

An Equilibrium Analysis of Marriage, Divorce and Risk-Sharing

By Dan Anderberg

University of Stirling, CEPR and CESifo

January 7, 2002

An Equilibrium Analysis of Marriage, Divorce and Risk-Sharing

Abstract

This paper considers marriage, divorce and reciprocity-based cooperation by couples in the form of sharing of earnings-risk. While risk sharing is one benefit to marriage it is also limited by divorce risk. With search in the marriage market there may be multiple equilibria differing not only in family formation and dissolution patterns but also in the role of marriage in providing informal insurance. Publicly provided earnings-insurance, despite potential equilibrium multiplicity, is shown to affect family formation and financial cooperation monotonically.

Keywords: Marriage, Divorce, Risk-Sharing

JEL Classification: J12, D11, D83, H30

I Introduction

Marriage as an institution has changed dramatically in the recent decades. The patterns are the same in many countries: age at first marriage has increased, cohabitation as a format for partnership has become more popular; but perhaps more controversial than any other indicator of change is the rapid increase in the number of divorces.

Several factors are likely to have contributed to this development, such as e.g. changing labour market conditions, contraceptives, and changes in the legislation surrounding divorce. However, many would also point out the potential role played by the welfare states. Indeed, paralleling the rapid growth of divorce rates has been the expansion of welfare state arrangements in most developed countries: a wide range of benefit systems, including unemployment insurance, benefits to lone parents, housing benefits etc. are likely to make it easier for individuals to cope on their own, and can therefore be expected to affect family formation- and dissolution patterns.

Taking a negative view it is conceivable that a main effect of public benefits and transfers is to crowd out private informal transfers and to make couples less willing to “stick it out”. Then if there are negative externalities associated with divorces – most notably on the children – there is a case for adopting a sceptical view. However, welfare state arrangements may also allow individuals to gain “financial independence” from their partners. Partnerships might then be formed and maintained, not for financial security, but for “love”.

The purpose of this paper is to suggest a simple framework for studying marriage, divorce and financial cooperation. A fundamental aspect of financial cooperation between partners is that it cannot, in general, be assumed to be legally enforced. Hence, more specifically, the focus of the paper is on the consequences of reciprocity-based financial cooperation in the form of voluntary sharing of earnings risk. The lack of legal enforcement is shown to have a number of potentially important effects.

Since risk sharing is a benefit to marriage, large responses to changes in welfare policy may obtain both in terms of marriage and divorce behavior and in terms of cooperation. By increasing the utility associated with being single (and hence the utility associated with divorce), publicly provided earnings insurance directly reduces the cooperation that a married couple can sustain; this in turn lowers the benefit to being married, further increasing the divorce risk, and again lowering the cooperation that can be sustained, and so on. A further “multiplier effect” can obtain through search complementarities in the marriage market: when an individual joins

the marriage market, this increases the probability of every other single individual of finding a new partner. Hence, if public earnings-insurance induces some individuals to choose singlehood rather than marriage, this in turn makes divorce even more attractive for others, leading to further divorces and less cooperation.

The first part of the paper sets up the basic model with a single couple who can cooperate financially, but who also have the option of divorcing. Match-quality varies over time and divorce occurs when the net benefit from remaining married becomes negative. The earnings of each partner fluctuate and the couple can smooth their individual consumption paths through voluntary transfers. The model is then closed by the introduction of a marriage market with search. Given "agglomeration" in search, multiple equilibria exhibiting qualitative differences may occur. While one equilibrium may exhibit a low "turnover" in the marriage market and a high degree of financial cooperation, another equilibrium will exhibit high "turnover" and less risk-sharing by partners. A natural question is which is better: a high- or a low-turnover equilibrium? This answer to this question is argued to be ambiguous when there is reciprocity-based cooperation.

The effects of publicly provided earnings-insurance are then examined. Public insurance is shown to have a monotonic effect on family formation patterns (though the possibility of multiple equilibria implies that levels may vary across economies). The analysis suggests that welfare spending will increase the fraction of single individuals, generally increase the rate of partnership dissolution, and possibly also the rate of partnership formation. The model makes no distinction between marriage and cohabitation, but in so far as divorce rates reflect partnership dissolution rates, the aggregate divorce rate is predicted to be positively associated with social spending.

Do these predictions make sense empirically? How do partnership formation and dissolution patterns vary with welfare spending in a cross-country comparison? Some rough indications can be obtained from Fig 1; the horizontal axis in each display measures social spending as percentage of GDP (in 1994). Focusing initially on entry and exit from first partnerships, Display 1 shows the fraction of "first partnerships that dissolve within five years". Display 2 shows the fraction of individuals who "enter a first partnership before the age of 25". Display 3 then shows the fraction individuals aged 35-39 who are "without current partners". Display 4 shows the aggregate divorce rate. Though the patterns are not overly strong, all displays

suggest monotonic relations to welfare spending.¹

FIG 1.

The model also predicts that publicly provided insurance will reduce financial cooperation by partners. Testing this implication involves comparing equilibrium outcomes across marriage markets with a focus on market-level variables. A tentative analysis of this prediction, using international survey data, is presented towards the end of the paper.

The current paper draws on number of different strands of literature. The literature on marriage and divorce was pioneered in the seminal papers by Becker (1973) and Becker, Landes and Michael (1977). The literature on marriage and divorce is surveyed in Weiss (1997). There is a growing literature on marriage markets with search. An early contribution is Mortensen (1988).

Recent work on has pointed out that inefficient allocations of family resources may obtain (see Lundberg and Pollak, 1994, 2001). The standard rationale for assuming away inefficiencies – e.g. by employing cooperative game theory or “unitary” models – is the belief that efficient allocations can be sustained through repeated interaction. However, not even repeated interaction can guarantee complete cooperation, particularly when there is a risk that the interaction may terminate. I model repeated interaction with (stochastic) endogenous break-ups and argue that inefficient risk sharing may obtain.

Models with marriage market search have recently been used to consider a wide range of issues. E.g. Burdett and Coles (1997) consider assortative mating. Burdett et al. (1999) consider how search while married can generate equilibrium multiplicity. Drewianka (2000) uses a model similar to the current one – albeit with a focus on relation-specific investments – to consider recent proposals for reforming the legislation surrounding marriage and divorce. Chiappori and Weiss (2000) considers a model with divorce settlements, and argue that contract externalities arise since the identities of any future partners are unknown at the time of the divorce. Burdett and Ehrmich (2001) consider out-of-wedlock fertility using a search-framework.

¹Social spending, obtained from the OECD Social Expenditure Database 1980-1997, includes old-age cash transfers, disability cash benefits, occupational injury and disease, sickness benefits, family cash benefits, unemployment compensation, early retirement benefits, as well as expenditures on health and housing. The data on partnership formation, dissolution and partnerships is from UNECE Population Activities Unit (and refer to the early- to mid 90s), supplemented with data from National Survey of Families and Households for the US, as reported in Bumpass and Sweet (1989). Aggregate divorce rates are from the UN 1997 Demographics Yearbook.

A rapidly growing literature – too extensive to be summarized here – considers the effect of welfare policies on family structure and welfare. Some of this literature is closely linked to the wider literature on intergenerational mobility. A very promising branch of this literature seeks to build comprehensive models of marriage, divorce, fertility, and human capital investments which can be used for public policy purposes. See in particular Aiyagari et al (2000), Greenwood et al. (2000a,b). Compared to these models, our aim is more modest and focuses more particularly on one aspect: the implications of reciprocity-based cooperation between partners.

Hess (2000) in an interesting recent study uses micro-level data to indirectly infer the importance of risk-sharing benefits by considering income streams and subsequent divorce behavior.

The current model also draws on the literature on self-enforcing risk-sharing. Two early contributions to this literature are Kimball (1988) and Coate and Ravallion (1993). Kocherlakota (1996) and Ligon et al. (1997) showed that optimal self-enforcing risk-sharing arrangements are generally not stationary (even when the underlying income-generating process is). Ligon et al. (1998) introduced savings, and Foster and Rosenzweig (2000) introduced altruism between the agents sharing risk. The current paper contributes to this literature by allowing for endogenous stochastic formation and breakups by partners sharing risk; on the other hand only “stationary” risk-sharing arrangements are considered.

The paper is organized as follows. Section II sets up the basic risk-sharing model. Section III studies the risk-sharing and divorce behavior of a given married couple. Section IV considers the effect of marginally extending public insurance while still ignoring family formation decisions (marriage and remarriage). Section V introduces the marriage market and Section VI considers the features of steady state equilibria. Section VII discusses a number of extensions while Section VIII presents a brief empirical illustration. Finally Section IX concludes.

II The Risk-Sharing Model

This section describes the basic risk-sharing model. A marriage market will be introduced in Section V, but for now the focus is on a given married couple. In order to keep the analysis tractable the model abstracts from a number of aspects, most notably relations-specific investments and idiosyncratic earnings-expectations. This leaves risk-sharing as the key reason for transfers between spouses, and it also allows us to analytically derive implications of publicly provided insurance. Some possible extensions are, however, discussed in Section VII.

Thus consider a given married couple; the current state of the marriage is described by a

match-quality variable. This variable captures feelings of “love”, and may also include other financial gains to marriage except risk-sharing. To capture the notion that “love comes and goes” the match-quality evolves stochastically as a simple Markov chain.

In each period, each partner receives a random income; for simplicity incomes are assumed to be uncorrelated across time and individuals. The couple can smooth their individual consumption paths through voluntary transfers. Transfers between the partners are, however, not legally enforced; hence they must be supported by expected reciprocity.

Since the couple may break up, divorce risk limits the sustainable levels of cooperation. Note, however, that since risk-sharing is also a benefit to marriage, the prospect of risk-sharing also influences the divorce decision. Each partner always has the option of walking away from the marriage (“no-fault divorce”) and will do so when the net benefit to continued marriage falls short of the benefit from divorce.²

Match Quality

Let $\mu \in \mathbb{R}$ denote “match-quality”, which evolves stochastically. For simplicity – and to focus on transfers between partners as a risk-sharing device – partners are assumed always to agree on the match-quality. There is a finite set of match qualities, $\mathcal{E} = \{\mu_0, \dots, \mu_N\}$; the set \mathcal{E} is ordered increasingly: $j > i$ implies $\mu_j > \mu_i$. Let γ_{ij} denote the probability that the match-quality will be μ_j next period given that it is μ_i in the current period.

