Plagiarism in student papers and cheating on exams
Results from surveys using special techniques for sensitive questions

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Outline

- Sensitive questions in survey research
- Some indirect approaches to elicit truthful answers
  - The Randomized Response Technique (RRT)
  - The Crosswise Model: A new alternative to RRT
  - Stata implementation of estimators
- Empirical application: Plagiarism and cheating on exams
  - Three preliminary studies
  - Main study from 2011
- A little bit of magic
- Conclusions
Eliciting truthful answers to sensitive questions – not an easy task

- Direct questioning (DQ) does often not work . . .
Eliciting truthful answers to sensitive questions – not an easy task

- Some examples for proportion of “liars” (respondents with a false negative response) in surveys that use direct questioning (estimates from validation studies):
  - Penal conviction: 42.5% (F2F, Wolter 2010)
  - Welfare and unemployment benefit fraud: 75% (F2F, van der Heijden et al. 2000)
  - Driving under influence: 54% (P&P, Locander et al. 1976)
  - Bankruptcy: 32% (P&P, Locander et al. 1976)
Asking the Embarrassing Question

By Allen H. Barton

University of Chicago

The pollster’s greatest ingenuity has been devoted to finding ways to ask embarrassing questions in non-embarrassing ways. We give here examples of a number of these techniques, as applied to the question, “Did you kill your wife?”

1. The Casual Approach:
   “Do you happen to have murdered your wife?”
Some traditional approaches to measurement

2. The Numbered Card:
   Would you please read off the number on this card which corresponds to what became of your wife?” (HAND CARD TO RESPONDENT)
   1. Natural death
   2. I killed her
   3. Other (What?)
   (GET CARD BACK FROM RESPONDENT BEFORE PROCEEDING!)

3. The Everybody Approach:
   “As you know, many people have been killing their wives these days. Do you happened to have killed yours?”

   …

8. Putting the question at the end of the interview.
The Randomized Response Technique (RRT)
(Warner 1965; Fox and Tracy 1986)

- Main principle: privacy protection through randomization (i.e. add random noise to the answers)

- A randomizing device, the outcome of which is only known to the respondent, decides whether . . .
  - the sensitive question has to be answered
  - or an automatic “yes” or “no” has to be given or a surrogate question has to be answered

- Since only the respondent knows the outcome of the randomization device, a “yes” cannot be interpreted as an admission of guilt.

- However, if the properties of the randomizing device are known, a prevalence estimate for the sensitive question can be derived.
Example (forced response RRT)

\[ \text{Pr(observed yes)} = \text{Pr(sensitive question)} \cdot \pi + \text{Pr(surrogate yes)} \]

\[ \Rightarrow \pi = \frac{\text{Pr(observed yes)} - \text{Pr(surrogate yes)}}{\text{Pr(sensitive question)}} \]
The Crosswise Model (CM): A new alternative to RRT
(Yu, Tian, and Tang 2008)

Very simply idea: Ask a sensitive question and a nonsensitive question and let the respondent indicate whether . . .

A the answers to the questions are the same (both “yes” or both “no”)
B the answers are different (one “yes”, the other “no”)

<table>
<thead>
<tr>
<th></th>
<th>nonsensitive question</th>
</tr>
</thead>
<tbody>
<tr>
<td>sensitive question</td>
<td>no</td>
</tr>
<tr>
<td></td>
<td>A</td>
</tr>
<tr>
<td>yes</td>
<td>B</td>
</tr>
</tbody>
</table>

Note: Questions must be uncorrelated and probability of “yes” must be unequal 0.5 for the nonsensitive question.
The Crosswise Model (CM): A new alternative to RRT
(Yu, Tian, and Tang 2008)

- Prevalence estimate:

\[ \Pr(A) = (1 - \pi) \cdot (1 - \Pr(\text{nonsensitive yes})) + \pi \cdot \Pr(\text{nonsensitive yes}) \]

\[ \Rightarrow \pi = \frac{\Pr(A) + \Pr(\text{nonsensitive yes}) - 1}{2 \cdot \Pr(\text{nonsensitive yes}) - 1} \]

- Note: Crosswise Model is formally identical to Warner’s original RRT model.
The Crosswise Model: Let’s practice

Two questions:

1. Is your mother’s birthday in January or February?
2. Did you ever falsify your data or results?
   (e.g. edit data points or delete observations so that hypothesis is confirmed, falsify entire dataset, invent or manipulate reported results)

Compare your answers: Are they the same or different?

