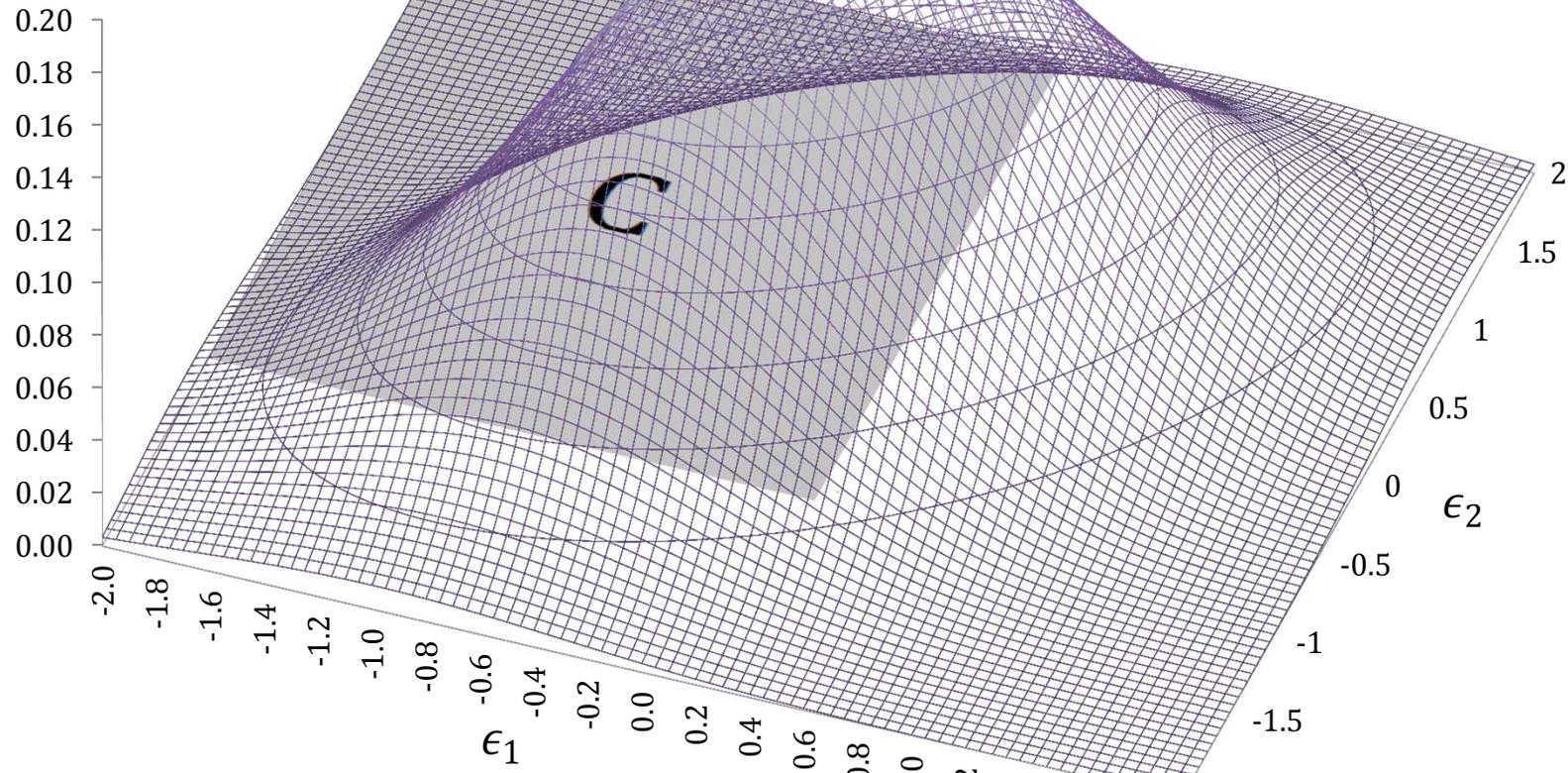
Case $\mathbf{y} = (0, 1)'$

$$\epsilon_1 < -\beta_1 x, \epsilon_2 > -\beta_2 x$$

$$C = \{(\epsilon_1, \epsilon_2) : \epsilon_1 < -\beta_1 x, \epsilon_2 > -\beta_2 x\}$$

$$\mathcal{L}(\beta_1, \beta_2, \boldsymbol{\Sigma}; \mathbf{y}_i | x_i) = \int_C \phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$$

$\phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$



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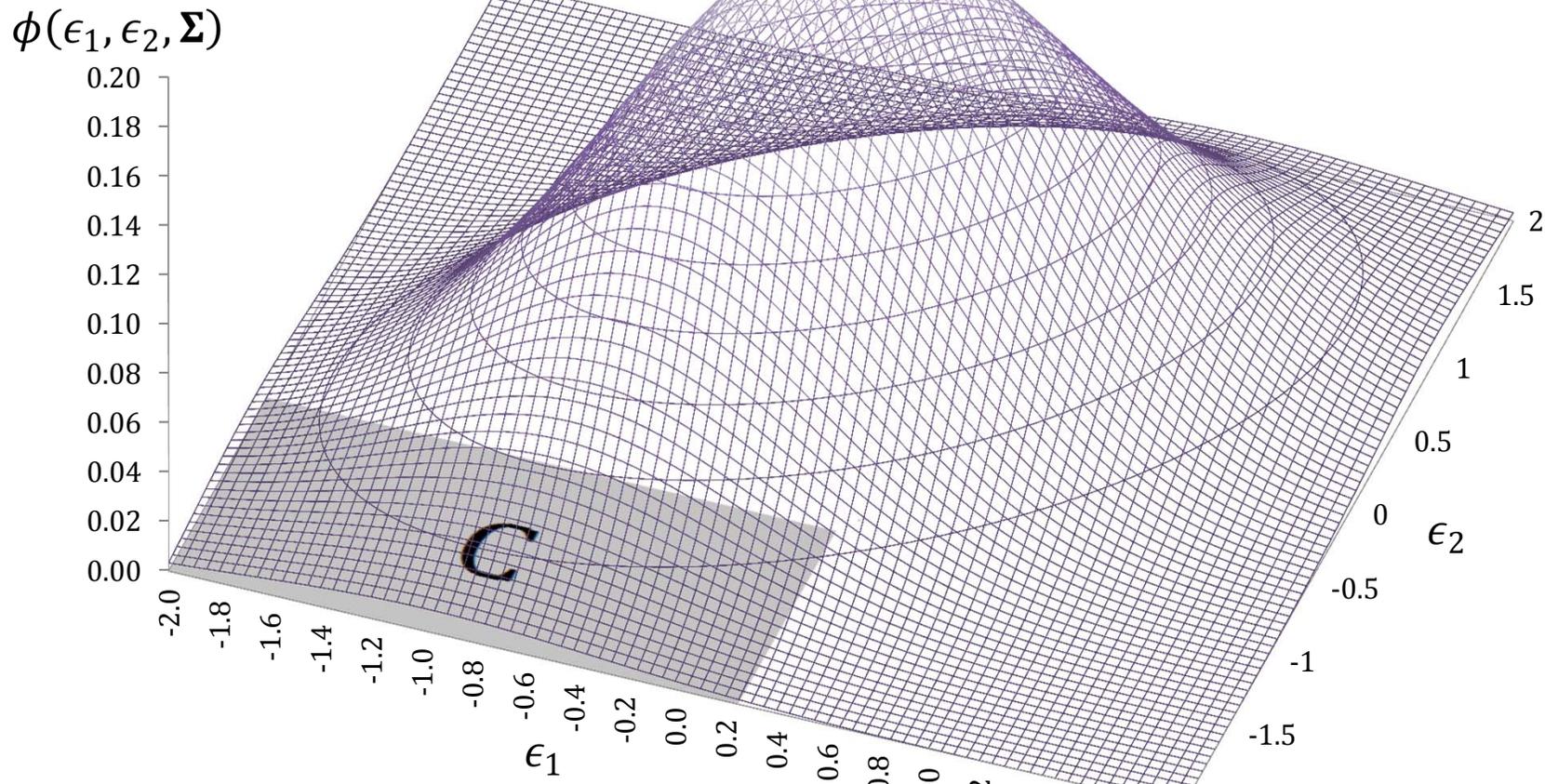
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Case $\mathbf{y} = (0, 0)'$