Assumption 1. The Markov transition matrix $\mathbf{P} = (\gamma_{ij})_{i,j=0}^N$ is regular and satisfies the following stochastic dominance condition: $\sum_{k=0}^j \gamma_{ik}$ is weakly decreasing in i for every j .

The stochastic dominance property ensures that a good match-quality tomorrow is more likely the better is the match-quality today; it thus ensures a degree of “persistence” of good and bad times.

Incomes and Consumption

Utility of consumption, $u(c)$, is increasing, concave and bounded. In each period each individual earns an income which is randomly drawn from a finite set of possible income levels

²The literature has put forward the idea that divorces are efficient in the sense that they occur when the utility from continued marriage falls short of the sum of the husband’s and wife’s outside opportunities. This efficiency result requires transferable utility and symmetric information (see Becker (1991) and Peters (1986)). In the current model outside opportunities are symmetric and common knowledge.

y^1, \dots, y^M . Earnings realizations are independent across time and individuals; the probability of an individual earning y^i in any period is denoted g_i . There are no savings.

Risk Sharing

Three assumptions about the risk-sharing will be made, all of which require some comments. First, risk-sharing is assumed to occur only between individuals who are currently married – all other relationships are assumed to be either too unstable or not to lend themselves easily to risk-sharing. Second, recent work on risk-sharing (in environments without breakups) has shown that it may be optimal to condition current transfers on past transfers.³ However, since the current analysis introduces endogenous breakups, it is natural to simplify in another dimension; thus transfers are assumed to be conditioned only on current income and match-quality. Third, love does not come in the form of “altruism” – only in the form of enjoyment of being together. Allowing the couple to care about each other’s consumption (or utility) is known to affect self-enforceable risk-sharing arrangements in two ways (see e.g. Foster and Rosenzweig, 2000). First it makes an individual more willing to make transfers to his/her partner, simply out of concern for the other person’s well-being. This relaxes the incentive constraints (see below) and enables more transfers. However, there is also a second effect, viz. an altruistic individual will voluntarily make transfers even when there is no implicit cooperation; this limits the threat of non-cooperation, which in turn tightens the incentive constraints. Avoiding the altruism formulation thus simplifies the analysis, and also avoids making assumptions about how altruistic feelings are affected by the act of divorce.

Given the current match-quality μ and consumption c , the current utility is simply $u(c) + \mu$. An individual’s consumption c may deviate from his/her current income y due to transfers to/from the partner. The partners agree on financial cooperation on a period-by-period basis – no long-term commitment is possible. This takes the following form: prior to learning the current income realizations, the partners agree that if one of them receives an income y^m and the other an income y^k , where $y^m > y^k$, the former should transfer a non-negative fraction, θ_{mk} , of the difference $y^m - y^k$ to the latter. Transfers are assumed to be symmetric (i.e. not depending on who receives y^m and who receives y^k). A risk-sharing agreement is thus an agreement, for one period, on how much to transfer, for each pair of asymmetric income realizations, from the partner with the higher income to the partner with the lower income. It

³See Kocherlakota (1996), Ligon et al. (1997) and Attanasio and Rios-Rull (2000).

is therefore fully described by a vector θ ,

$$\theta = (\theta_{mk})_{m>k} = (\theta_{21}; \theta_{31}; \theta_{32}; \dots; \theta_{m1}; \dots; \theta_{mmj-1}; \dots; \theta_{MMj-1}):$$

If half the earnings-difference is transferred, $\theta_{mk} = 1/2$, the partners enjoy the same consumption; larger transfers will never be relevant. θ can therefore be restricted to be in the set $A = [0; 1/2]^{M(M-1)/2}$.

The agreement θ determines each partner's expected utility from consumption in the current period; denote this expected utility $\hat{A}(\theta)$,

$$\hat{A}(\theta) = \sum_{m=1}^M g_m^2 u(y^m) + \sum_{m=2}^M \sum_{k=1}^{m-1} g_m g_k u(y^m_i + \theta_{mk} y^m_j + y^k + \theta_{mk} y^m_j + y^k) \quad (1)$$

Note that $\hat{A}(\theta)$ is maximized when all $\theta_{mk} = 1/2$; this simply reflects the fact that it would be optimal for the partners to share risk completely in each period. This may, however, not be incentive compatible, and in general the sustainable risk-sharing θ will depend on the current match-quality.

III The Decision Problem Facing a Married Couple

For now the value of being single – denoted $V(s)$ – will be treated as exogenous (and the same for both partners); later on $V(s)$ will be endogenized by the introduction of a marriage market. The focus will be on steady states; hence time will not be included as argument in the Bellman equations.

Incentive Compatible Plans

The timing within each period is as follows: ...rst the couple learns the current match-quality μ . Based on that observation they decide whether to stay together or to divorce. If they stay together they also decide on a risk-sharing agreement θ for that period; ...nally earnings are realized. In the beginning of each period the partners are identical and are assumed to choose among plans so as to maximize their common discounted stream of future expected utilities. However, some risk-sharing agreements may not be incentive compatible. It is therefore

necessary to consider the consequences of a partner failing to make an expected transfer. A failure to make an expected transfer is assumed to lead to an immediate divorce.⁴

A partner who is called upon to make a transfer must therefore be better off making the expected transfer – given that this leads to continued marriage – than unilaterally triggering divorce; since the incentive constraints are forward-looking and seeing as the transitions between match-qualities follow a Markov chain, the couple faces identical decision problems in any two periods where their match-quality is the same. A straightforward dynamic programming approach can therefore be adopted to characterize the couple’s optimal decision. For each $\mu \in \mathbb{R}$, define $V(\mu)$ as the maximal (common) discounted stream of future expected utility given the current match-quality μ ; also let $\beta \in (0, 1)$ denote the discount factor. Since the couple can either divorce – which would give the value $V(s)$ – or stay together, $V(\mu)$ must satisfy the following optimality equation: for all μ ,

$$V(\mu_i) = \max \left\{ V(s); \max_{\theta \in A_i} \left[\beta \sum_{j=0}^{\infty} \beta^j V(\mu_j) \right] + u(y^m) + \beta V(s) \right\} \quad (2)$$

The first term in the large brackets represents the value of breaking up. The second term represents the value of staying together; associated with this option is a choice of risk-sharing agreement θ . If the couple decides to stay together θ must also be such that no one will be better off, at any pair of income realizations, by unilaterally causing divorce through failing to make the agreed on transfer. The set of incentive compatible or “self-enforceable” risk-sharing agreements, A_i – a subset of the set of all possible risk-sharing agreements A – generally depends on the current match-quality μ_i . In particular A_i can be expected to be smaller the worse is the current match-quality, a conjecture that will be verified below.

The self-enforceability constraints are forward-looking and can be formulated as follows: the risk-sharing agreement θ is self-enforcing in state μ_i , i.e. $\theta \in A_i$, if and only if $\theta \in A$ and, for all $m > k$ such that $\theta_{mk} > 0$,⁵

$$u(y^m) + \beta \sum_{j=0}^{\infty} \beta^j V(\mu_j) \geq u(y^k) + \beta V(s) \quad (3)$$

The left hand side is the utility associated with making the prescribed transfer while the right hand side is the utility of deviating, unilaterally triggering divorce. Since the set of self-enforcing

⁴The literature on self-enforcing risk-sharing usually assumes that there is a reversion to the static no-transfer equilibrium. This captures the idea of broken trust.

⁵It is implicitly assumed that the deviating spouse loses the match-quality in the deviating period; this assumption is not crucial.

risk-sharing agreements depends on the current match-quality μ , so will in general the chosen α ; thus let $\alpha(\mu)$ denote the risk-sharing agreement adopted in state μ . Equation (2) together with (3) defines $V(\mu)$ as the solution to a functional equation. Next it is demonstrated that, under a sufficient condition, the functional equation has a unique solution and that $V(\mu)$ has some expected properties.

Risk-Sharing and Divorce

Since match-quality has a direct value it is a natural conjecture that the couple is better off the higher is the current match-quality. Since match-quality has a degree of persistence, and since a low enough match-quality can be expected to trigger divorce, a high current match-quality should be associated with a low future divorce risk. This in turn facilitates more cooperation. Thus it seems natural to conjecture that a high current match-quality is associated with a high current level of financial cooperation. This should further contribute to making the couple better off when the current match-quality is high.

Note however that since the scope for risk-sharing increases when the divorce risk decreases, staying together almost becomes self-justifying. To ensure uniqueness, a condition is imposed. It should be stressed that the condition, which imposes an upper bound on the value of risk-sharing, is only sufficient and, in most cases, probably far from necessary.⁶ Thus assume:

Assumption 2. (A bound on the value of risk-sharing). The following inequality holds:

$$\sum_{m=2}^{\infty} \alpha^m \sum_{k=1}^{\infty} g_m g_k \frac{\tilde{A} u^0(y^k)}{u^0(y^m)} \alpha^{k-1} < \frac{1 - \alpha}{\alpha};$$

Consider e.g. the case where there are only two income levels, y^1 and y^2 . The left hand side increases in the income difference $|y^2 - y^1|$ and in risk-aversion; moreover $g_1 g_2$ is maximized when the income variance is maximized. The inequality then places an upper bound on the weight placed on future utility, α . Suppose e.g. $u^0(y^1) = 2u^0(y^2)$ and $g_1 = g_2 = 1/2$; the condition then requires that $\alpha < 0.8$ while for other g_1 and g_2 the critical α is closer to unity.

The first conjecture can now be verified: the higher is the current match-quality, the better off is the couple.

⁶Indeed, it is a sufficient condition for Equation (2) and (3) to identify a mapping which satisfies Blackwell's sufficient condition for a contraction mapping, which in turn is sufficient the functional equation to have a unique solution (see the Appendix for details.)

Claim 1. The value function $V(\mu)$ exists, is unique, and is weakly increasing in μ .

Proof. See the Appendix.