- Write “A” if they are the same (both Yes or both No)
- Write “B” if they are different (one Yes, the other No)
Results of survey:

- 35 A and 16 B (N = 51)
- Design parameter: \( \Pr(\text{nonsensitive yes}) = \frac{2}{12} \)
- Prevalence estimate:
  \[
  \hat{\pi} = \frac{\Pr(A) + \Pr(\text{nonsensitive yes}) - 1}{2 \cdot \Pr(\text{nonsensitive yes}) - 1} = \frac{35/51 + 2/12 - 1}{2 \cdot 2/12 - 1} = 22.1\%
  \]
- Standard error:
  \[
  SE(\hat{\pi}) = \sqrt{\frac{Pr(A) \cdot (1 - Pr(A))}{N - 1 \cdot (2 \cdot Pr(\text{nonsensitive yes}) - 1)^2}} = \sqrt{\frac{35/51 \cdot (1 - 35/51)}{(51 - 1) \cdot (2 \cdot 2/12 - 1)^2}} = 9.8\%
  \]
- 95% confidence interval: [2.3%, 41.8%]
Generalized estimator for RRT and CM

- Let
  
  \( Y_i \) response (\( Y_i = 1 \) if “yes” in RRT or “A” in CM, else \( Y_i = 0 \))

  \( \lambda_i \) probability of \( Y_i = 1 \)

  \( \pi_i \) (unknown) prevalence of sensitive item

  \( p_i^w \) probability of being directed to the negated question in Warner’s RRT (or prevalence of nonsensitive item in CM)

  \( p_i^{yes} \) overall probability of surrogate “yes”

  \( p_i^{no} \) overall probability of surrogate “no”

- Then

  \[ \lambda_i = (1 - p_i^{yes} - p_i^{no})p_i^w\pi_i + (1 - p_i^{yes} - p_i^{no})(1 - p_i^w)(1 - \pi_i) + p_i^{yes} \]

  and hence

  \[ \pi_i = \frac{\lambda_i - (1 - p_i^{yes} - p_i^{no})(1 - p_i^w) - p_i^{yes}}{(2p_i^w - 1)(1 - p_i^{yes} - p_i^{no})} \]
Two Stata commands (available from the SSC Archive)

```
rrreg depvar [indepvars] [if] [in] [weight] [ , regress_options
    pwarner(#|varname) pyes(#|varname) pno(#|varname) ]
```

- Assumes $\pi_i = X_i'\beta$ and estimates $\beta$ using least squares with
  transformed response

$$
\tilde{Y}_i = \frac{Y_i - (1 - p_{i,\text{yes}} - p_{i,\text{no}})(1 - p_{i,w}) - p_{i,\text{yes}}}{(2p_{i,w} - 1)(1 - p_{i,\text{yes}} - p_{i,\text{no}})}
$$
Two Stata commands (available from the SSC Archive)

```
rrlogit depvar [indepvars] [if] [in] [weight] [ , logit_options
    pwarner(# | varname) pyes(# | varname) pno(# | varname) ]
```

- Assumes $\pi_i = \frac{e^{X_i'\beta}}{(1 + e^{X_i'\beta})}$ and estimates $\beta$ using maximum likelihood with

$$
\ln L = \sum_{i=1}^{n} \left\{ Y_i \ln(R_i) + (1 - Y_i) \ln(S_i) - \ln(1 + e^{X_i'\beta}) \right\}
$$

where

$$
R_i = c_i + q_i e^{X_i'\beta} \quad c_i = (1 - p^{yes}_i - p^{no}_i)(1 - p^w_i) + p^{yes}_i
$$

$$
S_i = (1 - c_i) + (1 - q_i) e^{X_i'\beta} \quad q_i = (1 - p^{yes}_i - p^{no}_i)p^w_i + p^{yes}_i
$$
Three preliminary studies on plagiarism
(Jann/Jerke/Krumpal forthcoming, Coutts/Jann/Krumpal/Näher forthcoming)

- **Study 1**
  - Web-Survey among student of ETH Zurich in 2005
  - Response rate 33 Percent
  - Comparing direct questioning (DQ) to forced response RRT

- **Study 2**
  - Web-Survey among students of the University of Konstanz in 2009
  - Response rate 24 Percent
  - Comparing direct questioning (DQ) to the Item Count Technique (ICT) (yet another technique; ask me if you want to know more)