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$$\mathcal{L}(\beta_1, \beta_2, \boldsymbol{\Sigma}; \mathbf{y}_i | x_i) = \int_C \phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$$



Mixed probit-uncensored

$$y_1^* = \beta_1 x + \epsilon_1$$

$$y_2^* = \beta_2 x + \epsilon_2$$

$$\mathbf{y} = (\mathbf{y}_1^*, \mathbf{1}\{y_2^* > 0\})'$$

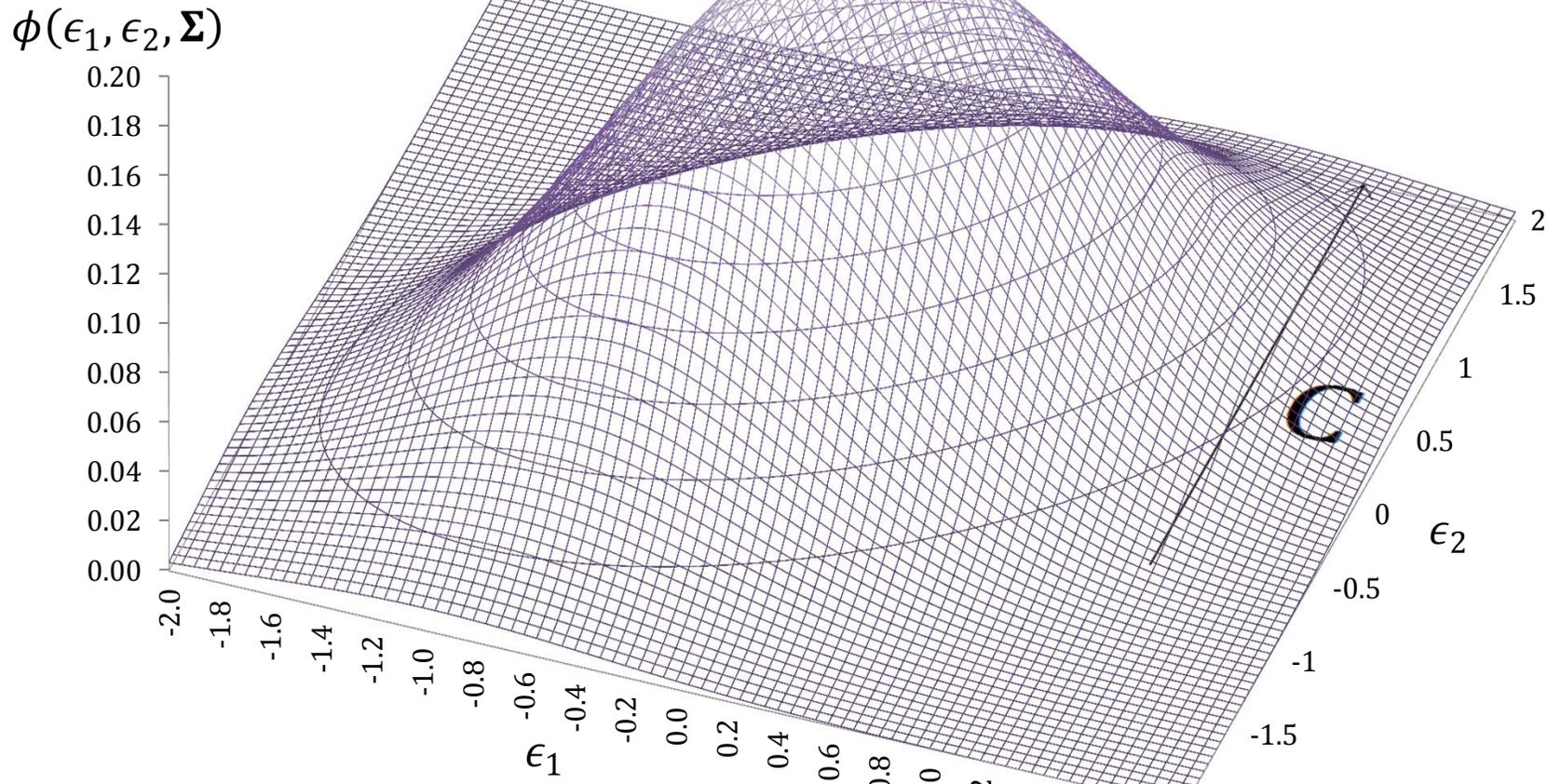
$$\boldsymbol{\epsilon} = (\epsilon_1, \epsilon_2)' \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$$

Case $\mathbf{y} = (y_1, 1)'$

$$\epsilon_1 = y_1 - \beta_1 x, \epsilon_2 > -\beta_2 x$$

$$C = \{(\epsilon_1, \epsilon_2) : \epsilon_1 = y_1 - \beta_1 x, \epsilon_2 > -\beta_2 x\}$$

$$\mathcal{L}(\beta_1, \beta_2, \boldsymbol{\Sigma}; \mathbf{y}_i | x_i) = \int_C \phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$$



Mixed probit-uncensored

$$y_1^* = \beta_1 x + \epsilon_1$$

$$y_2^* = \beta_2 x + \epsilon_2$$

$$\mathbf{y} = (y_1^*, \mathbf{1}\{y_2^* > 0\})'$$

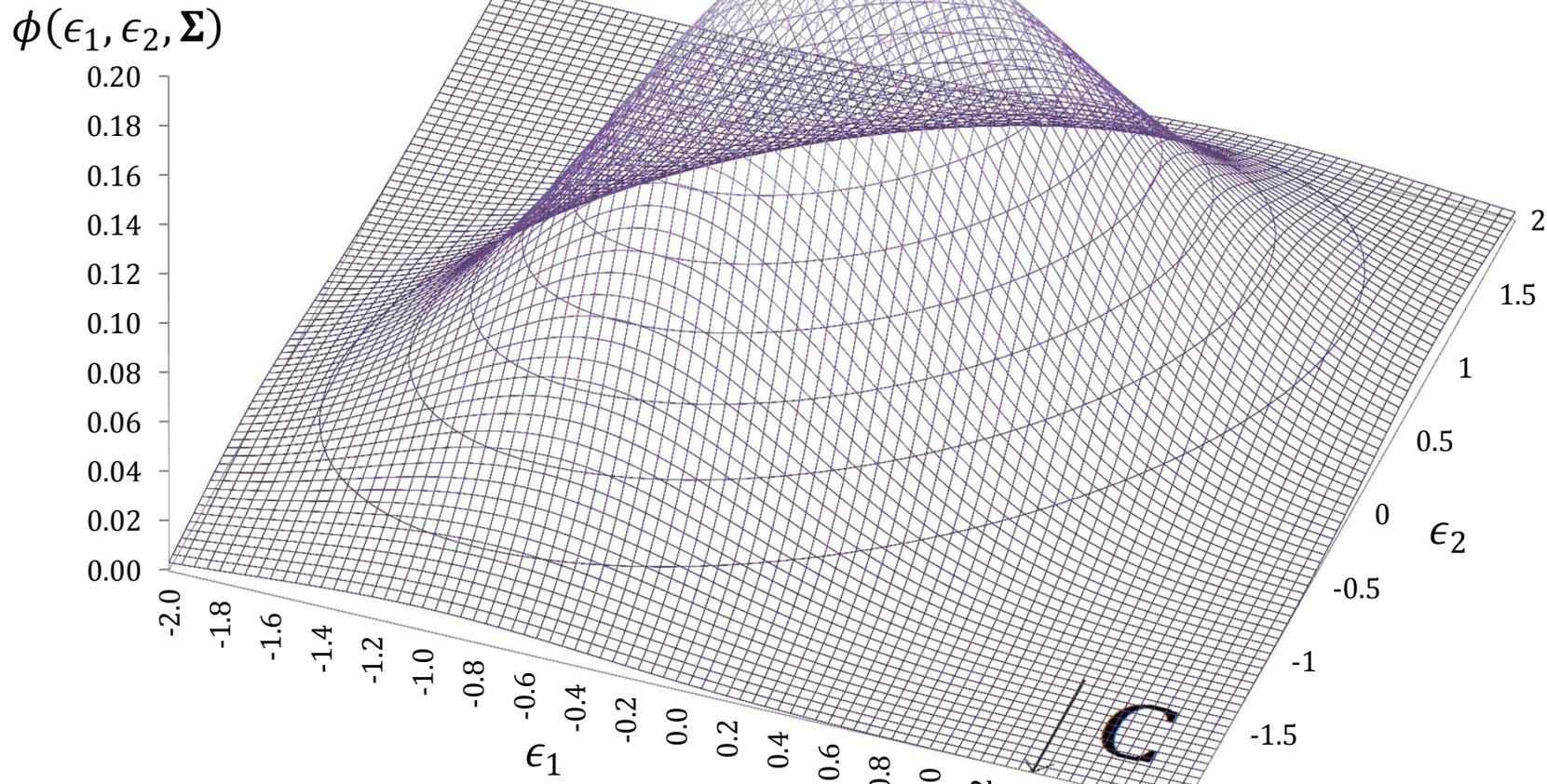
$$\boldsymbol{\epsilon} = (\epsilon_1, \epsilon_2)' \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$$