Knowing that $V(\mu)$ is increasing is sufficient to establish a cut-off rule for the divorce decision. Suppose that the variability in μ is large enough that there will be some match-qualities where the couple stays together and some where they break up; then by defining $\mu^* = \max\{\mu \in \mathcal{M} \mid V(\mu) = V(s)\}$ it follows that the couple stays together only as long as $\mu > \mu^*$. Combining the observation that divorce occurs at low match-qualities with the monotonicity of $V(\mu)$, and invoking the assumption that a high current match-quality is associated with high future match-qualities (Assumption 1) it can also be verified that risk-sharing is an increasing function of the current match-quality.

Claim 2. Given that the couple have not divorced prior to period t , the level of risk-sharing in that period is increasing in the current match-quality: $\mu_j > \mu_i$ implies $\sigma(\mu_j) > \sigma(\mu_i)$.⁷

Proof. See the Appendix.

The main results from this section is thus that the match-quality underlies both the divorce decision and the risk-sharing decision. The better is the current match-quality, the more the partners will cooperate financially, and, due to the persistence of the match-quality lower is the risk of imminent divorce. The model thus makes the very natural prediction that more current cooperation is negatively associated with future divorce risk.

IV A Partial Equilibrium Effect of Formal Insurance

This section provides, by ignoring the possibility of re-marriage, a partial equilibrium analysis of the impact of public insurance on divorce behavior and risk-sharing. Publicly provided insurance is shown to make the couple more prone to divorce in the sense that it expands the set of match-qualities where they divorce.

Two forces are at work. First, since formal insurance is more valuable to an individual who has no other insurance available, there is direct positive effect of formal insurance on the probability of divorce: intuitively formal insurance implies that an individual can afford to leave a relationship that has gone sour even if that means forgoing future access to informal risk-sharing. However, increasing the utility of being going-it-alone also reduces the cooperation

⁷ $\sigma(\mu_j) > \sigma(\mu_i)$ if every element in $\sigma(\mu_j)$ is at least as large as the corresponding element in $\sigma(\mu_i)$.

in states where the self-enforceability constraint binds, which further decreases the value of continued marriage.

To demonstrate these effects formally it is convenient to focus on the case where there are only two possible income levels y^1 and y^2 ; this simplifies the analysis and avoids making assumptions about the form of the publicly provided insurance: as long as formal insurance is based only on current individual income it reduces to a net transfer from individuals with high income, y^2 , to individuals with low income, y^1 . Thus let ζ denote the tax imposed on individuals with a current high income. By budget balance, the transfer to low-income individuals must equal $(g_2=g_1)\zeta$. Hence, given ζ , the net incomes are

$$y^2 = y^2 - \zeta; \quad \text{and} \quad y^1 = y^1 + \frac{g_2}{g_1}\zeta. \quad (4)$$

Consider then a marginal expansion of formal insurance from a situation with less than full insurance (i.e. initially $y^2 > y^1$). Note that, with formal insurance included, and using $M = 2$, the definition of the within-period expected consumption utility, $\bar{A}(\zeta)$, reduces to:

$$\begin{aligned} \bar{A}(\zeta) = & g_2^2 u^i(y^2 - \zeta) + g_2 g_1 u^i(y^2 - \zeta) \bar{\mu}^i(y^2 - \zeta, y^1 + \frac{g_2}{g_1}\zeta) \\ & + g_1 g_2 u^i(y^1 + \frac{g_2}{g_1}\zeta) \bar{\mu}^i(y^2 - \zeta, y^1 + \frac{g_2}{g_1}\zeta) + g_1^2 u^i(y^1 + \frac{g_2}{g_1}\zeta) \end{aligned} \quad (5)$$

Note also that the subscript on $\bar{\mu}$ has been dropped since, in the case with just two income levels, a risk-sharing agreement $\bar{\mu}$ reduces to a scalar.

To emphasize the impact of ζ , we can include ζ as a formal argument for the critical match-quality where divorce occurs, $\bar{p}(\zeta)$. The main result is that $\bar{p}(\zeta)$ is monotonic in ζ ; however, as the proof demonstrates, public insurance also crowds out private transfers: increasing ζ (weakly) decreases the absolute informal transfer (i.e. $\bar{\mu}^i(y^2 - \zeta, y^1 + \frac{g_2}{g_1}\zeta)$) in each state μ_i where the couple remain married.

Claim 3. Suppose that $M = 2$ and that remarriage is not possible. Then $\bar{p}(\zeta)$ is non-decreasing in ζ and the informal income transfer in each state is decreasing in ζ .

Proof. See the Appendix.

Publicly provided insurance is often suspected of crowding out private insurance coverage. But formal insurance can also crowd out less formal forms of insurance that occur within families. This was noted e.g. by Berry-Cullen and Gruber (2000) who argue that the existence of formal unemployment insurance may be a potentially important reason for why the literature on the

so-called added-worker effect typically finds relatively small effects. The current model focuses on direct transfers between partners as opposed to compensating income streams.

Yet the analysis suggests that crowding out can be pervasive: expanding formal insurance may crowd out informal risk-sharing at a high rate. The direct effect of public insurance is to make singlehood relatively more attractive by reducing the associated consumption variance; this effect reduces the sustainable level of cooperation for couples that remain married (by tightening the self-enforceability constraints). This in turn reduces the benefit to marriage, further increasing the relative attractiveness of singlehood, and so on.⁸

V Family Formation

The analysis in the previous section was only of a partial equilibrium nature in that family formation was ignored. To capture the feedback effect that can obtain through search complementarities in the marriage market, the model is now expanded to include family formation.

The Marriage Market

Consider an economy with a continuum (of unit measure) of infinitely lived identical individuals. Each individual is either married or single; let $S \in [0, 1]$ denote the number of single individuals. Single individuals search for new partners, and $\hat{A}(S)$ denotes the probability of finding a potential partner during a period of search (for simplicity a searching individual meets at most one potential partner during a period of search). The matching rate \hat{A} depends non-negatively on the number of searching individuals S , that is $\hat{A}'(S) \geq 0$ ("agglomeration"). Search is costless; however, a single individual has no one to share income with. Hence the expected utility during a period of search is $\hat{A}(0)$.

Potential partners meet at the end of a period of search; at the beginning of the next period they learn their initial match-quality. Based on the initial match-quality two newly matched individuals decide whether to form the partnership or to continue to search. Since the Markov process for the match-qualities is regular it has a limiting distribution F defined on \mathcal{E} . The density of F , denoted f , is strictly positive on \mathcal{E} and is uniquely defined through the following

⁸Di Tella and MacCulloch (1999) and Attanasio and Rios-Rull (2000) have recently considered the crowding out effect of formal insurance on voluntary risk-sharing (although without considering breakups).

equations:

$$f(\mu_i) = \sum_{k=0}^{\infty} \gamma_k f(\mu_k); \quad \text{and} \quad \sum_{i=0}^{\infty} f(\mu_i) = 1. \quad (6)$$

The distribution F is a natural candidate for the distribution of initial match-qualities. Since I assume that there are no specific marriage- or divorce costs, two newly matched individuals are in exactly the same position as a couple, with the same match-quality, that have been married for any arbitrary number of periods. Hence they will all adopt the same critical match-quality $\bar{\mu}$ when making their partnership formation/dissolution decisions.

The Choice of Critical Match-Quality

When remarriage is possible, the value of starting a period as single, $V(s)$, is endogenous. Formally, there is, in addition to (2) and (3) characterizing $V(\mu)$, an equation for $V(s)$. This equation has the following form:

$$V(s) = \bar{A}(0) + \beta \sum_{j=0}^{\infty} \gamma_j f(\mu_j) V(\mu_j) + (1 - \beta - \bar{A}) V(s); \quad (7)$$

where $\bar{A}(0)$ is the current expected utility and the bracketed term is the value of the continuation; $\bar{A}f(\mu)$ is the probability that the individual will find a new partner with initial match-quality μ and with probability $1 - \bar{A}$ no new potential partner is located during the period. The matching probability \bar{A} is taken as given by a searching individual even though it is endogenously determined by the aggregate behavior of the individuals.

Since $V(s)$ is now endogenous, the critical match-quality adopted by the agents when making partnership formation/dissolution decisions, $\bar{\mu}$, is best viewed as a function of the matching probability \bar{A} . To highlight this we can introduce the notation $\bar{\mu}(\bar{A})$. Clearly, the larger is \bar{A} , the easier it is to find new potential partners. This implies that singlehood becomes more attractive: why stay with a partner when the relationship has turned sour if it is easy to find a new partner? Equally, it makes sense to be "picky" when meeting a new potential partner. Thus, as \bar{A} increases, $\bar{\mu}(\bar{A})$, should, if anything, increase. But a high $\bar{\mu}$ also implies an increased divorce risk, which should reduce the financial cooperation that can be sustained; indeed the following result can be obtained:

Claim 4. The critical match-quality $\bar{\mu}(\bar{A})$ increases in \bar{A} and, moreover, risk-sharing $\bar{\mu}(\mu)$ decreases in \bar{A} for all $\mu > \bar{\mu}$.

Proof. See the Appendix.

Note that Claim 4 indicates a key strategic complementarity: the larger is the fraction of the individuals who are single, the more attractive it is for each individual to join the pool of single individuals.

Flow Equilibrium

The steady state is characterized by flow equilibrium. Moreover, the fraction of single individuals naturally depends positively on the critical match-quality $\bar{\mu}$. To see this, note that S has the alternative interpretation as the fraction of total time that an individual spends as single if the process is allowed to go on forever. Let $\tau(\mu) > 0$ denote the expected duration of a new marriage with initial quality μ . The expected time that a single individual is away from the pool of singles upon meeting a potential partner is then $\int_{\mu > \bar{\mu}} f(\mu) \tau(\mu) d\mu$, which naturally decreases in $\bar{\mu}$ (both since fewer meetings will result in marriages, and since the expected duration of every new marriage will be shorter). Using that $1/\bar{\mu}$ is the expected time until a potential partner is located the steady state fraction of single individuals must then satisfy:

$$S = \frac{1/\bar{\mu}}{1/\bar{\mu} + \int_{\mu > \bar{\mu}} f(\mu) \tau(\mu) d\mu} \quad (8)$$

Equation (8) implicitly and uniquely defines S as an increasing function of $\bar{\mu}$, henceforth denoted $S(\bar{\mu})$.