- **Study 3**
  - Classroom P&P survey at ETH Zurich, University of Leipzig, and LMU Munich in 2009
  - Comparing direct questioning (DQ) to the Crosswise Model (CM)
## Results (prevalence of plagiarism in percent)

<table>
<thead>
<tr>
<th>Study</th>
<th>DQ (N)</th>
<th>ICT (N)</th>
<th>Δ</th>
</tr>
</thead>
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<tr>
<td>Study 1</td>
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<td></td>
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<tr>
<td>in major papers</td>
<td>12.0 (2.0)</td>
<td>3.7 (4.0)</td>
<td>−8.3 (4.4)</td>
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<tr>
<td></td>
<td>N = 266</td>
<td>N = 495</td>
<td></td>
</tr>
<tr>
<td>in other papers</td>
<td>19.4 (1.4)</td>
<td>17.6 (2.4)</td>
<td>−1.8 (2.8)</td>
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<td></td>
<td>N = 826</td>
<td>N = 1521</td>
<td></td>
</tr>
<tr>
<td>Study 2</td>
<td>DQ (N = 396)</td>
<td>ICT (N = 846)</td>
<td>Δ</td>
</tr>
<tr>
<td>partial plagiarism</td>
<td>8.1 (1.4)</td>
<td>9.0 (4.0)</td>
<td>0.9 (4.2)</td>
</tr>
<tr>
<td>severe plagiarism</td>
<td>2.0 (0.7)</td>
<td>−4.0 (4.4)</td>
<td>−6.0 (4.5)</td>
</tr>
<tr>
<td>Study 3</td>
<td>DQ (N = 96)</td>
<td>CW (N = 310)</td>
<td>Δ</td>
</tr>
<tr>
<td>partial plagiarism</td>
<td>7.3 (2.7)</td>
<td>22.3 (5.5)</td>
<td>15.0 (6.1)</td>
</tr>
<tr>
<td>severe plagiarism</td>
<td>1.0 (1.0)</td>
<td>1.6 (5.0)</td>
<td>0.6 (5.1)</td>
</tr>
</tbody>
</table>
Summary of preliminary studies

- RRT and ICT do not seem to work well
- In particular, with the RRT, estimates of plagiarism are even lower than with direct questioning
- Reasons for the failure of RRT
  - difficulties understanding RRT, no trust in RRT
  - “self-protective no” bias
    - respondents who are not guilty are reluctant to give a “yes” answer and, hence, do not comply with the instructions
    - in RRT there is a “dominant strategy”: say “no”, no matter what
- The Crosswise Model works better
  - easier to understand
  - no obvious self-protective answering strategy
Main study

- Web-Survey among student of University of Bern and ETH Zurich in Spring 2011
- Response rate 33%
- Comparing direct questioning to three variants of RRT and two variants of the Crosswise Model
  - example
- Sensitive questions on
  - copying from other students in exam (copy)
  - using crib notes in exam (notes)
  - taking drugs to enhance performance on exam (drugs)
  - partial plagiarism (partial)
  - sever plagiarism/ghostwriting (severe)
- Team: Marc Höglinger, Ben Jann, Andreas Diekmann
Data analysis

prog mylincom
    ! version 1.0.1 08jul2009 Ben Jann
    syntax anything(everything equalok), Name(name) [ Level(passthru) ]
    lincom `anything´, `level´
    tempname b b1 V
    mat `b1´ = r(estimate)
    mat coln `b1´ = lincom:`name´
    mat `b´ = e(b)
    local eqs: coleq `b´, quoted
    local eqs: subinstr local eqs `"_"`` "main"`, all word
    mat coleq `b´ = `eqs´
    mat `b´ = `b´, `b1´
    mat `V´ = (e(V), J(rowsof(e(V)), 1, 0)) ///
        (J(1, colsof(e(V)), 0), r(se)^2)
    erepost b=`b´ V=`V´, rename
end

global sqvar copy notes drugs partial severe

forv i = 1/5 {
    local depvar: word `i´ of $sqvar
    di as res _n "==¿ depvar: `depvar´"
    rrreg `depvar´ DQ RRT CM, nocons hc2 pyes(pyesQ`i´) pno(pnoQ`i´) ///
        pwarner(pwarnQ`i´)
    mylincom RRT-DQ, name(RRT_DQ)
    mylincom CM-DQ, name(CM_DQ)
    eststo `depvar´
}
esttab, starkeep(lincom:) nonumb mti se b(1) compress transform(100*@ 100) ///
eqlab("Level" "Difference") coef(RRT_DQ "RRT - DQ" CM_DQ "CM - DQ")
## Results: Prevalence estimates