Case $\mathbf{y} = (y_1, \mathbf{0})'$

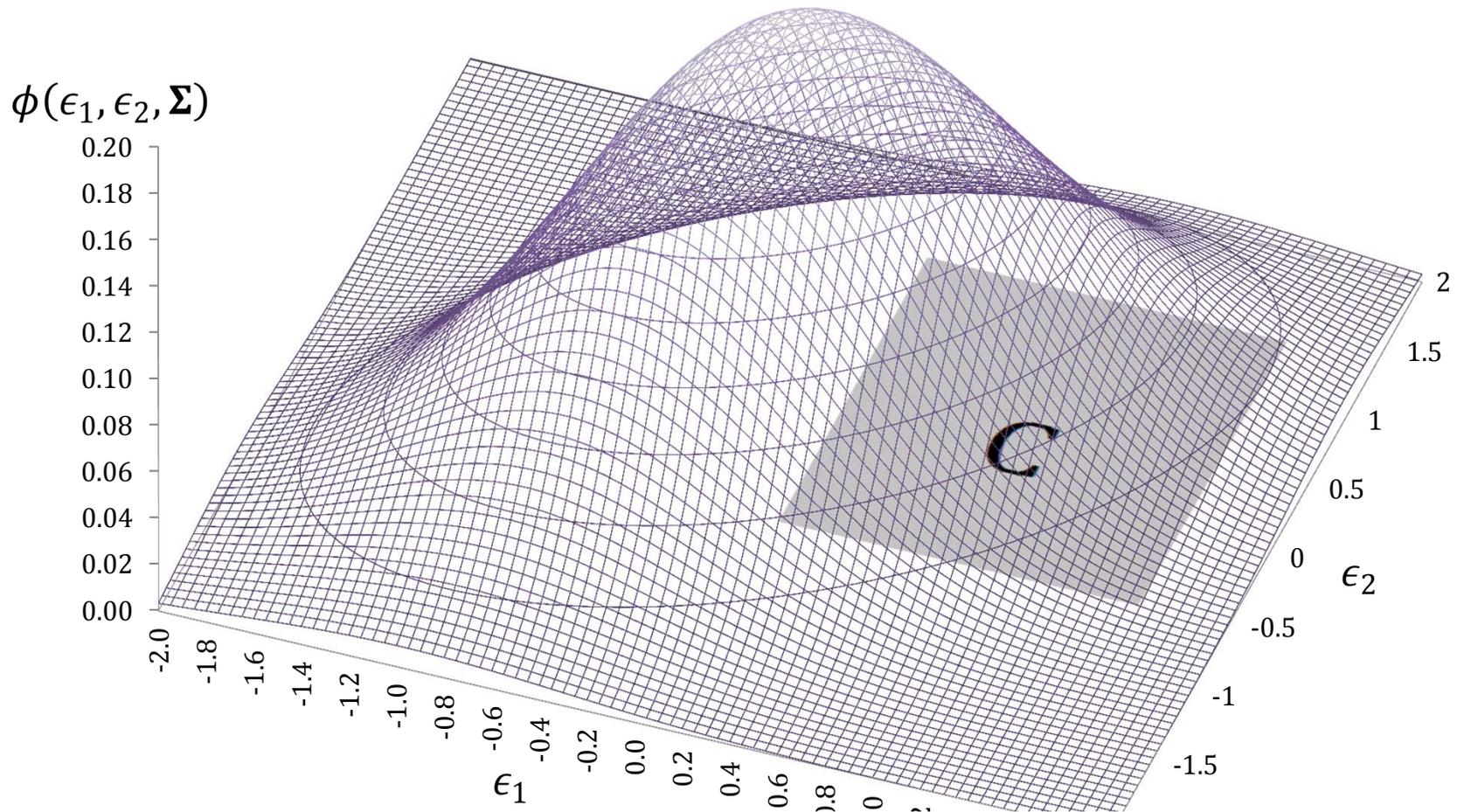
$$\epsilon_1 = y_1 - \beta_1 x, \epsilon_2 < -\beta_2 x$$

$$C = \{(\epsilon_1, \epsilon_2) : \epsilon_1 = y_1 - \beta_1 x, \epsilon_2 < -\beta_2 x\}$$

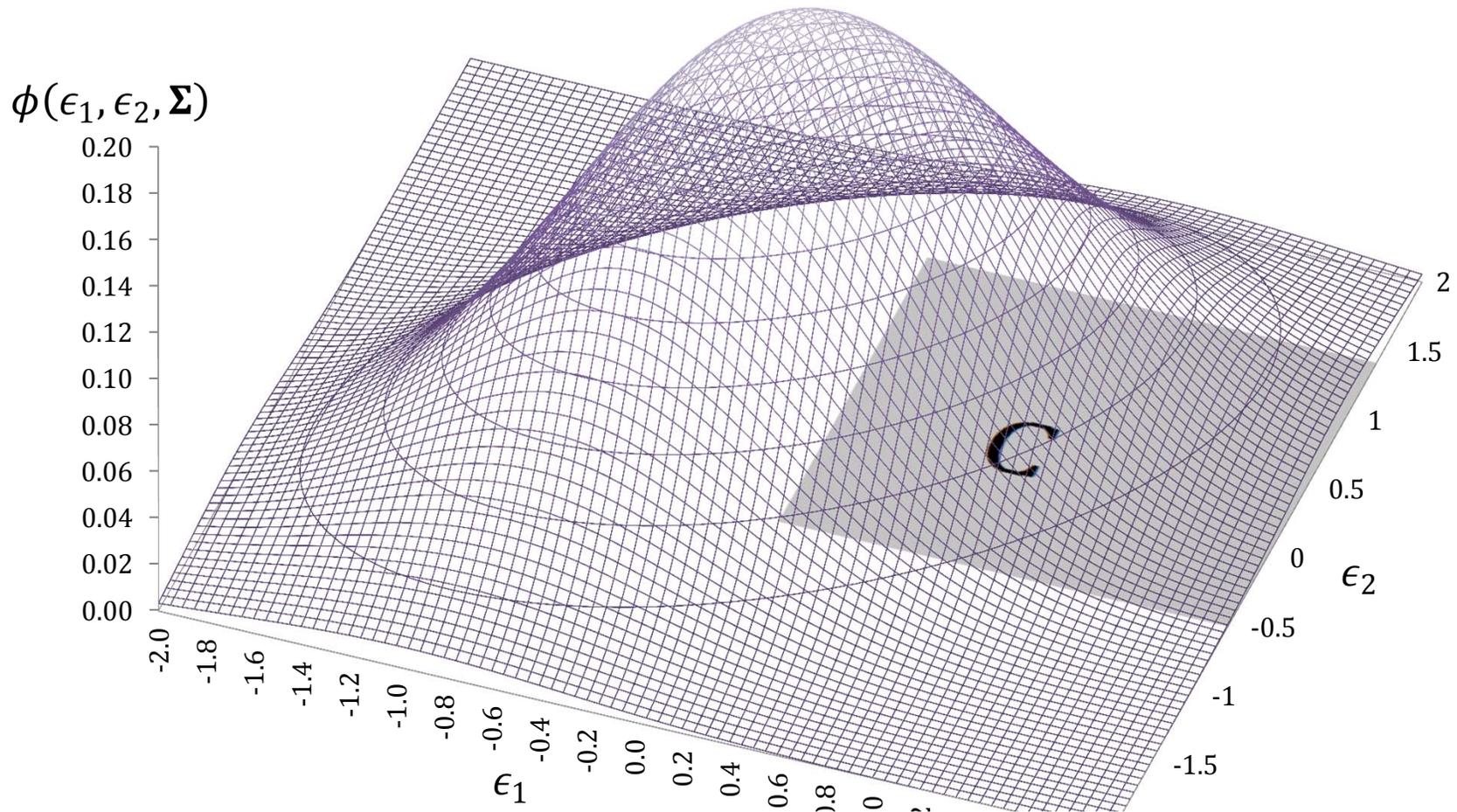
$$\mathcal{L}(\beta_1, \beta_2, \boldsymbol{\Sigma}; \mathbf{y}_i | x_i) = \int_C \phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$$