Lemma 5. The steady state fraction of single individuals in the economy, S , is increasing in the critical match-quality, $\bar{\mu}$.

Steady States

A steady state equilibrium is characterized by two conditions: flow equilibrium and individual rationality. Flow equilibrium was characterized just above, where it was found that the steady state fraction of single individuals, S , varied positively with the critical match-quality, $\bar{\mu}$. Individual rationality requires that $\bar{\mu}$ is privately optimal given the matching-rate $\bar{\mu}$, which in turn depends on S . A steady-state equilibrium can thus be characterized as a fixed-point for the mapping obtained by taking the composition of $S(\bar{\mu})$, $\bar{\mu}(S)$ and $\bar{\mu}(S)$; noting that all three components are non-decreasing (see Lemma 5 and Claim 4) it follows that the composite mapping $S(\bar{\mu}(\bar{\mu}(S)))$, which maps the unit interval into itself, is also non-decreasing. Hence, by

Tarski's fixed-point theorem, an equilibrium exists. Furthermore, given that there is sufficient "spread" in match-qualities, any equilibrium will naturally be "interior".⁹

VI Equilibrium Features

This section first looks at how the provision of formal insurance affects family-formation and cooperation in the general equilibrium setting. Then qualitative differences of multiple equilibria are investigated. Finally, the welfare properties of decentralized steady state equilibria are considered.

General Equilibrium Effects of Formal Insurance

The possibility of multiple equilibria offers a potential explanation for why countries with similar levels of social expenditures can have quite different family formation and dissolution patterns (as well as different attitudes to the role of the family in providing financial security). Moreover it allows this observation to be fully consistent with the claim that publicly provided insurance affects partnership formation and dissolution decisions, as well as the role of the family in providing financial security, in a monotonic fashion.

To demonstrate this, publicly provided insurance is now introduced into the general equilibrium model. Consider again the case with only two income levels, y^1 and y^2 ; net incomes given by (4) and ζ represents the generosity of public insurance. As usual with multiple equilibria, it is of interest to look at the "extremal equilibria". Thus let S_L and S_H denote the lowest- and the highest steady state fraction of single individuals respectively. To emphasize the impact of formal insurance, let ζ be an argument for the two bounds, i.e. $S_i(\zeta)$, $i = L; H$.

In Section IV it was noted that β was increasing in ζ when no remarriage was possible; the underlying reason was that formal insurance is more valuable to single individuals than to individuals who have access to informal risk-sharing. The same effect is at work in the general equilibrium context implying that β still tends to be increasing in ζ ; although a general proof is not available, simple sufficient conditions can be obtained. Consider e.g. the following "memoryless stochastic process" which allows the match-quality to be a continuous variable, $\underline{\mu} \leq \mu \leq \bar{\mu}$. Given any current match-quality μ , the probability that a shock occurs, which

⁹ If the best match-quality, μ_N , is positive such a match will never be rejected. Moreover, if the worst match-quality, μ_0 , is so bad that an individual would rather be single forever, there will always be some single individuals.

changes next period's match-quality, is $\mu \in (0, 1)$ (conversely, with probability $1 - \mu$ the match-quality remains μ). If a shock occurs the new match-quality is drawn from some distribution F , the density of which is strictly positive on \mathcal{E} . Similarly, all initial match-qualities are also drawn from F .¹⁰ For this process – which will be used more extensively below – it can be shown that a sufficient (but not necessary) condition for β to be increasing in ζ is that $\beta > \bar{A}$.¹¹ The condition $\beta > \bar{A}$ implies that match-qualities are not too stable.

Returning to the main model, suppose then that β increases in ζ , i.e. that the direct effect is to make singlehood more attractive. Due to the strategic complementarity in the individuals' decisions to join the singles-pool, the direct effect carries over to the general equilibrium setting. The set of equilibria thus moves monotonically "upwards" as illustrated by Fig 2. The direct effect of the increased public insurance is to (weakly) increase $\beta(\bar{A})$ for every \bar{A} ; a steady state exists wherever $\beta(\bar{A}(S))$ intersects the low-equilibrium condition $S = S - \beta$ (which is unaffected by ζ). Stated in precise terms:

Claim 6. Suppose that $M = 2$ and that β increases in ζ ; then $S_L(\zeta)$ and $S_H(\zeta)$ both increase in ζ .

Proof. See the Appendix.

FIG 2.

The model thus predicts that there is an underlying monotonic impact of an expansion of formal insurance of family formation and breakup behavior in the sense that the set of equilibria moves towards people spending more time as single.

Letting μ_1 denote the rate at which single individuals marry, $\mu_1 = \bar{A}(S) - 1 - F(\beta)$, and using μ_2 to denote the average rate at which married individuals divorce, low equilibrium implies $S = \mu_2 / (\mu_1 + \mu_2)$. Hence an expansion of formal insurance will increase the relative divorce rate μ_2 / μ_1 . More can be said about the absolute divorce rate if more specific stochastic processes are assumed. Consider e.g. the "memoryless" stochastic process introduced above. For this process

¹⁰ F is also the long-run distribution associated with the stochastic process.

¹¹ For this specific process, $\Phi(\mu) = V(\mu) - V(S)$ satisfies

$$\Phi(\mu) = \max \left\{ 0; \bar{A}(\mu) - \bar{A}(0) + \mu + \int_{\mu}^{\bar{A}} \Phi(\mu^0) dF + (1 - \mu) \Phi(\mu) \right\}$$

Using this a proof can be constructed along the lines of that of Claim 3.

the average divorce rate \bar{d} is equal to $\frac{1}{2}F(\beta)$. Hence, for this specific process, \bar{d} increases as long as β increases.

The impact of insurance on the marriage rate \bar{m} is more ambiguous; on the one hand, if insurance leads to an increase in β , this tends to decrease \bar{m} . But on the other hand, since S increases, so does the matching rate \bar{A} , generating an opposing effect on the marriage rate.

Turning to financial cooperation, recall that risk-sharing is decreasing in the matching rate \bar{A} ; then since insurance leads to an increase in S , and hence in \bar{A} , formal insurance leads to a reduction in cooperation in the sense that $\bar{\mu}$ decreases; in other words, the model exhibits crowding out also in the general equilibrium setting.

Qualitative Differences of Multiple Equilibria

The logic of the comparative static exercise carries over to a comparison of multiple equilibria. Thus consider an economy with at least two steady state equilibria; to avoid new notation consider the extremal equilibria, L and H , where $S_L < S_H$. Since $S_i = \bar{d}_i = (\bar{d}_i + \bar{m}_i)$, $i = L, H$ it follows that the relative aggregate divorce rate is higher in equilibrium H than in equilibrium L , i.e. $\bar{d}_H = \bar{m}_H > \bar{d}_L = \bar{m}_L$. Moreover, for the “memoryless” stochastic process the absolute divorce rate is higher in equilibrium H than in equilibrium L , $\bar{d}_H > \bar{d}_L$.¹²

Since, people spend more time as single (and generally marriages have shorter expected duration and divorce rates are higher) in equilibrium H than in equilibrium L it is natural to think of the former as a “high turnover” equilibrium and the latter as “low turnover” equilibrium. From the fact that the matching rate \bar{A} is higher in the high turnover equilibrium than in the low turnover equilibrium, it then also follows that, for any given current match-quality, μ , a married couple cooperates less in the form of risk-sharing in the high turnover equilibrium than in the low turnover equilibrium.

The model thus captures the idea that two fundamentally similar economies can sustain different equilibria where the people in one economy appear to be more “committed” to partnerships and cooperate more financially than the people in the other economy. Such differences can be expected to be reflected in measured attitudes towards the role of the family. In other words, the role of social norms may be to act as a coordination device under multiple equilibria.

¹²This follows since S is monotonically related to β (Lemma 5), and, for this specific process, $\bar{d} = \frac{1}{2}F(\beta)$. Hence $S_L < S_H$ implies $\bar{d}_L < \bar{d}_H$.

Welfare Aspects

In this short section I sketch two potentially important welfare implications of reciprocity-based cooperation. First, it is noted that while a standard search-externality would imply that there are too few single individuals in equilibrium, this conclusion can be overturned when expanding the pool of single individuals reduces the scope for cooperation by currently married couples. Second, it is noted that the presence of reciprocity-based cooperation may have implications for the desirability of variability in match-qualities.

A matched couple views their match as having an option value; this value determines the critical match-quality $\bar{\theta}$ adopted in a decentralized steady state. Consider e.g. the memoryless stochastic process described above. Each individual treats the matching rate \bar{A} as given. A couple then stays together as long as the value of doing so exceeds the value of breaking up, implying that $\bar{\theta}$ is characterized by $V(\bar{\theta}) = V(s)$. Manipulating this characterization, and using $r = (1 + \delta)^{-1}$ to denote the implicit "interest rate" corresponding to δ , yields the following more illuminating expression for $\bar{\theta}$,

$$\bar{A}(\bar{\theta}) + \beta \bar{A}(0) = (\bar{A} + \delta) \int_{\bar{\theta}}^{\infty} \frac{\mu \bar{A}(\mu) + \mu \bar{A}(\bar{\theta})}{r + \delta} dF \quad (9)$$

Note first that the integral on the right hand side is positive since the current utility (when married) is increasing in μ . The left hand side is the difference between the current utility from marriage at the critical match-quality $\bar{\theta}$ and the current utility from singlehood. This difference is then positive if and only if $\bar{A} > \delta$; in other words, if joining the singles-pool is a faster way to obtaining a new match-quality than remaining married, then an individual is willing to forego current utility in order to take that chance.