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<th>drugs</th>
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<td>17.5</td>
<td>8.8</td>
<td>3.4</td>
<td>2.5</td>
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<tr>
<td></td>
<td>(1.2)</td>
<td>(0.9)</td>
<td>(0.6)</td>
<td>(0.6)</td>
<td>(0.5)</td>
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<tr>
<td>RRT</td>
<td>19.6</td>
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<td>0.6</td>
<td>4.2</td>
<td>-0.6</td>
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<tr>
<td></td>
<td>(1.2)</td>
<td>(1.1)</td>
<td>(1.0)</td>
<td>(1.2)</td>
<td>(1.1)</td>
</tr>
<tr>
<td>CM</td>
<td>27.2</td>
<td>15.0</td>
<td>9.9</td>
<td>8.2</td>
<td>3.0</td>
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<td>(2.0)</td>
<td>(1.9)</td>
<td>(1.9)</td>
<td>(2.1)</td>
<td>(2.0)</td>
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<tr>
<td>RRT - DQ</td>
<td>2.1</td>
<td>3.9**</td>
<td>-2.8*</td>
<td>1.7</td>
<td>-2.1</td>
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<tr>
<td></td>
<td>(1.7)</td>
<td>(1.5)</td>
<td>(1.1)</td>
<td>(1.3)</td>
<td>(1.2)</td>
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<tr>
<td>CM - DQ</td>
<td>9.7***</td>
<td>6.2**</td>
<td>6.5***</td>
<td>5.7**</td>
<td>1.5</td>
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<tr>
<td></td>
<td>(2.3)</td>
<td>(2.1)</td>
<td>(2.0)</td>
<td>(2.2)</td>
<td>(2.1)</td>
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</tbody>
</table>

N          | 5726 | 5727  | 5711  | 4226    | 4224   |

Standard errors in parentheses
* p<0.05, ** p<0.01, *** p<0.001
Results: Determinants of sensitive behavior

<table>
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<th>notes</th>
<th>drugs</th>
<th>partial</th>
<th>severe</th>
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<tr>
<td>Perceived prevalence</td>
<td>0.053***</td>
<td>0.043***</td>
<td>0.042***</td>
<td>0.051***</td>
<td>0.072***</td>
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<tr>
<td></td>
<td>(0.003)</td>
<td>(0.003)</td>
<td>(0.006)</td>
<td>(0.006)</td>
<td>(0.014)</td>
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<tr>
<td>Perceived risk</td>
<td>-0.018***</td>
<td>-0.031***</td>
<td></td>
<td>0.002</td>
<td>-0.014</td>
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<tr>
<td></td>
<td>(0.005)</td>
<td>(0.008)</td>
<td></td>
<td>(0.005)</td>
<td>(0.016)</td>
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<td>Risk attitude</td>
<td>0.074*</td>
<td>0.088*</td>
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<td>0.104</td>
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<td></td>
<td>(0.032)</td>
<td>(0.036)</td>
<td>(0.064)</td>
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<td>RRT</td>
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<td>0.812*</td>
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<td>(0.172)</td>
<td>(0.334)</td>
<td>(0.350)</td>
<td>(0.679)</td>
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<td>CM</td>
<td>0.876***</td>
<td>0.774***</td>
<td>0.860**</td>
<td>1.515***</td>
<td>-0.711</td>
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<td></td>
<td>(0.178)</td>
<td>(0.212)</td>
<td>(0.322)</td>
<td>(0.377)</td>
<td>(1.895)</td>
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<td>-4.997***</td>
<td>-5.226***</td>
<td>-5.390***</td>
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<tr>
<td></td>
<td>(0.257)</td>
<td>(0.297)</td>
<td>(0.476)</td>
<td>(0.586)</td>
<td>(0.983)</td>
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N: 4956 4973 5270 3696 3162

Standard errors in parentheses
* p<0.05, ** p<0.01, *** p<0.001
Summary of main study

- Crosswise Model clearly outperforms direct questioning (if we are ready to accept the “more-is-better assumption”).
  - An exception is the last item (severe plagiarism), where prevalence is very low for all techniques.

- RRT, on the other hand, does not yield higher estimates than direct questioning
  - A reason might be the “self-protective no” bias, which prevents respondents to say “yes” if advised to do so by the randomizing device.
In variant 1, $\pi_\omega$ and $\gamma$ are identified (two equations, two unknowns). Variants 2 and 3 are not identified, I think (too many unknowns).