Bivariate ordered probit



Mixed probit-ordered probit



$-\delta y_1$

Likelihood for SUR models

- Computed separately for each observation
 - So model can vary by observation
 - Different censoring structure, different # of eqs
- For each obs, integral of multivariate normal distribution
- Over some “Cartesian” region
 - Point, line, ray, half-plane, quarter-plane, etc.
- Core of `cmp` does this math

Two-stage probit

$$y_1^* = \beta_1 x + \epsilon_1$$

$$y_2^* = \delta y_1 + \beta_2 x + \epsilon_2$$

$$\mathbf{y} = (\mathbf{1}\{y_1^* > 0\}, \mathbf{1}\{y_2^* > 0\})'$$

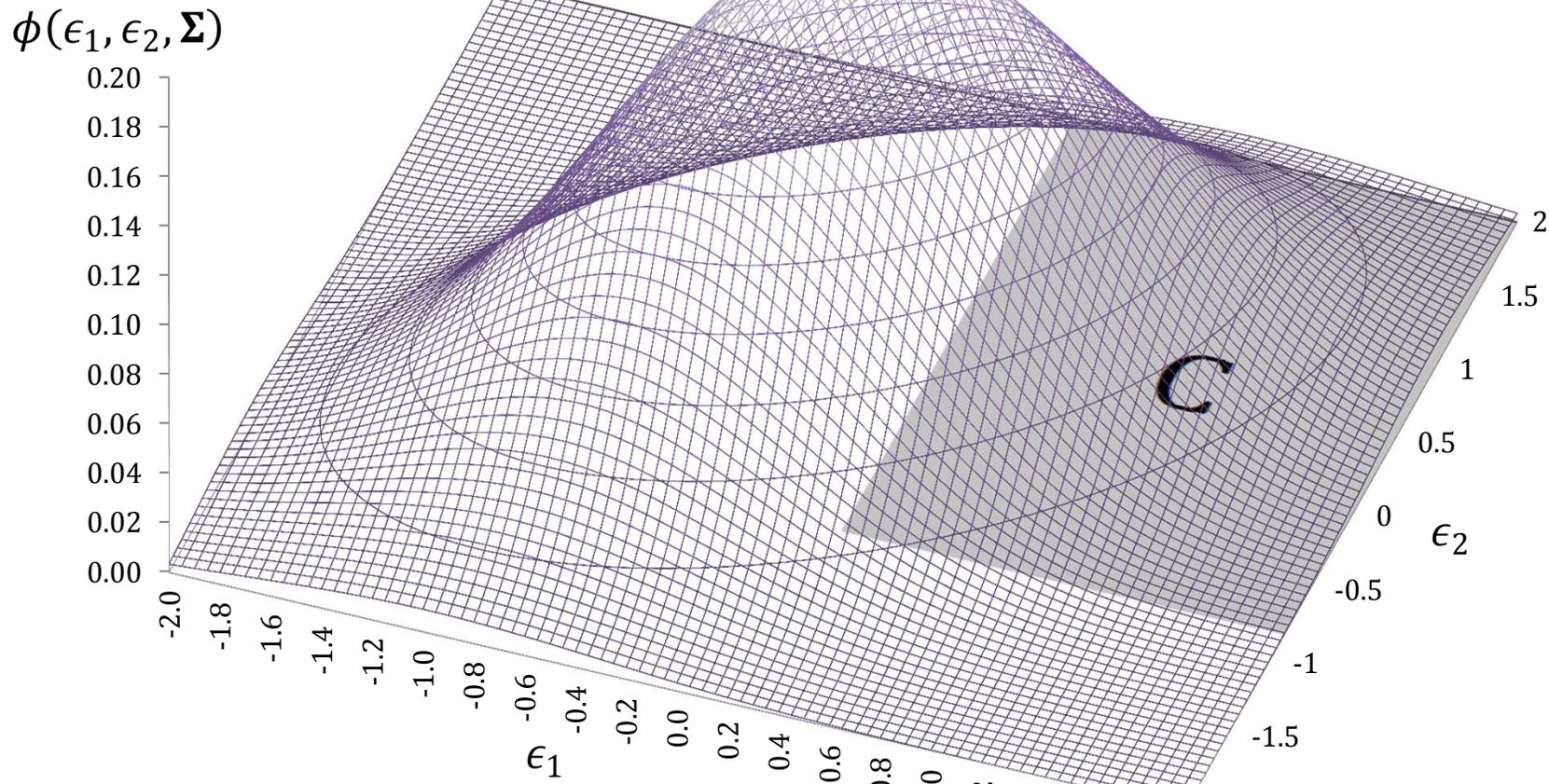
$$\boldsymbol{\epsilon} = (\epsilon_1, \epsilon_2)' \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$$

Case $\mathbf{y} = (1, 1)'$

$$\epsilon_1 > -\beta_1 x, \epsilon_2 > -\delta y_1 - \beta_2 x$$

$$C = \{(\epsilon_1, \epsilon_2) : \epsilon_1 > -\beta_1 x, \epsilon_2 > -\delta y_1 - \beta_2 x\}$$

$$\mathcal{L}(\beta_1, \beta_2, \boldsymbol{\Sigma}; \mathbf{y}_i | x_i) = \int_C \phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$$



- Same math!
 - Still integrate normal distribution over Cartesian region
- `cmp` automatically multi-stage estimator
 - Can do IV, Heckman selection, switching regressions...
- By same token
 - `biprobit` can be a two-stage estimator
 - `ivregress` can do SUR
- Wasn't widely appreciated
 - Greene (1998): “surprisingly” ... “seem not to be widely known”
 - Wooldridge (e-mail 2009): “I came to this realization somewhat late, although I've known it for a couple of years now.”
- Requires system to be “fully observed” and recursive.