If risk-sharing could be perfectly enforced, complete risk-sharing, $\theta_{mk} = 1=2$, would always obtain. In that case the characterization of the socially efficient cut-off match-quality would be identical to Equation (9), except with $\bar{A} + \delta \bar{A}(\bar{\theta})$ replacing $(\bar{A} + \delta)$, implying a higher cut-off match-quality $\bar{\theta}$; the extra term indicates the benefit to the searching individuals of expanding the pool of singletons (a standard search-externality – see e.g. Diamond, 1982). However, when risk-sharing cannot be legally enforced, the effect of expanding the pool of single individuals is to increase the matching rate \bar{A} which reduces the sustainable risk-sharing by currently married couples. In other words, the effect of expanding the pool of single individuals is also to make divorce privately more attractive, which, when cooperation is sustained by reciprocity, reduces

risk-sharing by currently married couples.¹³

Hence, contrary to the case where any amount of cooperation can be sustained, in this case a steady state equilibrium may have (locally) too few married individuals. Equally, if there are multiple equilibria these cannot be unambiguously welfare-ranked: whether a “high turnover” equilibrium (with a high average match-quality among married individuals and low levels of risk-sharing) or a “low turnover” equilibrium (with a lower average match-quality but higher levels of risk-sharing) is better cannot be determined on an a priori basis.

A second welfare implication of reciprocity-based cooperation concerns the variability in match-qualities. Suppose that risk-sharing could be enforced; in that case the value function $V(\mu)$ would be naturally convex. The convexity arises since the individuals have the option of leaving bad matches and instead joining the marriage market.¹⁴ In that case, an increase in the spread of the long-run distribution, F , of match-qualities implied by the stochastic process will generally be good for welfare. The intuition is simple: increasing the probability of good states and bad states, while keeping the average constant, will be good for welfare since the good states will be enjoyed while the bad states can be rejected.

In contrast, when risk-sharing is based on expected reciprocity, an increase in match-quality not only has a direct value, but also enhances risk-sharing. However, the second benefit to increased match-quality eventually diminishes since the value of insurance diminishes (due to V being concave) and, also, since at good-enough match-qualities, complete risk-sharing may be sustainable.

As a consequence, the value function $V(\mu)$ generally fails to be globally convex (but, rather, tends to have an inflection point). As a consequence, a mean-preserving spread in the long-run distribution of match-qualities, F , need not be welfare improving. Intuitively, if probability mass is (marginally) shifted from middle-range values the impact on the direct benefit is negligible, but the loss of utility due to reduced risk-sharing from the downward shifting of mass may not be matched by a corresponding increase in utility from increased risk-sharing associated with the upward shift of mass. Hence, with reciprocity-based cooperation, it may be beneficial to

¹³Note that even when risk-sharing is enforced, a married couple would become more prone to divorce when an additional individual joins the pool of singletons: this is simply to the strategic complementarity and would not constitute an externality. The difference here is that the utility of a married couple is negatively affected while remaining married.

¹⁴Consider e.g. the “memoryless” stochastic process described above; with enforced complete risk-sharing, $V(\mu)$ would be linearly increasing in μ above $\bar{\mu}$ and equal to $V(\bar{\mu})$ at all $\mu < \bar{\mu}$, thus making $V(\mu)$ globally convex.

welfare if match-qualities are less variable (that is, if there are less “booms and busts” in love) since this may promote valuable cooperation between partners.

VII Extensions

Duration Dependence and Learning A restrictive feature of the above framework is that partnership duration has no role to play, neither for partnership dissolution behavior nor for cooperation. Duration effects on the divorce-hazard has, however, been observed empirically (Weiss and Willis, 1997). Incorporating duration effects would require a different modeling strategy; an interesting alternative would be to abandon the Markov assumption and instead assume that, for each couple, there exists an underlying match-quality which is fixed, but which is only revealed over time (in the spirit of Jovanovic (1979)). A couple that is observed to have stayed together for a long time is then likely to strongly believe that their match-quality is good. Since in such a generalization, risk-sharing would be positively related to the couple’s current beliefs that their match is good, cooperation would tend to be positively related to duration.

Relation-Specific Investments Empirical evidence suggests that children and joint property stabilize marriages, causing the individual divorce hazard to drop over time (see Weiss and Willis, 1997). The above model abstracted from relation-specific investments in order to keep the analysis manageable. Incorporating relation-specific investments into stochastic models of marriage and divorce poses a great challenge; the incentives for investments hinges in an interesting way on the degree of irreversibility of such investments, and on the assumed stochastic structure (see Drewianka, 2000). E.g. in the current framework, it is conceivable that irreversible relation-specific investments (e.g. having a child, building a common network of friends etc.) may increase divorce costs. However, in the current framework “love never lasts forever” which limits the incentives for such investments.

One approach that is feasible within the current framework is to treat investments as occurring exogenously.¹⁵ If a stock investments increases the cost of divorce, it also affects divorce behavior and financial cooperation. So, for example, the presence of children can be expected to be associated with more financial cooperation. Similarly, investments, in so far as they are

¹⁵A simple formulation would be to include a benefit ϕ from marriage (in addition to the match-quality μ) which increases in duration $\phi(t + 1) > \phi(t)$. This would make the value function duration dependent, but the analysis and the results would still go through.

unobserved, may contribute to duration dependence in divorce behavior and ...nancial cooperation.

Persistent Income Shocks Allowing persistent income shocks could potentially lead to a number of interesting insights. When risk-sharing relies on reciprocity, an individual's willingness to make transfers to his/her partner hinges on an expectation that the favour may be reciprocated sometime in the future. However, if income shocks are strongly persistent, then the time until any favour can be expected to be reciprocated will be long. Hence it is possible that e.g. labour market institutions will have implications for the sustainability of ...nancial cooperation.

However, allowing persistent income shocks would also complicate the analysis. E.g. when meeting in the marriage market each single individual would be characterized by an individual-specific expected stream of future incomes. Thus it would be necessary to consider bargaining between newly matched individuals.

VIII An Tentative Empirical Investigation

The above analysis suggests that ...nancial cooperation between partners may be affected both by marriage market conditions and welfare spending. Testing this requires comparing outcomes across economies. While any such investigation must necessarily be treated with caution, this section presents a tentative analysis. The need to use cross-country data unfortunately rules out using micro-level data on individual consumption levels (as is typically used in the literature on intra-household allocations¹⁶). Instead I rely on a self-reported measure of cooperation obtained from the International Social Survey Programme (ISSP) 1994 survey on "Family and Changing Gender Roles". Couples were asked how they organize their incomes, in particular whether they "pool their incomes". One potential problem is that the interpretation of "income pooling" is not unequivocal, and may not be equivalent to "risk-pooling", especially if it reflects the outcome of a joint household decision making process involving specialization. What is of key importance for our purposes is that the reported measure captures an element of reciprocity-based cooperation. However, in the hope to capture risk-pooling rather than specialization, I focus on individuals who are full-time employed, and whose partners are also full-time employed.¹⁷

¹⁶See e.g. Browning and Chiappori (1996) and the references therein.

¹⁷Including part-time employed workers did not significantly affect the results.

I use two binary dependent variables: first whether the respondent has a “partner”, and second, if so, whether “incomes are pooled”. Both decisions depend on unobserved stochastic factors, e.g. match-quality, which can be expected to be correlated. Hence I use a bivariate probit model with sample selection which is estimated using the maximum likelihood method. Let $z^a = \beta^a x + \epsilon^a$ be a latent variable. The respondent “has a partner” ($z = 1$) if $z^a > 0$ and is “single” ($z = 0$) otherwise. For those individuals who have partners, let $q^a = \beta^a w + \epsilon^a$ be a second latent variable such that income pooling with the partner is “complete” ($q = 1$) if $q^a > 0$ and otherwise is “incomplete” ($q = 0$) (see below). $(\epsilon^a; \epsilon^q)$ has a bivariate normal distribution with zero means, unit variances and correlation ρ . The two sets of regressors, x and w , may overlap but need not be identical.

The Data

The ISSP 1994 survey was conducted in 22 countries; however, I restrict my attention to OECD countries; I also restrict the sample to individuals aged 20-65 who are full-time employed, and whose partners, if any, are also full-time employed. The final sample consists of 7,779 individuals from 16 countries: Australia, Austria, Canada, Czech Republic, Germany, Hungary, Ireland, Italy, Japan, Netherlands, New Zealand, Norway, Poland, Sweden, UK, and the US.¹⁸

I include a number of market-level variables. To capture the effect of welfare policy, I use social transfers as per cent of GDP. I include each country’s divorce rate (divorces per 1000 population). As measures of labour market institutions I use the female labour force participation rate and an employment protection legislation index calculated by the OECD which can range (continuously) from 0 to 6 with higher values representing stricter regulation.¹⁹

At the individual level age, age squared, gender (dummy for “male”) were included in both sets of regressors. For education, categories were used; the base category is “some secondary education” and dummies were included for “primary education or less” and “some university education”. In the income-pooling equation I use information on household earnings (log of net annual earnings in 1994 US Dollars). Since children, as a relation-specific investment, may

¹⁸Spain had to be eliminated due to lack of information on education.

¹⁹Social transfers were obtained from the OECD Social Expenditure Database 1980-1997 (See fn. 1). Supplementary information was obtained from the IMF Government Finance Yearbooks. Divorce rates are from the UN 1997 Demographics Yearbook. Female labour force participation rates were obtained from the ILO Yearbooks of Labour Statistics. The employment protection index is from OECD Employments Outlook, June 1999, Table 2.5, Overall strictness, version 1.

stabilize marriages I control for household size. I include dummies for prior divorce: though the theory does not predict any direct impact of prior divorce on financial cooperation, it is likely to be negatively related to the (unobserved) duration of the current relationship and can therefore be expected to be negatively associated with cooperation. Finally, dummies were included for “regular attendance to religious services” and self-employment.

The first dependent variable is a dummy which is unity if the respondent has a “partner” (spouse or steady life-partner). For those individuals with partners there is a second dependent variable constructed from the following question:

How do you and your spouse/partner organize the income that one or both of you receive?

The available answers were (i) “I manage all the money and give my partner his/her share”, (ii) “My partner manages all the money and gives me my share”, (iii) “We pool all the money and each take what we need”, (iv) “We pool some of the money and keep the rest separate”, (v) “We each keep our own money separate”. The most frequent answer was (iii); only about 12 percent of the answers fell in category (i) and (ii). I interpret (i) through (iii) as full financial cooperation or “full income pooling”, ($y = 1$) while (iv) and (v) is interpreted as “incomplete income pooling” ($y = 0$).