Let $\lambda_1$ and $\lambda_2$ be the observed proportion of "yes" answers in the two samples. An estimator for $\pi_\omega$ and $\gamma$ in variant 1 then is:

$$\hat{\gamma} = \lambda_1(1 - p_{\text{no}2}) + \lambda_2(p_{\text{no}1} - 1)$$

$$\pi_\omega = \lambda_1 - \hat{\gamma} p_{\text{yes}1} (1 - p_{\text{no}2}) - \hat{\gamma} p_{\text{yes}2} (1 - p_{\text{no}1})$$

**Main Assumptions:**

- **Monotonicity of social desirability:** Public opinion is always "no" if private opinion is "no"
- **No provocation:** Respondents do not say "yes" if advised to say "no"
A little bit of magic: Cheating detection in RRT

- Assuming that $\gamma$ and $\omega$ do not depend on $p^{\text{yes}}$ and $p^{\text{no}}$ (which may be justified if variation in $p$ is small) (and that $\gamma$ does not depend on the private opinion), this leads to the following log likelihood:

$$\ln L = \sum_{i=1}^{n} Y_i \ln(\ell_i) + (1 - Y_i) \ln(1 - \ell_i)$$

with

$$\ell_i = \pi_i \omega (1 - p_i^{\text{no}} - \gamma p_i^{\text{yes}}) + \gamma p_i^{\text{yes}}$$

- If $p^{\text{yes}}$ and $p^{\text{no}}$ are randomly varied between respondents, then $\pi_i \omega$ and $\gamma$ are identified.
A little bit of magic: Analysis

program define rrcheat lf
    args lnf theta1 cheat
    local p1 $rrcheat_pyes
    local p2 $rrcheat_pno
    quietly replace `lnf´ = cond($ML_y1, ///
    ln(`theta1´ * (1 - `p2´ - (1-`cheat´)*`p1´) + (1-`cheat´)*`p1´), ///
    ln(1 - (`theta1´ * (1 - `p2´ - (1-`cheat´)*`p1´) + (1-`cheat´)*`p1´)))
end
forv i = 1/5 {
    local depvar: word `i´ of $sqvar
    global rrcheat_pyes pyesQ`i´
    global rrcheat_pno pnoQ`i´
    ml model lf rrcheat lf (`depvar´: `depvar´ = ) /cheat if RRT==1
    ml maximize
    eststo `depvar´
}
esttab, nonumb nostar mti se b(1) transform(100*@ 100) ///
eqlab(none) coef(main:_cons "RRT adjusted" cheat:_cons "Cheaters")
### A little bit of magic: Results

<table>
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<tr>
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<th>copy</th>
<th>notes</th>
<th>drugs</th>
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<td>RRT adjusted</td>
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<td>16.7</td>
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<td>(5.6)</td>
<td>(6.6)</td>
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<td>Cheaters</td>
<td>-9.5</td>
<td>-3.6</td>
<td>88.9</td>
<td>54.3</td>
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Standard errors in parentheses

### Unadjusted results for comparison:

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<td>(0.6)</td>
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<td>3.0</td>
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<td>(1.9)</td>
<td>(1.9)</td>
<td>(2.1)</td>
<td>(2.0)</td>
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Methodological conclusions

- The Randomized Response Technique does not seem to be a good method for self-administered surveys. Although we put a lot of effort into pretesting and finding good implementations, no convincing evidence could be found that RRT yields more valid estimates than direct questioning.

- The Crosswise Model is a promising alternative, since it does not suffer from some of the deficiencies of the RRT ("self-protective no" bias, complexity).

- Improvement of RRT estimates is possible by correcting for cheating respondents who do not comply with the instructions. Such estimates, however, have low efficiency.
Substantive conclusions

- A substantial proportion of students have cheated on an exam (copying: 20 to 25 percent, crib notes: around 15 percent)
- Using drugs to enhance performance on exams is not uncommon (about 10 percent)
- Rates for partial plagiarism (using a passage from someone else’s work without providing proper citation) are about 10 percent. The prevalence of severe plagiarism (hand in someone else’s work) is about 2 percent.
- These numbers may not seem too high, but keep in mind:
  - There is lots of nonresponse, and probably mostly the “nice guys” participate.
  - Even with these low numbers we would expect about 200 papers a year containing plagiarism and about 40 papers, that are entirely falsified, at a small University with about 10000 Students.
Thank you for your attention!
References