Available models

- Uncensored
 - Tobit
 - Probit
 - Ordered probit
 - Interval
 - Multinomial probit
 - Rank-ordered probit (new)
-
- Truncation now allowed in all but last two

- But wait! (three years)
- ...there's more.
- No longer need:
 - fully observed
 - recursive

Two-stage probit with latent determinant

$$y_1^* = \beta_1 x + \epsilon_1$$

$$y_2^* = \gamma y_1^* + \beta_2 x + \epsilon_2$$

$$\mathbf{y} = (\mathbf{1}\{y_1^* > 0\}, \mathbf{1}\{y_2^* > 0\})'$$

$$\boldsymbol{\epsilon} = (\epsilon_1, \epsilon_2)' \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$$

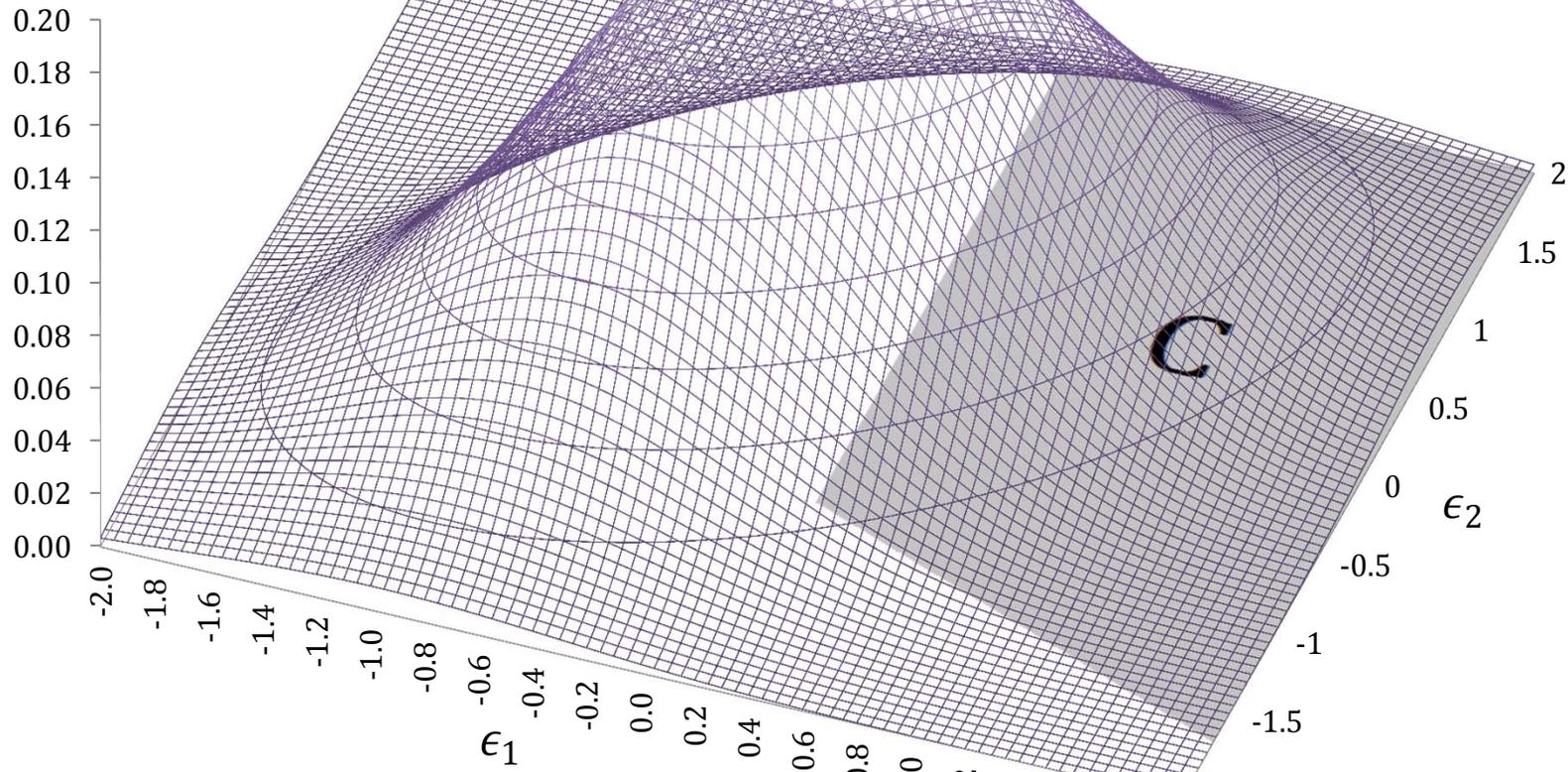
Case $\mathbf{y} = (1, 1)'$

$$\epsilon_1 > -\beta_1 x, \epsilon_2 > -\gamma y_1^* - \beta_2 x = -\gamma \theta_1 - \gamma \epsilon_1 - \beta_2 x$$

$$C = \{(\epsilon_1, \epsilon_2) : \epsilon_1 > -\beta_1 x, \epsilon_2 > -\gamma \theta_1 - \beta_2 x - \gamma \epsilon_1\}$$

$$\mathcal{L}(\beta_1, \beta_2, \boldsymbol{\Sigma}; \mathbf{y}_i | x_i) = \int_C \phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$$

$\phi(\epsilon_1, \epsilon_2, \boldsymbol{\Sigma})$



I don't know how to code that integral

Standard solution

$$y_1^* = \beta_1 x + \epsilon_1$$

$$y_2^* = \gamma y_1^* + \beta_2 x + \epsilon_2$$

$$= \gamma(\beta_1 x + \epsilon_1) + \beta_2 x + \epsilon_2 \text{ (substituting)}$$

$$= (\gamma\beta_1 + \beta_2)x + (\gamma\epsilon_1 + \epsilon_2) \text{ (rearranging)}$$

Define:

$$\pi_1 = \beta_1$$

$$\pi_2 = \gamma\beta_1 + \beta_2$$

$$\boldsymbol{\omega} = (\omega_1, \omega_2)' = (\epsilon_1, \gamma\epsilon_1 + \epsilon_2)'$$

Then the system is:

$$y_1^* = \pi_1 x + \omega_1$$

$$y_2^* = \pi_2 x + \omega_2$$

$$\mathbf{y} = (y_1^*, \mathbf{1}\{y_2^* > 0\})'$$

$$\boldsymbol{\omega} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Omega})$$

SUR system we know how to fit

Detail:

$$\begin{aligned} \boldsymbol{\Omega} = \mathbf{Cov}[\boldsymbol{\omega}] &= \begin{bmatrix} E[\epsilon_1 \epsilon_1] & E[\epsilon_1(\gamma\epsilon_1 + \epsilon_2)] \\ E[\epsilon_1(\gamma\epsilon_1 + \epsilon_2)] & E[(\gamma\epsilon_1 + \epsilon_2)(\gamma\epsilon_1 + \epsilon_2)] \end{bmatrix} \\ &= \begin{bmatrix} 1 & \gamma + \rho \\ \gamma + \rho & \gamma^2 + 2\gamma\rho + 1 \end{bmatrix} \end{aligned}$$

More abstractly...