Results

As a preliminary step I estimated the model using only the explanatory variables measured at the individual level together with a set of country-dummies.²⁰ This revealed significant country effects. E.g. people in the Nordic countries in particular were less likely to pool their incomes conditional on having a partner. I then proceeded by replacing the country-dummies with the aforementioned market-level variables. This produced no qualitative effects on the coefficients for the individually measured variables.

Column 2 of Table 1 shows the result for the “partner-equation”.²¹ The “marginal effects”

²⁰The results are available on request from the author.

²¹One must however keep in mind the limitations; since the key variables are only measured at the country-level the “effective number of observations” is obviously low. Also, if the individuals in the same country are affected by some common component not accounted for their error terms will be correlated; the effects of the country-level variables may then be measured less precisely than their t-ratios suggest.

in this case are calculated as traditional probit marginal effects:

$$\frac{\partial \Pr(z = 1|x)}{\partial x} = \hat{A} \phi(x);$$

where $\hat{A}(\cdot)$ is the pdf for the standard univariate normal distribution.

Age is naturally the most important factor explaining whether or not an individual has a partner. In line with the literature, education also matters. Relative to primary education, having secondary education increases the probability of having a partner. Further education on the other hand reduces the probability of having a partner, possibly because highly educated individuals tend to marry late. A prior divorce naturally negatively affects the probability of currently having a partner. The female participation rate comes out significant and may be picking up cultural factors when country-dummies are not employed.²²

Turning to the aggregate divorce rate and to social transfers, both have negative signs, but only the aggregate divorce rate is significant. The negative effect of the aggregate divorce rate suggests the presence of a “social multiplier” (Becker and Murphy, 2000) arising from strategic complementarities in marriage market behavior. Checking the predictions from the model did not reveal any strong geographical patterns of over- and under-predictions.

More interesting for our purposes is the “income-pooling equation”, the results for which are presented in column 1 of Table 1. The marginal effects reported in this case are the direct effects of the regressors in w on the conditional probability of pooling income,

$$\frac{\partial \Pr(q = 1|z = 1; x; w)}{\partial w} = \frac{\phi_2(x; w; \frac{1}{2})}{\phi(x)}; \quad (10)$$

where $\phi_2(\cdot; \cdot; \cdot)$ and $\phi(\cdot)$ are the cdfs for the bivariate and univariate standard normal distribution, respectively.

Age affects the probability of income pooling positively over most of the range; this is natural if there is learning about an underlying match-quality and/or unobserved relationship investments since “age” picks up the effect of (unobserved) duration. Higher education negatively affects income-pooling; this may be due to short duration. However, it may also be that education acts to stabilize the income streams reducing the need for risk-sharing. Income negatively affects income pooling.²³ Household size is positively associated with income pooling which is consistent with the stabilizing effect of relationship investments (and that kids appear when the partners perceive that their match-quality is good).

²²Its measured effect may also reflect the focus on full-time employed couples.

²³See Coate and Ravallion (1993) for comparative statics on reciprocity-based income-sharing.

Prior divorce, both for the respondent as well as the respondent's partners, strongly reduces the probability of income pooling. Though prior divorce plays no direct role in the theory above, it may play a role if there is learning and/or relation-specific investments since prior divorce signals short duration of the current relationship. The labour market variables – the female participation rate and employment protection – have no significant effects on income-pooling.

Turning to the main variables of interest, social transfers has a significant negative effect on income pooling, suggesting that public insurance indeed crowd out informal income-sharing within partnerships. The aggregate divorce rate also negatively affects income pooling: a couple living in an economy where breakups are more frequent would thus appear to be less likely to cooperate financially than an otherwise identical couple. Finally the estimated correlation coefficient $\frac{1}{2}$ is positive which is consistent with both decisions being driven by a common unobserved match-quality.²⁴

In conclusion the results – though best cautiously interpreted – seem consistent with the theory. Social spending appears to affect both family formation decisions, as well as financial cooperation by partners. In addition, the structure of the marriage market appears to affect financial cooperation with a high turnover reducing income pooling.

IX Conclusions

It is sometimes argued that welfare state arrangements break up families, partly because they take on some of the functions otherwise performed by the family. This paper constructs a stylized model of marriage and divorce in which partners can cooperate financially by sharing earnings-risk. Risk-sharing between partners is supported by expected reciprocity. Any policy change that increases the relative attractiveness of divorce will not only directly increase the individuals' proneness to choose singlehood, but will also reduce the level of cooperation between partners which in turn further reduces the relative attractiveness of marriage and so on. Hence the analysis suggests that the nature of reciprocity-based cooperation between partners implies that changes in welfare policies can have strong effects on both family formation and dissolution patterns, as well as on cooperation. The model thus highlights how reciprocity-based financial cooperation between partners can have important implications for policy-design.

The model further reconciles the observation that there is a fairly low correlation in a simple

²⁴A Wald test of the hypothesis that the equations are independent is rejected at the 8 percent level.

cross-country comparison of welfare spending and e.g. divorce rates with the claim that publicly provided insurance monotonically affects family formation and dissolution decisions. The mechanism underlying the last result is a standard complementarity in search in the marriage market. As a consequence an economy can sustain multiple steady state equilibria differing in the rate of turnover in the marriage market and in the role of partners in providing financial security.

A decentralized steady state equilibrium generally fails to be locally inefficient, and it is argued that the presence of reciprocity-based cooperation implies that the direction of inefficiency is ambiguous: an equilibrium may have too few married individuals.

Appendix

Proof of Claim 1. Let $B(\mathcal{E})$ denote the space of bounded real-valued functions on \mathcal{E} , and endow this space with the sup norm. Blackwell (1965) provides the following sufficient condition for $T : B(\mathcal{E}) \rightarrow B(\mathcal{E})$ to be a contraction mapping (with modulus $\bar{\alpha}$):

1. (monotonicity) $f, g \in B(\mathcal{E})$ and $f(\mu) \leq g(\mu)$ for all $\mu \in \mathcal{E}$ implies $(T_f)(\mu) \leq (T_g)(\mu)$ for all $\mu \in \mathcal{E}$, and
2. (discounting) there exists some $\bar{\alpha} \in (0, 1)$ such that $(T_{f+a})(\mu) \leq (T_f)(\mu) + \bar{\alpha}a$ for all $f \in B(\mathcal{E})$; $\mu \in \mathcal{E}$ and $a \geq 0$.

Define a mapping $T : B(\mathcal{E}) \rightarrow B(\mathcal{E})$ in the following way: for each $f \in B(\mathcal{E})$ let

$$(T_f)(\mu_i) = \max_{s \in \mathcal{S}} \left\{ V(s) ; \max_{\theta \in \mathcal{A}_i^f} \left[\lambda(\theta) + \mu_i + \sum_{j=0}^{\infty} \beta^j f(\mu_j) \right] \right\} \quad (A1)$$

where $\theta \in \mathcal{A}_i^f$ if and only if $\theta \in \mathcal{A}$ and, for all $m > k$ such that $\theta_{mk} > 0$,

$$u(y^m) + \theta_{mk} y^m + \sum_{j=0}^{\infty} \beta^j f(\mu_j) \leq u(y^k) + \mu_i + \sum_{j=0}^{\infty} \beta^j f(\mu_j) \leq u(y^m) + \mu_i + \sum_{j=0}^{\infty} \beta^j V(s) \quad (A2)$$

Lemma A.1. T is a contraction mapping.

Proof. The proof uses Blackwell's sufficiency conditions. Consider first "monotonicity". From Equation (A2) it follows that $f(\mu) \leq g(\mu)$ for all $\mu \in \mathcal{E}$ implies $\mathcal{A}_i^g \subseteq \mathcal{A}_i^f$ for all i ; "monotonicity" then follows immediately from Equation (A1).

Consider then "discounting". Let $i_i(a) \sim \max_{\theta \in 2A_i^{f+a}} \hat{A}(\theta) + \pm a$ and note that $(T_{f+a})(\mu_i) = \max_{s \in S} fV(s); i_i(a) + K_i g$ where K_i does not depend on a . If $i_i(a)$ then always grows at a rate less than unity, "discounting" holds. Totally differentiating (1) and substituting for $d^{\otimes_{mk}} da$ using (A2) yields

$$i_i'(a) = \pm \sum_{m=2}^{\infty} \sum_{k=1}^{\infty} \mathbb{1}_{mk} g_m g_k \frac{u^0(y^k) + \otimes_{mk} i y^m i y^k}{u^0(y^m i \otimes_{mk} (y^m i y^k))} i^{-1} + \pm;$$

where $\mathbb{1}_{mk} = 1$ if the mk 'th incentive constraint is relaxed by the increase in a and else is zero. Thus, since $\otimes_{mk} \in [0; 1=2]$ and $y^m > y^k$,

$$i_i'(a) < \pm \sum_{m=2}^{\infty} \sum_{k=1}^{\infty} g_m g_k \frac{u^0(y^k)}{u^0(y^m)} i^{-1} + \pm < (1 i^{-1}) + \pm = 1,$$

where the second inequality follows from Assumption 2. #

Since T is a contraction mapping it has a unique fixed point $f^* \in B(\epsilon)$ and furthermore, $T_h^n \rightarrow f^*$ as $n \rightarrow \infty$ for any $h \in B(\epsilon)$ ("the method of successive approximations"). From (2) and (3) $T_V = V$; thus V exists and is unique. Then apply the method of successive approximations: define $V_0(\mu) = \mu$ for all $\mu \in \mathcal{E}$. For $n \geq 1$, V_n is recursively defined: $V_n = T_{V_{n-1}}$ (implying that $V_n = T_{V_0}^n$). Since \mathcal{E} is ordered increasingly $V_0(\mu_j)$ increases in j . Assume then that $V_{n-1}(\mu_j)$ also increases in j . Using stochastic dominance (Assumption 1), $\prod_{j=0}^n \mathbb{1}_{ij} V_{n-1}(\mu_j)$ then increases in i , whereby $i > j$ implies $A_j^{V_{n-1}} \leq A_i^{V_{n-1}}$. Consequently

$$V_n(\mu_i) = \max_{s \in S} fV(s); \max_{\theta \in 2A_i^{V_{n-1}}} \hat{A}(\theta) + \mu_i + \pm \sum_{j=0}^n \mathbb{1}_{ij} V_{n-1}(\mu_j),$$

increases in i . By induction on n , $V_n(\mu_i)$ increases in i for all n , whereby $V(\mu_i) = \lim_{n \rightarrow \infty} V_n(\mu_i)$ also increases in i . Since \mathcal{E} is ordered increasingly, this is equivalent to $V(\cdot)$ being increasing in μ . ■