Rewrite:
$$\begin{bmatrix} y_1^* \\ y_2^* \end{bmatrix}' = \begin{bmatrix} y_1^* \\ y_2^* \end{bmatrix}' \begin{bmatrix} 0 & \gamma \\ 0 & 0 \end{bmatrix} + x[\beta_1 \quad \beta_2] + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \end{bmatrix}'$$

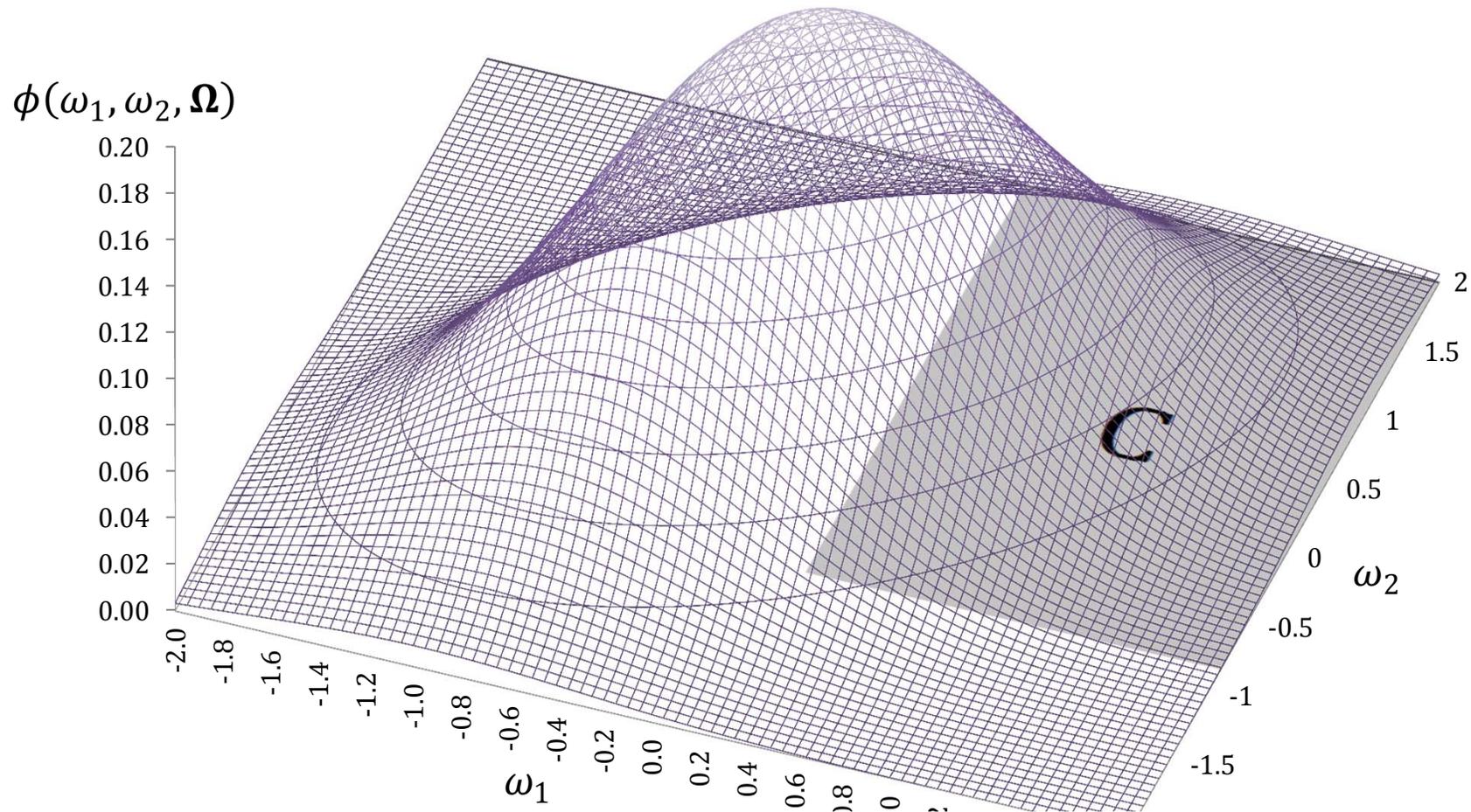
or:
$$\mathbf{y}^{*'} = \mathbf{y}^{*'} \mathbf{\Gamma} + x\mathbf{B} + \boldsymbol{\epsilon}'$$

Solve:
$$\begin{aligned} \mathbf{y}^{*'} &= x\mathbf{B}(\mathbf{I} - \mathbf{\Gamma})^{-1} + \boldsymbol{\epsilon}'(\mathbf{I} - \mathbf{\Gamma})^{-1} \\ &\equiv x\mathbf{\Pi} + \boldsymbol{\omega}' \end{aligned}$$

where:
$$\mathbf{\Pi} = \mathbf{B}(\mathbf{I} - \mathbf{\Gamma})^{-1} \quad \boldsymbol{\omega} = (\mathbf{I} - \mathbf{\Gamma})^{-1'} \boldsymbol{\epsilon}$$

and thus:
$$\begin{aligned} \boldsymbol{\Omega} &= E[\boldsymbol{\omega}\boldsymbol{\omega}'] = E[(\mathbf{I} - \mathbf{\Gamma})^{-1'} \boldsymbol{\epsilon}\boldsymbol{\epsilon}'(\mathbf{I} - \mathbf{\Gamma})^{-1}] \\ &= (\mathbf{I} - \mathbf{\Gamma})^{-1'} E[\boldsymbol{\epsilon}\boldsymbol{\epsilon}'](\mathbf{I} - \mathbf{\Gamma})^{-1} \\ &= (\mathbf{I} - \mathbf{\Gamma})^{-1'} \boldsymbol{\Sigma}(\mathbf{I} - \mathbf{\Gamma})^{-1} \end{aligned}$$

Multiplying system by $(\mathbf{I} - \mathbf{\Gamma})^{-1}$ transforms error space to make \mathcal{C} Cartesian



```
sysuse auto, clear
```

```
cmp setup
```

```
* Fully observed
```

```
cmp (price = foreign ) (foreign = mpg), ind($cmp_cont $cmp_probit)
```

```
* With latent determinant
```

```
cmp (price = foreign#) (foreign = mpg), ind($cmp_cont $cmp_probit)
```

Math still works if Γ not triangular...

Rewrite:
$$\begin{bmatrix} y_1^* \\ y_2^* \end{bmatrix}' = \begin{bmatrix} y_1^* \\ y_2^* \end{bmatrix}' \begin{bmatrix} 0 & \gamma_{12} \\ \gamma_{21} & 0 \end{bmatrix} + x[\beta_1 \quad \beta_2] + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \end{bmatrix}'$$

or:
$$\mathbf{y}^{*'} = \mathbf{y}^{*'} \Gamma + x\mathbf{B} + \boldsymbol{\epsilon}'$$

- References to linear functions need not be recursive
- Full simultaneity OK
- But references to censored variables must still be recursive

```
webuse klein, clear
```

```
cmp setup
```

```
* similar:
```

```
reg3 (consump      wagepriv  wagegovt) ///  
      (wagepriv    consump    govt capital1)
```

```
cmp   (consump = wagepriv# wagegovt) ///  
      (wagepriv = consump# govt capital1), ///  
      ind($cmp_cont $cmp_cont) nolr tech(dfp) qui
```

Random effects and coefficients

- Hierarchical/multilevel
- At each level, can be correlated within and across equations
 - (or not)
- Handled using simulation or (adaptive) quadrature
 - Adaptive quadrature not available yet with more than two levels

```
webuse union, clear
```

```
* Random efect:
```

```
xtprobit union age grade c.south##c.year  
cmp      (union = age grade c.south##c.year || idcode:), ///  
         ind($cmp_probit) tech(dfp) nolr
```

```
* Random coefficient too, potentially correlated with random effect:
```

```
cmp      (union = age grade c.south##c.year || idcode: grade), ///  
         ind($cmp_probit) tech(dfp) nolr
```

```
* Random coefficient too, independent of random effect:
```

```
cmp      (union = age grade c.south##c.year || idcode: grade, cov(ind)), ///  
         ind($cmp_probit) tech(dfp) nolr intpoints(8)
```

```
* Random coefficient only:
```

```
cmp      (union = age grade c.south##c.year || idcode: grade, nocons), ///  
         ind($cmp_probit) tech(dfp) nolr intpoints(8)
```

Probit with random effect

$$y_{it}^* = \beta x_{it} + v_i + \epsilon_{it}$$

$$y_{it} = \mathbf{1}\{y_{it}^* > 0\}$$

$$v_i \sim \mathcal{N}(0, \sigma_v^2)$$

$$\epsilon_{it} | v_i \sim \mathcal{N}(0, \sigma^2)$$

To compute likelihood, need to integrate probability density for v_i, ϵ_{it} over all values of *both* that are together compatible with the y_{it} we observe.

For each group i , integrate over possible values of random effect—weighting by probability of effect of that value—of product of resulting likelihoods of the observations in the group:

$$\mathcal{L}(\beta, \sigma_v^2, \sigma^2; \mathbf{y}_i | \mathbf{x}_i) = \int_{-\infty}^{\infty} \phi(v_i, \sigma_v^2) \left[\prod_t \mathcal{L}(\beta, \sigma^2; y_{it} | x_{it}, v_i) \right] dv_i$$

Inner likelihood computed as in non-random effects models:

$$\mathcal{L}(\beta, \sigma^2; \mathbf{y}_i | \mathbf{x}_i, v_i) = \int_C \phi(\epsilon_{it}, \sigma^2)$$

Numerical method I: Quadrature

- Evaluate function at ~ 4 – 16 special points with special weights
- Traditional strength: computationally cheap
- Weak if integrated function is a narrow bump between special points; Rabe-Hesketh, Skrondal, Pickles 2002)
- Adaptive quadrature fixes by searching for the bump; slower
- Evaluations increase exponentially in dimensions of integration
 - Integrating over 3-D effect can take 4^3 – 16^3 evaluations
 - BUT “sparse grids” can help (Heiss & Winschel 2008)

Univariate nodes:

X_1



X_2

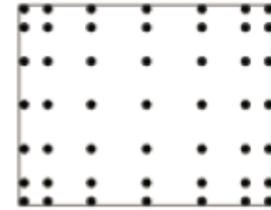


X_3



Product rule:

$X_3 \otimes X_3$



$X_1 \otimes X_2$



$X_1 \otimes X_3$



$X_2 \otimes X_1$



$X_2 \otimes X_2$



$X_3 \otimes X_1$



Sparse grid:

$X_{2,3}$



Source: Heiss & Winschel 2008

Numerical method I: Simulation

- Sample integrated function at 50, 100, 500 points...
- Far more CPU time than quadrature
 - Maybe not more than adaptive quadrature?
- Extends naturally to higher dimensions
 - Multidimensional Halton sequences (Cappellari & Jenkins 2006, Drukker & Gates 2006)
 - Can “scramble” (Kolenikov 2012)
 - Can “step”: first estimate coarsely with 10 points, then refine with 20, 40...

cmp offers (adaptive) quadrature and simulation

- But adaptive quadrature not ready for models with >2 levels

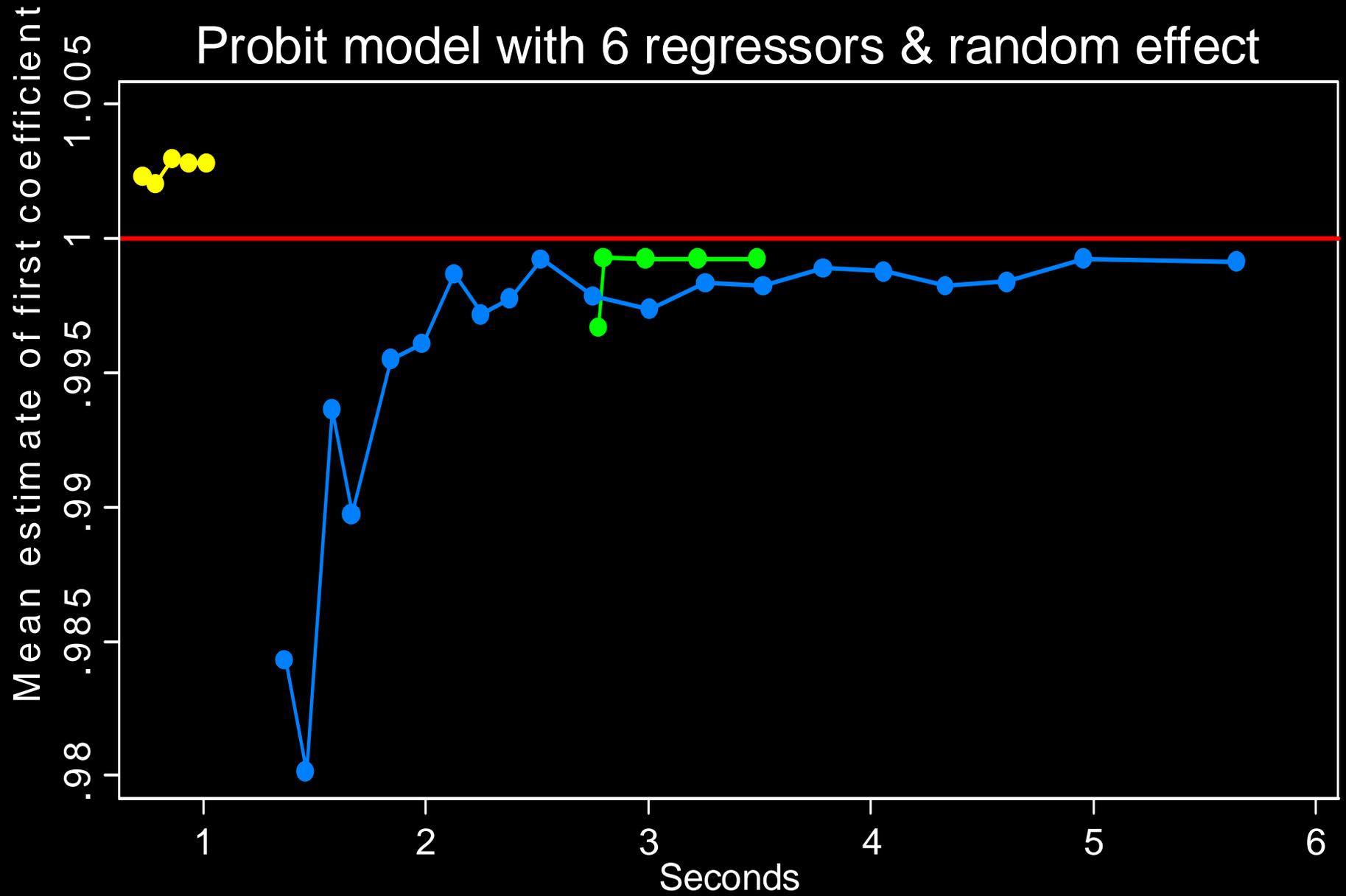
Computation trick: pseudo-linear form

- In Stata ML, new “lf1” evaluator type meets my plea
- Computationally efficient when
 - Evaluator program computes 1st derivatives analytically, but not 2nd
 - Model has linear components $\theta_j = \mathbf{x}' \boldsymbol{\beta}_j$
 - Evaluator does $\frac{\partial \ln \mathcal{L}}{\partial \theta_j}$. Stata does $\frac{\partial \ln \mathcal{L}}{\partial \boldsymbol{\beta}_j} = \mathbf{x}' \frac{\partial \ln \mathcal{L}}{\partial \theta_j}$.
 - Stata computes 2nd derivatives numerically: just one evaluator call per θ_j , not per element of $\boldsymbol{\beta}_j$.
- Requires log likelihood (LL) to have “linear form”
 - Full LL is sum of LLs for each observation
 - Not true in random effects models: LL is a trait of groups
- Trick. For each observation, evaluator provides:
 - Fake LL \equiv full LL / # of observations
 - Real $\frac{\partial \ln \mathcal{L}}{\partial \theta_j}$
 - Correctly, Stata still does $\frac{\partial \ln \mathcal{L}}{\partial \boldsymbol{\beta}_j} = \mathbf{x}' \frac{\partial \ln \mathcal{L}}{\partial \theta_j}$