Proof of Claim 2. Since the incentive constraints are independent of each other (see Equation (3)) A_i can be expressed as follows: $A_i = \mathcal{E}_{m>k} [0; \otimes_{itmk}]$ where each \otimes_{itmk} is an upper bound in the range $[0; 1=2]$. Furthermore, trivially $\otimes(\mu_i) = \otimes_i \sim (\otimes_{itmk})_{m>k}$. Using Claim 1 and stochastic dominance (Assumption 1), $i > j$ implies $A_j \leq A_i$ whereby $\otimes_j \leq \otimes_i$. ■

Proof of Claim 3. The Claim follows if $\Phi(\mu) \sim V(\mu) \leq V(s)$ decreases in ζ for all $\mu \in \mathcal{E}$. The argument is by induction. Consider an n -period approximation to original problem: Suppose the couple must divorce after n periods, but can divorce at any time before that. (The reader might argue that no risk-sharing can be sustained if the horizon is known to be finite. However, that

relies on a subgame perfection argument that does not invalidate the approximation.) Let $V_n(\mu)$ denote the value of being in state μ with a maximum of n periods remaining; then as the horizon n goes to infinity its impact will vanish. Since no remarriage is possible $V(s) = \bar{A}(0) = (1 - \beta)V(s)$.

When $n = 0$, $V_0(\mu) = V(s)$ for all $\mu \in \mathcal{E}$, and $\Phi_0(\mu) = V_0(\mu) - V(s)$ trivially (weakly) decreases in ζ for all $\mu \in \mathcal{E}$. Assume then that $\Phi_{n-1}(\mu)$ decreases in ζ for all $\mu \in \mathcal{E}$. For $n \geq 1$, $V_n(\mu)$ satisfies the following recursive definition:

$$V_n(\mu_i) = \max_{\theta \in \mathcal{A}_i^{V_{n-1}}} \left\{ V(s); \max_{\theta \in \mathcal{A}_i^{V_{n-1}}} \bar{A}(\theta) + \mu_i + \beta \sum_{j=0}^J \frac{1}{2} \theta_{ij} V_{n-1}(\mu_j) \right\}; \quad (A3)$$

where $\theta \in \mathcal{A}_i^{V_{n-1}}$ if and only if $\theta \in \mathcal{A}$ and, for all $m > k$ such that $\theta_{mk} > 0$,

$$u(y^2) - \theta_{ij} u(y^2) - \theta_{ij} u(y^1) \leq u(y^2) + \mu_i + \beta \sum_{j=0}^J \frac{1}{2} \theta_{ij} \Phi_{n-1}(\mu_j) \leq 0; \quad (A4)$$

Using that $(1 - \beta)V(s) = \bar{A}(0)$, it follows that

$$\Phi_n(\mu_i) = \max_{\theta \in \mathcal{A}_i^{V_{n-1}}} \left\{ 0; \bar{A}_n^s(\mu_i) - \bar{A}(0) + \mu_i + \beta \sum_{j=0}^J \frac{1}{2} \theta_{ij} \Phi_{n-1}(\mu_j) \right\}; \quad (A5)$$

where $\bar{A}_n^s(\mu_i) = \max_{\theta \in \mathcal{A}_i^{V_{n-1}}} \bar{A}(\theta)$. Thus if $\bar{A}_n^s(\mu_i) - \bar{A}(0)$ decreases in ζ , $\Phi_n(\mu_i)$ will also decrease in ζ . If complete risk-sharing is sustainable, $\bar{A}_n^s(\mu_i) = \bar{A}(1=2)$; but $\bar{A}(1=2) - \bar{A}(0)$ naturally decreases in ζ since the additional formal insurance is more valuable when no risk-sharing is available. Suppose then that risk-sharing is incentive constrained implying that (A4) holds with equality. An increase in ζ decreases y^2 as well as $\Phi_{n-1}(\mu_j)$ for every j ; from (A4) (using $u'' < 0$) the self-enforceability constraint is therefore tightened, forcing a reduction in the absolute transfer $\theta_{ij} y^2 - \theta_{ij} y^1$. $\bar{A}_n^s(\mu_i) - \bar{A}(0)$ then decreases in ζ also due to the crowding out effect on private risk-sharing. By induction on n , $\Phi_n(\mu)$ then decreases in ζ for every $\mu \in \mathcal{E}$ and n . Letting n go to infinity $\Phi(\mu) = \lim_{n \rightarrow \infty} \Phi_n(\mu)$ also decreases in ζ for each $\mu \in \mathcal{E}$. ■

Proof of Claim 4. The proof uses that \bar{A} affects $V(\mu)$ only through $V(s)$. Thus start by treating $V(s)$ as parametrically given and note that:

Lemma A.2. $\Phi(\mu) = V(\mu) - V(s)$ and $\Phi(\mu)$ decreases in $V(s)$ for all $\mu \in \mathcal{E}$.

Proof. The proof uses the same n -period approximation as used in the proof of Claim 3. For $n = 0$, $V_0(\mu) = V(s)$ for all $\mu \in \mathcal{E}$ while for $n \geq 0$, $V_n(\mu)$ is defined recursively as in (A3), where now $\theta \in \mathcal{A}_i^{V_{n-1}}$ if and only if $\theta \in \mathcal{A}$ and, for all $m > k$ such that $\theta_{mk} > 0$,

$$u(y^m) - \theta_{mk} u(y^m) - \theta_{mk} u(y^k) \leq u(y^m) + \mu_i + \beta \sum_{j=0}^J \frac{1}{2} \theta_{ij} \Phi_{n-1}(\mu_j) \leq 0; \quad (A6)$$

$\Phi_0(\mu) = V_0(\mu) \pm V(s)$ trivially (weakly) decreases in $V(s)$ for all μ . Assume then that $\Phi_{n-1}(\mu)$ decreases in $V(s)$ for all μ . From (A6) the set $A_i^{V_{n-1}}$ then decreases in $V(s)$. Note that

$$\Phi_n(\mu_i) = \max_{\substack{0; \\ \text{max}_{\text{in } A_i^{V_{n-1}}} \text{max}_{\text{in } A_i^{V_{n-1}}}}} \left[\bar{A}(\mu) + \mu_i \pm \sum_{j=0}^{N-1} \frac{1}{2} \Phi_{n-1}(\mu_j) \pm (1 \pm \epsilon) V(s) \right] \quad (\text{A7})$$

Hence $\Phi_n(\mu_i)$ decreases in $V(s)$. Moreover, as noted in the proof of Claim 2, $A_i^{V_{n-1}}$ can be expressed as the cross-product of $M(M-1)/2$ intervals where the set of upper bounds is the optimal risk-sharing agreement. Then since $A_i^{V_{n-1}}$ decreases in $V(s)$ for all i it follows that $\bar{A}_n(\mu)$ decreases (component by component) in $V(s)$ for all μ . By induction on n it follows that $\Phi_n(\mu)$ as well as $\bar{A}_n(\mu)$ decreases in $V(s)$ for all n and μ . Letting $n \rightarrow \infty$ the result follows. #

Lemma A.3. $V(s)$ increases in \bar{A} .

Proof. Subtracting $V(s)$ from both sides of (7) shows that:

$$\frac{(1 \pm \epsilon) V(s) \pm \bar{A}(0)}{\pm \sum_{j=0}^{N-1} \frac{1}{2} \Phi(\mu_j)} = \bar{A} \quad (\text{A8})$$

holds identically. Noting that l.h.s. increases in $V(s)$ (by Lemma A.2) the result follows. #

Combining Lemma A.3 and Lemma A.2 and noting that divorce is optimal whenever $\Phi(\mu) = 0$, monotonicity of \bar{p} and $\bar{A}(\mu)$ in \bar{A} follows. ■

Proof of Claim 6. Since \bar{E} is discrete and each component of the composite mapping $S \rightarrow \bar{p}(\bar{A}(\bar{c}))$ are increasing, the composite mapping is an increasing step-function and is hence “continuous but for upward jumps”. Furthermore, since \bar{p} increases in \bar{c} the composite mapping increases in \bar{c} ; then from Milgrom and Roberts (1994, Corollary 1) it follows that the lowest and the highest fixed point, $S_L(\bar{c})$ and $S_H(\bar{c})$, both increase in \bar{c} . ■

References

- Aiyagari, S. R., Greenwood, J. & Guner, N. (2000), ‘On the state of the union’, *Journal of Political Economy* 108, 213–244.
- Attanasio, O. & Rios-Rull, J.-V. (2000), ‘Consumption smoothing in island economies: Can public insurance reduce welfare?’, *European Economic Review* 44, 1225–1258.
- Becker, G. S. (1973), ‘A theory of marriage: Part I’, *Journal of Political Economy* 81, 813–846.