Computing time for Hessian

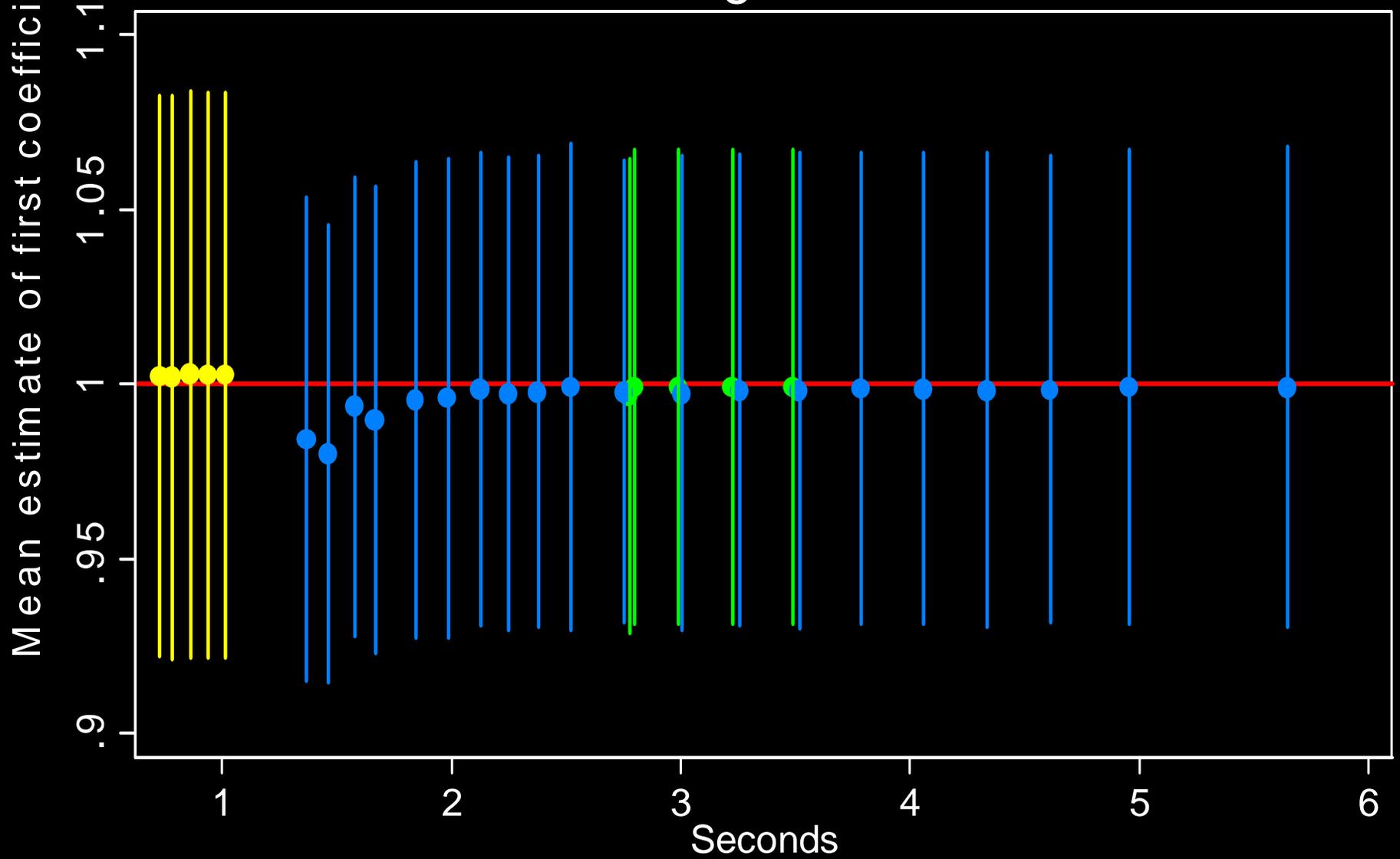
- `gllamm`
 - quadratic in # of parameters
- `cmp`
 - linear in # of linear components + auxiliary parameters

Probit model with 6 regressors & random effect



—●— Quadrature —●— Adaptive quadrature —●— Simulation

Probit model with 6 regressors & random effect

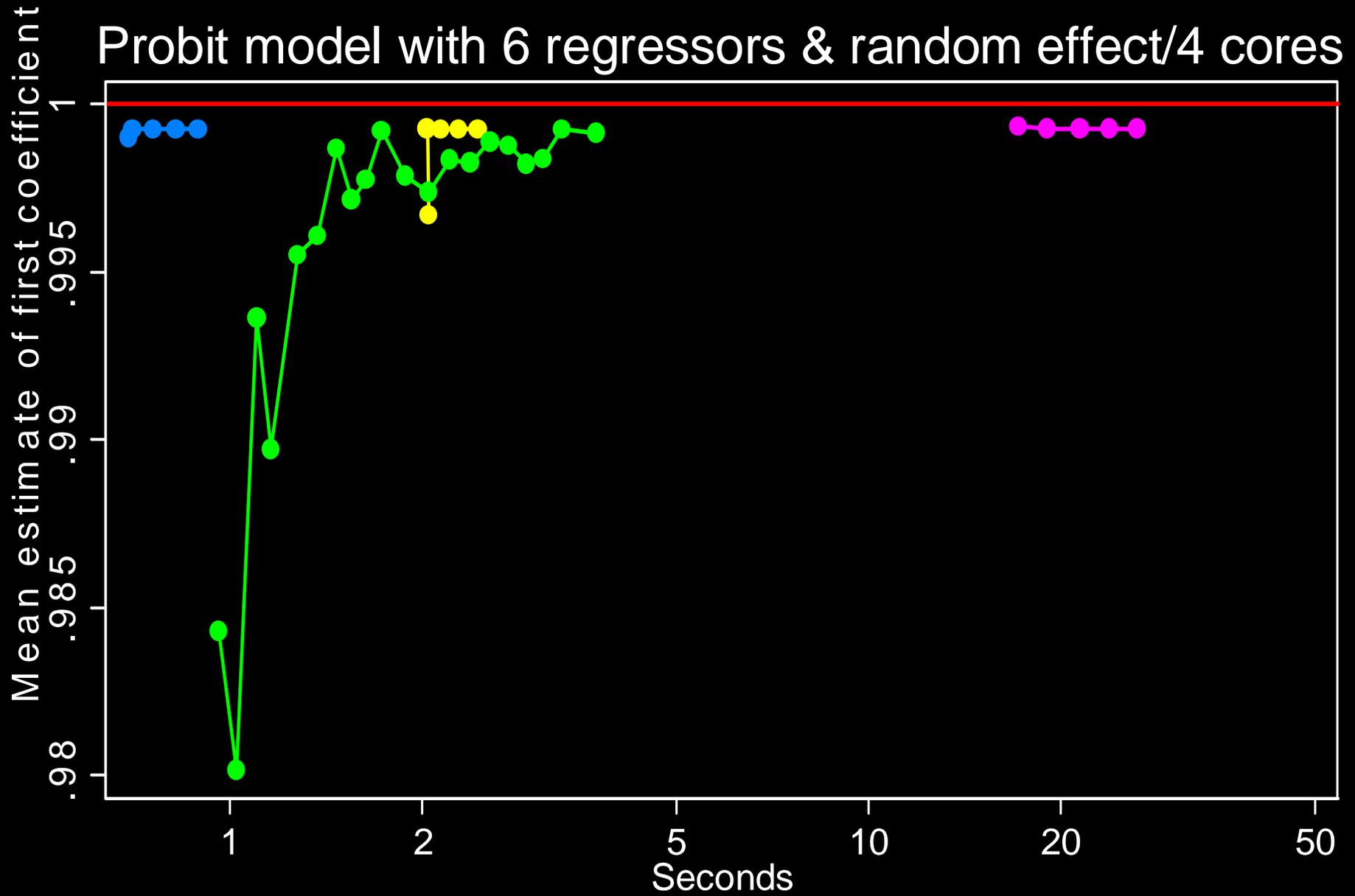


● Quadrature

● Adaptive quadrature

● Simulation

Probit model with 6 regressors & random effect/4 cores



—●— cmp adapt —●— cmp sim —●— xtprobit —●— gllamm adapt