- Becker, G. S. (1991), *A Treatise on the Family*, 2nd edn, Harvard University Press, Cambridge, Mass.
- Becker, G. S., Landes, E. & Michael, R. (1977), 'An economic analysis of marital instability', *Journal of Political Economy* 85, 1141–1187.
- Becker, G. S. & Murphy, K. (2000), *Social Economics: Market Behavior in a Social Environment*, Belknap Press of Harvard University Press, Cambridge, MA.
- Berry-Cullen, J. & Gruber, J. (2000), 'Does unemployment insurance crowd out spousal labor supply?', *Journal of Labor Economics* 18, 546–572.
- Blackwell, D. (1965), 'Discounted dynamic programming', *Annals of Mathematical Statistics* 36, 226–235.
- Browning, M. & Chiappori, P. A. (1996), 'Efficient intra-household allocations: A general characterization and empirical tests', *Econometrica* 66, 1241–1278.
- Bumpass, L. L. & Sweet, J. A. (1989), 'National estimates of cohabitation', *Demography* 26, 615–625.
- Burdett, K. & Coles, M. (1997), 'Marriage and class', *Quarterly Journal of Economics* 112, 141–168.
- Burdett, K. & Ermisch, J. (2001), 'Matching in the marriage market and non-marital childbearing'. Working Paper No 10/2001, Centre for Household, Income, Labour and Demographic Economics (CHILD).
- Burdett, K., Imai, R. & Wright, R. (1999), 'Unstable relationships'. University of Essex, Mimeo.
- Chiappori, P. A. & Weiss, Y. (2000), 'Marriage contracts and divorce: An equilibrium analysis'. University of Chicago, Mimeo.
- Coate, S. & Ravallion, M. (1993), 'Reciprocity without commitment: Characterization and performance of informal insurance arrangements', *Journal of Development Economics* 40, 1–24.
- Diamond, P. A. (1982), 'Aggregate demand management in search equilibrium', *Journal of Political Economy* 90, 881–894.

- DiTella, R. & MacCulloch, R. (1999), 'Informal family insurance and the design of the welfare state', *Economic Journal* . Forthcoming.
- Drewianka, S. (2000), 'Proposed reforms of marital institutions'. University of Chicago, Mimeo.
- Foster, A. D. & Rosenzweig, M. R. (2000), 'Imperfect commitment, altruism and the family: Evidence from transfer behavior in low-income rural areas', *Review of Economics and Statistics* . Forthcoming.
- Greenwood, J., Guner, N. & Knowles, J. (2000a), 'A macroeconomic analysis of marriage, fertility, and the distribution of income'. University of Pennsylvania, Mimeo.
- Greenwood, J., Guner, N. & Knowles, J. (2000b), 'Women on welfare: A macroeconomic analysis', *American Economic Review: Papers and Proceedings* 90, 383–388.
- Hess, G. D. (2001), 'Marriage and consumption insurance: What's love got to do with it?'. Oberlin College, Mimeo.
- Jovanovic, B. (1979), 'Job-matching and the theory of turnover', *Journal of Political Economy* 87, 972–990.
- Kimball, M. (1988), '"Farmers" cooperatives as behavior toward risk', *American Economic Review* 78, 224–232.
- Kocherlakota, N. R. (1996), 'Implications of efficient risk sharing without commitment', *Review of Economic Studies* 63, 595–609.
- Ligon, E., Thomas, J. P. & Worrall, T. (1997), 'Informal insurance arrangements in village economies'. Keele University, Department of Economics, Working Paper 97/08.
- Ligon, E., Thomas, J. P. & Worrall, T. (1998), 'Mutual insurance, individual savings and limited commitment'. University of St Andrews, Mimeo.
- Lundberg, S. & Pollak, R. A. (1994), 'Noncooperative bargaining models of marriage', *American Economic Review (Papers and Proceedings)* 84, 132–137.
- Lundberg, S. & Pollak, R. A. (2001), 'Efficiency in marriage'. NBER Working Paper No. W8642.
- Milgrom, P. & Roberts, J. (1994), 'Comparing equilibria', *American Economic Review* 84, 441–459.

- Mortensen, D. T. (1988), 'Matching: Finding a partner for life or otherwise', *American Journal of Sociology* 94, S215–S240.
- Peters, E. (1986), 'Marriage and divorce: Informational constraints and private contracting', *American Economic Review* 74, 437–454.
- Weiss, Y. (1997), The formation and dissolution of families: Why marry? Who marries whom? And what happens upon divorce?, in M. R. Rosenzweig & O. Stark, eds, 'Handbook of Population and Family Economics', Elsevier Science B.V., Amsterdam.
- Weiss, Y. & Willis, R. (1997), 'Match quality, new information and marital dissolution', *Journal of Labor Economics* 15, S293–S329.

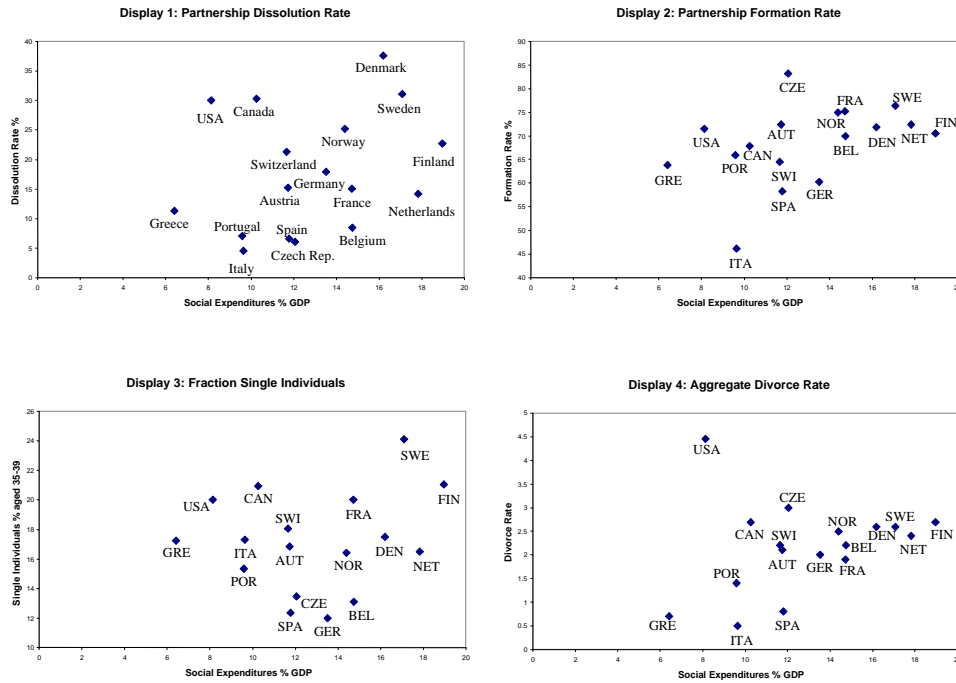


Figure 1: Social expenditures (percent of GDP in 1994), and measures of partnership formation and dissolution rates, partnership frequency, and aggregate divorce rate.

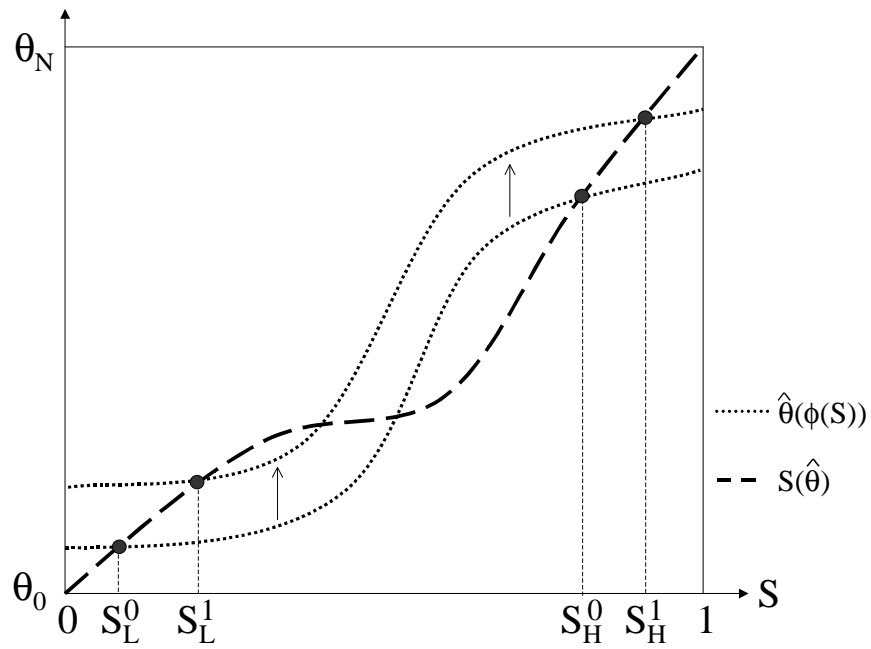


Figure 2: An increase in public insurance increases the steady state fraction of single individuals.

	Income Pooling		Partner	
	Coeff.	Marg. Eff.	Coeff.	Marg. Eff.
Gender (male)	0:028 (0:058)	0:011	i 0:138 (0:035)	i 0:052
Age	0:135 (:048)	0:057	0:270 (0:010)	0:102
Age Squared	i 0:001 (0:0006)	i 0:0006	i 0:003 (0:0001)	i 0:001
Primary Education	0:001 (0:065)	0:0006	i 0:133 (0:044)	i 0:051
University (male)	i 0:137 (0:059)	i 0:057	i 0:030 (0:046)	0:011
University (female)	i 0:274 (0:064)	i 0:114	i 0:299 (0:047)	i 0:113
Self-employed	0:042 (0:056)	0:017	0:043 (0:045)	0:016
Log Household-Earnings	i 0:125 (0:040)	i 0:052	i	i
Household Size 3-4	0:319 (0:063)	0:133	i	i
Household Size >4	0:490 (0:098)	0:205	i	i
Divorced	i 0:702 (0:137)	i 0:293	i 0:961 (0:042)	i 0:365
Partner Divorced	i 0:130 (0:063)	i 0:054	i	i
Religious Services	0:044 (0:039)	0:018	0:016 (0:031)	0:006
Social Transfers % of GDP	i 0:016 (0:007)	i 0:007	i 0:007 (0:0057)	i 0:003
Aggregate Divorce Rate	i 0:075 (0:032)	i 0:031	i 0:066 (0:025)	i 0:025
Female Partication Rate	i 0:006 (0:005)	i 0:003	0:011 (0:002)	0:004
Employment Protection	0:023 (0:027)	0:01	0:024 (0:023)	0:009
Constant	i 1:366 (1:523)		i 5:488 (0:221)	
Censored obs: 3058	Wald Test of Indep. Eq.: $\hat{A}^2(1) = 3:24$			
Uncensored obs: 4721	Log L = -6949.75			

Table 1: Effects on the probability of having a partner and the probability of pooling income given a partner. Estimation by bivariate probit with sample selection. Robust standard error in paranthesis.